# **Bayesian Deep Neural Networks**

Bayesian Neural Networks

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• A neural network with a single hidden layer with  $N_H$  units:

$$f(\mathbf{x}) = b + \sum_{j=1}^{N_H} v_j h(\mathbf{x}; \mathbf{u}_j) \in \mathbb{R}$$

where **x** is an input,  $h(\cdot; \mathbf{u}_i)$  is *j*-th hidden node.

- It is known that a neural network one hidden layer is a **universal** approximator as  $N_H \rightarrow \infty$  [1].
  - It can approximate any smooth function on a compact support under mild conditions.
  - This property also holds for a polynomial function or Gaussian process regression.

Neural network converges to a Gaussian process:

$$f(\mathbf{x}) = b + \sum_{j=1}^{N_H} v_j h(\mathbf{x}; \mathbf{u}_j) \in \mathbb{R}$$

- Suppose that
  - $b \sim \mathcal{N}(0, \sigma_b^2)$  and  $v_j \sim \mathcal{N}(0, \sigma_v^2)$
  - **u**<sub>j</sub> are independently and identically distributed
  - $\sigma_v^2$  scales as  $w^2/N_H$  then

$$\mathbb{E}(f(\mathbf{x})) = 0$$

$$\mathbb{E}(f(\mathbf{x})f(\mathbf{x}')) = \sigma_b^2 + \sum_{j=1}^{N_H} \sigma_v^2 \mathbb{E}_u(h(\mathbf{x}; \mathbf{u}_j)h(\mathbf{x}'; \mathbf{u}_j))$$

$$= \sigma_b^2 + w^2 \mathbb{E}_u(h(\mathbf{x}; \mathbf{u}_j)h(\mathbf{x}'; \mathbf{u}_j))$$

• By the central limit theorem,  $f(\mathbf{x})$  converges to a Gaussian process as  $N_H \to \infty$ .

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$$= \sigma_b^2 + w^2 \mathbb{E}_u(h(\mathbf{x}; \mathbf{u}_j)h(\mathbf{x}'; \mathbf{u}_j))$$

• By the central limit theorem,  $f(\mathbf{x})$  converges to a Gaussian process as  $N_H \to \infty$ .

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• If  $h(\mathbf{x}; \mathbf{u}) = \text{erf}(u_0 + \sum_{j=1}^{N_H} u_j x_j)$  and  $\mathbf{u} \sim \mathcal{N}(0, \Sigma)$ , then the covariance function of the neural network is

$$k_{NN}(\mathbf{x}, \mathbf{x}') = \frac{2}{\pi} \sin^{-1} \left( \frac{2\bar{\mathbf{x}}^T \Sigma \bar{\mathbf{x}}'}{\sqrt{(1 + 2\bar{\mathbf{x}}^T \Sigma \bar{\mathbf{x}})(1 + 2\bar{\mathbf{x}}'^T \Sigma \bar{\mathbf{x}}'}} \right)$$

where 
$$\bar{\mathbf{x}} = [1, x_1, ..., x_d]^T$$
.

First four papers are optimizing the SAME objectives.

Length

Minimizing the Description

• Geoffrey Hinton and Drew van Camp, 1993

- Minimum description length principle
  - When fitting models to data, it is alway possible to fit the training data better by using a more complex model.
  - The minimum description length principle asserts the the best model is the one that minimizes the combined cost of:
    - 1. describing the model
    - 2. describing the misfit between the model and the data
  - One can think in terms of:
    - 1. a sender who can see both the input and output
    - 2. a receiver who can only see the input.

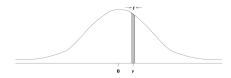
- 1. The sender first fits a neural network to the complete set of training data.
- 2. Then sends the weights of the network to the receiver.
- 3. For each training data, the sender also sends the discrepancy between the network output and the correct output.
- 4. By adding the discrepancy to the output of the network, the receiver can generate exactly the correct output.

