Lecture 13 Case study: Developing a Trading Strategy

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Outline

Information Ratio

- Case study
 - From two trading signals to full-fledged fund trading strategy

- Appendix: Multivariate Volatility Models
 - MGARCH
 - * CCC-GARCH
 - ⋆ DCC-GARCH

Information Ratio

The information ratio (IR) is a standard performance metric:

$$IR_i = \frac{E\left[R_i - R_{benchmark}\right]}{\sigma\left[R_i - R_{benchmark}\right]}.$$

Let $R_{i,t}$ be excess returns to fund i and $R_{benchmark} = \beta_i MKT_t$

• Estimate α_i and β_i from the usual regression

$$R_{i,t} = \alpha_i + \beta_i' F_t + \varepsilon_{i,t}$$
.

Then:

$$IR_{i} = \frac{E\left[R_{i,t} - \beta_{i}MKT_{t}\right]}{\sigma\left[R_{i,t} - \beta_{i}MKT_{t}\right]} = \frac{\alpha_{i}}{\sigma\left(\varepsilon_{i,t}\right)}$$

Information Ratio and maximal Sharpe Ratio

Assume the fund's information ratio is 0.3 and the market Sharpe ratio is 0.4

• What is the maximal Sharpe ratio one could achieve by combining the fund and the market?

$$\max SR = \sqrt{SR_{MKT} + IR_{Fund}} = \sqrt{0.3^2 + 0.4^2} = 0.5.$$

- Math is simplest case of MVE math from prior lecture note.
- Notice that if $IR \neq 0$ it is possible to increase Sharpe ratio
 - ► Typically, though, we can't short sell a fund...

This is why α is interesting

- It means you can improve your Sharpe ratio relative to your benchmark.
- This is valuable and can therefore justify higher fees

Mini-Case:

From trading signals to trading strategy

Initial idea

- You have an idea about how to choose stocks that outperform existing benchmark portfolios
- In particular, you use a combination of textual analysis, stock prices, and social media-based data to come up with two trading signals for each stock in your trading universe
 - A valuation signal (of fundemental value): val_{i,t}
 - **2** A sentiment signal (of shorter-term trend): $trend_{i,t}$
- Our task is to:
 - See if the signals have any information
 - If so, implement an efficient portfolio strategy trading based on these signals
- Caveat: For this illustrative exercise, we only do in-sample testing

Unconditional returns to simple trading strategies

- A natural starting point is to sort into portfolios based on the signals
- One could sort into decile portfolios for each signal (20 portfolios in total), for instance, and look at average return for each decile portfolio
 - ▶ Valid approach
 - Can see if there is significant spread in average returns to portfolio 10 vs portfolio 1 for each signal
 - Can also spot nonlinear effects (perhaps U-shaped average return pattern across portfolios)
 - However, cannot tell from this if the signals have marginal value relative to each other and other known signals

Unconditional returns to simple trading strategies

- Regression-based portfolio sorts: Fama-MacBeth regressions
- For each time t, run:

$$R_{i,t}^e = \lambda_{0,t} + \lambda_{1,t} \textit{val}_{i,t-1} + \lambda_{2,t} \textit{trend}_{i,t-1} + \lambda_{3,t}' Z_{i,t-1} + \epsilon_{i,t}$$

where $Z_{i,t-1}$ are firm signals such as industry-dummies, market beta, etc.

- We know that $\lambda_{1,t}$ and $\lambda_{2,t}$ are the time t returns to trading strategies using the two new signals with regression-implied portfolio weights known at time t-1.
- Test if λ_1 and λ_2 (average returns to the new trading strategies) are significantly different from zero
- This is similar to how Fama and French constructs alternative factors, such as HML and Size.

