

Ellipse and Gaussian Distribution

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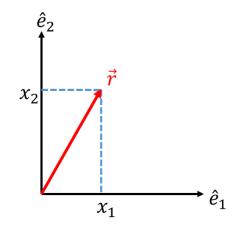


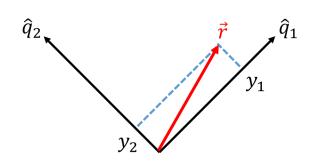
Coordinates



Coordinates with Basis

• Basis $\{\hat{e}_1 \ \hat{e}_2\}$ or basis $\{\hat{q}_1 \ \hat{q}_2\}$





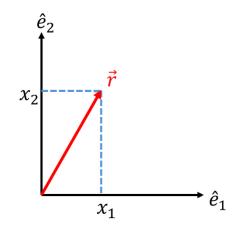
$$\overrightarrow{r}_I = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}$$
: coordinate of \overrightarrow{r} in basis $\{\hat{e}_1 \ \hat{e}_2\}$ $(=I)$

$$\overrightarrow{r}_{I} = \begin{bmatrix} x_{1} \\ x_{2} \end{bmatrix} : \text{coordinate of } \overrightarrow{r} \text{ in basis } \{\hat{e}_{1} \ \hat{e}_{2}\} \ (=I)$$

$$\overrightarrow{r}_{Q} = \begin{bmatrix} y_{1} \\ y_{2} \end{bmatrix} : \text{coordinate of } \overrightarrow{r} \text{ in basis } \{\hat{q}_{1} \ \hat{q}_{2}\} \ (=Q)$$

Coordinate Transformation

• Basis $\{\hat{e}_1 \; \hat{e}_2\}$ or basis $\{\hat{q}_1 \; \hat{q}_2\}$



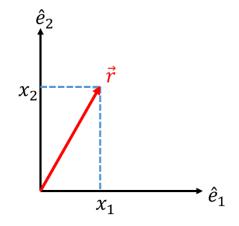
$$\hat{q}_2$$
 y_1
 \hat{q}_1

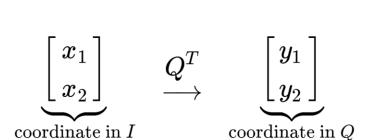
$$egin{aligned} \overrightarrow{r} &= x_1 \hat{e}_1 + x_2 \hat{e}_2 = y_1 \hat{q}_1 + y_2 \hat{q}_2 \ &= \left[egin{aligned} \hat{e}_1 & \hat{e}_2 \end{array}
ight] egin{bmatrix} x_1 \ x_2 \end{bmatrix} = \left[egin{aligned} \hat{q}_1 & \hat{q}_2 \end{array}
ight] egin{bmatrix} y_1 \ y_2 \end{bmatrix} \end{aligned}$$

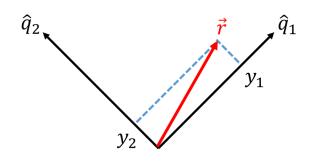
$$egin{aligned} \Longrightarrow egin{bmatrix} x_1 \ x_2 \end{bmatrix} &= Q egin{bmatrix} y_1 \ y_2 \end{bmatrix} \ \Longrightarrow egin{bmatrix} y_1 \ y_2 \end{bmatrix} &= Q^{-1} egin{bmatrix} x_1 \ x_2 \end{bmatrix} &= Q^T egin{bmatrix} x_1 \ x_2 \end{bmatrix} \end{aligned}$$

