

# **Probabilistic Machine Learning**

Prof. Seungchul Lee Industrial AI Lab.



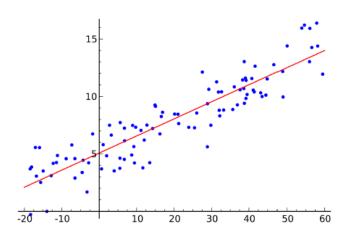
#### **Outline**

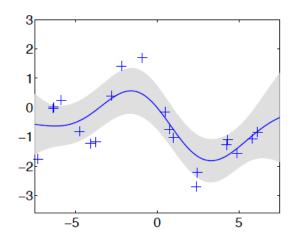
- Probabilistic Linear Regression
- Probabilistic Classification
- Probabilistic Clustering
- Probabilistic Dimension Reduction

# **Frequentist View of Linear Regression**



### **Probabilistic Linear Regression**





• Inference idea

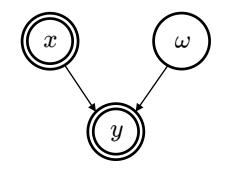
$$P(X \mid \theta) = \text{Probability [data | pattern]}$$

data = underlying pattern + independent noise

- Change your viewpoint of data
  - Generative model

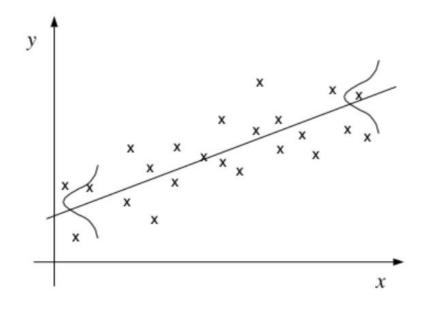
# **Generative Model: Regression**

$$y = \hat{y} + arepsilon = \omega^T x + arepsilon, \quad arepsilon \sim \mathcal{N}(0, \sigma^2)$$



$$D = \{(x_1, y_1), (x_2, y_2), \cdots, (x_m, y_m)\}$$

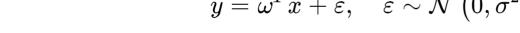
$$P\left(y\mid x;\omega,\sigma^{2}
ight)=\mathcal{N}\left(\omega^{T}x,\sigma^{2}
ight)$$

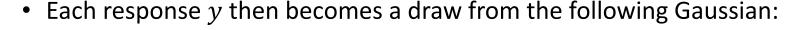


## **Probabilistic Linear Regression**

- Given observed data  $D = \{(x_1, y_1), (x_2, y_2), \dots, (x_m, y_m)\},\$
- We want to estimate the weight vector  $\omega$
- Each response generated by a linear model plus Gaussian noise

$$y = \omega^T x + arepsilon, \quad arepsilon \sim \mathcal{N}\left(0, \sigma^2
ight)$$

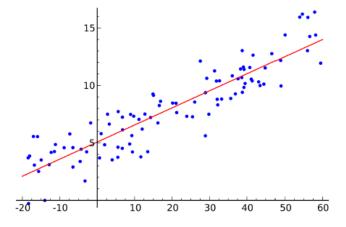




$$y \mid x \sim \left(\omega^T x, \sigma^2
ight)$$

Probability of each response variable

$$P(y \mid x \, ; \omega) = \mathcal{N}\left(\omega^T x, \sigma^2
ight) = rac{1}{\sqrt{2\pi\sigma^2}} \mathrm{exp}igg(-rac{1}{2\sigma^2}ig(y-\omega^T xig)^2igg)$$



## **Maximum Likelihood Estimation (MLE)**

- Estimate parameters  $\theta = (\omega, \sigma^2)$  such that maximize the likelihood given a generative model
  - Likelihood

$$\mathcal{L} = P(D \mid \theta) = P(D; \theta)$$

$$\hat{\theta}_{MLE} = \underset{\theta}{\operatorname{argmax}} P(D; \theta)$$

Log-likelihood:

$$egin{aligned} \ell(\omega,\sigma) &= \log \mathcal{L}(\omega,\sigma) = \log P(D\,;\omega,\sigma^2) \ &= \log P(Y\mid X\,;\omega,\sigma^2) \ &= \log \prod_{i=1}^m P\left(y_i\mid x_i\,;\omega,\sigma^2
ight) \ &= \sum_{i=1}^m \log P\left(y_i\mid x_i\,;\omega,\sigma^2
ight) \ &= \sum_{i=1}^m \log rac{1}{\sqrt{2\pi\sigma^2}} \mathrm{exp}\left(-rac{\left(y_i-\omega^Tx_i
ight)^2}{2\sigma^2}
ight) \ &= \sum_{i=1}^m \left\{-rac{1}{2} \mathrm{log}(2\pi\sigma^2) - rac{\left(y_i-\omega^Tx_i
ight)^2}{2\sigma^2}
ight\} \end{aligned}$$

