

DSA 503: Statistical Inference

- Estimation
- Test of Hypothesis

A population consists of all elements; individuals, items or objects whose characteristics are being studied.

sample: A representative part of population

Ex: Blood in the body is population

The small part of it is sample. [For blood testing]

- There are some cons to work with population
 - Time consuming
 - Expensive
 - Sometimes infeasible.

Parameters:

Population characteristics

Ex: Population mean (μ)

Statistics:

Sample characteristics

Ex: Sample mean (\bar{x})

Not statistics

The population distribution is the probability distribution of the population data.

- ↳ Normal
- ↳ Binomial
- ↳ Poisson

Relative frequency = 1

Frequency distribution

$$\sum P(x) = 1$$

$$0 \leq \text{Probability} \leq 1$$

Probability range

Probability Distribution

↳ ALWAYS for population

Sampling Distribution

↳ probability distribution of statistic

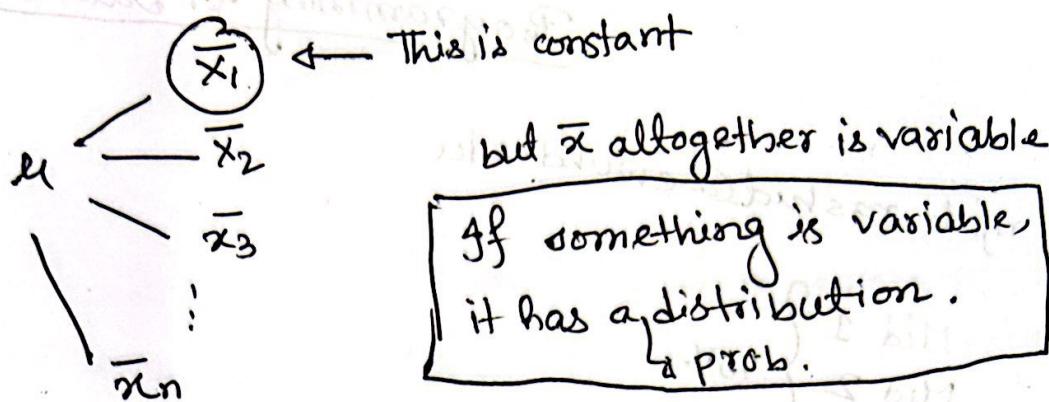
probability distribution of \bar{x} .

Population 5, Sample 3,

Can be taken in 5C_3 ways

$${}^nC_r = \frac{n!}{r!(n-r)!}$$

$${}^5C_3 = \frac{5!}{3!2!} = \frac{5 \cdot 4 \cdot 3 \cdot 2 \cdot 1}{3 \cdot 2 \cdot 1 \cdot 2 \cdot 1} = 10$$



All sampling distributions are probability distribution, but all prob. dist. are not sampling distribution.

