

Chapter Five

Ryan Honea

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Exercise One

Question

Consider the single response variable Y with $Y \sim \text{Bin}(n, \pi)$.

Solution

(a): Find the Wald statistic $(\hat{\pi} - \pi)^T \mathcal{J}(\hat{\pi} - \pi)$, where $\hat{\pi}$ is the maximum likelihood estimator of π and \mathcal{J} is the information.

Solution: The most likely estimator is $\frac{Y}{n}$, and as this is a single response variable, we know the Wald statistic is $(\hat{\pi} - \pi)^2 \mathcal{J}$. We also saw from Example 5.22 that $\mathcal{J} = \frac{n}{\pi(1-\pi)}$.

So, we have $W = \left(\frac{Y}{n} - \pi\right)^2 \left(\frac{n}{\pi(1-\pi)}\right)$

(b): Verify that the Wald statistic is the same as the score statistic $\mathbf{U}^T \mathcal{J}^{-1} \mathbf{U}$ in this case (see Example 5.22).

Solution:

$$\begin{aligned} \frac{U^2}{\mathcal{J}} &= \frac{(Y - n\pi)^2}{n\pi(1-\pi)} \\ W &= \left(\frac{Y}{n} - \pi\right)^2 \left(\frac{n}{\pi(1-\pi)}\right) \\ &= \left(\frac{Y^2}{n^2} - \frac{2Y\pi}{n} + \pi^2\right) \left(\frac{n}{\pi(1-\pi)}\right) \\ &= \frac{1}{n^2} (Y^2 - 2Yn\pi + \pi^2 n^2) \left(\frac{n}{\pi(1-\pi)}\right) \\ &= \frac{1}{n^2} (Y - n\pi)^2 \left(\frac{n}{\pi(1-\pi)}\right) \\ &= \frac{(Y - n\pi)^2}{n\pi(1-\pi)} = \frac{U^2}{\mathcal{J}} \end{aligned}$$

(c): Find the deviance

$$2[l(\hat{\pi}; y) - l(\pi; y)].$$

$$\begin{aligned} l(\hat{\pi}; Y) &= \log \binom{n}{Y} + Y \log \hat{\pi} + (n - Y) \log(1 - \hat{\pi}) \\ l(\pi; Y) &= \log \binom{n}{Y} + Y \log \pi + (n - Y) \log(1 - \pi) \\ 2[l(\hat{\pi}; y) - l(\pi; y)] &= 2 \left[Y \log \left(\frac{\hat{\pi}}{\pi} \right) + (n - Y) \log \left(\frac{1 - \hat{\pi}}{1 - \pi} \right) \right] \end{aligned}$$

(d): For large samples, both the Wald/score statistic and the deviance approximately have the $\chi^2(1)$ distribution. For $n = 10$ and $y = 3$, use both statistics to assess the adequacy of the models:

- (1) $\pi = 0.1$
- (2) $\pi = 0.3$
- (3) $\pi = 0.5$

Do the two statistics lead to the same conclusions?

Solution:

I use the following functions to find my results.

```
wald <- function(y, n, p) {  
  return (((y - n*p)^2.0)/(n*p*(1.0-p)))  
}  
  
dev <- function(y, n, p) {  
  p_hat = y/n  
  return (2*(y*log(p_hat/p) + (n - y)*log((1-p_hat)/(1-p))))  
}
```

(1)	$W = 4.4$	$D = 3.073$
(2)	$W = 0$	$D = 0$
(3)	$W = 1.6$	$D = 1.65$

The 95th percentile of the χ^2 test is 3.84 and so our rejection region is $D > 3.84$ or $W > 3.84$. In the case of (1), W and D disagree, where W leads us to reject our hypothesis that $\pi = 0.1$ while D leads us to fail to reject. For (2) and (3), we reject the null hypothesis for both D and W .

Exercise Two

Question

Consider a random sample Y_1, \dots, Y_N with the exponential distribution

$$f(y_i, \theta_i) = \theta_i \exp(-y_i \theta_i)$$

Derive the deviance by comparing the maximal model with different values for θ_i for each Y_i and the model with $\theta_i = \theta$ for all i .

Solution

Note that $y_i = 1/\theta_i$ and $\hat{y}_i = 1/\theta$.

$$\begin{aligned} l(f(y_i; \theta_i)) &= \sum_{i=1}^n (\log \theta_i - y_i \theta_i) \\ l(f(y_i; \theta)) &= \sum_{i=1}^n (\log \theta - y_i \theta) \\ 2[l(f(y_i; \theta_i)) - l(f(y_i; \theta))] &= 2 \left[\sum_{i=1}^n \log \theta_i - y_i \theta_i - \log \theta + y_i \theta \right] \\ &= 2 \left[\sum_{i=1}^n \log(1/y_i) - 1 - \log(1/\hat{y}_i) + (y_i/\hat{y}_i) \right] \\ &= 2 \left[\sum_{i=1}^n \log \left(\frac{\hat{y}_i}{y_i} \right) + \frac{y_i}{\hat{y}_i} - 1 \right] \\ &= 2 \left[\sum_{i=1}^n \frac{y_i}{\hat{y}_i} - \log \left(\frac{y_i}{\hat{y}_i} \right) - 1 \right] \end{aligned}$$

Exercise Four

Question

For the leukemia survival data in 4.2:

Solution

(a): Use the Wald statistic to obtain an approximate 95% confidence interval for the parameter β_1 .

(b): By comparing the deviances for two appropriate models, test the null hypothesis $\beta_2 = 0$ against the alternative hypothesis $\beta_2 \neq 0$. What can you conclude about the use of the initial white blood cell count as a predictor of survival time.