# **Maximum Likelihood Estimation (MLE)**

Maximum likelihood solution:

$$\begin{split} \log \mathcal{L}(\omega, \sigma) &= \sum_{i=1}^{m} \left\{ -\frac{1}{2} \log \left( 2\pi \sigma^{2} \right) - \frac{\left( y_{i} - \omega^{T} x_{i} \right)^{2}}{2\sigma^{2}} \right\} \\ \hat{\omega}_{MLE} &= \arg \max_{\omega} \log P(D; \omega, \sigma^{2}) \\ &= \arg \max_{\omega} \ -\frac{1}{2\sigma^{2}} \sum_{i=1}^{m} \left( y_{i} - \omega^{T} x_{i} \right)^{2} \\ &= \arg \min_{\omega} \frac{1}{2\sigma^{2}} \sum_{i=1}^{m} \left( y_{i} - \omega^{T} x_{i} \right)^{2} \\ &= \arg \min_{\omega} \sum_{i=1}^{m} \left( y_{i} - \omega^{T} x_{i} \right)^{2} \end{split}$$

- Big lesson
  - It is equivalent to the least-squares objective function for linear regression (amazing!)
  - In least squares, we implicitly assume that noise is Gaussian distributed

## **Compute MLE for Linear Regression**

$$egin{aligned} \mathcal{L}(\omega,\sigma) &= P\left(y_1,y_2,\cdots,y_m \mid x_1,x_2,\cdots,x_m; \ oldsymbol{\omega},\sigma
ight) \ &= \prod_{i=1}^m P\left(y_i \mid x_i; \ \omega,\sigma
ight) \ &= rac{1}{\left(2\pi\sigma^2
ight)^{rac{m}{2}}} \mathrm{exp}igg(-rac{1}{2\sigma^2} \sum_{i=1}^m (y_i - \omega^T x_i)^2igg) \end{aligned}$$



# **Compute MLE for Linear Regression**

$$\mathcal{L}(\omega, \sigma) = \frac{1}{(2\pi\sigma^2)^{\frac{m}{2}}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^m (y_i - \omega^T x_i)^2\right)$$

$$\ell = -\frac{m}{2} \log 2\pi - m \log \sigma - \frac{1}{2\sigma^2} \sum_{i=1}^m (y_i - \omega^T x_i)^2$$

$$\frac{d\ell}{d\omega} = -2X^T Y + 2X^T X \omega = 0 \implies \omega_{MLE} = (X^T X)^{-1} X^T Y$$

$$\frac{d\ell}{d\sigma} = -\frac{m}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^m (y_i - \omega^T x_i)^2 = 0 \implies \sigma_{MLE}^2 = \frac{1}{m} \sum_{i=1}^m (y_i - \omega^T x_i)^2$$

- Big lesson
  - It is equivalent to the least-squares objective function for linear regression (amazing!)
  - In least squares, we implicitly assume that noise is Gaussian distributed

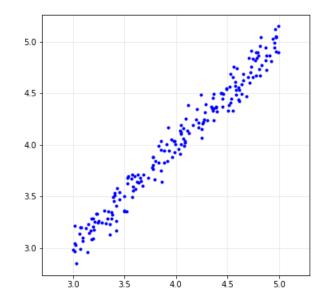


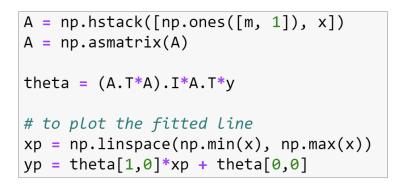
### **Linear Regression: A Probabilistic View**

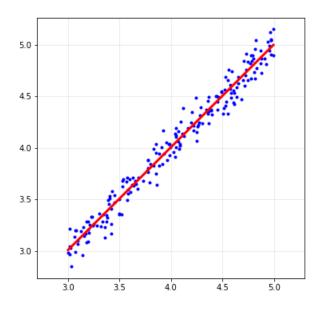
```
m = 200

a = 1
x = 3 + 2*np.random.uniform(0,1,[m,1])
noise = 0.1*np.random.randn(m,1)

y = a*x + noise;
y = np.asmatrix(y)
```









