

Response Space	Type of Modeling	g return	Example alg
$y \in \mathbb{R}$	regression	$\hat{y} \in y$	OLS
$y = \{c_1, c_2, \dots, c_K\}$	classification	$\hat{y} \in y$	KNN
$K=2, y = \{0, 1\}$	binary classification	$\hat{y} \in y$	SVM
$y \in \mathbb{R}_{\geq 0}$	survival	$\hat{y} \in y$	Weibull regression
$y \in \{0, 1, 2, \dots\}$	count	$\hat{y} \in y$	Poisson regression
$y \in (0, 1)$	proportion	$\hat{y} \in y$	Beta regression
$y = \{c_1, c_2, \dots, c_K\}$	probability estimation	$\hat{p} := \begin{matrix} P(Y=c_1 \vec{x}) \\ P(Y=c_2 \vec{x}) \\ \vdots \\ P(Y=c_K \vec{x}) \end{matrix}$	Multi-logit Regression
$K=2, y = \{0, 1\}$	probability estimation	$\hat{p} := P(Y=1 \vec{x})$	Logistic Regression
$y \in \{c_1, \dots, c_K\}$ Ordinal	probability estimation		Proportional odds model
$\text{If } y = \{0, 1\} \text{ for all } u,$			

$$y = t(\vec{z})$$

$$= f(\vec{x}) + \delta \quad \text{where } \delta \in \{0, -1, +1\}$$

$$= h^*(\vec{x}) + \varepsilon \quad \text{where } \varepsilon \in \{0, -1, +1\}$$

$$= g(\vec{x}) + e \quad \text{where } e \in \{0, -1, +1\}$$

How do we build a probability estimation model? Naively, $g_0 = \bar{y}$

$$\Leftrightarrow Y \sim \text{Bern}(t(\vec{z}))$$

We now view Y as a realization from a random variable (bernoulli).

We will assume there exists a function $f_{pr}(\vec{x}): \mathbb{R}^{p+1} \rightarrow (0,1)$

and this function is the best guess of the probability $P(Y=1|xvec)$ you can create with $xvec$.

$$Y \sim \text{Bern}(f_{pr}(\vec{x}) + \underbrace{t(\vec{z}) - f_{pr}(\vec{x})}_{f_{pr}})$$

$$\Rightarrow Y \sim \text{Bern}(f_{pr}(\vec{x})). \quad f_{pr} \text{ is the model we want to find.}$$

Let's assume that all the data (all the n observations) in D are independently realized.

$$P(D) = P(Y_1=y_1, Y_2=y_2, \dots, Y_n=y_n \mid \vec{x}_1, \dots, \vec{x}_n)$$

$$= \prod_{i=1}^n P(Y_i=y_i \mid \vec{x}_i)$$

$$= \prod_{i=1}^n f_{pr}(\vec{x}_i)^{y_i} (1-f_{pr}(\vec{x}_i))^{1-y_i}$$

$V \sim \text{Bern}(b)$
 $\sim \theta^V (1-\theta)^{1-V}$

Now we want to "fit" f_{pr} using our data (learning from data paradigm). How? Is this even possible? NO. We cannot fit arbitrary functions in any dimension. We need a set of candidate functions that we can fit. Call that \mathcal{H}_{pr} . Each element in this set maps $\mathbb{R}^{p+1} \rightarrow (0,1)$. How about:

$$\mathcal{H}_{pr} = \{ \vec{w} \cdot \vec{x} : \vec{w} \in \mathbb{R}^{p+1} \} ?$$

We can't use this since it returns values outside $(0,1)$, the space of legal probabilities. But... we really like $wvec \cdot xvec$ because (1) easy to interpret and we have lots of intuition about it from all of our previous modeling we've done and (2) monotonic in each of the x_j 's. How we do we have our cake and eat it too?

We need a function that takes $wvec \cdot xvec$ and maps it into the space $(0,1)$ i.e. $\phi: \mathbb{R} \rightarrow (0,1)$ which is called a "link function" I think because it links the two spaces (the reals and the prob's). We restrict the link function to be strictly increasing. Thus,

$$\mathcal{H}_{pr} = \{ \phi(\vec{w} \cdot \vec{x}) : \vec{w} \in \mathbb{R}^{p+1} \}$$

These types of models are called "generalized linear models" (glm) because they retain $wvec \cdot xvec$ (the linear model) but then manipulate it in some way. Which link function should we use? There are three common ones. In order of use:

- ① Logistic / logit: $\phi(u) := \frac{e^u}{1+e^u} = \frac{1}{1+e^{-u}}$. Note: $1-\phi(u) = \frac{1}{1+e^u}$
- ② Probit: $\phi(u) := \Phi(u)$ i.e. the CDF of the std. normal.
- ③ Complementary Log-Log (cloglog)

$$\phi(u) = 1 - e^{-e^u} \Rightarrow 1 - \phi(u) = e^{-e^u} \Rightarrow \ln(1 - \phi(u)) = -e^u$$

$$\Rightarrow -\ln(1 - \phi(u)) = e^u \Rightarrow u = \ln(-\ln(1 - \phi(u)))$$

$\underbrace{\ln(-\ln(1 - \phi(u)))}_{\text{complement}}$

Let's employ the logistic link function:

$$\mathcal{H} = \{ \frac{1}{1+e^{-\vec{w} \cdot \vec{x}}} : \vec{w} \in \mathbb{R}^{p+1} \}$$

What is A ? How to get $g \in \mathcal{H}$?