Unconditional returns to simple trading strategies

- We know have time-series excess returns to two trading strategies $\{\lambda_{1,t}\}_{t=1}^T$ and $\{\lambda_{2,t}\}_{t=1}^T$
- ullet To continue, we will from here on assume these happen to correspond to HML_t and MOM_t
 - ▶ The value and momentum factors from Kenneth French's webpage
 - ► Recall:

	HML	MOM
$\mathbb{E} r_{t+1}^e$	0.039	0.064
σ	0.100	0.156
SR	0.383	0.411
ρ	-0.171	

Constructing a simple joint trading strategy

• Use the unconditional moments to form the unconditional mean-variance efficient portfolio based on these two sub-portfolios

$$w = k \begin{bmatrix} 0.1^2 & -0.171 \times 0.1 \times 0.156 \\ -0.171 \times 0.1 \times 0.156 & 0.156^2 \end{bmatrix}^{-1} \begin{bmatrix} 0.039 \\ 0.064 \end{bmatrix}$$
$$= k \begin{bmatrix} 4.74 \\ 3.14 \end{bmatrix}$$

Find k be setting portfolio variance to 0.15^2 :

$$0.15^{2} = k^{2} \begin{bmatrix} 4.74 \\ 3.14 \end{bmatrix}' \begin{bmatrix} 0.1^{2} & -0.171 \times 0.0156 \\ -0.171 \times 0.0156 & 0.156^{2} \end{bmatrix} \begin{bmatrix} 4.74 \\ 3.14 \end{bmatrix}$$
$$= k^{2} \times 0.385$$

Constructing a simple joint trading strategy (cont'd)

• Solving for k, we have:

$$k = \sqrt{\frac{0.15^2}{0.385}} = 0.24$$

and so the per-period portfolio weights are:

$$w = 0.24 \times \left[\begin{array}{c} 4.74 \\ 3.14 \end{array} \right] = \left[\begin{array}{c} 1.14 \\ 0.75 \end{array} \right]$$

- So, for a \$10M portfolio, put \$11.4M in HML, \$7.5M in MOM
 - ▶ If HML and MOM are long-short, zero-investment portfolios, this means long \$11.4M in value financed by shorting the same amount in growth, and long \$7.5M in winners, financed by the same amount short in losers. Finally, put \$10M in risk-free rate.
 - Or, if HML and MOM are long-only versions, you need to borrow \$8.9M in the risk-free rate to finance these positions.
 - Maintain these portfolio weights by rebalancing each month

Constructing a simple joint trading strategy (cont'd)

 The expected return and Sharpe ratio of this baseline trading strategy are then:

$$E\left[R^{MVE}\right] = \begin{bmatrix} 1.14 & 0.75 \end{bmatrix} \begin{bmatrix} 0.039 \\ 0.064 \end{bmatrix} = 9.25\%.$$
 $SR\left(R^{MVE}\right) = \frac{9.25\%}{15\%} = 0.62$

- Let's see if we can improve on this baseline by attempting to estimate
 - The conditional expected returns
 - ► The conditional covariance matrix

Estimating conditional expected returns

- Our tool for this is forecasting regressions.
 - ► Need predictive variables
 - Thus, we need extra instruments/signals to implement conditional strategies
 - We will keep it simple and estimate a VAR(1) on HML and MOM returns:

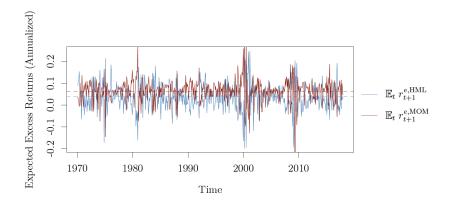
$$\left[\begin{array}{c} \textit{HML}_t \\ \textit{MOM}_t \end{array}\right] = \left[\begin{array}{c} \phi_{01} \\ \phi_{02} \end{array}\right] + \left[\begin{array}{c} \phi_{11} & \phi_{12} \\ \phi_{21} & \phi_{22} \end{array}\right] \left[\begin{array}{c} \textit{HML}_{t-1} \\ \textit{MOM}_{t-1} \end{array}\right] + \left[\begin{array}{c} \epsilon_{\textit{HML},t} \\ \epsilon_{\textit{MOM},t} \end{array}\right]$$

• Implementing this in R (see CCLE), we get:

$$\left[\begin{array}{c} \hat{\phi}_{01} \\ \hat{\phi}_{02} \end{array} \right] = \left[\begin{array}{c} 0.30 \\ 0.63 \end{array} \right], \ \left[\begin{array}{cc} \hat{\phi}_{11} & \hat{\phi}_{12} \\ \hat{\phi}_{21} & \hat{\phi}_{22} \end{array} \right] = \left[\begin{array}{cc} 0.17 & -0.01 \\ -0.08 & 0.06 \end{array} \right], \ \begin{array}{c} R_{HML}^2 = 2.9\% \\ R_{MOM}^2 = 0.7\% \end{array}$$

