$$\int_{0}^{\infty} \int_{0}^{\infty} \int_{0$$

Let's step it up a notch. Let's allow for many different datasets boldD where X (n x p) matrix of feature values is constant but y1,...,yn are drawn from the rv's Y1,..., Yn meaning δ 1, ..., δ n draws from Δ 1, ..., Δ n. We still want to find the MSE just at x* Yn meaning $\delta 1, ..., \delta n$ are

=> Because the D's are random => g becomes random (g is fixed only with a fixed D).

fixed only with a fixed D).

$$E_{\chi}[17] = 17$$

$$M_{5}E(x_{c}) = E_{\Delta_{1}, \Delta_{2}, ..., \Delta_{n}, \Delta_{k}} \left[\left(\bigvee_{k} - g(\widehat{x}_{k}) \right)^{2} \middle| X_{i}, \widehat{x}_{k} \right]$$

$$= E_{\Delta_{1}, ..., \Delta_{i}, \Delta_{k}} \left[Y_{i}^{2} \right] - 2 E_{\Delta_{1}, ..., \Delta_{n}, \Delta_{k}} \left[Y_{k} g(x_{k}) \right] + E_{\Delta_{1}, ..., \Delta_{n}, \Delta_{k}} \left[g(x_{k})^{2} \right]$$

$$= E_{\Delta_{k}} \left[Y_{k}^{2} \right] - 2 E_{\Delta_{k}} \left[Y_{k} \right] E_{\Delta_{1}, ..., \Delta_{n}, \Delta_{k}} \left[g(x_{k}) \right] + E_{\Delta_{1}, ..., \Delta_{n}} \left[g(x_{k})^{2} \right]$$

$$= \int_{\mathbb{R}^{2}} \left[f(x_{k}) - E[g(x_{k})]^{2} + V_{a} F[g(x_{k})] \right]$$

$$= G^{2} + \left(f(x_{k}) - E[g(x_{k})]^{2} + V_{a} F[g(x_{k})] \right]$$

= 02 + Bies(g&x) 72 + Ver[g&x] Let's step it up two more notches in one shot (a) Let x1, ..., xn trandom realizations from P(X), thus X is random and (b) x^* is a

random realization from P(X) as well.

MSE

dataset-dataset.

$$= E_{\chi} \left[\sigma^2 + B_{in} f_g \right]^2 + V_{av} [g]$$

$$= \sigma^2 + E_{\chi} \left[B_{in} f_g \right]^2 \right] + E_{\chi} [V_{av} [g]]$$
This is the general bias-variance decomposition formula. This is not computable in practice since you only have one D and one g. It is a theoretical formula.

 $M5E = E_{X_1, ..., X_{k_1} X_{k_1}} [m_5 E(x_k)]$ previous result

Is there a "bias-variance tradeoff"? Yes and no... It is not a zero sum game. If you make smart modeling decisions both bias AND variance decrease. For one A, varying complexity of H

complexity optimal complexity

Overfit models have low bias since they are flexible and can fit f very snugly. Overfit models have high variance since they fit the y's snugly which means they fit the
$$\delta$$
's snugly and the δ 's change dataset-dataset. Underfit models have high bias since they cannot express f's complexity. Underfit models have low variance since they have few parameters whose estimates are stable

MSE

For any A, pick a large H (to intentionally overfit). What does MSE look like over n?

do anything about). Let's consider a metaalgorithm called "model averaging". You fit g_1, g_2, ..., g_M and then you ship their average,
$$g_{\text{avg}} := \frac{g_1 + g_2 + ... + g_M}{\mathcal{M}}$$

As n-->infinity, Var[g] --> 0. We are only limited by Bias which we can reduce by increasing complexity in curlyH (sigsq we can't

MSE = 62 + Ex [Diss[fam]] + Ex [Var/gray]

What is the MSE of g_avg?

$$= \sigma^{2} + \mathbb{E}_{X} \left[\left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{n} \left(\int_{-g_{1} + \dots + g_{m}}^{g_{1} + \dots + g_{m}} \right)^{2} \right] + \mathbb{E}_{X} \left[\sqrt{$$

Although it's impossible, can we decrease their dependence? Yes. Enter Leo Breiman in 1994. Here was his big breakthrough. Consider taking samples of n rows of D with replacement (this is called a "nonparametric bootstrap sample". You can prove (in Math 241) that approximately 2/3 of the rows of D appear in each sample and 1/3 are missing. Do this M times.

sample and 1/3 are missing. Do this M times.

$$D_{(1)} = 5 \text{ mple}(D), \quad D_{(2)} = 5 \text{ mple}(D), \dots, \quad D_{(N)} = 5 \text{ mple}(D)$$
All these bootstrap-sample D_m's are a little bit different from each other since they have different observations missing

All these bootstrap-sample D_m's are a little bit different from each other since they have different observations missing and different observations duplicated. Then, you fit your models using the same A and H,

$$g_1 = \mathcal{A}\left(\mathcal{D}_{\mathcal{O}_1},\mathcal{H}\right), \ g_2 = \mathcal{A}\left(\mathcal{D}_{\mathcal{C}_2},\mathcal{H}\right), \dots, \ g_{\mathcal{D}_1} = \mathcal{A}\left(\mathcal{D}_{\mathcal{C}_2},\mathcal{H}\right)$$
 these models will be less dependent on one another than if you didn't use boostrap samples of D. So when you average,

 $g_{\text{avg}} = \frac{1}{m} \sum_{b=1}^{m} f_b$ You should be able to bring MSE down. This meta-procedure is called boostrap aggregation = "bagging". We will calculate the reduction from bagging in MSE next class.