Coordinate Transformation

• Coordinate change to basis of $\{\hat{q}_1 \ \hat{q}_2\}$



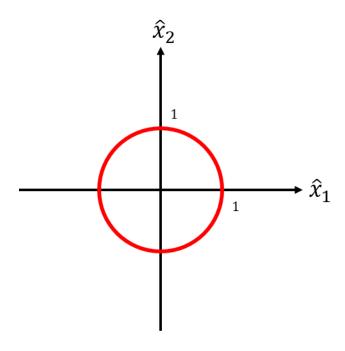




$$egin{aligned} \Longrightarrow egin{aligned} egin{aligned} & \Longrightarrow egin{aligned} egin{aligned} x_1 \ x_2 \end{bmatrix} = Q egin{bmatrix} y_1 \ y_2 \end{bmatrix} = Q^{-1} egin{bmatrix} x_1 \ x_2 \end{bmatrix} = Q^T egin{bmatrix} x_1 \ x_2 \end{bmatrix} \end{aligned}$$



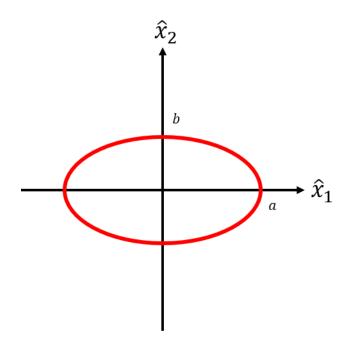
• Unit circle



$$x_1^2 + x_2^2 = 1 \implies$$

$$\left[egin{array}{cc} x_1 & x_2 \end{array}
ight] \left[egin{array}{cc} 1 & 0 \ 0 & 1 \end{array}
ight] \left[egin{array}{cc} x_1 \ x_2 \end{array}
ight] = 1$$

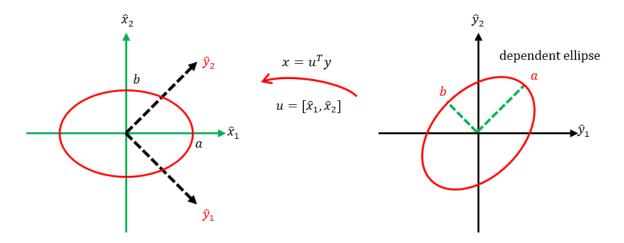
• Independent ellipse



$$egin{aligned} rac{x_1^2}{a^2} + rac{x_2^2}{b^2} &= 1 \implies [egin{array}{ccc} x_1 & x_2 \end{bmatrix} egin{bmatrix} rac{1}{a^2} & 0 \ 0 & rac{1}{b^2} \end{bmatrix} egin{bmatrix} x_1 \ x_2 \end{bmatrix} &= 1 \ \implies [egin{array}{ccc} x_1 & x_2 \end{bmatrix} \Sigma_x^{-1} egin{bmatrix} x_1 \ x_2 \end{bmatrix} &= 1 \end{aligned}$$

where
$$\Sigma_x^{-1}=egin{bmatrix}rac{1}{a^2}&0\0&rac{1}{b^2}\end{bmatrix},\ \Sigma_x=egin{bmatrix}a^2&0\0&b^2\end{bmatrix}$$

- Dependent ellipse (Rotated ellipse)
 - Coordinate changes

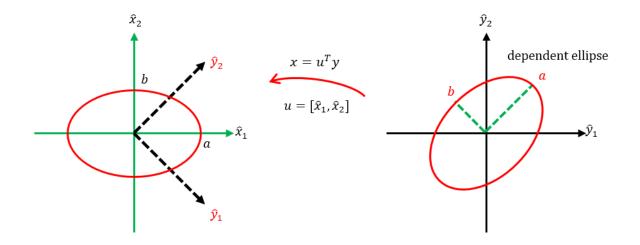


$$egin{bmatrix} x_1 \ x_2 \end{bmatrix} = u^T egin{bmatrix} y_1 \ y_2 \end{bmatrix}, & x = u^T y \ ux = y \end{bmatrix}$$

• Now we know in basis $\{\hat{x}_1, \hat{x}_2\} = I$

$$x^T \Sigma_x^{-1} x = 1 \quad ext{and} \quad \Sigma_x = egin{bmatrix} a^2 & 0 \ 0 & b^2 \end{bmatrix}$$

