

ST758, Homework 2

Due Oct 15, 2013

1. (a) Read in the file ‘longley.dat’ on course webpage, which has the response \mathbf{y} in the first column and six explanatory variables in the other columns.
- (b) Compute the 6×6 sample covariance matrix and call it \mathbf{V} .
- (c) Compute the 6×6 correlation coefficient matrix \mathbf{C} from \mathbf{V} . What do you observe in \mathbf{C} ?
- (d) Partition the matrix \mathbf{V} as

$$\mathbf{V} = \begin{pmatrix} \mathbf{V}_{11} & \mathbf{V}_{12} \\ \mathbf{V}_{21} & \mathbf{V}_{22} \end{pmatrix},$$

where the blocks \mathbf{V}_{ij} have size 3×3 . Compute $\mathbf{V}_{22} - \mathbf{V}_{21}\mathbf{V}_{11}^{-1}\mathbf{V}_{12}$, using Cholesky decomposition.

- (e) Compute $\mathbf{V}_{22} - \mathbf{V}_{21}\mathbf{V}_{11}^{-1}\mathbf{V}_{12}$ again, using sweeping.
 - (f) Assume linear model $\mathbf{y} \sim N(\mathbf{X}\boldsymbol{\beta}, \sigma^2\mathbf{I})$. Include an intercept in your model and compute the regression coefficients $\hat{\boldsymbol{\beta}}$, their standard errors, variance estimate $\hat{\sigma}^2$, fitted values $\hat{\mathbf{y}}$, and residuals $\hat{\mathbf{e}}$ using three methods – Cholesky, QR, and sweeping. Please compute them directly; you can use other “black-box” function, e.g. `lm()`, only to check.
 - (g) Let $\mathbf{X}_i \in \mathbb{R}^{n \times i}$ contain the first i columns of the design matrix. That is $\mathbf{X}_1 = \mathbf{1}_n$, $\mathbf{X}_2 = (\mathbf{1}_n, \mathbf{x}_1)$, $\mathbf{X}_3 = (\mathbf{1}_n, \mathbf{x}_1, \mathbf{x}_2)$, and so on. Compute $\mathbf{y}^T \mathbf{P}_{\mathbf{X}_1} \mathbf{y}$, $\mathbf{y}^T (\mathbf{P}_{\mathbf{X}_i} - \mathbf{P}_{\mathbf{X}_{i-1}}) \mathbf{y}$, $i = 2, \dots, 7$, and $\mathbf{y}^T (\mathbf{I}_n - \mathbf{P}_{\mathbf{X}}) \mathbf{y}$, where $\mathbf{P}_{\mathbf{A}}$ denotes the orthogonal projection onto the column space of a matrix \mathbf{A} . (Hint: don’t ever think about forming these projection matrices.)
2. Let $\mathbf{X} \in \mathbb{R}^{n \times p}$, $n \geq p$, be the design matrix in linear regression. The least squares solution is given by

$$\hat{\boldsymbol{\beta}} = \min_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_2^2.$$

- (a) Show that $\hat{\boldsymbol{\beta}}$ is a least squares solution if and only if it satisfies the normal equation $\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} = \mathbf{X}^T \mathbf{y}$.
- (b) If \mathbf{X} has rank $r < p$, the least squares solution is not unique. Show that any vector $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^- \mathbf{X}^T \mathbf{y}$, where $(\mathbf{X}^T \mathbf{X})^-$ is any generalized inverse, is a least squares solution, and the residual vector $\mathbf{y} - \mathbf{X} \hat{\boldsymbol{\beta}}$ is invariant to the choice of the generalized inverse.
- (c) Assume \mathbf{X} has rank $r < p$ and the *QR decomposition with (column) pivoting* yields

$$\mathbf{X} = \mathbf{Q} \begin{pmatrix} \mathbf{R}_{11} & \mathbf{R}_{12} \\ \mathbf{0}_{(n-r) \times r} & \mathbf{0}_{(n-r) \times (p-r)} \end{pmatrix} \mathbf{\Pi}^T, \quad (1)$$

where $\mathbf{Q} \in \mathbb{R}^{n \times n}$ is orthogonal, $\mathbf{R}_{11} \in \mathbb{R}^{r \times r}$ is upper triangular with positive diagonal entries, and $\mathbf{\Pi} \in \mathbb{R}^{p \times p}$ is a permutation matrix. Show that

$$\mathbf{X}^- = \mathbf{\Pi} \begin{pmatrix} \mathbf{R}_{11}^{-1} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} \end{pmatrix} \mathbf{Q}^T$$

is a generalized inverse of \mathbf{X} . Is this generalized inverse the Moore-Penrose inverse of \mathbf{X} ? The `qr()` function in R performs QR decomposition with (column) pivoting. Study the documentation of `qr()` and do the following on the Longley data in Q1.

