

Let  $\{Y_n\}_{n=1}^{\infty}$  be uncorrelated random variables in  $L^2$ .

I.e.  $\text{Cov}(Y_n, Y_m) = 0$  for  $n \neq m$

$$E[\overset{\circ}{Y}_n \overset{\circ}{Y}_m] = \langle \overset{\circ}{Y}_n, \overset{\circ}{Y}_m \rangle_{L^2} \quad \overset{\circ}{Y}_n = Y_n - E[Y_n]$$

i.e.  $\{\overset{\circ}{Y}_n\}$  is an orthogonal seq. in  $L^2$ .

Prop: If  $\{Y_n\}_{n=1}^{\infty}$  are uncorrelated, and  $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$ ,  
then  $\sum_{n=1}^{\infty} (Y_n - E[Y_n])$  converges in  $L^2$ .

Pf.  $\text{Var}(Y_n) = \text{Cov}(Y_n, Y_n) = E[\overset{\circ}{Y}_n^2] = \|\overset{\circ}{Y}_n\|_{L^2}^2$   
 $\overset{\circ}{S}_n = \sum_{j=1}^n \overset{\circ}{Y}_j \quad \|\overset{\circ}{S}_n - \overset{\circ}{S}_m\|_{L^2} = \left\| \sum_{j=m+1}^n \overset{\circ}{Y}_j \right\|_{L^2} = \sum_{j=m+1}^n \|\overset{\circ}{Y}_j\|_{L^2}^2 \rightarrow 0 \text{ as } n, m \rightarrow \infty,$   
 $\therefore \overset{\circ}{S}_n = \sum_{j=1}^n (Y_j - E[Y_j])$  is Cauchy in  $L^2$ .  
 $\therefore$   $\overset{\uparrow}{\text{convergent}}$ . ///

We would like to upgrade this convergence from  $L^2$  to a.s. There's no reason for that to be true in general (orthogonal sums in  $L^2$  usually fail to converge a.s.), unless we also upgrade the orthogonality to super-orthogonality  
— i.e. independence.

**Theorem:** (Kolmogorov's Convergence Criterion)

Let  $\{Y_n\}_{n=1}^{\infty}$  be independent  $L^2$  random variables.

If  $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$ , then  $\sum_{n=1}^{\infty} (Y_n - E[Y_n])$  converges a.s.

In particular, if in addition  $\sum_{n=1}^{\infty} E[Y_n] < \infty$ , then  $\left\{ \sum_n Y_n = \sum_n Y_n + \sum_n E[Y_n] \right\}$  converges a.s. and in  $L^2$ .

## Maximal Inequalities

Let  $\{Y_n\}_{n=1}^{\infty}$  be independent r.v.'s, with  $E[Y_n] = 0$

Set  $S_n = Y_1 + \dots + Y_n$ . If  $Y_n \in L^2$ , then Markov

tells us

$$P(|S_n| \geq \varepsilon) \leq \frac{1}{\varepsilon^2} E[S_n^2] = \frac{1}{\varepsilon^2} \sum_{j=1}^n E[Y_j^2]$$

But what can we say about

$$S_n^* := \max_{1 \leq j \leq n} |S_j|$$

Turns out: the Markov conclusion still applies.

Theorem: (Kolmogorov's Maximal Inequality)

With  $Y_n, S_n$  as above,

$$P(S_n^* \geq \varepsilon) \leq \frac{1}{\varepsilon^2} E[S_n^2 \mathbb{1}_{S_n^* \geq \varepsilon}] \quad \forall \varepsilon > 0$$

$$\leq \frac{1}{\varepsilon^2} E[S_n^2] = \frac{1}{\varepsilon^2} \sum_{j=1}^n E[Y_j^2].$$

Pf. Fix  $\varepsilon > 0$ , and set  $\tau := \inf \{j \in \mathbb{N} : |S_j| \geq \varepsilon\}$   
 $(\inf \emptyset := \infty)$

$T$  is a random variable. Note that

$$\{\tau = j\} = \{ |S_1| < \varepsilon, |S_2| < \varepsilon, \dots, |S_{j-1}| < \varepsilon, |S_j| \geq \varepsilon \}$$

Now,  $\{S_n^* \geq \varepsilon\} = \left\{ \max_{1 \leq j \leq n} |S_j| \geq \varepsilon \right\} = \{\exists j \in [n] : |S_j| \geq \varepsilon\} = \{\tau \leq n\}$

**Notation:**  $E[X : A] := E[X|A]$

$$\therefore \mathbb{E}[S_n^2 : S_n \geq \varepsilon] = \mathbb{E}[S_n^2 : \tau \leq n] = \sum_{j=1}^n \mathbb{E}[S_n^2 : \tau = j]$$

$\uparrow$

$$\mathbb{1}_{\{\tau \leq n\}} = \sum_{j=1}^n \mathbb{1}_{\{\tau = j\}}$$

Now a trick:

$$S_n^2 = (S_j + S_n - S_j)^2 = S_j^2 + (S_n - S_j)^2 + 2S_j(S_n - S_j)$$

$$\mathbb{E}[S_n^2 : S_n^* \geq \varepsilon] = \sum_{j=1}^n \mathbb{E}\left[(S_j^2 + (S_n - S_j)^2 + 2S_j(S_n - S_j)) \mathbb{1}_{\{\tau=j\}}\right]$$