• Then, we can find Σ_{γ} such that



$$y^T \Sigma_y^{-1} y = 1 \quad \text{and} \quad \Sigma_y = ?$$

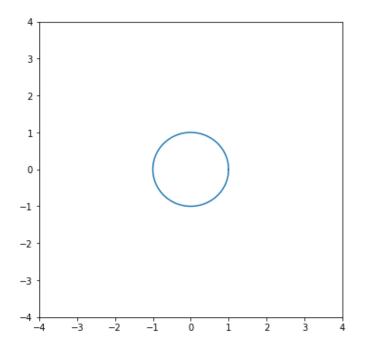
$$\implies x^T \Sigma_x^{-1} x = y^T u \Sigma_x^{-1} u^T y = 1 \quad (\Sigma_y^{-1} : \text{similar matrix to } \Sigma_x^{-1})$$

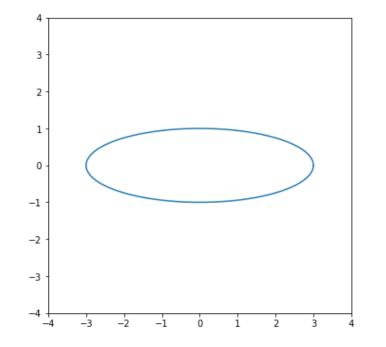
$$\Sigma_y^{-1} = u \Sigma_x^{-1} u^T \quad \text{or} \quad \Sigma_y = u \Sigma_x u^T$$

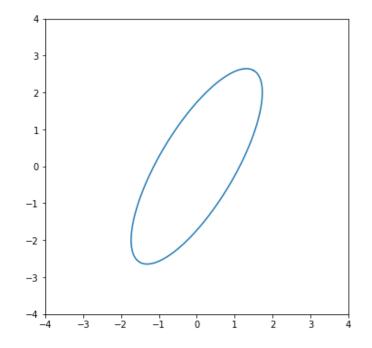
```
theta = np.arange(0,2*np.pi,0.01)
x1 = np.cos(theta)
x2 = np.sin(theta)
```

```
x1 = 3*np.cos(theta);
x2 = np.sin(theta);
```

```
\begin{array}{l} \text{u = np.array}([[1/2, \text{-np.sqrt(3)/2}], [\text{np.sqrt(3)/2}, 1/2]])} \\ \text{X = np.array}([\text{x1, x2}]) \\ \\ \text{u = np.asmatrix(u)} \\ \text{X = np.asmatrix(X)} \\ \text{y = u*X} \\ \\ \text{print(u)} \\ \end{array} \\ u = [\hat{x}_1 \ \hat{x}_2] = \begin{bmatrix} \frac{1}{2} & -\frac{\sqrt{3}}{2} \\ \frac{\sqrt{3}}{2} & \frac{1}{2} \end{bmatrix} \\ \text{print(u)} \\ \end{array}
```









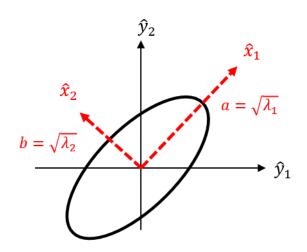
Question (Reverse Problem)

- Given Σ_{ν}^{-1} (or Σ_{ν}), how to find a (major axis) and b (minor axis) or
- How to find the proper matrix u
- Eigenvectors of Σ

$$A = S\Lambda S^T \qquad ext{where } S = [v_1 \ v_2] ext{ eigenvector of } A, ext{ and } \Lambda = egin{bmatrix} \lambda_1 & 0 \ 0 & \lambda_2 \end{bmatrix}$$

here,
$$\Sigma_y = u \Sigma_x u^T = u \Lambda u^T$$
 where $u = \begin{bmatrix} \hat{x}_1 & \hat{x}_2 \end{bmatrix}$ eigenvector of Σ_y , and $\Lambda = \begin{bmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{bmatrix} = \begin{bmatrix} a^2 & 0 \\ 0 & b^2 \end{bmatrix}$

