Aggregation of Information about the Cross Section of Stock Returns: a Latent Variable Approach

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Outline

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Introduction

Overview

- Light et al. (2017) propose a new approach for predicting stock returns on individual stocks from a large number firm characteristics.
- ▶ The expected returns are treated as latent variables, and the firm characteristics are assumed to link to expected returns through one or more common latent factors.
- ► They use partial least squares (PLS) estimator to estimate expected returns and considered 26 firm characteristics.
- ► They argue that their PLS approach outperforms alternative techniques (Factor Analysis, PCA, Fama-MacBeth Regression, Rank-based approach).

Introduction

List of 26 Firm Characteristics (4+7+9+2+2+2=26!)

Value Related	Return	Growth	Profitability	Distress	Others (2)
(4)	Related (7)	Related (9)	(2)	Related (2)	
	omarket beta (B) omoretum (MOM) olong-term reversal (LTR) oshort-term reversal (STR) oidiosyncratic volatility (idVOL) oMax daily return over past month (MAX) observed dicosyncratic skewness (EIS)	 ♦ total asset growth (AG) ♦ abnormal capital investments (CI) ♦ investment growth (IG) ♦ investment-to-capital ratio (I/K) ♦ investment-to-assets ratio (I/A) ♦ accurals (ACC) ♦ net operating assets (NOA) ♦ net stock issues (NS) ♦ composite stock issuance (L) 	oreturns on equity (ROE) oreturns on assets (ROA)	◇leverage (LV) ◇O-score (O)	oturnover (TO) oanalysts' forecasts dispersion (D)

Baseline Specification: Single Latent Variable

Consider a set of N stocks whose characteristics are observed in at least two periods. The best predictor of returns on stock i at time t is expected return $\mu_{it} = E\left[R_{it+1}|\mathcal{F}_t\right]$ where \mathcal{F}_t denotes public information.

The realized return on stock i can be written as

$$R_{it+1} = \underbrace{\mu_{it}}_{ ext{unobserved factor in the firm characteristic space}} + \varepsilon_{it+1}$$

where $E\left[\varepsilon_{it+1}|\mathcal{F}_{t}\right]=0$ and unexpected returns ε_{it+1} are assumed to be orthogonal to \mathcal{F}_{t} . ε_{it+1} and ε_{jt+1} can be correlated.

Baseline Specification: Single Latent Variable

Econometricians lack some information contained in \mathcal{F}_t and do not know μ_{it} . Instead they observe firm characteristics $X_{it}^a, a=1,\ldots,A$ such that $\{X_{it-s}, s\geq 0, a=1,\ldots,A, i=1,\ldots,N\}\subseteq \mathcal{F}_t$. Assume at each time period X_{it}^a have zero cross-sectional means and unit variances (demean, standardize). Assume

$$X_{it}^{a} = \delta_{t}^{a} \left(\mu_{it} - \bar{\mu}_{t} \right) + u_{it}^{a}$$

where δ^a_t measures the sensitivity of characteristics a to expected returns, and $\bar{\mu}_t$ is the cross-sectional average of expected returns at time t. Denote the sample cross-sectional variance and covariance as \overline{Var} and \overline{Cov} respectively, and reserve the sample cross-sectional variance and covariance in the characteristic space as \overline{Var} and \overline{Cov} .

Identifying Assumptions

Assumption 1 (Distribution of expected returns) In each period t,

$$\overline{\mu}_t = \frac{1}{N} \sum_{i=1}^N \mu_{it} \xrightarrow{p} \mu_t \text{ and } \overline{Var} \left(\mu_{it} \right) = \frac{1}{N} \sum_{i=1}^N \left(\mu_{it} - \overline{\mu}_t \right)^2 \xrightarrow{p} V_t \text{ as } N \to \infty \text{ where } V_t > 0.$$

Assumption 2 (Distribution of characteristic loadings) In each period t,

$$\widetilde{\delta}_t = \frac{1}{A} \sum_{a=1}^A \delta_t^a \overset{p}{\to} \delta_t \text{ and } \widetilde{Var} \left(\delta_t^a \right) = \frac{1}{A} \sum_{a=1}^A \left(\delta_t^a - \widetilde{\delta}_t \right)^2 \overset{p}{\to} \Lambda_{t,t} \text{ as } A \to \infty \text{ where } \Lambda_{t,t} > 0.$$