## **Linear Regression: A Probabilistic View**

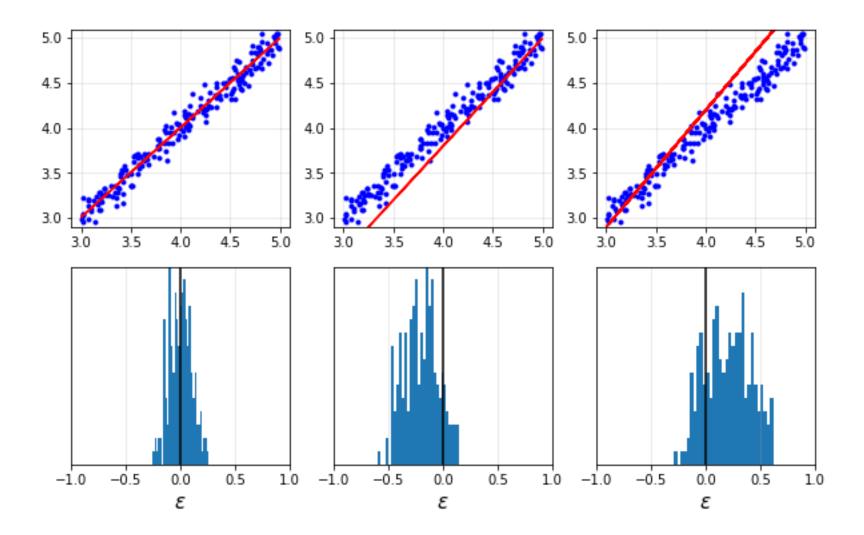
#### Demonstrate

$$arepsilon \sim \mathcal{N}\left(0, \sigma^2
ight)$$

```
yhat0 = theta[1,0]*x + theta[0,0]
err0 = yhat0 - y

yhat1 = 1.2*x - 1
err1 = yhat1 - y

yhat2 = 1.3*x - 1
err2 = yhat2 - y
```





### **Linear Regression: A Probabilistic View**

- Demonstrate
  - samples are independent

```
a0x = err0[1:]

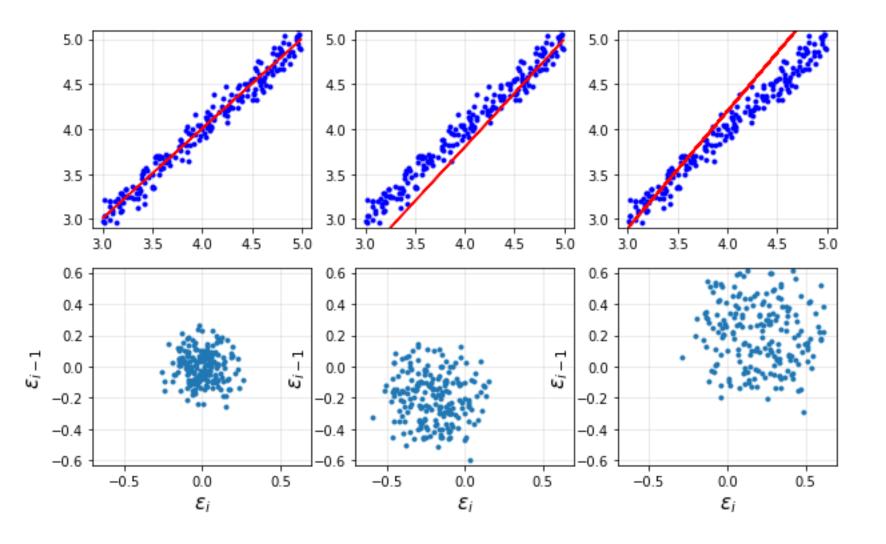
a0y = err0[0:-1]

a1x = err1[1:]

a1y = err1[0:-1]

a2x = err2[1:]

a2y = err2[0:-1]
```



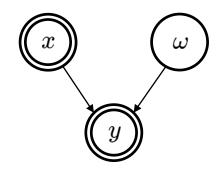


# **Bayesian View of Linear Regression**



# **Generative Model: Regression**

$$y = \hat{y} + arepsilon = \omega^T x + arepsilon, \quad arepsilon \sim \mathcal{N}(0, \sigma^2)$$