Question (Reverse Problem)



eigen-analysis
$$\begin{cases} \Sigma_y \hat{x}_1 = \lambda_1 \hat{x}_1 \\ \Sigma_y \hat{x}_2 = \lambda_2 \hat{x}_2 \end{cases} \implies \Sigma_y [\hat{x}_1 \quad \hat{x}_2] = [\hat{x}_1 \quad \hat{x}_2] \begin{bmatrix} \lambda_1 \quad 0 \\ 0 \quad \lambda_2 \end{bmatrix}$$

$$\Sigma_y u = u \Lambda \ \Sigma_y = u \Lambda u^T = u \Sigma_x u^T$$

$$a=\sqrt{\lambda_1} \ x=u^T y \qquad \qquad b=\sqrt{\lambda_2}$$

$$\left[egin{array}{c} x_1 \ x_2 \end{array}
ight] = u^T \left[egin{array}{c} y_1 \ y_2 \end{array}
ight] \qquad egin{array}{c} ext{major axis} = \hat{x}_1 \ ext{minor axis} = \hat{x}_2 \end{array}$$

Question (Reverse Problem)

```
D, U = np.linalg.eig(Sy)

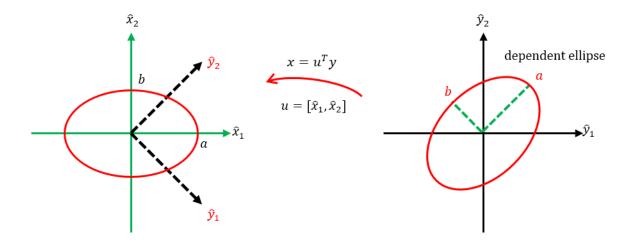
idx = np.argsort(-D)
D = D[idx]
U = U[:,idx]

print(D)
print(np.diag(D))
print(U)
```

```
[ 9. 1.]
[[ 9. 0.]
[ 0. 1.]]
[[-0.5 -0.8660254]
[-0.8660254 0.5 ]]
```



Summary



- Independent ellipse in $\{\hat{x}_1, \hat{x}_2\}$
- Dependent ellipse in $\{\hat{y}_1, \hat{y}_2\}$
- Decouple
 - Diagonalize
 - Eigen-analysis