• Total cost = coding the data misfits + coding the weights

#### Coding the data misfits

- If the data misfits are real numbers, an infinite amount of information is needed.
- So we assume that they are finely quantized using intervals of fixed width t.
- The coding theorem tells us that:
  - if a sender and a receiver agreed on a probability distribution that assigns a probability mass p(Δy) to each possible quantized data misfit.
  - then we can code the misfit using  $-\log_2 p(\Delta y)$  bits.
  - Example) If  $\Delta y$  quantizes the misfit into eight parts and the probability mass is assigned equally, then  $p(\Delta y) = \frac{1}{8}$  and we need  $-\log_2(\frac{1}{8}) = 3$  bits.

#### · Coding the data misfits



• We assume that the data misfits are encoded by assuming that they are drawn from a zeros-mean Gaussian:

$$p(d^{c} - y^{c}) = t \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left[\frac{-(d^{c} - y^{c})^{2}}{2\sigma^{2}}\right].$$

Using an optimal code, the description length of a data misfit,
 d − y, in units of log<sub>2</sub>(·) bits (called 'nats') is:

$$-\log p(d^c - y^c) = -\log t + \log \sqrt{2\pi} + \log \sigma + \frac{(d^c - y^c)^2}{\sigma^2}$$

### Coding the data misfits

Using an optimal code, the description length of a data misfit,
 d − y, in units of log<sub>2</sub>(·) bits (called 'nats') with N data is:

$$DL = -N \log t + N \log \sigma + \sum_{c} \frac{(d^{c} - y^{c})^{2}}{\sigma^{2}}$$

- $\sigma^*$  that minimizes DL is  $\sigma^* = \sqrt{\frac{1}{N} \sum_c \frac{(d^c y^c)^2}{\sigma^2}}$ .
- Substituting  $\sigma^*$ , we get

$$DL^* = -N \log t + \frac{N}{2} \log \left[ \frac{1}{N} \sum_{c} (d^c - y^c)^2 \right] + \frac{N}{2}$$

• The final data misfit cost becomes:

$$C_{\text{data-misfit}} = kN + \frac{N}{2} \log \left[ \frac{1}{N} \sum_{c} (d^{c} - y^{c})^{2} \right]$$

where k is a constant that depends on t.

#### Coding the weights

• We assume that the weights of the network are finely quantized and come from a zero-mean Gaussian distribution

$$p(w) = t \frac{1}{\sqrt{2\pi\sigma_w^2}} \exp(-\frac{w^2}{2\sigma_w^2}):$$

• Then, the description length of the weights becomes:

$$DL = \frac{1}{2\sigma_w^2} \sum_j w_j^2$$

- Total description length
  - The total description length (TDL) is the sum of
    - 1. coding the data misfits
    - 2. coding the weights.
  - Ignoring k which related to the quantization level t, we get

$$TDL = \sum_{i} \frac{1}{\sigma^{2}} \sum_{c} (d_{i}^{c} - y_{i}^{c})^{2} + \frac{1}{2\sigma_{w}^{2}} \sum_{j} w_{j}^{2}.$$

• Surprisingly, the is just the standard weight-decay method.

## Noisy weights

#### • The 'bits back' argument

- Suppose that the sender and receiver have an agreed Gaussian prior P(w).
- After learning, the sender has a Gaussian posterior Q(w).
- We will show that the amount of required bits to communicate the posterior distribution of a weight is equal to the KL-distance from P to Q:

$$G(P,Q) = \int Q(w) \log \frac{Q(w)}{P(w)} dw.$$

• Note that  $D_{KL}(P||Q) = \int p(x) \log \frac{p(x)}{q(x)} dx$ .

## **Noisy weights**

#### The 'bits back' argument

- 1. To communicate the noisy weights, the sender first collapses the weights drawn from the posterior Q(w) to pick a precise value within the tolerance t.
- 2. Then the sender sends each weight for Q(w) by coding them using the prior P(w) where the communication cost is:

$$C(w) = \log t - \log P(w)$$

where  $\log t$  is for quantization and  $\log P(w)$  is for coding w with the prior. The sender also communicates the data-misfits.

