

Theorem: (Cramér's Large Deviation Theorem)

Let  $\{X_n\}_{n=1}^{\infty}$  be iid. exponentially integrable random variables.

Let  $\overset{*}{X}_n = X_n - \mathbb{E}[X_n]$ , and  $\overset{*}{S}_n = \overset{*}{X}_1 + \dots + \overset{*}{X}_n$ . For any  $\varepsilon > 0$ ,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log P(\overset{*}{S}_n > n\varepsilon) = -\Psi_{\overset{*}{X}_1}^*(\varepsilon) \leq 0.$$

We have already proved that

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log P(\overset{*}{S}_n > n\varepsilon) \leq -\Psi_{\overset{*}{X}_1}^*(\varepsilon).$$

Thus, it suffices to prove that

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log P(\overset{*}{S}_n > n\varepsilon) \geq -\Psi_{\overset{*}{X}_1}^*(\varepsilon).$$

For the sake of readability, set  $\Psi(t) := \Psi_{\overset{*}{X}_1}(t)$   $M(t) := M_{\overset{*}{X}_1}(t)$   
 $\Psi^*(\varepsilon) := \Psi_{\overset{*}{X}_1}^*(\varepsilon)$

Observations about  $\Psi^*$ :

- Since  $\Psi = \log M$  is smooth, and convex,

$$\Psi^*(\eta) = \sup_{t \geq 0} (\eta t - \Psi(t)) = \eta t_\eta - \Psi(t_\eta) \quad \text{for some maximizer } t_\eta > 0.$$

$$0 = \frac{d}{dt} (\eta t - \Psi(t)) \Big|_{t=t_\eta} = \eta - \Psi'(t_\eta)$$

$$\Psi'(t_\eta) = \eta$$

Useful trick:

Given exponentially integrable  $Y$  (w.r.t.  $(\Omega, \mathcal{F}, P)$ ),

for each  $t$  define a new probability measure  $P_t$  on  $(\Omega, \mathcal{F})$ :

$$dP_t = \frac{e^{tY}}{M_Y(t)} dP$$

$$\int_{\Omega} dP_t = \int \frac{e^{tY}}{M(t)} dP = \frac{1}{M(t)} \int e^{tY} dP = 1.$$

$\underbrace{\mathbb{E}[e^{tY}] = M(t)}$

In the proof, we will use independent samples from  $P_t$

with  $t = t_m$  (maximizer defining  
 $\Psi_{X_1}^*(\eta)$ )

Prop. For all  $\varepsilon > 0$ ,

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log P(\overset{\circ}{S}_n > n\varepsilon) \geq -\underline{\Psi}^*(\varepsilon)$$
$$= \underline{\Psi}(t_\varepsilon) - \varepsilon t_\varepsilon$$

Pf. Any maximizer  $t_\varepsilon$  of  $t \mapsto \varepsilon t - \underline{\Psi}(t)$  is a solution of

$$\underline{\Psi}'(t) = \varepsilon.$$

$$\frac{d}{dt} \log M(t) \stackrel{II}{=} \frac{M'(t)}{M(t)} = \frac{\mathbb{E}[e^{t\overset{\circ}{X}_1}]}{M(t)} = \int \overset{\circ}{X}_1 \frac{e^{t\overset{\circ}{X}_1}}{M(t)} dP$$

II

I.e.  $t = t_\varepsilon$  is characterized by the requirement that

$$\mathbb{E}_{P_t}[\overset{\circ}{X}_1] = \varepsilon.$$

Let  $\{W_n\}_{n=1}^{\infty}$  be i.i.d. random variables with

$$W_n \stackrel{d}{=} \overset{\circ}{X}_1 * P_t = \text{Law}_{P_t}(\overset{\circ}{X}_1)$$

$$\mu_{W_n}(B) = P_t(\overset{\circ}{X}_1 \in B) = \int_{\overset{\circ}{X}_1^{-1}(B)} \frac{e^{tx_1}}{M(t)} dP$$

$$\mathbb{E}_{P_t}[\overset{\circ}{X}_1] = \varepsilon$$

$$\Rightarrow \mathbb{E}_P[W_n] = \varepsilon$$

$$= \int_B \frac{e^{tx}}{M(t)} \mu_{\overset{\circ}{X}_1}(dx)$$

$$\therefore \mu_{W_n}(dx) = \frac{e^{tx}}{M(t)} \mu_{\overset{\circ}{X}_1}(dx)$$