$$egin{aligned} x &= u^T y \ u &= [\ \hat{x}_1 & \hat{x}_2] \end{aligned}$$

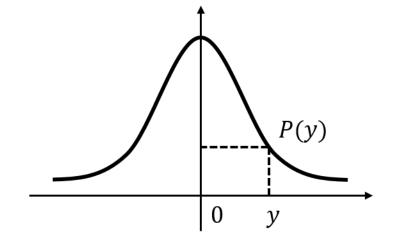
Gaussian Distribution



Standard Univariate Normal Distribution

$$P_{Y}\left(Y=y
ight)=rac{1}{\sqrt{2\pi}}\mathrm{exp}igg(-rac{1}{2}y^{2}igg)$$

$$rac{1}{2}y^2 = {
m const} \implies {
m prob.\ contour}$$

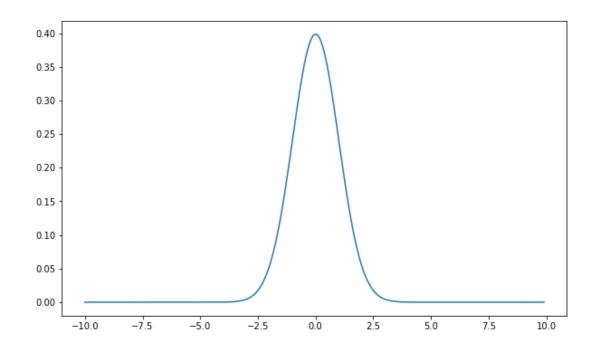


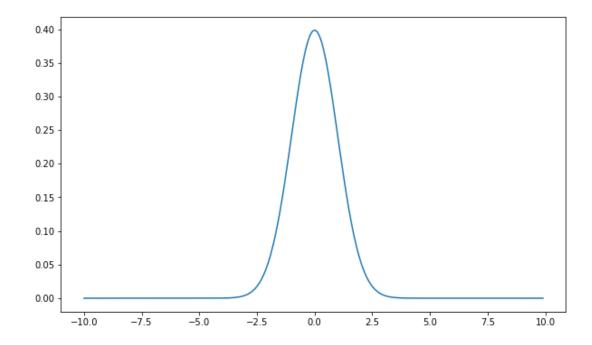


Standard Univariate Normal Distribution

```
y = np.arange(-10,10,0.1)
ProbG = 1/np.sqrt(2*np.pi)*np.exp(-1/2*y**2)
```

from scipy.stats import norm
ProbG2 = norm.pdf(y)



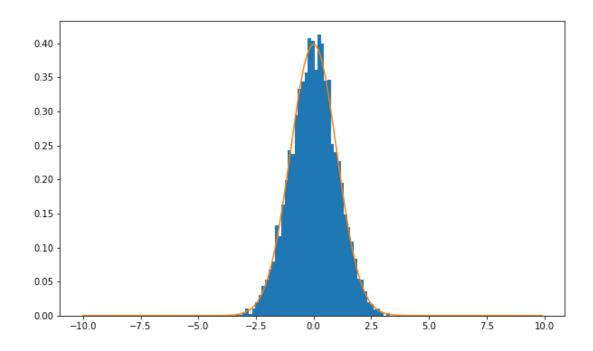




Standard Univariate Normal Distribution

```
x = np.random.randn(5000,1)

plt.figure(figsize=(10,6))
plt.hist(x, bins=51, normed=True)
plt.plot(y, ProbG2, label='G2')
plt.show()
```



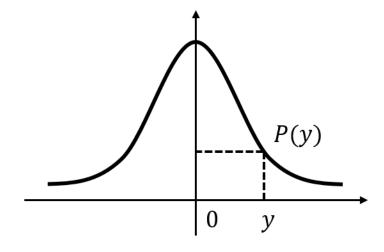


Univariate Normal distribution

• Gaussian or normal distribution, 1D (mean μ , variance σ^2)

$$N(x;\,\mu,\sigma) = rac{1}{\sqrt{2\pi}\sigma} \mathrm{exp}igg(-rac{1}{2}rac{(x-\mu)^2}{\sigma^2}igg)$$

$$egin{aligned} x &\sim N\left(\mu,\sigma^2
ight) \ \implies P_Y\left(y
ight) = P_X\left(x
ight), \quad y = rac{x-\mu}{\sigma} \ P_X\left(X = x
ight) &\sim \exp\left(-rac{1}{2}igg(rac{x-\mu}{\sigma}igg)^2
ight) \ &= \exp\left(-rac{1}{2}rac{\left(x-\mu
ight)^2}{\sigma^2}
ight) \end{aligned}$$



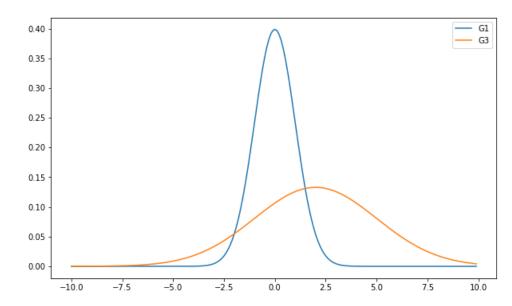
Univariate Normal distribution

```
mu = 2
sigma = 3

x = np.arange(-10, 10, 0.1)

ProbG3 = 1/(np.sqrt(2*np.pi)*sigma) * np.exp(-1/2*(x-mu)**2/(sigma**2))

plt.figure(figsize=(10,6))
plt.plot(y,ProbG, label='G1')
plt.plot(x,ProbG3, label='G3')
plt.legend()
plt.show()
```