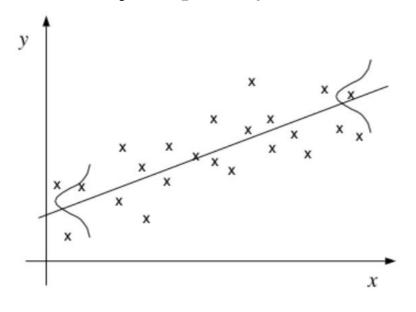
$$\omega_1 = \begin{bmatrix} \omega_0 \\ \omega_1 \end{bmatrix}$$

$$\omega_0$$

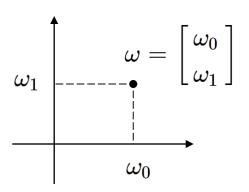
$$D = \{(x_1,y_1), (x_2,y_2), \cdots, (x_m,y_m)\}$$

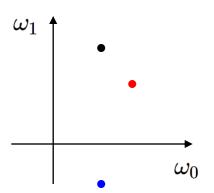
$$P\left(y\mid x;\omega,\sigma^{2}
ight)=\mathcal{N}\left(\omega^{T}x,\sigma^{2}
ight)$$

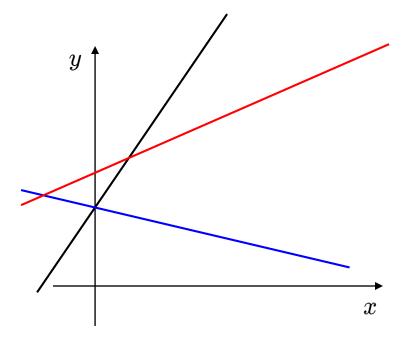
$$y = \omega_1 x + \omega_0 + \varepsilon$$



# Meaning of $\omega$

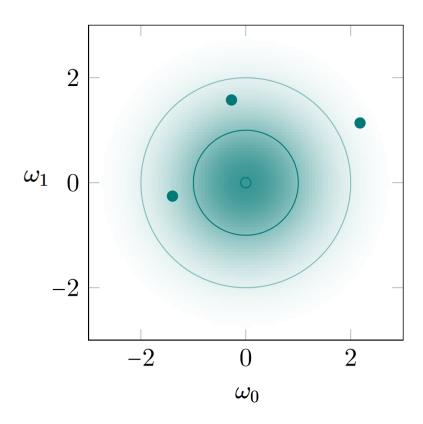


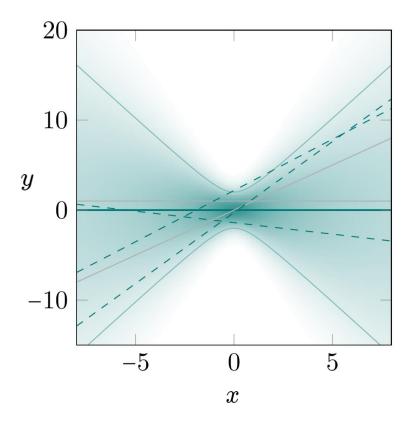




### Prior on $\omega$

• Suppose to assume a Gaussian prior distribution over the weight vector  $\boldsymbol{\omega}$ 

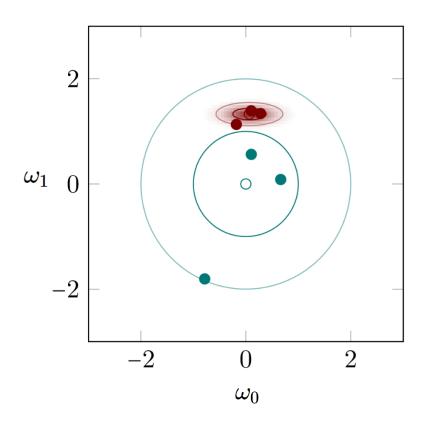


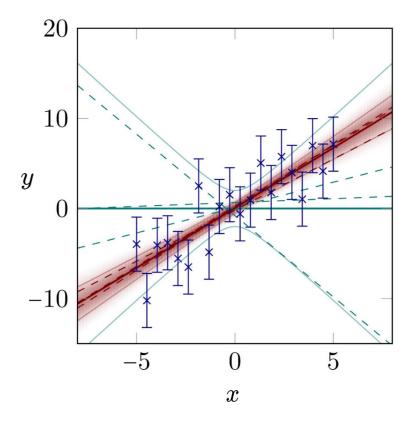




### Prior on $\omega$

• Suppose to assume a Gaussian prior distribution over the weight vector  $\boldsymbol{\omega}$ 