- (d) Add an extra column (sum of the intercept and the first predictor) to the original design matrix such that $\mathbf{X}_{\text{new}} = (\mathbf{1}_n, \mathbf{x}_1, \mathbf{1}_n + \mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_6)$. Compute a least squares solution using the generalized inverse in (c) obtained from QR. Compare the least squares solution, fitted values, and residuals to those obtained in Q1(f).
- (e) Given the singular value decomposition (SVD)

$$\mathbf{X} = \mathbf{U} \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \mathbf{0}_{r \times (p-r)} \\ \mathbf{0}_{(n-r) \times r} & \mathbf{0}_{(n-r) \times (p-r)} \end{pmatrix} \mathbf{V}^\top,$$

where $\mathbf{U} \in \mathbb{R}^{n \times n}$ and $\mathbf{V} \in \mathbb{R}^{p \times p}$ are orthogonal, and $\boldsymbol{\Sigma}_{11} \in \mathbb{R}^{r \times r}$ is diagonal with positive diagonal entries. Show that

$$\mathbf{X}^+ = \mathbf{V} \begin{pmatrix} \boldsymbol{\Sigma}_{11}^{-1} & \mathbf{0}_{r \times (n-r)} \\ \mathbf{0}_{(p-r) \times r} & \mathbf{0}_{(p-r) \times (n-r)} \end{pmatrix} \mathbf{U}^\top$$

is the Moore-Penrose inverse of \mathbf{X} .

- (f) Show that $\hat{\boldsymbol{\beta}} = \mathbf{X}^+ \mathbf{y}$ is a least squares solution and has the minimum ℓ_2 norm among all least squares solutions.
- (g) Redo part (d) but using the Moore-Penrose inverse obtained from SVD. Compare the least squares solution, fitted values, and residuals to those obtained in Q2(d).

3. Let $\mathbf{X} \in \mathbb{R}^n$ be a random vector with i.i.d. standard normal entries.

- (a) Show that \mathbf{X} has density

$$f_{\mathbf{X}}(\mathbf{x}) = \left(\frac{1}{2\pi} \right)^{n/2} e^{-\mathbf{x}^\top \mathbf{x} / 2}.$$

- (b) Any affine transformation $\mathbf{Y} = \mathbf{A}\mathbf{X} + \boldsymbol{\mu}$ of \mathbf{X} , where $\mathbf{A} \in \mathbb{R}^{m \times n}$ and $\boldsymbol{\mu} \in \mathbb{R}^m$, is called a multivariate normal random vector. Assume $\text{rank}(\mathbf{A}) = m$. Show that

$$\mathbf{E}(\mathbf{Y}) = \boldsymbol{\mu}, \quad \text{Var}(\mathbf{Y}) = \mathbf{A}\mathbf{A}^\top = \boldsymbol{\Omega},$$

and the density of \mathbf{Y} is

$$f_{\mathbf{Y}}(\mathbf{y}) = \left(\frac{1}{2\pi} \right)^{m/2} |\det(\boldsymbol{\Omega})|^{-1/2} e^{-(\mathbf{y} - \boldsymbol{\mu})^\top \boldsymbol{\Omega}^{-1} (\mathbf{y} - \boldsymbol{\mu}) / 2}.$$

- (c) Suppose \mathbf{Y} is partitioned as two sub-vectors $\mathbf{Y} = (\mathbf{Y}_1^\top, \mathbf{Y}_2^\top)^\top$ and $\boldsymbol{\mu} = (\boldsymbol{\mu}_1^\top, \boldsymbol{\mu}_2^\top)^\top$ and $\boldsymbol{\Omega} = \begin{pmatrix} \boldsymbol{\Omega}_{11} & \boldsymbol{\Omega}_{12} \\ \boldsymbol{\Omega}_{21} & \boldsymbol{\Omega}_{22} \end{pmatrix}$ are partitioned compatibly. Show that the conditional distribution of \mathbf{Y}_2 given \mathbf{Y}_1 is normal with mean and variance

$$\begin{aligned} \mathbf{E}(\mathbf{Y}_2 \mid \mathbf{Y}_1) &= \boldsymbol{\Omega}_{21} \boldsymbol{\Omega}_{11}^{-1} (\mathbf{Y}_1 - \boldsymbol{\mu}_1) + \boldsymbol{\mu}_2 \\ \text{Var}(\mathbf{Y}_2 \mid \mathbf{Y}_1) &= \boldsymbol{\Omega}_{22} - \boldsymbol{\Omega}_{21} \boldsymbol{\Omega}_{11}^{-1} \boldsymbol{\Omega}_{12}. \end{aligned}$$

Let $\mathbf{Y} \sim N(\mathbf{0}_m, \boldsymbol{\Omega}_m)$, where $\boldsymbol{\Omega}_m = (r^{|i-j|})_{i,j}$. Use $m = 10$ and $r = 0.9$ in below.