

$$1. (a) \text{ LIBOR Market Model: } dL_i(t) = \sigma_i L_i(t) dW^{i+1}(t) \\ \Rightarrow L_i(t) = L_i(0) e^{-\frac{1}{2}\sigma_i^2 t + \sigma_i W^{i+1}(t)}$$

\therefore The forward LIBOR rate $D(t, T_i) = (1 + \Delta_i L_i(T_i, T_{i+1})) D(t, T_{i+1})$

$$\therefore L_t(T_i, T_{i+1}) = \frac{1}{\Delta_i} \cdot \frac{D(t, T_i) - D(t, T_{i+1})}{D(t, T_{i+1})}$$

denote $L_i(t) = L_t(T_i, T_{i+1})$, $D_i(t) = D(t, T_i)$

$$\Rightarrow \Delta_i L_i(t) = \frac{D_i(t) - D_{i+1}(t)}{D_{i+1}(t)},$$

this is a ratio of marketed assets.

If we take the discount bond $D_{i+1}(t)$ as numeraire, then under the martingale measure \mathbb{Q}^{i+1} associated with the numeraire $D_{i+1}(t)$, the process $\Delta_i L_i(t)$ must be a martingale.

(b) The Libor market model is defined as

$$dL_i(t) = \sigma_i L_i(t) dW^{i+1}(t)$$

where $W^{i+1}(t)$ is a standard Brownian motion under the risk neutral measure \mathbb{Q}^{i+1} , associated to the zero coupon discount bond $D_{i+1}(t) = D(t, T_{i+1})$.

The solution to the Libor market model is given by

$$L_i(T) = L_i(0) e^{-\frac{\sigma_i^2}{2} T + \sigma_i W_T^{i+1}}$$

Let V_T denote the value of the financial contract at time t . Under the martingale measure, we have

$$\frac{V_0}{D_{i+1}(0)} = \mathbb{E}^{i+1} \left[\frac{V_T}{D_{i+1}(T)} \right]$$

$$\begin{aligned} V_0 &= D(0, T_{i+1}) \mathbb{E}^{i+1} [\Delta_i \sqrt{L_i(T)}] \\ &= D(0, T_{i+1}) \Delta_i \mathbb{E}^{i+1} \left[\sqrt{L_i(0)} \cdot e^{-\frac{\sigma_i^2}{4} T + \frac{1}{2} \sigma_i W_T^{i+1}} \right] \\ &= D(0, T_{i+1}) \Delta_i \sqrt{L_i(0)} \cdot e^{-\frac{\sigma_i^2}{4} T + \frac{1}{8} \sigma_i^2 T} \\ &= D(0, T_{i+1}) \Delta_i \sqrt{L_i(0)} e^{-\frac{1}{8} \sigma_i^2 T} \end{aligned}$$

$$(c) \because V_T = 1_{K_1 \leq L_i(T) \leq K_2}$$

$$\therefore V_0 = D(0, T_{i+1}) \mathbb{E}^{i+1} [1_{K_1 \leq L_i(T) \leq K_2}]$$

Then, we apply CDF: the probability that the forward LIBOR rate $L_i(T_i)$ is within the range $[K_1, K_2]$ under the T_{i+1} forward measure is given by:

$$\begin{aligned} \mathbb{P}^{T_{i+1}}(K_1 \leq L_i(T) \leq K_2) &= \mathbb{P}^{T_{i+1}}(\ln K_1 \leq \ln L_i(T) \leq \ln K_2) \\ &= \mathbb{P}^{T_{i+1}}(\ln K_1 \leq \ln L_i(0) - \frac{1}{2} \sigma_i^2 T + \sigma_i W_T^{i+1} \leq \ln K_2) \\ &= \mathbb{P}^{T_{i+1}}\left(\frac{\ln K_1 - \ln L_i(0) + \frac{1}{2} \sigma_i^2 T}{\sigma_i \sqrt{T}} \leq \frac{\sigma_i W_T^{i+1}}{\sigma_i \sqrt{T}} \leq \frac{\ln K_2 - \ln L_i(0) + \frac{1}{2} \sigma_i^2 T}{\sigma_i \sqrt{T}}\right) \end{aligned}$$

$\therefore W_T^{i+1}$ is a standard Brownian motion, we recognize

$$\frac{W_T^{i+1}}{\sqrt{T}} = Z \sim N(0, 1)$$

$$\begin{aligned} \therefore \mathbb{P}^{T_{i+1}}(K_1 \leq L_i(T) \leq K_2) &= \mathbb{P}^{T_{i+1}}\left(\frac{\ln \frac{K_1}{L_i(0)} + \frac{1}{2} \sigma_i^2 T}{\sigma_i \sqrt{T}} \leq Z \leq \frac{\ln \frac{K_2}{L_i(0)} + \frac{1}{2} \sigma_i^2 T}{\sigma_i \sqrt{T}}\right) \\ &= \Phi(d_2) - \Phi(d_1) \end{aligned}$$

$$\therefore V_0 = D(0, T_{i+1}) [\Phi(d_2) - \Phi(d_1)],$$

$$\text{where } d_1 = \frac{\ln \frac{K_1}{L_i(0)} + \frac{1}{2} \sigma_i^2 T}{\sigma_i \sqrt{T}}$$

$$d_2 = \frac{\ln \frac{K_2}{L_i(0)} + \frac{1}{2} \sigma_i^2 T}{\sigma_i \sqrt{T}}$$