Also, for consecutive periods t-1 and t

$$\widetilde{\textit{Cov}}\left(\delta_{t-1}^{\textit{a}}, \delta_{t}^{\textit{a}}\right) = \frac{1}{\textit{A}} \sum_{t=1}^{\textit{A}} \left(\delta_{t-1}^{\textit{a}} - \widetilde{\delta}_{t-1}\right) \left(\delta_{t}^{\textit{a}} - \widetilde{\delta}_{t}\right) \overset{\textit{p}}{\rightarrow} \Lambda_{t-1,t} \text{ as } \textit{A} \rightarrow \infty, \text{ where } \Lambda_{t-1,t} > 0.$$

Identifying Assumptions

Assumption 3 (Orthogonality of errors and expected returns) In each period t and for each characteristic $a,\ a=1,\ldots,A,$

$$\overline{\textit{Cov}}\left(\mu_{it}, \textit{u}_{it}^{\textit{a}}\right) = \frac{1}{\textit{N}} \sum_{i=1}^{\textit{N}} \left(\mu_{it} - \overline{\mu}_{t}\right) \left(\textit{u}_{it}^{\textit{a}} - \overline{\textit{u}}_{t}^{\textit{a}}\right) \overset{\textit{p}}{\rightarrow} 0 \text{ as } \textit{N} \rightarrow \infty.$$

Assumption 4 (Orthogonality of errors and characteristic loadings) In each period t and for each stock i, i = 1, ..., N,

$$\widetilde{\textit{Cov}}\left(\delta_{t-1}^{\textit{a}},\textit{u}_{\textit{it}}^{\textit{a}}\right) = \frac{1}{\textit{A}} \sum_{\textit{a}=1}^{\textit{A}} \left(\delta_{t-1}^{\textit{a}} - \tilde{\delta}_{t-1}\right) \left(\textit{u}_{\textit{it}}^{\textit{a}} - \tilde{\textit{u}}_{\textit{it}}\right) \overset{\textit{p}}{\rightarrow} 0 \text{ as } \textit{A} \rightarrow \infty.$$

Estimation Procedure: Motivation

- ▶ The main problem of an econometrician is to estimate μ_{it} by filtering them from the observable characteristics X_{it}^a .
- Light et al. (2017) use a procedure that belongs to a family of partial least squares (PLS) algorithms.
- The PLS framework is actively used in computational chemistry and behavioral sciences, and recently in applications in finance:
 - Kelly and Pruitt (2013): extract market expectations from a cross-section of variation ratios;
 - Huang et al. (2015): construct an investor sentiment index from several sentiment proxies;
 - Giglio et al. (2016): extend PLS to quantile regressions for aggregation of 19 measures of systemic risk.

Estimation Procedure: Implementation

The PLS-based estimation of baseline model at time t can be implemented in two steps:

- ▶ **Step 1**: Run separate cross-sectional regressions of R_{it} , $i=1,\ldots,N$, on each individual firm characteristic X_{it-1}^a , $i=1,\ldots,N$ for $a=1,\ldots,A$ and denote the obtained slopes a λ_t^a .
- ▶ **Step 2**: For each firm i, i = 1, ..., N run a regression of X_{it}^a on λ_t^a , a = 1, ..., A, and denote the obtained slopes as $\hat{\mu}_{it}$.

Extension: Multiple Latent Variables - Model

Instead of a single latent factor, the firms characteristics are linked to expected returns through L common latent factors $\beta_{lit}, l=1,\ldots,L$. The expected return on stock i at time t is related to the latent factor as

$$\mu_{it} = \sum_{l=1}^{L} \beta_{lit} \gamma_l$$

where $\gamma_I, I=1,\ldots,L$ are loadings on various factors. Introducing a $1\times L$ matrix of factor values B_{it} and an $L\times 1$ matrix of loadings Γ , the realized return on stock i at time t+1 can be written as

$$R_{it+1} = B_{it}\Gamma + \varepsilon_{it+1}$$
.

Unexpected returns ε_{it+1} are independent from all variables available at time t, $E\left[\varepsilon_{it}|\mathcal{F}_{t}\right]=0$ and matrices B_{it} and Γ are unobservable to econometricians.