Statistical Inference

Estimation

Point

Interval

test

→ MCA

→ Proportion

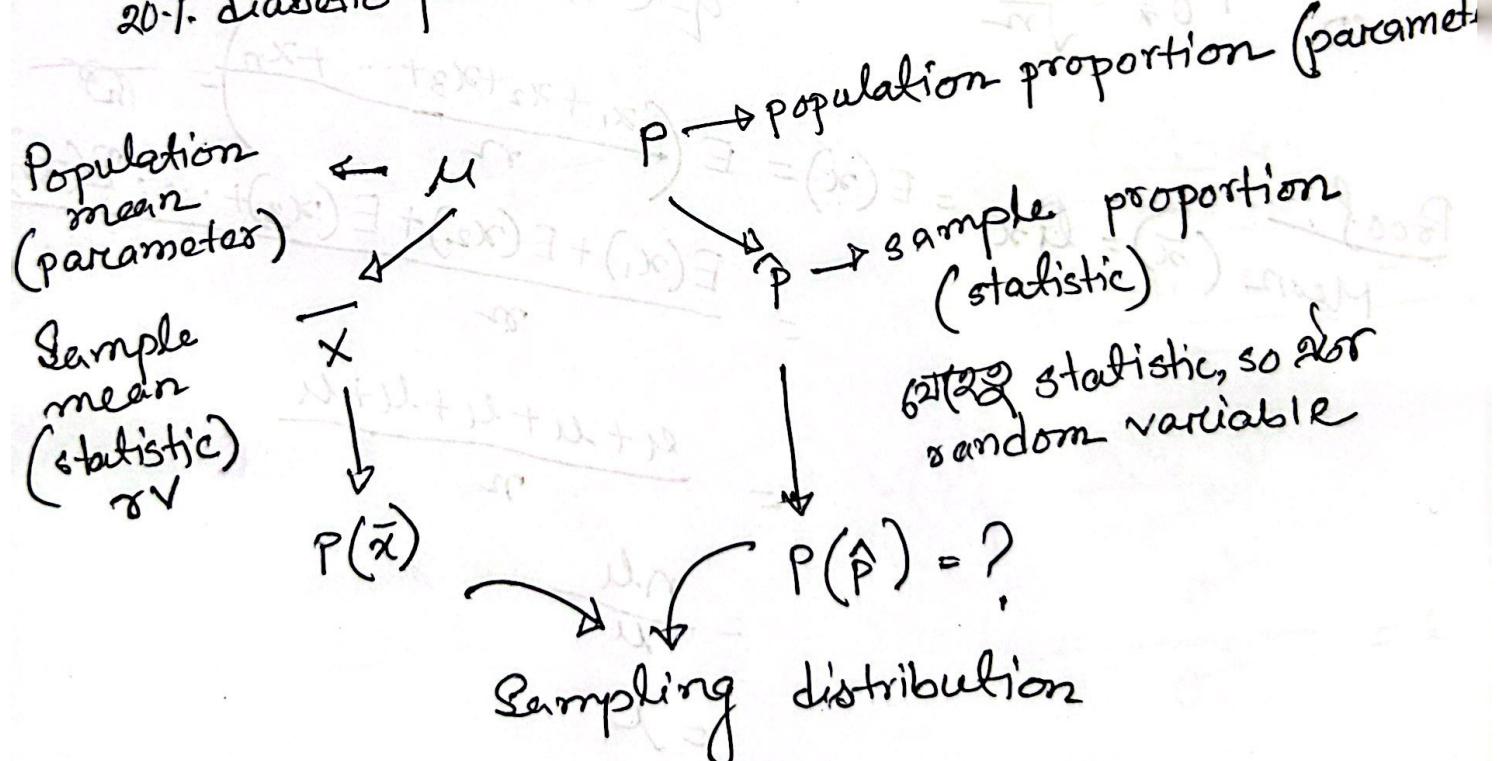
→ Chi square

→ ANOVA F

→ Non-parametric test

$$P = \frac{X}{N} \quad \text{and} \quad \hat{P} = \frac{x}{n}$$

Q1. diabetic patients in a class, the proportion is 0.2.



\hat{P} is a point estimate.

\hat{P}	$f_{\text{sample}}(\hat{P})$
0.33	3 $3/10 = 0.3$
0.67	6 $6/10 = 0.6$
1	1 $1/10 = 0.1$

$$\sum P(\hat{P}) = 1$$

if $x \sim N(\mu, \sigma^2)$
 $\Rightarrow \bar{x} \sim N(\mu, \frac{\sigma^2}{n})$

	\bar{x}	\hat{P}
Mean	$E(\bar{x}) = \mu$	$E(\hat{P}) = p$
Variance	$\sigma_{\bar{x}}^2 = \frac{\sigma^2}{n}$	$\sigma_{\hat{P}}^2 = \frac{pq}{n}$
SD.	$\sigma_{\bar{x}} = \sqrt{\frac{\sigma^2}{n}}$	$\sigma_{\hat{P}} = \sqrt{\frac{pq}{n}}$

$$V(cx) = c^2 V(x)$$

$$E(cx) = c E(x)$$

Proof:

$$\text{Mean } (\bar{x}) = E(\bar{x}) = E(x) = E\left(\frac{x_1 + x_2 + x_3 + \dots + x_n}{n}\right) = \frac{E(x_1) + E(x_2) + E(x_3) + \dots + E(x_n)}{n}$$

$$= \frac{q\mu + q\mu + q\mu + q\mu + q\mu}{n}$$

$$= \frac{5q\mu}{n}$$

$$= \frac{nq\mu}{n}$$

$$= q\mu$$

Bonus

Proof:

$$V(\bar{x}) = \sigma_{\bar{x}}^2 = V(\bar{x}) = V\left(\frac{x_1 + x_2 + \dots + x_n}{n}\right)$$

