

Quick Rev

Sampling distribution (χ^2 , t & F distribution)

Content Covered:

- χ^2 - distribution
- t - distribution
- F - distribution.

Note that:

* Let $x_1, x_2, x_3, \dots, x_n$ be a random sample from a population with mean μ and Variance σ^2 , M.g.f. $M_X(t)$. Then

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim N(0, 1) \text{ as } n \rightarrow \infty.$$

* If population is itself Normal then

$$\bullet \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim N(0, 1) \text{ irrespective of large or small } n.$$

$$\Rightarrow \bar{X} - \mu \sim N\left(0, \frac{\sigma^2}{n}\right)$$

$$\Rightarrow \bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right).$$

* Population $A \sim N(\mu_1, \sigma_1^2)$ Population $B \sim N(\mu_2, \sigma_2^2)$

$$\Rightarrow \bar{X}_1 \sim N\left(\mu_1, \frac{\sigma_1^2}{n_1}\right) \quad \bar{X}_2 \sim N\left(\mu_2, \frac{\sigma_2^2}{n_2}\right)$$

$$\bar{X}_1 - \bar{X}_2 \sim N\left(\mu_1 - \mu_2, \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)$$

↓
Variance always gets added.

$$\frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0, 1)$$

Note that: ^{or} If populations are not Normal then this is valid for large n . i.e.

i.e., $n > 30$.

Chi-square distribution:

A Continuous random variable X is $\chi^2_{(n)}$ if its p.d.f. is given by

$$f(x) = \begin{cases} \frac{1}{2^{n/2} \sqrt{\pi}} e^{-x/2} x^{n/2-1} & x > 0 \\ 0 & \text{o/w} \end{cases}$$

$$\int_0^{\infty} f(x) dx = \int_0^{\infty} \frac{1}{2^{n/2} \sqrt{\pi}} e^{-x/2} x^{n/2-1} dx$$

$$= \frac{1}{2^{n/2} \sqrt{\pi}} \int_0^{\infty} e^{-x/2} x^{n/2-1} dx$$

$$\therefore \int_0^{\infty} e^{-qx} x^{n-1} dx = \frac{\Gamma(n)}{q^n}$$

$$= \frac{1}{2^{n/2} \sqrt{\pi}} \frac{\sqrt{\pi}}{(1/2)^{n/2}} = 1$$

So X is basically $\text{Gamma}(n/2, 1/2)$

$$X \sim G(\alpha, \beta)$$

$$f(x) = \frac{\beta^\alpha e^{-\beta x} x^{\alpha-1}}{\Gamma(\alpha)} \quad x > 0.$$

$$E(X) = \frac{\alpha}{\beta}$$

$$\text{Var}(X) = \frac{\alpha}{\beta^2}$$

$$M_X(t) = \left(1 - \frac{t}{\beta}\right)^{-\alpha}; t < \beta.$$

$$\Rightarrow E(X) = \frac{\binom{n}{2}}{\frac{1}{2}} = n$$

$$\Rightarrow \text{Var}(X) = \frac{\binom{n}{2}}{\left(\frac{1}{4}\right)} = 2n$$

$$\Rightarrow M_X(t) = (1-2t)^{-n/2} \quad |t| < \frac{1}{2}.$$

i.e., If $X \sim \chi^2_{(n)}$ then

$$E(X) = n, \quad \text{Var}(X) = 2n$$

$$\& \quad M_X(t) = (1-2t)^{-n/2}; \quad |t| < \frac{1}{2}.$$

Results: If $Y \sim N(0,1)$

$$\text{then } X = Y^2 \\ \sim \chi^2_{(1)}$$

If $X_i \sim \chi^2_{(1)}$, all X_i are independent

$$\Rightarrow \sum_{i=1}^n X_i \sim \chi^2_{(n)}$$



$$P(X \leq x) = F_X(x)$$

$$F_X(x) = P(Y^2 \leq x)$$

$$= P(|Y| \leq \sqrt{x})$$

$$= P(-\sqrt{x} \leq Y \leq \sqrt{x})$$

$$\therefore \gamma \sim N(0, 1)$$

$$f(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}$$

$$= P(-\sqrt{x} \leq \gamma \leq \sqrt{x})$$

$$= \int_{-\sqrt{x}}^{\sqrt{x}} f(y) dy$$

$$= \int_{-\sqrt{x}}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy$$

$$\Rightarrow F_X(x) = \int_{-\sqrt{x}}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy$$

we know that

$$\frac{d}{dx}(F_X(x)) = f(x)$$

we know by Leibnitz integral rule

$$\begin{aligned} \frac{d}{dx} \left(\int_{a(x)}^{b(x)} f(x, t) dt \right) &= f(x, b(x)) \frac{d}{dx}(b(x)) \\ &\quad - f(x, a(x)) \frac{d}{dx}(a(x)) \\ &\quad + \int_{a(x)}^{b(x)} \frac{\partial}{\partial x} f(x, t) dt. \end{aligned}$$

In particular

$$\frac{d}{dx} \int_{a(x)}^{b(x)} f(t) dt = f(b(x)) \frac{d}{dx}(b(x)) - f(a(x)) \frac{d}{dx}(a(x)).$$

Therefore,

$$\frac{d}{dx} F_X(x) = f(x) = \frac{d}{dx} \left(\int_{-\sqrt{x}}^{\sqrt{x}} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy \right)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{x}{2}} \frac{d}{dx}(\sqrt{x})$$

$$- \frac{1}{\sqrt{2\pi}} e^{-\frac{x}{2}} \frac{d}{dx}(-\sqrt{x})$$

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x}{2}} \left(\frac{1}{2\sqrt{x}} \right) + \frac{1}{\sqrt{2\pi}} e^{-\frac{x}{2}} \left(\frac{1}{2\sqrt{x}} \right)$$

$$f(x) = \frac{1}{\sqrt{2\pi} \sqrt{x}} e^{-\frac{x}{2}}$$

$$f(x) = \frac{1}{2^{\frac{k}{2}} \sqrt{k}} e^{-\frac{x}{2}} x^{\frac{k}{2}-1} \quad 0 < x < \infty$$

$$X \sim \chi^2_{(1)} \quad \because \left(\frac{X-4}{6} \right)^2 \sim \chi^2_{(1)}$$

\therefore square of a standard normal $\sim \chi^2_{(1)}$

$$\left(\frac{X_i - \mu}{\sigma}\right)^2 \sim \chi^2_{(1)}$$

$$\sum_{i=1}^n \left(\frac{X_i - \mu}{\sigma}\right)^2 \sim \chi^2_{(n)}$$

and $\text{Gamma}\left(\frac{1}{2}, \frac{1}{2}\right) = \chi^2_{(1)}$

Results: If

$X_i \sim N(0, 1)$ be independent
random variables

then $Y_n = \sum_{i=1}^n X_i^2 \sim \chi^2_{(n)}$

Result: If \bar{X} and S^2 are the mean and
variance of a random sample of size n

from a normal population with mean μ
and standard deviation σ , then \bar{X} and

S^2 are independent and

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{(n-1)}$$