Extension: Multiple Latent Variables - Model

The firm characteristics are aligned with stock returns because they are proxies for the unobservable latent factors: the demeanded and standardized characteristics X_{it}^a , $a=1,\ldots,A$ of firm i at time t are related to B_{it} as

$$X_{it}^{a} = \left(B_{it} - \overline{B}_{t}\right)\delta^{a} + u_{it}^{a}$$

where Δ^a , $a=1,\ldots,A$ are $L\times 1$ matrices that determine the sensitivity of characteristics to the factors, \overline{B}_t is the cross-sectional average of the factors, and u^a_{it} are components of characteristics that are unrelated to stock returns and that can be arbitrary correlated among themselves.

Methodology Extension: Multiple Latent Variables Identifying assumptions

Assumption 1' (Distribution of latent factors) In each period t, $\overline{Var}(B_{it}) \stackrel{p}{\to} V_t$ as $N \to \infty$ and the matrix V_t is positive definite.

Assumption 2' (Distribution of characteristic loadings) $\widetilde{Var}(\Delta^a) \overset{p}{\to} \Lambda$ as $A \to \infty$ and the matrix Λ is positive definite.

Assumption 3' (Orthogonality of errors and latent factors) In each period t and for each characteristic $a, a = 1, \ldots, A, \overline{Cov}\left(B_{it}, u_{it}^a\right) \stackrel{p}{\to} 0$ as $N \to \infty$.

Assumption 4' (Orthogonality of errors and characteristic loadings) In each period t and for each stock i, $i=1,\ldots,N$, $\widetilde{Cov}\left(\Delta^a,u_{it}^a\right)\overset{p}{\to}0$ as $A\to\infty$.

Methodology Extension: Multiple Latent Variables

Estimation Procedure

The three-pass regression filter (3PRF) procedure is adopted from Kelly and Pruitt (2015):

- ▶ **Step 1:** Run cross-sectional regressions of the characteristics X_{it-1}^a , $a=1,\ldots,A$ on the proxy variables Z_{it}^l , $l=1,\ldots,L$ and retain the $L\times 1$ matrices of slopes λ_t^a .
- ▶ **Step 2:** For each stock i, i = 1, ..., N, run regressions of X_{it-1}^a and X_{it}^a on λ_t^a in the characteristics space and denote the obtained matrices of slopes are \hat{B}_{it-1} and \hat{B}_{it} .
- ▶ Step 3: Run a cross-sectional regression of R_{it} on $\hat{B}_{it-1}, i=1,\ldots,N$ and denote the obtained matrix of slopes as $\hat{\Gamma}_t$. Compute an estimate of expected returns as $\hat{\mu}_{it} = \hat{B}_{it}\hat{\Gamma}_t$.

- standard sources: CRSP monthly files, Compustat
 Fundamentals annuals files
- exchanges: NYSE, AMEX, NASDAX common stocks
- exclude financial firms, adjust for delisting, exclude stocks below \$5 at the end of the previous month
- use accounting data in calendar year t from financial statements with the fiscal year end in year t-1
- construct 26 firm characteristic variables

Empirical Analysis

Individual characteristics: individual characteristic are informative about future stock returns.

Table 1 Individual characteristics

Panel.	A: equa	ıl-weigh	ted port	folios									
	S	\mathbf{B}/\mathbf{M}	\mathbf{E}/\mathbf{P}	C/P	В	MOM	LTR	STR	IdVol	MAX	EIS	ROE	ROA
Means	-0.16	0.98	0.84	0.95	-0.46	1.59	-0.55	-1.04	-1.02	-0.96	-0.65	-0.07	1.54
Stds	3.71	4.55	3.75	4.26	7.33	5.86	3.59	4.97	6.43	5.97	3.47	2.16	4.30
t-stats	-1.01	4.88	5.11	5.08	-1.43	6.15	-3.50	-4.76	-3.59	-3.65	-3.26	-0.69	7.63
	AG	CI	IG	I/K	I/A	ACC	NOA	NS	ι	LV	О	то	D
Means	-0.73	-0.25	-0.43	-0.55	-0.68	-0.52	-0.60	-0.34	-0.70	0.70	-0.64	-0.56	-0.72
Stds	2.86	1.60	1.98	4.59	2.53	2.32	3.62	2.30	3.62	5.17	3.11	4.79	3.92
t-stats	-5.82	-3.54	-4.95	-2.75	-6.08	-5.15	-3.79	-3.40	-4.40	3.09	-4.33	-2.65	-3.83
Panel	B: valu	e-weigh	ted port	folios									
	S	B/M	E/P	C/P	В	MOM	LTR	STR	IdVol	MAX	EIS	ROE	ROA
Means	-0.17	0.58	0.62	0.49	-0.19	1.34	-0.26	-0.20	-0.99	-0.62	-0.73	-0.07	0.90
Stds	4.02	4.99	4.97	5.08	7.46	7.34	5.07	6.19	7.18	7.13	5.19	3.53	5.29
t-stats	-0.94	2.66	2.83	2.20	-0.57	4.16	-1.17	-0.72	-3.12	-1.98	-2.43	-0.44	3.62
	AG	CI	IG	I/K	I/A	ACC	NOA	NS	ι	LV	О	то	D
Means	-0.37	-0.32	-0.29	-0.35	-0.49	-0.38	-0.28	-0.57	-0.59	0.51	-0.12	-0.27	-0.40
Stds	3.70	3.27	3.43	6.43	3.15	3.95	4.04	2.97	4.19	5.74	4.22	5.15	5.15
t-stats	-2.28	-2.23	-1.93	-1.25	-3.55	-2.17	-1.57	-4.35	-3.19	2.00	-0.58	-1.19	-1.61