Poisson Distribution

$$= \frac{V(x_1) + V(x_2) + V(x_3) + \dots + V(x_n)}{n^2}$$

$$= \frac{\sigma^2 + \sigma^2 + \dots + \sigma^2}{n^2}$$

$$= \frac{n\sigma^2}{n^2}$$

$$V(\bar{x}) = \frac{\sigma^2}{n}$$

All \rightarrow population

Shape needs to be understood.

Shape depends on SD.

Normal Distribution

$$z = \frac{x - \mu}{\sigma}$$

Mean

Moving to center of stemites and rotamites

$$E(z) = \frac{E(x) - \mu}{\sigma}$$

$$= \frac{\mu - \mu}{\sigma} = \frac{0}{\sigma} = 0$$

$$V(z) = \frac{V(x) - 0}{\sigma^2}$$

$$= \frac{0^2}{\sigma^2} = 1$$

$np > 5$ and $nq > 5$

RADIANT
PHARMACEUTICALS

ONCE-MONTHLY
Boniva®
ibandronate acid

t -distribution is used when SD of population is not known.

$$t: t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$$

Chi-square: Sum of square of ind. std normal variables

F:

23-Jun-23

Estimator: The statistic which is used to estimate a population

The value of estimator is estimate.

Minimum variance unbiased estimator (MVUE)

Qualities

good estimator

Unbiased

Consistent

Low MSE (Mean Square Error)

Low MVUE

MLE → Maximum Likelihood Estimate

Poisson Distribution → For discrete variables

↳ No of accidents

$$\mu = \lambda$$
$$\bar{x} = \bar{Y}$$

$$\hat{\theta}_{MLE} = \bar{Y}$$

$$\hat{\theta}_{MOM} = \hat{\theta}_{MLE} = \hat{\theta}_{LSE} = \bar{Y}$$

Point Estimate

Interval Estimate

Margin of error

Population Mean
(μ)

\bar{x} Population

$$\bar{x} \pm ME \Rightarrow \bar{x} \pm 2\sigma_{\bar{x}}$$

$$2\sigma_{\bar{x}} = Z \frac{\sigma}{\sqrt{n}}$$

Population Proportion
Sample

Population Proportion

Sample

$$\bar{x} \pm t \sigma_{\bar{x}} = \bar{x} \pm t \frac{s}{\sqrt{n}}$$

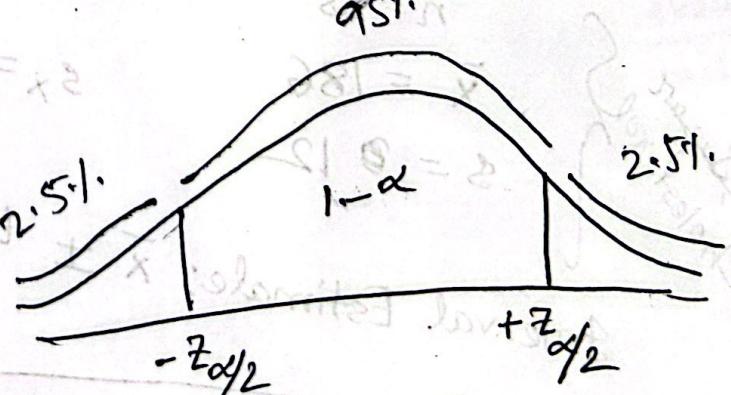
$$t \sigma_{\bar{x}} = t \frac{s}{\sqrt{n}}$$

$$Z = \frac{\bar{x} - \mu}{\sigma_{\bar{x}}}$$

$$Z \sigma_{\bar{x}} = \bar{x} - \mu$$

$$\mu = \bar{x} - Z \sigma_{\bar{x}}$$

Confidence Interval (Proof)



$$P \left[-Z \leq \frac{\bar{x} - \mu}{\sigma_{\bar{x}}} \leq Z \right] = 1 - \alpha$$