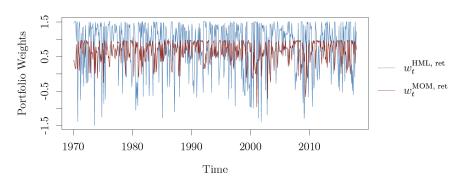
Expected Returns from the VAR(1)

- The time-series of annualized expected monthly returns is below
- Lots of variation, economically, even though R^2 s were small



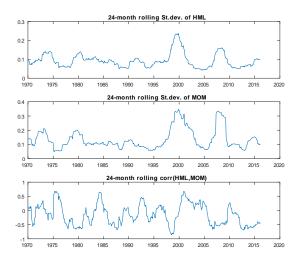
MVE Portfolio Weights

- Using the expected returns from the VAR, along with a constant covariance matrix of the residuals, we obtain portfolio weights
 - ▶ As before, set volatility of portfolio to be 15% annualized
 - HML was most predictable and therefore sees biggest changes in portfolio weights
 - ► Notice economically large changes in weights, despite modest R²s in forecasting regressions
 - Full sample Sharpe ratio for this strategy is 0.78 > 0.61 (from no expected return timing)



Adding Volatility Timing

 Recall that a cursory look at the data indicates substantial time-variation in the conditional covariance matrix



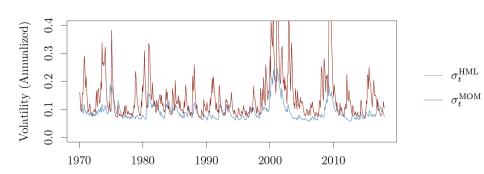
Estimating a DCC-GARCH(1,1)

• We will use the ccgarch-package in R

```
#Run GARCH on residuals
a <- c(0.003, 0.005)
A \leftarrow diag(c(0.2,0.3))
B < -diag(c(0.75, 0.6))
dcc.para <- c(0.01.0.98)
dcc.GARCH = dcc.estimation(inia = a, iniA = A, iniB = B, ini.dcc = dcc.para,
        dvar = data.dt[complete.cases(data.dt), .(eps.log.HML, eps.log.MOM)].
        model = "extended"
data.dt[!is.na(shift(log.HML, 1)), `:=`(sigma.HML = sqrt(dcc.GARCH$h[, 1]),
                      sigma.MOM = sqrt(dcc.GARCH$h[, 2]),
                      rho = dcc.GARCH$DCC[, 2],
                      eta.HML = dcc.GARCHSstd.resid[. 1].
                      eta.MOM = dcc.GARCHSstd.resid[, 2]
                      )]
```

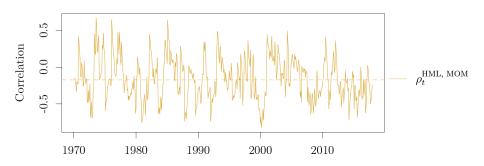
Estimated conditional volatilities

- Lots of time-variation, persistent processes
- Momentum has most time-variation in vol (least in expected return)



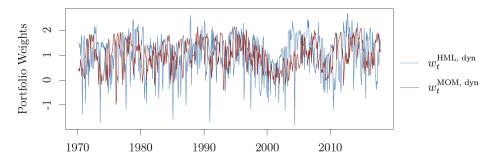
Estimated conditional correlation

- Conditional correlation between HML and MOM
- LOTS of time-variation here (-0.7 to 0.7), likely large effects on MVE weights



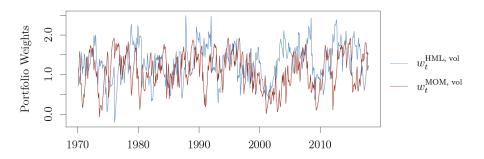
MVE weights from vol and return timing

- More variation than when only return timing
- The vol component is also more persistent (though this is likely a feature of the predictors for returns only being lagged return)



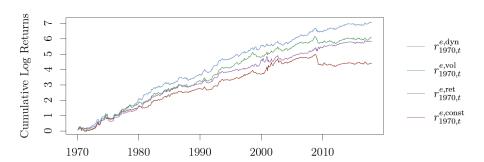
Additional strategy: Vol timing only

- Also run case where no attempt at forecasting return, only the conditional covariance matrix
- I.e., only vol timing. Less frequent trading, no shorting (pretty much).