Why not find the $wvec$ that provides us the highest probability?

$$A: \vec{b} := \underset{\vec{w} \in \mathbb{R}^{p+1}}{\text{argmax}} \left\{ \underbrace{\prod_{i=1}^n \left(\frac{1}{1+e^{-\vec{w} \cdot \vec{x}_i}} \right)^{y_i} \left(\frac{1}{1+e^{\vec{w} \cdot \vec{x}_i}} \right)^{1-y_i}}_{P(D)} \right\}$$

In OLS, we took the derivative and set it equal to zero to solve for $bvec$ and we found an analytical solution. However, there is no analytical solution here. You need to use a computer.

$$\vec{\nabla} P(D) \stackrel{\text{set}}{=} \vec{0}_{p+1} \quad \text{and approximate}$$

Usually this is done with "gradient descent". Computing $bvec$ is called "running a logistic regression". Once this is done... we can predict using

$$\hat{p} = g_{pr}(\vec{x}) = \phi(\vec{b} \cdot \vec{x}) = \frac{1}{1+e^{-\vec{b} \cdot \vec{x}}} \quad \text{hopefully close to } f_{pr}(\vec{x})$$

$$\hat{p}(Y=1|\vec{x})$$

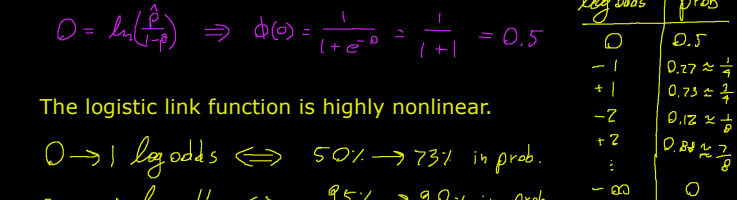
What is the interpretation of the slope coefficients (the entries in the b -vec)?

$$\hat{p} = \frac{1}{1+e^{-\vec{b} \cdot \vec{x}}} \Rightarrow \frac{1}{\hat{p}} = 1+e^{-\vec{b} \cdot \vec{x}} \Rightarrow \frac{1}{\hat{p}} - 1 = e^{-\vec{b} \cdot \vec{x}}$$

$$\Rightarrow \frac{1-\hat{p}}{\hat{p}} = e^{-\vec{b} \cdot \vec{x}} \Rightarrow \ln\left(\frac{1-\hat{p}}{\hat{p}}\right) = -\vec{b} \cdot \vec{x} \Rightarrow \vec{b} \cdot \vec{x} = \ln\left(\frac{\hat{p}}{1-\hat{p}}\right)$$

$$\text{Odds} = \frac{\hat{p}}{1-\hat{p}} \quad \underbrace{\ln\left(\frac{\hat{p}}{1-\hat{p}}\right)}_{\text{log-odds}}$$

$\Rightarrow b_j$ is the change in the log-odds of $Y=1$ if x_j increases by 1.



$$0 = \ln\left(\frac{\hat{p}}{1-\hat{p}}\right) \Rightarrow \phi(0) = \frac{1}{1+e^0} = \frac{1}{1+1} = 0.5$$

The logistic link function is highly nonlinear.

$$0 \rightarrow 1 \text{ log odds} \Leftrightarrow 50\% \rightarrow 73\% \text{ in prob.}$$

$$3 \rightarrow 4 \text{ log odds} \Leftrightarrow 95\% \rightarrow 98\% \text{ in prob.}$$

Probability estimation models predict probabilities but we observe labels (i.e. 0 or 1). The true probabilities f_{pr} are unobserved! We need a metric called a "scoring rule" S that can compare a p -hat value to a y value.

A "proper scoring rule" $S(p\text{-hat}, y)$ is one where:

$$\forall i \quad f_{pr}(\vec{x}_i) = \underset{\hat{p}}{\text{argmax}} \{ S(\hat{p}_i, y_i) \}$$

We will study two proper scoring rules:

- ① Brier score (1950). Let $s_i := -(y_i - \hat{p}_i)^2 \leq 0$

$$\bar{s} := \frac{1}{n} \sum_{i=1}^n s_i \leq 0$$

- ② Log scoring rule. Let $s_i := y_i \ln(\hat{p}_i) + (1-y_i) \ln(1-\hat{p}_i) \leq 0$

$$\bar{s} = \frac{1}{n} \sum s_i \leq 0$$

These scores are used as an "R^2" of the model (but they're not between 0 and 1) in a conceptual sense. The closer to zero, the better the probability estimation model.