RADIANT
PHARMACEUTICALS

ONCE-MONTHLY
Bonova®
ibandronic acid

Population Proportion

Interval Estimate

Margin of error

$$\hat{p} \pm 2 \sqrt{\frac{\hat{p}\hat{q}}{n}}$$

$$2\hat{s}_p = 2 \sqrt{\frac{\hat{p}\hat{q}}{n}}$$

$$df = n - 1$$

proportion
always

Sample size
25, 50, 100

Sample proportion
draw
inaccurate

$$Z = \frac{\hat{p} - p}{\hat{s}_p}$$

$$p = \hat{p} - Z\hat{s}_p$$

$$P = \hat{p} \pm 2\hat{s}_p$$

$$n = 25$$

$$\bar{x} = 145$$

$$\sigma_{\bar{x}} = 35$$

$$\sigma = 35$$

ALL SUCH TEXTBOOKS

population SD ~~for~~

Preferably, just look
at for Confidence
interval

$$n = 25$$

$$\bar{x} = 186$$

$$s = 12$$

$$s_x = \frac{s}{\sqrt{n}} = \frac{12}{\sqrt{25}} \quad df = 24$$

For these maths, always
use the two-tail table.

Interval Estimate: $\bar{x} \pm t s_x$

Sugar
Cholesterol

* If sigma not given, ALWAYS go for the sigma

because population sigma not such sample

503

7-Jul-23

- Sir reviewed previous classes (Point & Interval Estimates if σ is known/unknown & for proportion)
- Sir did the textbook math again
- Sir " " Dr. Moore cholesterol math again
 - ↳ If we construct a lot of samples, 95% of the parameters will follow under this interval estimate)

Confidence level যাত্রের interval বুঝ হবে।
means Margin of error বুঝ হবে।

$$\bar{x} \pm z \sigma_{\bar{x}}$$

$$\text{for } 95\%, z = 1.96$$

$$\text{for } 99\%, z = 2.58$$

For increasing confidence interval, the margin of error increases.

proportion এর large sample কির অন্ত?

বিশ্ব কীভাবে মুক্ত হয়ে large sample?

$$n\hat{p} > 5 \text{ and } n\hat{q} > 5$$

P এর আসর
কোথা থাই।

Example math of slide (Religion)

$$n = 1000$$

$$\hat{p} = 0.44; \hat{q} = 0.56$$

a) Point estimate of P is $\hat{p} = 0.44$

b) CI = 99%.

$$\text{Interval Estimate} = \hat{p} \pm z_{SP} = \hat{p} \pm \sqrt{\frac{\hat{p}\hat{q}}{n}}$$

Margin of error

Alcohol Math

$$n = 2401, \hat{p} = 0.25, \hat{q} = 0.75$$

$$n\hat{p} = 2401 \times 0.25$$

$$a) \hat{p} = 0.25$$

$$b) CI = \hat{p} \pm z_{SP} = \hat{p} \pm 2\sqrt{\frac{\hat{p}\hat{q}}{n}}$$

Quiz on 14th July

Method of Moment (MOM)

$$E(x') = mr'$$

r^{th} population

First raw moment,
always population mean

$Y_i \sim \text{Poisson}(\lambda)$

mean
 $= E(Y)$

Mean = Variance = λ

By MOM,
 $\mu_1' = m'$
 $E(Y) = \frac{1}{n} \sum_{i=1}^n Y_i$
 $\bar{x} = \bar{Y}$ (Sample Mean)

poisson \rightarrow count data
 ↓
 (variable)
 for discrete data; i.e.
 No of accidents

$$\hat{\lambda}_{MOM} = \bar{Y}$$

Method of Maximum Likelihood (ML)

$$L(\theta_1, \theta_2, \dots, \theta_k; y_1, \dots, y_n) = \prod_{i=1}^n f(y_i; \theta_1, \theta_2, \dots, \theta_k)$$

↓
 population parameter

y যেহেতু n এক্ষয়ক, তাহলে n সম্পর্ক ফাংশন আছে,

$$f(y_1; \theta_1, \theta_2, \theta_3, \dots, \theta_k)$$

$$f(y_2; \theta_1, \theta_2, \theta_3, \dots, \theta_k)$$

$$f(y_3; \theta_1, \theta_2, \theta_3, \dots, \theta_k)$$

আজ্ঞা estimate করতে চাহিএ θ

↓ Said as Likelihood of
 θ factor & y factor

Data has to be independent

RADIANT
PHARMACEUTICALS

ONCE-MONTHLY
Boniva®
ibandronic acid

Likelihood is product of marginal probability function.