Cumulative Log Returns for four strategies

- Strategies are: No timing, Er timing, Vol timing, both Er and Vol timing (dynamic)
- Ranking is: Both Er and Vol, only Vol, only Er, no timing
- Caveat: recall, these are in-sample tests



Summary statistics for four strategies

- Timing of return and variance-covariance matrix adds substantially to Sharpe ratios
 - Marginal SR increase: $IR = \sqrt{0.943^2 0.613^2} = 0.717$
 - ► Caveat: all in-sample and ignoring transaction costs
 - ▶ Trading is hazardous to your wealth

	Dynamic	Volatility-Timing	Return-Timing	Constant
$\mathbb{E} r_{t+1}^e$	0.148	0.127	0.122	0.092
σ	0.157	0.154	0.156	0.150
SR	0.943	0.824	0.781	0.613

	Correlation Matrix			
	Dynamic	Volatility-Timing	Return-Timing	Constant
Dyn.	1.000	0.816	0.873	0.667
Vol.		1.000	0.625	0.855
Ret.			1.000	0.592
Const.				1.000

On timing strategies

- In terms of MVE portfolio weights, the ones we calculated move too much
 - A lot of trading comes with a lot of trading costs (hazardous to your wealth)

- In practice, out-of-sample metrics lead to "shrinkage" of weights so they move less
 - Likely benefit is smaller than what we found in-sample
 - Market-timing is not easy
 - Neverthess, this should still be a part of your portfolio optimization

Appendix:

Multivariate Volatility Modeling

The Conditional Covariance Matrix

• Consider a vector of asset returns:

$$\mathbf{r}_t = \boldsymbol{\mu}_t + \boldsymbol{\varepsilon}_t$$

where all variables are $N \times 1$ vectors

• The conditional means (modelled from, e.g., a VAR) are:

$$\mu_t = E_{t-1} \left[\mathbf{r}_t \right]$$

• Denote the conditional $N \times N$ covariance matrix as:

$$\mathbf{H}_t = E_{t-1} \left[\mathbf{\varepsilon}_t \mathbf{\varepsilon}_t' \right]$$

• A mean-variance optimizer would choose $N \times 1$ portfolio weights:

$$\mathbf{w}_{t-1} = k \mathbf{H}_t^{-1} \boldsymbol{\mu}_t$$

The Conditional Covariance Matrix (cont'd)

ullet Define the mean-zero, unit variance N imes 1 vector of i.i.d. shocks

$$\eta_{t} \sim WN(\mathbf{0}_{N}, I_{N})$$
.

This distribution is typically chosen to be Normal or Student's t

• We can then write:

$$oldsymbol{arepsilon}_t = \mathbf{H}_t^{1/2} oldsymbol{\eta}_t$$

as

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MGARCH

• Multivariate GARCH-type models take as their starting point

$$\mathbf{r}_t = \boldsymbol{\mu}_t + \mathbf{H}_t^{1/2} \boldsymbol{\eta}_t$$

where $\mathbf{H}_t^{1/2}$ is the Cholesky decomposition of \mathbf{H}_t

- ullet These models specify the dynamics of $oldsymbol{\mathsf{H}}_t$
 - ▶ Needs to be positive definite for each t
 - ▶ Note: size of this matrix is of order N^2 , so need to keep N small for good performance

Useful decomposition

We can write:

$$\mathbf{H}_t = \mathbf{D}_t \mathbf{R}_t \mathbf{D}_t$$

where the diagonal "standard deviations"-matrix is

$$\mathbf{D}_t = \left[egin{array}{cccc} \sqrt{\sigma_{1,t}^2} & 0 & \cdots & 0 \\ 0 & \sqrt{\sigma_{2,t}^2} & \cdots & 0 \\ dots & dots & \ddots & dots \\ 0 & 0 & \cdots & \sqrt{\sigma_{N,t}^2} \end{array}
ight]$$

and the symmetric correlation matrix is

$$\mathbf{R}_t = \left[egin{array}{cccc} 1 &
ho_{12,t} & \cdots &
ho_{1N,t} \
ho_{21,t} & 1 & \cdots &
ho_{2N,t} \ dots & dots & \ddots & dots \
ho_{N1,t} &
ho_{N2,t} & \cdots & 1 \end{array}
ight]$$

Are individual portfolio StDevs and Corrs time-varying?