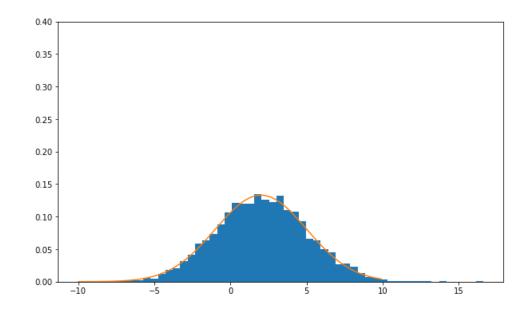




Univariate Normal distribution

```
x = mu + sigma*np.random.randn(5000,1)

plt.figure(figsize=(10,6))
plt.hist(x, bins=51, normed=True)
plt.plot(y,ProbG2, label='G2')
plt.ylim([0,0.4])
plt.show()
```





Multivariate Gaussian Models

• Similar to a univariate case, but in a matrix form

$$egin{aligned} Nig(x;\,\mu,\Sigmaig) &= rac{1}{(2\pi)^{rac{n}{2}}|\Sigma|^{rac{1}{2}}} \mathrm{exp}igg(-rac{1}{2}(x-\mu)^T\Sigma^{-1}\,(x-\mu)igg) \ \mu &= \mathrm{length}\,\,n\,\,\mathrm{column}\,\,\mathrm{vector} \ \Sigma &= n imes n\,\,\mathrm{matrix}\,\,(\mathrm{covariance}\,\,\mathrm{matrix}) \ |\Sigma| &= \mathrm{matrix}\,\,\mathrm{determinant} \end{aligned}$$

- Multivariate Gaussian models and ellipse
 - Ellipse shows constant Δ^2 value...

$$\Delta^2 = (x-\mu)^T \Sigma^{-1} (x-\mu)$$

Two Independent Variables

$$egin{aligned} P\left(X_{1}=x_{1},X_{2}=x_{2}
ight) &= P_{X_{1}}\left(x_{1}
ight)P_{X_{2}}\left(x_{2}
ight) \\ &\sim \exp\left(-rac{1}{2}rac{\left(x_{1}-\mu_{x_{1}}
ight)^{2}}{\sigma_{x_{1}}^{2}}
ight) \cdot \exp\left(-rac{1}{2}rac{\left(x_{2}-\mu_{x_{2}}
ight)^{2}}{\sigma_{x_{2}}^{2}}
ight) \\ &\sim \exp\left(-rac{1}{2}\left(rac{x_{1}^{2}}{\sigma_{x_{1}}^{2}}+rac{x_{2}^{2}}{\sigma_{x_{2}}^{2}}
ight)
ight) \end{aligned}$$

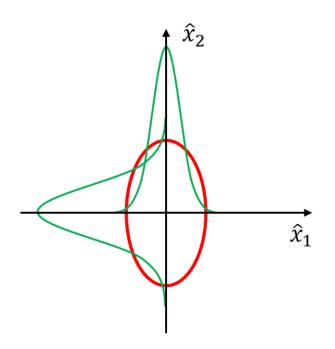
In a matrix form

$$P(x_1) \cdot P(x_2) = rac{1}{Z_1 Z_2} \exp \left(-rac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu)
ight)$$

$$\left(x = \left[egin{array}{c} x_1 \ x_2 \end{array}
ight], \quad \mu = \left[egin{array}{c} \mu_1 \ \mu_2 \end{array}
ight], \quad \Sigma = \left[egin{array}{c} \sigma_{x_1}^2 & 0 \ 0 & \sigma_{x_2}^2 \end{array}
ight]
ight)$$