$\downarrow$

$$\mathbb{E}\left[(S_n - S_j) \cdot S_j \mathbb{1}_{\{\tau=j\}}\right] = \mathbb{E}[S_n - S_j] \mathbb{E}[S_j : \tau=j]$$

$\sigma(Y_{j+1}, \dots, Y_n)$        $\sigma(Y_1, \dots, Y_j)$

$$= \sum_{j=1}^n (\mathbb{E}[S_j^2 : \tau=j] + \mathbb{E}[(S_n - S_j)^2 : \tau=j])$$

$$\geq \sum_{j=1}^n \mathbb{E}[S_j^2 : \tau=j]$$

$$\left( \mathbb{E}[S_j^2 \mathbb{1}_{\{\tau=j\}}] \subseteq \{\tau=j\} \setminus \{ |S_j| > \varepsilon \} \right)$$

$$S_j^2 \mathbb{1}_{\{\tau=j\}} \geq \varepsilon^2 \mathbb{1}_{\{\tau=j\}}$$

$$\geq \sum_{j=1}^n \mathbb{E}[\varepsilon^2 \mathbb{1}_{\{\tau=j\}}] = \varepsilon^2 \sum_{j=1}^n P(\tau=j) = \frac{\varepsilon^2}{\varepsilon^2} P(\tau \leq n) = \varepsilon^2 P(S_n^* \geq \varepsilon) \quad //$$

Theorem: (Kolmogorov's Convergence Criterion)

Let  $\{Y_n\}_{n=1}^{\infty}$  be independent  $L^2$  random variables.

If  $\sum_{n=1}^{\infty} \text{Var}(Y_n) < \infty$ , then  $\sum_{n=1}^{\infty} Y_n$  converges a.s.

Pf. Let  $S_n = \sum_{j=1}^n Y_j$ . For  $m < n$ ,  $S_n - S_m = Y_{m+1} + \dots + Y_n$ .

Apply Kolmogorov's Maximal Inequality:

$$P\left(\max_{m < j \leq n} |S_j - S_m| \geq \frac{\varepsilon}{2}\right) \leq \frac{1}{(\varepsilon/2)^2} E[(S_n - S_m)^2] = \frac{4}{\varepsilon^2} \sum_{j=m+1}^n E[Y_j^2] = \frac{4}{\varepsilon^2} \sum_{j=m+1}^n \text{Var}(Y_j)$$

Now let  $n \rightarrow \infty$ ; we have

$$\overbrace{P\left(\sup_{j \geq m} |S_j - S_m| \geq \frac{\varepsilon}{2}\right)} \leq \frac{4}{\varepsilon^2} \sum_{j=m+1}^{\infty} \text{Var}(Y_j) \rightarrow 0 \text{ as } m \rightarrow \infty,$$

Further,  $\sup_{j, k \geq m} |S_j - S_k| = \sup_{j, k \geq m} |S_j - S_m + S_m - S_k| \leq \sup_{j \geq m} |S_j - S_m| + \sup_{k \geq m} |S_k - S_m|$

$$= 2 \sup_{j \geq m} |S_j - S_m|$$

$$\left\{ \sup_{j, k \geq m} |S_j - S_k| \geq \varepsilon \right\} \subseteq \left\{ 2 \sup_{j \geq m} |S_j - S_m| \geq \varepsilon \right\}$$
$$\Rightarrow P(\ ) \leq P(\ ) \rightarrow 0 \text{ as } m \rightarrow \infty.$$

∴ We have proven that,  $\forall \varepsilon > 0$ ,

$$P\left(\sup_{j,k \geq m} |S_j - S_k| > \varepsilon\right) \rightarrow 0 \text{ as } m \rightarrow \infty$$

I.e. the random variables  $S_m := \sup_{j,k \geq m} |S_j - S_k| \rightarrow_p 0$ .

But  $S_m \downarrow$  and  $S_m \geq 0$ , ∴  $S := \lim_{m \rightarrow \infty} S_m$  exists (surely).  $\begin{cases} S = 0, \\ S > 0. \end{cases}$

$$\therefore \sup_{j,k \geq m} |S_j - S_k| \rightarrow 0 \text{ a.s. as } m \rightarrow \infty$$

i.e.  $\{S_j\}_{j=1}^\infty$  is a.s. Cauchy.

$$\sum_{k=1}^j Y_k \quad \therefore \text{a.s. convergent.}$$

E.g. Let  $\{X_n\}_{n=1}^\infty$  be iid Rademacher r.v.'s:  $P(X_n = \pm 1) = \frac{1}{2}$ .

Does the series  $\sum_{n=1}^\infty \frac{X_n}{n}$  converge? [ $\pm \sum_{n=1}^\infty \frac{1}{n}$  diverges;  $\sum_{n=1}^\infty \frac{(-1)^n}{n}$  converges]

$$Y_n = \frac{X_n}{n} \text{ resp., } L^2, \quad \tilde{Y}_n = Y_n \quad \text{Var}(Y_n) = E[\tilde{Y}_n^2] = \frac{1}{n^2} E[X_n^2] = \frac{1}{n^2}$$

$$\therefore \sum_{n=1}^\infty \text{Var}(Y_n) < \infty,$$

$$\therefore \sum_{n=1}^\infty \tilde{Y}_n = \sum_{n=1}^\infty \frac{X_n}{n} \text{ converges a.s.}$$