$$L(\theta; y) = \prod_{i=1}^n f(y_i; \theta) \rightarrow \text{For single parameter}$$

$y_i \sim_{\text{iid}} \text{Poisson}(\theta)$

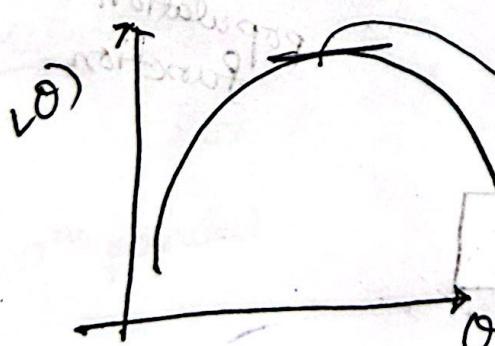
y_i follows poisson distribution with θ parameter

$$f(y_i; \theta) = \frac{e^{-\theta} \theta^{y_i}}{y_i!}$$

Defn $L(\theta; y) = \prod_{i=1}^n f(y_i; \theta)$

$$\prod_{i=1}^n e^{-\theta} = e^{-\theta} \cdot e^{-\theta} \dots e^{-\theta} = e^{-n\theta}$$

$$\prod_{i=1}^n \theta^{y_i} = \theta^{y_1} \cdot \theta^{y_2} \dots \theta^{y_n} = \theta^{y_1 + y_2 + \dots + y_n}$$



$\frac{dy}{dx} \rightarrow \times \text{प्रथम घातके } y \text{ का rate of change}$

$\frac{dL(\theta)}{d\theta} \rightarrow \theta \text{ प्रथम घातके Likelihood}$

Likelihood maximum क्या है?

Likelihood is not a probability function, it just estimates which one of the parameters is more likely.

$y_i \sim_{iid}$ Bernoulli(θ)
 ↓
 Binomial
 but
 $n=1$

Probability of success

if $y_i \sim_{iid}$ Binomial(θ)

for Binomial, $f(y_i, \theta) = n C y_i \theta^{y_i} (1-\theta)^{n-y_i}$

For Bernoulli, $f(y_i, \theta) = \theta^{y_i} (1-\theta)^{1-y_i}$

$$L = \prod_{i=1}^4 f(y_i; \theta) = (\theta^{y_1} (1-\theta)^{1-y_1}) \cdot (\theta^{y_2} (1-\theta)^{1-y_2}) \cdot (\theta^{y_3} (1-\theta)^{1-y_3}) \cdot (\theta^{y_4} (1-\theta)^{1-y_4})$$

$$= \prod_{i=1}^4 \theta^{y_i} (1-\theta)^{1-y_i}$$

(0,0,0,0)

Case-I:

$$= \theta^{\sum y_i} (1-\theta)^{4-\sum y_i}$$

$$\text{For } \theta = \theta_1 = 0.3 \Rightarrow L = (0.3)^{\sum y_i} (0.7)^{4-\sum y_i} = 0.2401$$

$$\theta_2 = 0.8 \Rightarrow L = (0.8)^{\sum y_i} (0.2)^{4-\sum y_i} = 0.0016$$

$\theta = 0.3$ is more likely.

Case III: $(1, 1, 1, 0)$

(a) all normal \rightarrow iid \rightarrow
known to fit distribution

Binomial
law
 $f = m$

(b) Binomial \rightarrow iid \rightarrow y_i

iid = independent
& identically
distributed

HW
Construct likelihood for

$y_i \sim$ iid Binomial (n, p) known

$y_i \sim$ iid Poisson (λ)

$y_i \sim$ iid Normal (μ, σ^2)

Normal:

$$f(y_i; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y_i - \mu)^2}$$

$-\infty < y_i < \infty$

$-\infty < \mu < \infty$

$\sigma^2 > 0$

Mean = μ

Variance = σ^2

Binomial:

$$f(y; p) = {}^m C_y p^y (1-p)^{m-y}$$

Mean = mp

Variance = mpq