

# Genome-wide scan for linear mixed models

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## 1 Single block

Let

$$\mathbf{y} \sim \mathcal{N}(\mathbf{F}\boldsymbol{\alpha}; \mathbf{K})$$

be the marginal likelihood of a LMM with a single covariate block  $\mathbf{F}$ . The maximum likelihood estimation of the fixed effects is

$$\hat{\boldsymbol{\alpha}} = (\mathbf{F}^\top \mathbf{K}^\dagger \mathbf{F})^\dagger \mathbf{F}^\top \mathbf{K}^\dagger \mathbf{y}.$$

In practice, the method of least squares can be used to solve the above equation without explicitly finding pseudoinverses.

## 2 Double block

Let

$$\mathbf{y} \sim \mathcal{N}(\mathbf{F}\boldsymbol{\alpha} + \mathbf{G}\boldsymbol{\beta}; \mathbf{K})$$

be the marginal likelihood of a LMM with two covariate blocks  $\mathbf{F}$  and  $\mathbf{G}$ . We want to solve

$$\begin{bmatrix} \hat{\boldsymbol{\alpha}} \\ \hat{\boldsymbol{\beta}} \end{bmatrix} = (\mathbf{X}^\top \mathbf{X})^\dagger \mathbf{X}^\top \mathbf{K}^{\frac{1}{2}\dagger} \mathbf{y}$$

for

$$\mathbf{L} = \mathbf{K}^{\frac{1}{2}\dagger} \mathbf{F}, \quad \mathbf{R} = \mathbf{K}^{\frac{1}{2}\dagger} \mathbf{G} \quad \text{and} \quad \mathbf{X} = \begin{bmatrix} \mathbf{L} & \mathbf{R} \end{bmatrix}.$$

We have

$$(\mathbf{X}^\top \mathbf{X})^\dagger = \begin{bmatrix} (\mathbf{L}^\top \mathbf{L})^\dagger + (\mathbf{L}^\top \mathbf{L})^\dagger (\mathbf{L}^\top \mathbf{R}) \mathbf{W}^\dagger (\mathbf{L}^\top \mathbf{R})^\top (\mathbf{L}^\top \mathbf{L})^\dagger & -(\mathbf{L}^\top \mathbf{L})^\dagger (\mathbf{L}^\top \mathbf{R}) \mathbf{W}^\dagger \\ -\mathbf{W}^\dagger (\mathbf{L}^\top \mathbf{R})^\top (\mathbf{L}^\top \mathbf{L})^\dagger & \mathbf{W}^\dagger \end{bmatrix},$$

from Eq. 3 of [1], where  $\mathbf{W} = \mathbf{R}^\top \mathbf{R} - \mathbf{R}^\top \mathbf{L} (\mathbf{L}^\top \mathbf{L})^\dagger \mathbf{L}^\top \mathbf{R}$ .

Defining  $\mathbf{A} = \mathbf{L}^\top \mathbf{L}$  and  $\mathbf{B} = \mathbf{L}^\top \mathbf{R}$  leads us to

$$(\mathbf{X}^\top \mathbf{X})^\dagger \mathbf{X}^\top = \begin{bmatrix} \mathbf{A}^\dagger \mathbf{L}^\top + \mathbf{A}^\dagger \mathbf{B} \mathbf{W}^\dagger \mathbf{B}^\top \mathbf{A}^\dagger \mathbf{L}^\top - \mathbf{A}^\dagger \mathbf{B} \mathbf{W}^\dagger \mathbf{R}^\top \\ -\mathbf{W}^\dagger \mathbf{B}^\top \mathbf{A}^\dagger \mathbf{L}^\top + \mathbf{W}^\dagger \mathbf{R}^\top \end{bmatrix}.$$

Finally,

$$(\mathbf{X}^\top \mathbf{X})^\dagger \mathbf{X}^\top \mathbf{K}^{\frac{1}{2}\dagger} \mathbf{y} = \begin{bmatrix} \mathbf{A}^\dagger \mathbf{F}^\top \mathbf{K}^\dagger \mathbf{y} + \mathbf{A}^\dagger \mathbf{B} \mathbf{W}^\dagger \mathbf{B}^\top \mathbf{A}^\dagger \mathbf{F}^\top \mathbf{K}^\dagger \mathbf{y} - \mathbf{A}^\dagger \mathbf{B} \mathbf{W}^\dagger \mathbf{G}^\top \mathbf{K}^\dagger \mathbf{y} \\ \mathbf{W}^\dagger \mathbf{G}^\top \mathbf{K}^\dagger \mathbf{y} - \mathbf{W}^\dagger \mathbf{B}^\top \mathbf{A}^\dagger \mathbf{F}^\top \mathbf{K}^\dagger \mathbf{y} \end{bmatrix}.$$

A robust implementation of the above equation has to: (i) associate matrix multiplications in such a way that a sequence of  $\mathbf{K}^\dagger \dots \mathbf{K}^\dagger$  is avoided; and (ii) handle low-rank matrices  $\mathbf{W}$ . A better association of matrix multiplications is given by

$$(\mathbf{X}^\top \mathbf{X})^\dagger \mathbf{X}^\top \mathbf{K}^{\frac{1}{2}\dagger} \mathbf{y} = \begin{bmatrix} \mathbf{A}^\dagger \mathbf{F}^\top \mathbf{K}^\dagger \mathbf{y} + \mathbf{A}^\dagger \mathbf{B} \mathbf{W}^\dagger (\mathbf{B}^\top \mathbf{A}^\dagger \mathbf{F}^\top \mathbf{K}^\dagger \mathbf{y} - \mathbf{G}^\top \mathbf{K}^\dagger \mathbf{y}) \\ \mathbf{W}^\dagger (\mathbf{G}^\top \mathbf{K}^\dagger \mathbf{y} - \mathbf{B}^\top \mathbf{A}^\dagger \mathbf{F}^\top \mathbf{K}^\dagger \mathbf{y}) \end{bmatrix}.$$

$\mathbf{W}^\dagger$  can be found via economic SVD decomposition.

## References

- [1] Charles A Rohde. "Generalized inverses of partitioned matrices". In: *Journal of the Society for Industrial and Applied Mathematics* 13.4 (1965), pp. 1033–1035.