

- The generalized voltage v_k
 - $v_k \triangleq \Pr[\text{A random walk starting at } k \text{ visits } j \text{ before } i]$
- Calculating $\mathbf{v} \triangleq (v_1 \ v_2 \ \dots \ v_n)^\top$
 - $v_i = 0$ and $v_j = 1$
 - $\forall k \neq i, j, v_k = \sum_l p_{kl} \cdot v_l$
 - Split $\mathbf{P} = \begin{pmatrix} \hat{\mathbf{P}} & \mathbf{c}_i & \mathbf{c}_j \\ \mathbf{r}_i^\top & 0 & p(i, j) \\ \mathbf{r}_j^\top & p(j, i) & 0 \end{pmatrix}$ and $\mathbf{v} = (\hat{\mathbf{v}} \ 0 \ 1)^\top$
 - Then $\hat{\mathbf{v}} = \hat{\mathbf{P}}\hat{\mathbf{v}} + \mathbf{r}_j \Rightarrow \hat{\mathbf{v}} = (\mathbf{I} - \hat{\mathbf{P}})^{-1}\mathbf{r}_j$
- $\text{ep}(i \rightarrow j) = \sum_k p_{ik} \cdot v_k = \mathbf{r}_i^\top (\mathbf{I} - \hat{\mathbf{P}})^{-1} \mathbf{c}_j + p(i, j)$
- Computing all $\text{ep}(i \rightarrow j)$ requires $\Theta(n^2)$ matrix inversions

Theorem. *Let*

$$\mathbf{Q} = [q(i, j)] \triangleq (\mathbf{I} - c\mathbf{P})^{-1}.$$

$\forall i \neq j$, there is

$$\text{ep}(i \rightarrow j) = \frac{q(i, j)}{q(i, i)q(j, j) - q(i, j)q(j, i)}.$$

- Fast solution to all-pair proximities
 - Compute $\mathbf{Q} = (\mathbf{I} - c\mathbf{P})^{-1}$
 - For all pair of nodes i and j , compute $\text{Prox}(i, j) = \frac{q(i, j)}{q(i, i)q(j, j) - q(i, j)q(j, i)}$
- Time complexity $\Theta(1 \text{ matrix inversion}) + \Theta(n^2)$
- When only the proximity of one pair of nodes
 - Only need two columns of \mathbf{Q}
 - Taylor expansion
$$(\mathbf{I} - c\mathbf{P})^{-1} = \mathbf{I} + c\mathbf{P} + (c\mathbf{P})^2 + \dots$$
 - Computing i -th column of \mathbf{Q}

$$\mathbf{Q}\mathbf{e}_i = (\mathbf{I} - c\mathbf{P})^{-1}\mathbf{e}_i = \mathbf{e}_i + c\mathbf{P}\mathbf{e}_i + (c\mathbf{P})^2\mathbf{e}_i + \dots$$
- Time complexity $\Theta(t(n + m))$, where t is the number of iterations

Lemma. *The expected time r_i for a random walk starting at node i to return to i is the reciprocal of the stationary probability of i . That is*

$$r_i = \frac{1}{\pi_i}.$$

- Intuitively
 - A long walk always ends up in stationary distribution π
 - Suppose the walk length is T , then the expected number of times it visits i is $\pi_i \cdot T$
 - The average length between two visits is $\frac{T}{\pi_i \cdot T} = \frac{1}{\pi_i}$
 - A rigorous proof requires the Strong Law of Large Numbers

Theorem. *The probability that a random walk starting at node i visits j before returning to i , which equals $\text{ep}(i \rightarrow j)$, satisfies*

$$\text{ep}(i \rightarrow j)c(i, j) = \frac{1}{\pi_i},$$

where $c(i, j)$ is the commute time between i and j .

- Proof
 - Consider a random walk w starting at i , and random variables
 - * X = the first time w returns to i
 - * Y = the first time w returns to i after visiting j
 - By definition $E(X) = \frac{1}{\pi_i}$ and $E(Y) = c(i, j)$
 - Clearly $X \leq Y$, and $\Pr[X = Y] = p \triangleq \text{ep}(i \rightarrow j)$
 - * $E(Y - X) = p \cdot 0 + (1 - p) \cdot E(Y) = (1 - p)c(i, j)$
 - Also $E(Y - X) = E(Y) - E(X) = c(i, j) - \frac{1}{\pi_i}$
- $\text{ep}(i \rightarrow j) + \text{ep}(j \rightarrow i) = \frac{1}{c(i, j)} \left(\frac{1}{\pi_i} + \frac{1}{\pi_j} \right)$
- Recall that $h(i, j)$ is small whenever π_j is large, which is bad for personalization
 - Sarkar et al. (2008) alleviate this by restricting the length of random walk
 - Tong et al. (2007) alleviate this by reducing the dependence on stationary distribution