

How many times do you need to toss a fair coin to get 10 Heads?

If we toss it n times, $\mathbb{E}(\#\text{Heads}) = \mathbb{E}\left[\sum_{j=1}^n X_j\right] = \sum_{j=1}^n \mathbb{E}[X_j] = \frac{n}{2}$.
 $X_j \stackrel{d}{=} \text{Ber}\left(\frac{1}{2}\right)$

So, we want $\frac{n}{2} = 10$, $n = 20$... But that's not really answering the question.

$\{X_1, X_2, \dots\}$ iid. $\text{Ber}\left(\frac{1}{2}\right)$ $\mathcal{F}_n = \sigma(X_1, X_2, \dots, X_n)$

$\tau_k = \inf\{n \geq 0 : \sum_{j=1}^n X_j = k\}$

The question is : what is $\mathbb{E}[\tau_{10}]$?

$\{\tau_k = n\} = \left\{ \sum_{j=1}^n X_j = k, \sum_{j=1}^{n-1} X_j = k-1 \right\} \in \mathcal{F}_n$.

stopping time.

Theorem: (Wald's Identity)

Let $\{X_n\}_{n=1}^{\infty}$ be iid. random variables, and let τ be a stopping time relative to $\mathcal{F}_n = \sigma(X_1, \dots, X_n)$. If $f \geq 0$, or if $f(X_n) \in L^1$ and $E[\tau] < \infty$, then

$$E\left[\sum_{n=1}^{\tau} f(X_n)\right] = E[f(X_1)] E[\tau].$$

Pf. First, assume $f \geq 0$.

$$E\left[\sum_{n=1}^{\tau} f(X_n)\right] = E\left[\sum_{n=1}^{\infty} f(X_n) \mathbb{1}_{n \leq \tau}\right] \stackrel{\text{Fubini/Tonelli}}{\downarrow} \sum_{n=1}^{\infty} E[f(X_n) \mathbb{1}_{n \leq \tau}].$$

$$\{n \leq \tau\} = \{\tau \leq n-1\}^c$$

$$\in \mathcal{F}_{n-1}$$

$$\therefore \text{by Doob-Dynkin} \\ \mathbb{1}_{n \leq \tau} = F_n(X_1, \dots, X_{n-1})$$

$$\begin{aligned} & \sum_{n=1}^{\infty} \sum_{k \geq n} \underbrace{E[\mathbb{1}_{\tau \geq k}]}_{P(\tau \geq k)} \\ &= \sum_{k=1}^{\infty} \sum_{n \leq k} P(\tau \geq k) \\ &= \sum_{k=1}^{\infty} k P(\tau \geq k) \\ &= \sum_{k=1}^{\infty} k \underbrace{E[\mathbb{1}_{\tau \geq k}]}_{E[\mathbb{1}_{n \leq \tau}]} \\ &= E[f(X_1)] \sum_{n=1}^{\infty} E[\mathbb{1}_{n \leq \tau}] \end{aligned}$$

Repeat, apply to $|f|$.

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E.g. Roll a die, yielding some value $D \in \{1, 2, 3, 4, 5, 6\}$.

Now roll the die D times, and add up the D values.

What's the expected sum of these D rolls?

$D_0 = D$, $\{D_n\}_{n=0}^{\infty}$ iid die rolls $F_n = \sigma(D_0, D_1, \dots, D_n)$.

$$D_n \stackrel{d}{=} \text{Unif}\{1, \dots, 6\}.$$

$\{D = k\} \in F_0 \subseteq F_k \quad \therefore D$ is a stopping time.

$$\mathbb{E}\left[\sum_{j=0}^D D_j\right] = \mathbb{E}[D_0] \mathbb{E}[D] = \mathbb{E}[D]^2$$

$$\mathbb{E}[D_0] + \mathbb{E}\left[\sum_{j=1}^D D_j\right]$$

$$\hookrightarrow = \mathbb{E}[D]^2 - \mathbb{E}[D]$$

$$= 8.75.$$

Eg. (Gambler's Ruin, revisited)

Let $(X_n)_{n \geq 1}$ be a random walk on \mathbb{Z} , $P(X_{n+1} = k+1 | X_n = k) = p \in (0, 1)$.

Then we can construct it as $X_n = \sum_{k=1}^n \zeta_k$ $\{\zeta_k\}_{k \geq 1}$ iid, $\zeta_k \stackrel{d}{=} p\delta_1 + (1-p)\delta_{-1}$

How long does it take, starting at 0, to reach $k \neq 0$?

$$X_0 = 0$$

$$\tau = \inf\{n \geq 1 : X_n = k\}$$

$$k = E[X_\tau] = E\left[\sum_{n=1}^{\tau} \zeta_n\right] = E[\zeta_1]E[\tau] = (p - (1-p))E[\tau].$$

$$\therefore E[\tau] = \frac{k}{p - (1-p)} \quad ! \quad \text{But this is } < \infty \text{ if } k, p - \frac{1}{2} \text{ have opposite signs.}$$

→ If $E[\tau] \not< \infty$, then $k = E[\zeta_1]E[\tau]$

$p = \frac{1}{2} :$

With $p = \frac{1}{2}$, we can conclude $E[\tau] = \infty$.