This table reports time-series averages, standard deviations, and r-statistics of hedge returns defined as a difference in monthly returns on top and bottom decile portfolios formed individually by 26 firm characteristics. The decile portfolios are equal-weighted in panel A and value-weighted in panel B. All variables are named in Section 2.1 | and their construction is described in Appendix B. Average returns and their standard deviations are reported in percentage points. The sample is from 1970 to 2012 for all characteristics except for EIS, ROA, O, and D, for which it starts in 1987, 1975, 1976, and 1977, respectively.

Empirical Analysis

Filtered expected returns (AFER): high expected returns are generated by AFER without a commensurate increase in the volatility of returns.

Table 2
Returns on decile AFER portfolios

	EW portfolios									VW portfolios												
	1	2	3	4	5	6	7	8	9	10	(10-1)	1	2	3	4	5	6	7	8	9	10	(10-1)
Means	-0.10	0.43	0.70	1.00	1.06	1.33	1.38	1.46	1.57	1.59	1.69	0.10	0.62	0.74	0.78	0.90	1.06	0.97	1.10	1.15	1.31	1.21
Stds	8.02	7.09	6.45	5.98	5.66	5.49	5.48	5.63	5.91	6.82	7.05	8.00	6.93	6.15	5.88	5.46	5.27	5.22	5.50	5.82	6.76	7.60
t-stats	-0.28	1.38	2.45	3.79	4.25	5.48	5.71	5.89	6.04	5.30	5.45	0.28	2.04	2.72	3.00	3.73	4.56	4.23	4.54	4.49	4.39	3.61
Panel B	: averagii	ng of λ_t^a	over pa	st 5 yea	rs																	
	EW portfolios												VW	portfo	lios							
	1	2	3	4	5	6	7	8	9	10	(10-1)	1	2	3	4	5	6	7	8	9	10	(10-1)
Means	-0.30	0.49	0.82	1.07	1.15	1.32	1.39	1.41	1.48	1.58	1.88	-0.27	0.34	0.66	0.74	0.96	1.05	1.00	1.16	1.23	1.14	1.41
Stds	7.85	7.11	6.55	6.09	5.78	5.54	5.32	5.13	5.16	5.88	4.68	8.60	7.46	6.76	6.03	5.69	5.20	4.90	4.81	4.72	5.27	6.53
t-stats	-0.87	1.55	2.84	4.00	4.51	5.41	5.95	6.24	6.50	6.10	9.12	-0.72	1.05	2.21	2.79	3.83	4.57	4.65	5.47	5.93	4.91	4.91
Panel C	: averagii	ng of λ_t^a	over pa	st 10 ye	ars																	
		EW portfolios													VW	portfo	lios					
	1	2	3	4	5	6	7	8	9	10	(10-1)	1	2	3	4	5	6	7	8	9	10	(10-1)
Means	-0.35	0.48	0.84	1.09	1.18	1.34	1.35	1.43	1.46	1.59	1.94	-0.17	0.24	0.49	0.79	0.99	0.94	1.02	1.03	1.06	1.20	1.37
Stds	8.19	7.35	6.72	6.20	5.83	5.45	5.13	4.90	4.81	5.58	4.49	8.84	7.56	6.69	6.14	5.49	5.10	4.56	4.31	4.17	4.79	6.51
t-stats	-0.97	1.47	2.84	3.99	4.61	5.58	5.95	6.62	6.89	6.49	9.82	-0.44	0.73	1.66	2.92	4.08	4.17	5.09	5.41	5.77	5.68	4.78
Panel D	: averagii	ng of λ_t^a	over al	l previo	us mon	ths																
	EW portfolios							VW portfolios														
	1	2	3	4	5	6	7	8	9	10	(10-1)	1	2	3	4	5	6	7	8	9	10	(10-1)
Means	-0.38	0.43	0.81	1.04	1.16	1.33	1.39	1.42	1.53	1.68	2.06	-0.28	0.27	0.55	0.85	0.95	1.00	1.06	1.14	1.10	1.28	1.56
Stds	8.49	7.60	6.91	6.23	5.84	5.42	5.09	4.81	4.64	5.16	5.02	8.93	7.58	6.58	5.85	5.36	4.93	4.47	4.14	4.20	4.94	6.52
t-stats	-1.02	1.29	2.65	3.81	4.50	5.58	6.21	6.69	7.48	7.38	9.31	-0.72	0.80	1.89	3.29	4.01	4.62	5.39	6.25	5.96	5.89	5.45