3. Since the receiver has correct output and misfits, he can also recover the exact same posterior probability Q(w). Hence, following information is redundant:

$$R(w) = \log t \log Q(w)$$

4. So the true expected description length is determined by taking an expectation under Q(w):

$$G(P,Q) = \int Q(w) \log \frac{Q(w)}{P(w)} dw.$$

## Noisy weights

- The 'bits back' argument
  - The 'bits back' argument tells us that when we have a prior P(w)
     on weights and the sender has a posterior Q(w), the expected
     description length for a noisy (random) weights is:

$$G(P,Q) = \int Q(w) \log \frac{Q(w)}{P(w)} dw$$

where  $G(P,Q) = D_{KL}(Q||P)$ .

• Note that the goal of most of variational inference problems for Bayesian neural networks is to find Q(w) that minimizes  $D_{KL}(Q||P)$  given some prior P(w).

**Ensemble Learning in Bayesian** 

**Neural Network** 

#### Bayesian neural network

• Consider a two-layer network with *H* hidden units:

$$f(x,w) = \sum_{i=1}^{H} v_i \rho(u^T x)$$

where  $\rho(a)$  is the 'erf' (cumulative Gaussian) function

$$\rho(a) = \sqrt{\frac{2}{\pi}} \int_0^a \exp(-s^2/2) ds.$$

• We assume the output has a measurement noise whose variance is  $\beta^{-1}$  and the likelihood of w and  $\beta$  becomes:

$$P(D|w,\beta) = \frac{\exp(-\beta E_D)}{Z_D}$$

where  $Z_D$  is a normalizing factor and  $E_D$  is the training error:

$$E_D(w) = \frac{1}{2} \sum_i (f(x^i, w) - t^i)^2.$$

- Bayesian neural network
  - Likelihood:

$$P(D|w,\beta) = \frac{\exp(-\beta E_D)}{Z_D}$$

where  $Z_D$  is a normalizing factor and  $E_D$  is the training error:

$$E_D(w) = \frac{1}{2} \sum_i (f(x^i, w) - t^i)^2.$$

• Prior:

$$p(w|A) = \frac{\exp(-E_W(w))}{Z_P}$$

where  $E_W(w) = \frac{1}{2} w^T A w$ .

• Posterior:

$$p(w|D,\beta,A) = \frac{1}{Z_F} \exp(\beta E_D(w) - E_W(w)).$$

- Bayesian neural network
  - Given a posterior distribution:

$$p(w|D,\beta,A) = \frac{1}{Z_F} \exp(-\beta E_D(w) - E_W(w)),$$

the prediction for a new input are given by **integration** over the posterior distribution.

• For example, the predictive mean is given by

$$\langle f(x) \rangle = \int f(x,w) P(w|D,\beta,A) dw.$$

### Laplace's method

 We approximate the posterior distribution with a Gaussian distribution:

$$p(w|D, \beta, A) = \exp(-\phi(w)).$$

ullet Then, expand  $\phi$  around the **mode** of the distribution

$$w_* = \arg\min\phi(w)$$

so that

$$\phi(w) \approx \phi(w_*) + \frac{1}{2}(w - w_*)^T H(w - w_*)$$

where  $H = \nabla \nabla \phi(W)|_{w_*}$ .

• The expected value of f,  $\langle f(x) \rangle = \int f(x, w) P(w|D, \beta, A) dw$  can be evaluated by making a further local linearization of f(x, w).

#### Markov chain Monte Carlo method

 The central idea is to replace integrals weighted by the posterior by finite sums, so that

$$\int p(w|D,\beta,A)g(w)dw \approx \frac{1}{m}\sum_{i=1}^{m}g(w_i)$$

where  $w_i$  are drawn from the posterior.

- Examples of MCMC:
  - Metropolis-Hastings algorithm: generates a random walk using a function that is proportional to the proposal density.
  - Gibbs sampling: Sample from the conditional distributions. For example, if we want to sample from p(x<sub>1</sub>, x<sub>2</sub>), use p(x<sub>1</sub>|x<sub>2</sub>) and p(x<sub>2</sub>|x<sub>1</sub>).