$$\therefore \mu_{\overset{\circ}{X}_1}(dx) = M(t) e^{-tx} \mu_{W_n}(dx)$$

Then for measurable  $f: \mathbb{R}^n \rightarrow [0, \infty]$ ,

$$\begin{aligned} \mathbb{E}[f(\overset{\circ}{X}_1, \dots, \overset{\circ}{X}_n)] &= \int_{\mathbb{R}^n} f(x_1, \dots, x_n) \mu_{\overset{\circ}{X}_1}(dx_1) \cdots \mu_{\overset{\circ}{X}_n}(dx_n) \\ &= \int_{\mathbb{R}^n} f(x_1, \dots, x_n) M(t)^n e^{-t(x_1 + \cdots + x_n)} \mu_{W_n}(dx_1) \cdots \mu_{W_n}(dx_n) \\ &= M(t)^n \mathbb{E}[f(W_1, \dots, W_n) e^{-t(w_1 + \cdots + w_n)}] \end{aligned}$$

Let  $T_n = W_1 + \dots + W_n$ .

$\therefore$  We've shown that, for measurable  $f: \mathbb{R}^n \rightarrow [0, \infty]$ ,

$$\mathbb{E}[f(\overset{\circ}{X}_1, \dots, \overset{\circ}{X}_n)] = M(t)^n \mathbb{E}[f(W_1, \dots, W_n) e^{-t T_n}]$$

Apply this with  $f(x_1, \dots, x_n) = \mathbb{1}_{x_1 + \dots + x_n > n\epsilon}$ .

$$\begin{aligned} P(\overset{\circ}{S}_n > n\epsilon) &\approx M(t)^n \mathbb{E}[e^{-t T_n} \mathbb{1}_{T_n \geq n\epsilon}] \\ &\quad \Downarrow \\ &\quad \mathbb{1}_{n\epsilon \leq T_n} \leq n(\epsilon + \delta) \quad \delta > 0 \\ &\geq M(t)^n e^{-tn(\epsilon+\delta)} P(n\epsilon \leq T_n \leq n(\epsilon+\delta)) \\ &= e^{n(\bar{\Phi}(t) - t\epsilon)} e^{-nt\delta} P(n\epsilon \leq T_n \leq n(\epsilon+\delta)) \\ &= e^{-n\bar{\Phi}^*(\epsilon)} e^{-nt\delta} P(n\epsilon \leq T_n \leq n(\epsilon+\delta)) \end{aligned}$$

$$P(\overset{\circ}{S}_n > n\epsilon) \geq e^{-n\bar{\Psi}^*(\epsilon)} e^{-nt_\epsilon \delta} P(n\epsilon \leq T_n \leq n(\epsilon+\delta))$$

$$\therefore \liminf_{n \rightarrow \infty} \frac{1}{n} \log P(\overset{\circ}{S}_n > n\epsilon) \geq -\bar{\Psi}^*(\epsilon) - t_\epsilon \delta$$

$$+ \liminf_{n \rightarrow \infty} \frac{1}{n} \log P(n\epsilon \leq T_n \leq n(\epsilon+\delta)) \quad (\star)$$

for any  $\delta > 0$ .

To complete the proof, we want to let  $\delta \downarrow 0$ . This means we have to show that  $(\star)$  is well-behaved as  $n \rightarrow \infty$ .

Note:  $E[w_n] = \epsilon$ , so  $E[T_n] = E[w_1 + \dots + w_n] = n\epsilon$ .

$$\begin{aligned} \therefore P(n\epsilon \leq T_n \leq n(\epsilon+\delta)) &\geq P(0 \leq T_n - n\epsilon \leq n\delta) \\ &= P(0 \leq \overset{\circ}{T}_n \leq n\delta) = P(0 \leq \frac{\overset{\circ}{T}_n}{\sqrt{n \text{Var}(w_1)}} \leq \sqrt{\frac{n}{\text{Var}(w_1)}} \delta) \end{aligned}$$

\* Flash Forward \*

$$\text{CLT} : P\left(\frac{\overset{\circ}{T}_n}{\sqrt{n \text{Var}(w_1)}} \geq 0\right) \xrightarrow{n \rightarrow \infty} \int_0^\infty \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \frac{1}{2}$$

$\rightarrow \infty$

///