2.(a) The numeraire security associated with the risk-neutral measure $\mathbb{Q}^{n+1, N}$ is the present value of a basis point (PVBp) or equivalently the annuity $P_{n+1, N}(t)$, which is the price of zero-coupon bond maturing at T_{n+1} .

$$(b) V_{n, N}^{dig}(0) = P_{n+1, N}(T) S_{n, N}(T) \mathbb{1}_{S_{n, N}(T) > K}$$

$$\frac{V_{n, N}^{dig}(0)}{P_{n+1, N}(0)} = \mathbb{E}^{n+1, N} \left[\frac{V_{n, N}^{dig}(T)}{P_{n+1, N}(T)} \right]$$

$$\begin{aligned} V_{n, N}^{dig}(0) &= P_{n+1, N}(0) \mathbb{E}^{n+1, N} [S_{n, N}(T) \mathbb{1}_{S_{n, N}(T) > K}] \\ &= P_{n+1, N}(0) \mathbb{E}^{n+1, N} [S_{n, N}(0) e^{-\frac{1}{2} \sigma_{n, N}^2 T + \sigma_{n, N} W^{n+1, N}(T)} \mathbb{1}_{S_{n, N}(T) > K}] \end{aligned}$$

The condition $S_{n, N}(T) > K$ translate to:

$$S_{n, N}(0) e^{-\frac{1}{2} \sigma_{n, N}^2 T + \sigma_{n, N} W^{n+1, N}(T)} > K$$

$$\Rightarrow \ln S_{n,N}(0) - \frac{1}{2} \sigma_{n,N}^2 T + \sigma_{n,N} \sqrt{T} Z > \ln K$$

$$Z > \frac{\ln K - \ln S_{n,N}(0) + \frac{1}{2} \sigma_{n,N}^2 T}{\sigma_{n,N} \sqrt{T}}$$

$$Z > \frac{\ln\left(\frac{K}{S_{n,N}(0)}\right) + \frac{1}{2} \sigma_{n,N}^2 T}{\sigma_{n,N} \sqrt{T}} = \chi^*$$

Integral the expectation :

$$\begin{aligned} \mathbb{E}^{n+1,N}[S_{n,N}(T) \mathbb{1}_{S_{n,N} > K}] &= S_{n,N}(0) \int_{\chi^*}^{\infty} e^{-\frac{1}{2} \sigma_{n,N}^2 T + \sigma_{n,N} \sqrt{T} x} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= S_{n,N}(0) \cdot \frac{1}{\sqrt{2\pi}} \int_{\chi^*}^{\infty} e^{-\frac{1}{2} \sigma_{n,N}^2 T + \sigma_{n,N} \sqrt{T} x - \frac{x^2}{2}} dx \\ &= S_{n,N}(0) \cdot \frac{1}{\sqrt{2\pi}} \int_{\chi^*}^{\infty} e^{-\frac{1}{2} (x - \sigma_{n,N} \sqrt{T})^2} dx \\ &= S_{n,N}(0) \cdot [\Phi(\infty) - \Phi(\chi^* - \sigma_{n,N} \sqrt{T})] \\ &= S_{n,N}(0) \cdot [\Phi(-\chi^* + \sigma_{n,N} \sqrt{T})] \\ &= S_{n,N}(0) \cdot \Phi\left(\frac{\ln\left(\frac{S_{n,N}(0)}{K}\right) + \frac{1}{2} \sigma_{n,N}^2 T}{\sigma_{n,N} \sqrt{T}}\right) \end{aligned}$$

$$\therefore V_{n,N}^{\text{dig}}(0) = P_{n+1,N}(0) S_{n,N}(0) \Phi\left(\frac{\ln\left(\frac{S_{n,N}(0)}{K}\right) + \frac{1}{2} \sigma_{n,N}^2 T}{\sigma_{n,N} \sqrt{T}}\right)$$

(c) The forward swap rate at a martingale under the annuity measure $Q^{n+1,N}$, not the standard risk-neutral measure Q . Converting between these measures introduces a convexity adjustment due to lognormality of the swap rate. The expected value Q is not simply $S_{n,N}(0)$ but requires an adjustment term $e^{\frac{1}{2} \sigma_{n,N}^2 T}$.