### Ensemble learning

- This method is called a variational inference, nowadays.
- We introduce a variational distribution Q(w) that approximates the posterior.
- We start with a marginal likelihood:

$$\ln p(D|\beta, A) = \ln \int p(D|w, \beta)p(w|A)dw$$

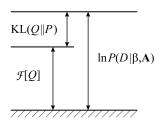
$$= \ln \int \frac{p(D|w, \beta)p(w|A)}{Q(w)}Q(w)dw$$

$$\geq \int \ln \left\{ \frac{p(D|w, \beta)p(w|A)}{Q(w)} \right\}Q(w)dw$$

$$= \mathcal{F}[Q].$$

•  $\mathcal{F}[Q]$  only requires likelihood  $p(D|w,\beta)$  and prior p(w|A).

• Ensemble learning



•  $\mathcal{F}[Q]$  is a lower bound of log marginal likelihood In  $p(D|\beta, A)$ .

## Ensemble learning

• Show that 
$$\ln p(D|\beta, A) = KL(Q||P) + \mathcal{F}[Q]$$
:

1.  $\ln p(D|\beta, A) = KL(Q||P) + \int \ln \left\{ \frac{p(D|w,\beta)p(w|A)}{Q(w)} \right\} Q(w) dw$ 

2.  $= -\int \ln \frac{p(w|D,\beta,A)}{Q(w)} Q(w) dw + \int \ln \left\{ \frac{p(D|w,\beta)p(w|A)}{Q(w)} \right\} Q(w) dw$ 

3.  $= \int Q(w) \left[ \ln \frac{p(D|w,\beta)p(w|A)}{p(w|D,\beta,A)} \right] dw$ 

4.  $= \int Q(w) \left[ \ln \frac{p(D,w,\beta,A)}{p(w|D,\beta,A)} \right] dw$ 

5.  $= \int Q(w) \left[ \ln \frac{p(w|D,\beta,A)p(D|\beta,A)}{p(w|D,\beta,A)} \right] dw$ 

6.  $= \int Q(w) \ln p(D|\beta,A) dw$ 

7.  $= p(D|\beta,A)$ 

• More fancy if we go the reverse direction.

#### Ensemble learning

- What we have seen so far is that if we maximize  $\mathcal{F}[Q]$  then it is equivalent to saying that we are maximizing the lower bound of the marginal log likelihood  $\ln p(D|\beta, A)$ .
- \( \mathcal{F}[Q] \) is often calls evidence lower bound (ELBO) and heavily used
  in formulating variational auto-encoders (VAEs).
- Furthermore,

$$\mathcal{F}[Q] = \int \ln \left\{ \frac{p(D|w,\beta)p(w|A)}{Q(w)} \right\} Q(w)dw$$

$$= \int \left( \ln p(D|w,\beta) + \ln \frac{p(w|A)}{Q(w)} \right) Q(w)dw$$

$$= \int \ln p(D|w,\beta)Q(w)dw - D_{KL}(Q(w)||p(w|A)$$

- ullet As our goal is to maximize  $\mathcal{F}[Q]$ , we are trying to
  - 1. maximize the marginal likelihood p(D)
  - 2. reduce the gap between p(w|D) and q(w).
  - 3. keep the variational distribution Q(w) close to our prior p(w|A).

#### Ensemble Learning

- In this paper, Q(w) is a Gaussian with mean  $\bar{w}$  and covariance C.
- Main contribution of this paper is to find an efficient update rule for optimizing Q(w) w.r.t.  $\bar{w}$  and C using matrix calculus and the fact that  $\rho(a)$  is the 'erf' function.
- Note that the previous paper (Hinton and Camp, 1993) assumes that C is a diagonal matrix.

#### Introduction

- "In the eighteen years since variational inference was first proposed for neural networks (Hinton and Camp, 1993) it has not seen widespread use.
- We believe this is largely due to the difficulty of deriving analytical solutions to the required integrals over the variational posteriors."
- "The approach taken here is to forget about analytical solutions and search instead for variational distributions whose expectation values (and derivatives thereof) can be efficiently approximated with numerical integration."