# Maximum-a-Posteriori (MAP)

- No prior information or uniform distribution on  $\omega$  leads to MLE
- Suppose to assume a Gaussian prior distribution over the weight vector  $\omega$ 
  - (Make sure you understand what it means)
  - Assume  $E[\omega] = 0$  for simplicity

$$P(\omega) \sim \mathcal{N}\left(0, \Sigma
ight) = \mathcal{N}\left(0, \lambda^{-1} I
ight) = rac{1}{(2\pi)^{D/2}} \mathrm{exp}igg(-rac{\lambda}{2}\omega^T\omegaigg)$$

- Excellent explanation by Philipp Henning
  - <a href="https://www.youtube.com/watch?v=50Vgw11qn0o">https://www.youtube.com/watch?v=50Vgw11qn0o</a>



#### **Posterior**

- Posterior probability
  - Bayes rule

$$P(\omega \mid D) = \frac{P(D \mid \omega)P(\omega)}{P(D)}$$

Log posterior probability

$$\log P(\omega \mid D) = \log rac{P(D \mid \omega)P(\omega)}{P(D)} = \log P(D \mid \omega) + \log P(\omega) - \underbrace{\log P(D)}_{ ext{constant}}$$

Maximize log posterior probability

## Maximum-a-Posteriori (MAP)

$$\begin{split} \hat{\omega}_{MAP} &= \arg\max_{\omega} \log P(\omega \mid D) \\ &= \arg\max_{\omega} \left\{ \log P(D \mid \omega) + \log P(\omega) \right\} \\ &= \arg\max_{\omega} \left\{ \sum_{i=1}^{m} \left\{ -\frac{1}{2} \log(2\pi\sigma^{2}) - \frac{\left(y_{i} - \omega^{T} x_{i}\right)^{2}}{2\sigma^{2}} \right\} - \frac{D}{2} \log(2\pi) - \frac{\lambda}{2} \omega^{T} \omega \right\} \\ &= \arg\min_{\omega} \frac{1}{2\sigma^{2}} \sum_{i=1}^{m} \left( y_{i} - \omega^{T} x_{i} \right)^{2} + \frac{\lambda}{2} \omega^{T} \omega \\ & \text{ (ignoring constants and changing max to min)} \end{split}$$

• For  $\sigma = 1$  (or some constant) for each input, it's equivalent to the regularized least-squares objective (amazing!)

$$\hat{\omega}_{MAP} = rg\min_{\omega} \left\{ \sum_{i=1}^{m} \left( y_i - \omega^T x_i 
ight)^2 + \lambda \omega^T \omega 
ight\}$$

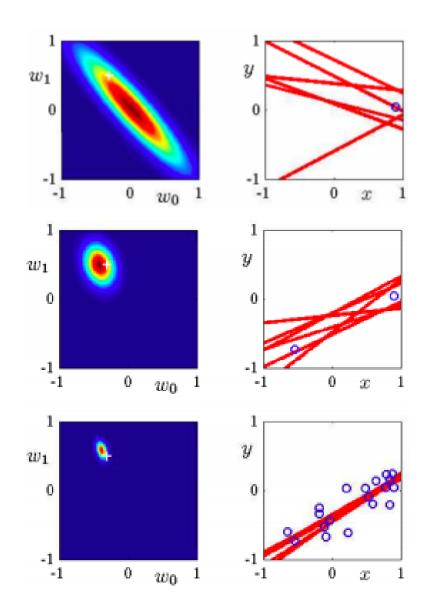
• Big lesson: MAP =  $l_2$  norm regularization

### **MAP Illustration**

• One observation

• Two observations

• 20 observations



## **Summary: MLE vs MAP**

• MLE solution:

$$\hat{\omega}_{MLE} = rg \min_{\omega} rac{1}{2\sigma^2} \sum_{i=1}^m \left(y_i - \omega^T x_i
ight)^2$$

MAP solution:

$$\hat{\omega}_{MAP} = rg \min_{\omega} rac{1}{2\sigma^2} \sum_{i=1}^m \left(y_i - \omega^T x_i
ight)^2 + rac{\lambda}{2} \omega^T \omega^T$$