Consider monthly data from 1970 to 2017 for the Fama-French HML (value) and MOM (momentum) factors

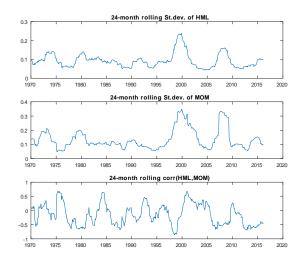
Annualized summary statistics

	HML	MOM
$\mathbb{E} r_{t+1}^e$	0.039	0.064
σ	0.100	0.156
SR	0.383	0.411
ρ	-0.171	

Are individual portfolio StDevs and Corrs time-varying?

Plot 24-month rolling (annualized) st.devs. and correlation

- Yes! These moments look strongly time-varying
- MVE portfolio not likely to have constant portfolio weights



MGARCH - Specifications

- The diagonal elements of \mathbf{D}_t can be obtained from N univariate GARCH models run on each element in \mathbf{r}_t
- ullet The correlation matrix is tricker, in part as each correlation must be between -1 and 1 and the diagonal has to equal 1
- Two standard models:
 - CCC-GARCH
 - ★ Constant conditional correlations
 - OCC-GARCH
 - ★ Dynamic conditional correlations

CCC-GARCH

Define:

$$v_t = \mathbf{D}_t^{-1} \varepsilon_t$$
.

which implies that

$$v_t \sim WN(0, \mathbf{R}_t)$$
.

CCC-GARCH simply assumes that the correlation matrix is constant, thus:

$$\mathbf{R}_t = \mathbf{R} = \frac{1}{T} \sum_{t=1}^{T} \mathbf{v}_t \mathbf{v}_t'.$$

where each diagonal element of \mathbf{D}_t is obtained from N univariate GARCH processes

DCC-GARCH(1,1)

- This model is due to Engle and Sheppard (2001), it is a benchmark model of time-varying conditional covariance matrix estimation
 - ▶ We still have problem if *N* is not small
- Main idea, put structure (AR(1)-like structure) on how conditional correlations move over time
 - ightharpoonup Ensure well-behaved so positive definite conditional correlation matrix and each element between -1 and 1
- Decompose correlation matrix to achieve this:

$$\begin{array}{rcl} \mathbf{R}_t & = & \mathbf{Q}_t^{*-1} \mathbf{Q}_t \mathbf{Q}_t^{*-1}, \\ \mathbf{Q}_t & = & (1-a-b) \, \overline{\mathbf{Q}} + a \boldsymbol{\nu}_{t-1} \boldsymbol{\nu}_{t-1}' + b \mathbf{Q}_{t-1} \\ \overline{\mathbf{Q}} & = & \frac{1}{T} \sum_{t=1}^T \boldsymbol{\nu}_t \boldsymbol{\nu}_t' \\ a & > & 0, \quad b > 0, \quad a+b < 1 \end{array}$$

DCC-GARCH (cont'd)

• Denote the ij'th element in \mathbf{Q}_t as $q_{ij,t}$. Then:

$$\mathbf{Q}_t^* = \left[egin{array}{cccc} \sqrt{q_{11,t}} & 0 & \cdots & 0 \ 0 & \sqrt{q_{22,t}} & \cdots & 0 \ dots & dots & \ddots & dots \ 0 & 0 & \cdots & \sqrt{q_{NN,t}} \end{array}
ight]$$

- The likelihood function depends on the chosen distribution for the shocks (e.g., Normal), but is otherwise found in a way similar to the ARMA likelihood
 - Assume initial shocks equal unconditional average (in particular, need \mathbf{Q}_0 to be positive definite)
 - Use multivariate probability density function
 - This is all coded up in R/MatLab, so not covered here, though we will implement this in mini-case to follow