#### Variational inference

• The goal of the variational inference is to fit Q(w) by minimizing variational free energy  $\mathcal{F}]Q]$  w.r.t. parameter  $\beta$ :

$$\mathcal{F} = -\left\langle \ln \left[ \frac{p(D|w)p(w|\alpha)}{Q(w|\beta)} \right] \right\rangle_{w \sim Q(\beta)}$$

#### Minimum description length

• The variational free energy (to be minimized) can be reinterpreted as:

$$\mathcal{F} = -\left\langle \ln \left[ \frac{p(D|w)p(w|\alpha)}{Q(w|\beta)} \right] \right\rangle_{w \sim Q(\beta)}$$
$$= \langle L^{N}(w,D) \rangle_{w \sim Q(\beta)} + D_{KL}(Q(\beta)||P(\alpha))$$

where  $L^N = -\ln p(D|w) = -\sum_{(x,y)\in D} \ln p(y|x,w)$  is the negative log likelihood.

• The first term in  $\mathcal{F}$  is called the error loss:

$$L^{E}(\beta,D)=\langle L^{N}(w,D)\rangle_{w\sim Q(\beta)}.$$

ullet The second term in  ${\mathcal F}$  is called the complexity loss:

$$L^{C}(\alpha, \beta) = D_{KL}(Q(\beta)||P(\alpha)).$$

which can be realized with bits-back coding.

#### Total loss

• The total loss is:

$$L(\alpha,\beta,D) = L^{E}(\beta,D) + L^{C}(\alpha,\beta)$$

where 
$$L^{E}(\beta, D) = \langle L^{N}(w, D) \rangle_{w \sim Q(\beta)}$$
 and  $L^{C}(\alpha, \beta) = D_{KL}(Q(\beta) || P(\alpha)).$ 

ullet The negative log likelihood  $L^N$  is:

$$L^{N}(w, D) = -\ln p(D|w) = -\sum_{(x,y)\in D} \ln p(y|x, w)$$

#### Choice of distributions

- Delta posterior
  - The simplest posterior  $Q(\beta)$  that assigns probability 1 to a particular set of weights w.
  - We recover ordinary maximum likelihood training.
  - If the prior is a Laplace distribution with µ = 0, then we have L1 regularization.
  - If the prior is a Gaussian distribution with  $\mu$  = 0, then we have L2 regularization.

#### Choice of distributions

- Gaussian posterior
  - Suppose  $Q(\beta) \sim \mathcal{N}(\mu, \sigma)$ .
  - We approximate  $L^{E}(\beta, D)$  as:

$$L^{E}(\beta, D) \approx \frac{1}{S} \sum_{k=1}^{S} L^{N}(w^{k}, D)$$

where  $w^k$  drawn independently from  $Q(\beta)$ .

• We also utilize following identities of Gaussian differentiations:

$$\nabla_{\mu} \langle V(a) \rangle_{a \sim \mathcal{N}} = \langle \nabla_{a} V(a) \rangle_{a \sim \mathcal{N}} \text{ and } \nabla_{\Sigma} \langle V(a) \rangle_{a \sim \mathcal{N}} = \frac{1}{2} \langle \nabla_{a} \nabla_{a} V(a) \rangle_{a \sim \mathcal{N}}$$

- This paper is the first paper to formally use a variational inference on Bayesian neural networks.
- This paper assumes that the posterior distribution is a Gaussian distribution where the covariance matrix is diagonal and presented efficient gradients for optimizing the mean and variance.

- Weight uncertainty in neural networks
  - All weights in the neural networks are represented by probability distributions over passible values, rather than having a single fixed value.
  - Instead of training a single network, the propose method trains an ensemble of networks, where each network has its weights drawn from a shared, learnt probability distribution.

- Being Bayesian by backpropagation
  - The goal is to minimize the variational free energy:

$$\begin{split} \mathcal{F}(D,\theta) &= D_{KL}(q(w|\theta)||p(w)) - \mathbb{E}_{q(w|\theta)}[\log p(D|w)] \\ &= \int q(w|\theta) \log \frac{q(w|\theta)}{p(w)} dw - \int q(w|\theta) \log p(D|w) dw \\ &= \int q(w|\theta) [\log q(w|\theta) - \log p(W) - \log p(D|w)] dw \end{split}$$

ullet We can approximate  $\mathcal{F}(D, heta)$  as

$$\mathcal{F}(D, \theta) \approx \log q(w|\theta) - \log p(W) - \log p(D|w)$$

where w is drawn from  $q(w|\theta)$ .