Two Independent Variables



$$rac{x_1^2}{\sigma_{x_1}^2} + rac{x_2^2}{\sigma_{x_2}^2} = c \quad ext{(ellipse)}$$

$$egin{bmatrix} \left[egin{array}{ccc} x_1 & x_2
ight] \left[egin{array}{ccc} rac{1}{\sigma_{x_1}^2} & 0 \ 0 & rac{1}{\sigma_{x_2}^2} \end{array}
ight] \left[egin{array}{ccc} x_1 \ x_2 \end{array}
ight] = c \qquad (\sigma_{x_1} < \sigma_{x_2}) \ \end{array}$$

Summary in a matrix form

$$N\left(0,\Sigma_{x}
ight)\sim\exp\!\left(-rac{1}{2}x^{T}\Sigma_{x}^{-1}x
ight)$$

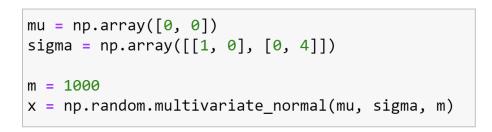
$$N\left(\mu_x, \Sigma_x
ight) \sim \exp\!\left(-rac{1}{2}(x-\mu_x)^T\Sigma_x^{-1}\left(x-\mu_x
ight)
ight)$$

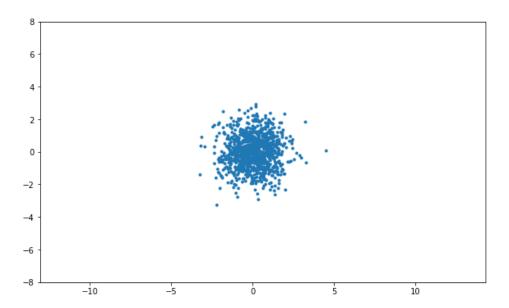
Two Independent Variables

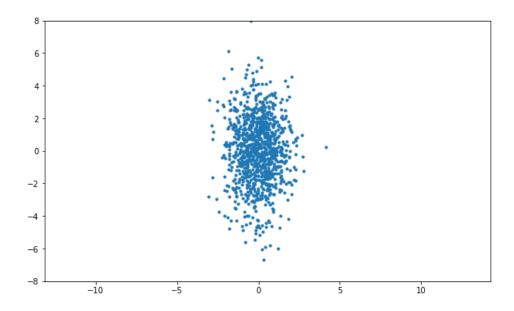
```
mu = np.array([0, 0])
sigma = np.eye(2)

m = 1000
x = np.random.multivariate_normal(mu, sigma, m)
print(x.shape)

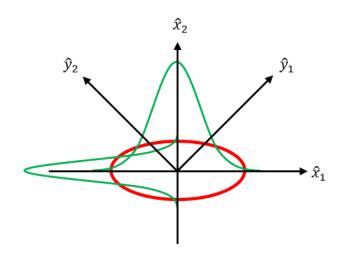
(1000, 2)
```

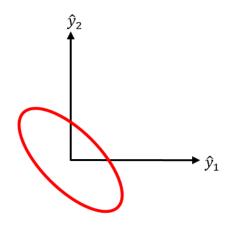












• Compute $P_Y(y)$ from $P_X(x)$

$$P_X(x) = P_Y(y) \; ext{ where } \; x = \left[egin{array}{c} x_1 \ x_2 \end{array}
ight], y = \left[egin{array}{c} y_1 \ y_2 \end{array}
ight]$$

• Relationship between y and x

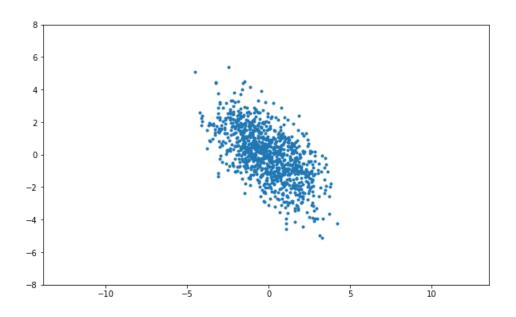
$$x = \left[egin{array}{cc} \hat{x}_1 & \hat{x}_2
ight]^T y = u^T y$$

$$egin{aligned} x^T \Sigma_x^{-1} x &= y^T u \Sigma_x^{-1} u^T y = y^T \Sigma_y^{-1} y \ dots & \Sigma_y^{-1} &= u \Sigma_x^{-1} u^T \ & o \Sigma_y &= u \Sigma_x u^T \end{aligned}$$