- Take-home messages:
  - MLE estimation of a parameter leads to unregularized solutions
  - MAP estimation of a parameter leads to regularized solutions
  - The prior distribution acts as a regularizer in MAP estimation
- Note: for MAP, different prior distributions lead to different regularizers
  - Gaussian prior on  $\omega$  regularizes the  $l_2$  norm of  $\omega$
  - Laplace prior  $exp(-C\|\omega\|_1)$  on  $\omega$  regularizes the  $l_1$  norm of  $\omega$

## **Probabilistic Classification**



### **Probabilistic** Classification

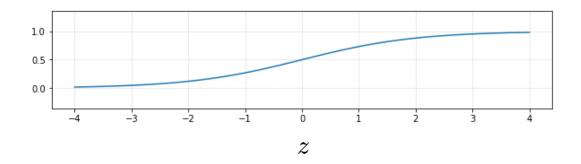
- We want to predict the label probabilities
  - E.g.,  $P(y = +1 | x, \omega)$ : the probability that the label is  $P(y | x, \omega)$
  - In a sense, it is our confidence in the predicted label +1

## **Probabilistic Linear Classification**

- Probabilistic classification models allow us do that (y = -1/+1)
- Consider the following function in a compact expression

$$P(y \mid x, \omega) = \sigma\left(y\omega^T x
ight) = rac{1}{1 + \exp(-y\omega^T x)}$$

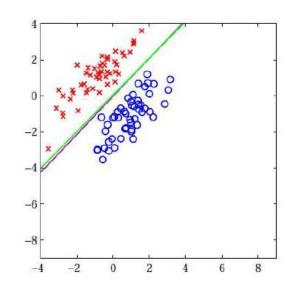
•  $\sigma$  is the logistic function which maps all real number into (0,1)



## **Logistic Regression**

- What does the decision boundary look like for logistic regression?
- At the decision boundary labels -1/+1 becomes equiprobable

$$P(y=+1\mid x,\omega) = P(y=-1\mid x,\omega)$$
  $\dfrac{1}{1+\exp(-\omega^T x)} = \dfrac{1}{1+\exp(\omega^T x)}$   $\exp(-\omega^T x) = \exp(\omega^T x)$   $\omega^T x = 0$ 



- The decision boundary is therefore linear ⇒ logistic regression is a linear classifier
- Note: it is possible to kernelize and make it nonlinear

### **Maximum Likelihood Solution**

- Goal: want to estimate  $\omega$  from the data  $D = \{(x_1, y_1), \dots, (x_m, y_m)\}$
- Log-likelihood:

$$egin{aligned} \ell(\omega) &= \log \mathcal{L}(\omega) = \log P(D \mid \omega) \ &= \log P(Y \mid X, \omega) \ &= \log \prod_{i=1}^m P(y_i \mid x_i, \omega) \ &= \sum_{i=1}^m \log P(y_i \mid x_i, \omega) \ &= \sum_{i=1}^m \log rac{1}{1 + \exp(-y_i \omega^T x_i)} \ &= \sum_{i=1}^m - \log \left[ 1 + \exp(-y_i \omega^T x_i) 
ight] \end{aligned}$$

## **Maximum Likelihood Solution**

Maximum Likelihood Solution:

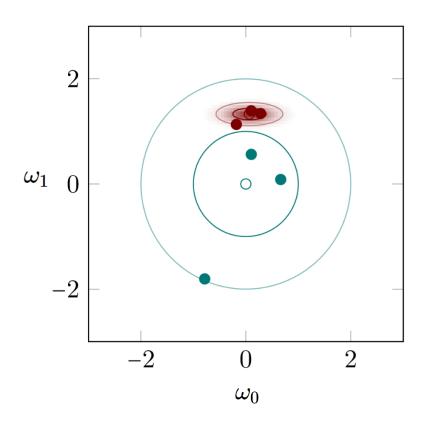
$$\hat{\omega}_{MLE} = rg \max_{\omega} \log \mathcal{L}(\omega) = rg \min_{\omega} \sum_{i=1}^{m} \log igl[ 1 + \expigl( -y_i \omega^T x_i igr) igr]$$