This table shows time-series averages, standard deviations, and -statistics of monthly equal-weighted (EW) and value-weighted (VW) stock teurns on decile portfolios formed by sorting firms on the aggregate filtered expected returns AFER. The columns (10–1) report the statistics for the difference in returns on the top and bottom portfolios. The variable AFER is constructed without time-series averaging of λ_i^{μ} over the most recent 5 years (in panel B), with averaging of λ_i^{μ} over the most recent 10 years (in panel C), and with averaging of λ_i^{μ} over all previous months (in panel D). The sample is from January 1970 to December 2012. The portfolios are rebalanced monthly, All returns and standard deviations are reported in percentage points.

Empirical Analysis

Filtered expected returns in subsamples

Table 5 Filtered expected returns in subsamples

		Sample periods												
	He	dge retu												
	(1)	(2)	(3)	(1)	(2)	(3)								
EW	2.06	2.15	1.92	9.31	10.12	4.21								
VW	1.56	1.55	1.58	5.45	5.11	2.83								

	Size portfolios							Idiosyncratic volatility portfolios								
	Hedge returns			t-stats			Не	dge retu	ırns	t-stats						
	(1)	(2)	(3)	(1)	(2)	(3)	(1)	(2)	(3)	(1)	(2)	(3)				
EW	2.52	1.75	1.15	11.31	6.81	4.56	0.95	1.94	2.66	6.30	10.37	11.68				
VW	2.45	1.68	0.93	10.11	6.50	3.53	1.03	1.68	1.94	4.57	6.12	5.28				

This table reports the average differences in monthly equal-weighted (EW) and value-weighted (VW) returns on top and bottom decile AFER portfolios and their t-statistics in various subsamples. To denote subsamples we use the following notation. Sample periods: (1) full sample (January 1970–December 2012), (2) early sample (January 1970–December 1995), (3) late sample (January 1996–December 2012); size portfolios: (1) microcap, (2) small, (3) large; idiosyneratic volatility portfolios: (1) low volatility, (2) medium volatility, (3) high volatility. In all panels AFER is constructed by averaging λ_T^p from all previous months. All returns are reported in percentage points.

Alternative aggregation techniques

Advantage of PLS-based regression

- ► Factor analysis vs PLS: common factor in the characteristics produces much smaller hedge returns than *AFER* does.
- PCA vs PLS: the first principle component is a poor predictor of future returns, AFER statistically produce higher hedge returns.
- Fama-MacBeth vs PLS: regression based estimators of expected returns are likely to be less efficient; they are imprecise and even nonexistent when the number of characteristics is large; potential multicollinearity problems when characteristics are highly correlated.
- Rank-based approach vs PLS: theoretical and empirical advantage.
- ► Large cap stocks: the superior performance of the PLS-based aggregation technique is robust in the subsample of liquid large cap stocks.

Conclusion

Light et al. (2017) They find that the hedge returns produced by expected returns filtered from 26 characteristics are large, highly statistically significant, and exceed the hedge returns produced by each characteristic individually, which suggests:

- 1. there is commonality in the considered asset pricing anomalies
- 2. the relations between characteristics and returns are not a statistical fluke
- individual characteristics are likely to be diverse signals about expected returns

Reference

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