- Σ_x : covariance matrix of x
- Σ_{v} : covariance matrix of y
- If u is an eigenvector matrix of Σ_y , then Σ_x is a diagonal matrix

```
mu = np.array([0, 0])
sigma = 1./2.*np.array([[5, -3], [-3, 5]])

m = 1000
x = np.random.multivariate_normal(mu, sigma, m)
```





Remark

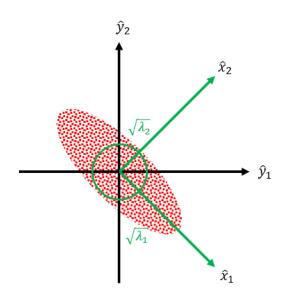
 $x \sim N(\mu_x, \Sigma_x)$ and y = Ax + b affine transformation

$$\implies y \sim N(\mu_y, \Sigma_y) = N(A\mu_x + b, A\Sigma_x A^T)$$

 $\implies y \; ext{ is also Gaussian with } \; \mu_y = Ax + b, \; \; \Sigma_y = A\Sigma_x A^T$

Decouple using Covariance Matrix

• Given data, how to find Σ_y and major (or minor) axis (assume $\mu_y=0$)



$$\Sigma_y = egin{bmatrix} ext{var}(y_1) & ext{cov}(y_1,y_2) \ ext{cov}(y_2,y_1) & ext{var}(y_2) \end{bmatrix}$$

eigen-analysis
$$\Sigma_x^{-1} = \begin{bmatrix} rac{1}{\sqrt{\lambda_1}^2} & 0 \\ 0 & rac{1}{\sqrt{\lambda_2}^2} \end{bmatrix}$$
 $\Sigma_y \hat{x}_1 = \lambda_1 \hat{x}_1$ $\Sigma_y \hat{x}_2 = \lambda_2 \hat{x}_2$ $\Sigma_x = \begin{bmatrix} \sqrt{\lambda_1}^2 & 0 \\ 0 & \sqrt{\lambda_2}^2 \end{bmatrix}$

$$egin{aligned} \Sigma_y \left[egin{array}{ccc} \hat{x}_1 & \hat{x}_2
ight] &= \left[egin{array}{ccc} \hat{x}_1 & \hat{x}_2
ight] \left[egin{array}{ccc} \lambda_1 & 0 \ 0 & \lambda_2 \end{array}
ight] & y = ux \implies u^T y = x \ &= \left[egin{array}{ccc} \hat{x}_1 & \hat{x}_2
ight] \Sigma_x & \left[egin{array}{ccc} \hat{x}_1 & \hat{x}_2 \end{array}
ight] = u \end{aligned}$$

$$\Sigma_y = u \Sigma_x u^T$$

Decouple using Covariance Matrix

```
S = np.cov(x.T)
print ("S = \n", S)

S =
  [[ 2.59216411 -1.54924881]
  [-1.54924881    2.54567035]]
```

```
D, U = np.linalg.eig(S)

idx = np.argsort(-D)
D = D[idx]
U = U[:,idx]

print ("U = \n", U)
print ("D = \n", D)
```

```
xp = np.arange(-10, 10)

plt.figure(figsize=(10,6))
plt.plot(x[:,0],x[:,1],'.')
plt.plot(xp, U[1,0]/U[0,0]*xp, label='u1')
plt.plot(xp, U[1,1]/U[0,1]*xp, label='u2')
plt.axis('equal')
plt.ylim([-8, 8])
plt.legend()
plt.show()
```