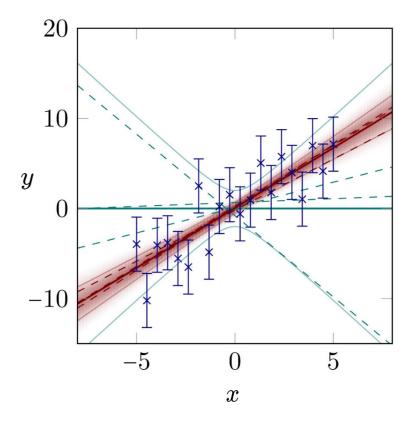
- No closed-form solution exists, but we can do
  - CVXPY (we did it)
  - Gradient descent on  $\omega$

$$egin{aligned} 
abla_{\omega} \log \mathcal{L}(\omega) &= \sum_{i=1}^m -rac{1}{1+\exp(-y_i\omega^Tx_i)} \expigl(-y_i\omega^Tx_iigr)(-y_ix_iigr) \ &= \sum_{i=1}^m rac{1}{1+\exp(y_i\omega^Tx_i)} y_ix_i \end{aligned}$$

### Prior on $\omega$

• Suppose to assume a Gaussian prior distribution over the weight vector  $\boldsymbol{\omega}$ 



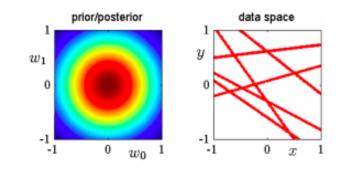




### **Maximum-a-Posteriori Solution**

• Let's assume a Gaussian prior distribution over the weight vector  $\omega$ 

$$P(\omega) = \mathcal{N}\left(0, \lambda^{-1}I
ight) = rac{1}{(2\pi)^{D/2}} \mathrm{exp}igg(-rac{\lambda}{2}\omega^T\omegaigg)$$



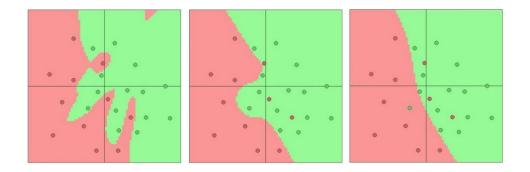
Maximum-a-Posteriori Solution:

$$\begin{split} \hat{\omega}_{MAP} &= \arg\max_{\omega} \log P(\omega \mid D) \\ &= \arg\max_{\omega} \{\log P(D \mid \omega) + \log P(\omega) - \underbrace{\log P(D)}_{\text{constant}} \} \\ &= \arg\max_{\omega} \{\log P(D \mid \omega) + \log P(\omega) \} \\ &= \arg\max_{\omega} \left\{ \sum_{i=1}^{m} -\log \left[1 + \exp\left(-y_{i}\omega^{T}x_{i}\right)\right] - \frac{D}{2}\log(2\pi) - \frac{\lambda}{2}\omega^{T}\omega \right\} \\ &= \arg\min_{\omega} \sum_{i=1}^{m} \log \left[1 + \exp\left(-y_{i}\omega^{T}x_{i}\right)\right] + \frac{\lambda}{2}\omega^{T}\omega \end{split}$$
 (ignoring constants and changing max to min)

• Big lesson: MAP =  $l_2$  norm regularization

### **Maximum-a-Posteriori Solution**

- Q: What does regularizer do in a classifier?
- A: Nonlinear classifier gives more intuitive explanation



- No closed-form solution exists but we can do gradient descent on  $\omega$ 
  - See "<u>A comparison of numerical optimizers for logistic regression</u>" by Tom Minka on optimization techniques (gradient descent and others) for logistic regression
  - (both MLE and MAP)

# **Summary: MLE vs MAP**

MLE solution:

$$\hat{\omega}_{MLE} = rg\min_{\omega} \sum_{i=1}^m \logigl[1 + \expigl(-y_i\omega^T x_iigr)igr]$$

MAP solution:

$$\hat{\omega}_{MAP} = rg \min_{\omega} \sum_{i=1}^{m} \log igl[ 1 + \expigl( -y_i \omega^T x_i igr) igr] + rac{\lambda}{2} \omega^T \omega^T$$

- Take-home messages (we already saw these before)
  - MLE estimation of a parameter leads to unregularized solutions
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- Note: For MAP, different prior distributions lead to different regularizers
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# **Probabilistic** Clustering

• will not cover in this course



## **Probabilistic** Dimension Reduction

• will not cover in this course



### **Summary**

- *Probabilistic* Linear Regression
- Probabilistic Classification
- Probabilistic Clustering
- Probabilistic Dimension Reduction