- Being Bayesian by backpropagation
  - Use  $f(w, \theta) = \log q(w|\theta) \log p(W) \log p(D|w)$  for training  $q(w|\theta)$ .
  - This is it.

- Log marginal likelihood
  - = ELBO (variational free energy) +  $D_{KL}(q(w|\theta)||p(w|D))$
- Note that p(w|D) can hardly be computed analytically.

Derivation:

$$\ln p(D) = \int \ln p(D)q(w|\theta)dw$$

$$= \int q(w|\theta) \ln \frac{p(D)p(w|D)}{p(w|D)}dw$$

$$= \int q(w|\theta) \ln \frac{p(w,D)}{p(w|D)}dw$$

$$= \int q(w|\theta) \ln \frac{p(D|w)p(w)}{p(w|D)}dw$$

$$= \int q(w|\theta) \ln \frac{q(w|\theta)p(D|w)p(w)}{q(w|\theta)p(w|D)}dw$$

$$= \int q(w|\theta) \ln \frac{q(w|\theta)p(D|w)p(w)}{q(w|\theta)p(w|D)}dw$$

$$= \int q(w|\theta) \ln \frac{q(w|\theta)}{p(w|D)}dw + \int q(w|\theta) \ln \frac{p(D|w)p(w)}{q(w|\theta)}dw$$

$$= D_{KL}(q(w|\theta)||p(w|D)) + \mathcal{F}[q]$$

where  $\mathcal{F}[q]$  is the variational free energy or ELBO.

- The variational free energy or ELBO can further experessed as:
- Derivation:

$$\mathcal{F}[q] = \int q(w|\theta) \ln \frac{p(D|w)p(w)}{q(w|\theta)} dw$$

$$= \int q(w|\theta) \ln p(D|w) dw + \int q(w|\theta) \ln \frac{p(w)}{q(w|\theta)} dw$$

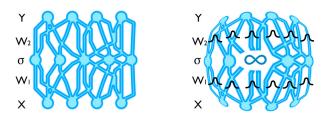
$$= \mathbb{E}_{q(w|\theta)} [\ln p(D|w)] - D_{KL}(q(w|\theta)||p(w))$$

$$= \text{likelihood under } q - \text{prior fitting term}$$

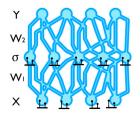
- ullet We try to maximize  $\mathcal{F}[q]$  to
  - 1. maximize the marginal likelihood p(D)
  - 2. reduce the gap between p(w|D) and q(w).
  - 3. keep the variational distribution Q(w) close to our prior p(w|A).

Dropout as a Bayesian

**Approximation** 



**Figure 1:** A neural network and Bayesian neural network. (Copyright to Yarin Gal's blog.)



**Figure 2:** A Bayesian neural network with dropout. (Copyright to Yarin Gal's blog.)

Goal of a Bayesian neural network is to find a posterior distribution:

$$p(W|X,Y) \propto P(Y|X,W)P(W)$$

where X and Y are the input and output training data and W is a set of parameters of our interest.

• Once we have p(W|X, Y), the output  $y_*$  at unseen  $\mathbf{x}_*$  is predicted as:

$$p(y_*|\mathbf{x}_*) = \int_W p(y|\mathbf{x}_*, W) p(W|X, Y) dW.$$

• In practice, none of them is tractable. ③

- Variational inference is often used to handle this issue.
- Instead of finding p(W|X,Y), optimize a variational distribution  $q_{\theta}(W)$  by minimizing

$$\mathsf{KL}(q_{\theta}(W)||p(W|X,Y)) = \int q_{\theta}(W) \log \frac{q_{\theta}(W)}{p(W|X,Y)} dW.$$

- Of course computing  $KL(q_{\theta}(W)||p(W|X,Y))$  is not tractable as well.  $\odot$
- However, we can compute the lower-bound called evidence lower-bound (ELBO) instead!

$$\mathsf{ELBO} = \int q_{\theta}(W) \log p(Y|X, W) dW - \mathsf{KL}(q_{\theta}(W)||p(W))$$

• In computing ELBO, we do not need to know p(W|X,Y).  $\odot$ 

$$\mathsf{ELBO} = \int q_{\theta}(W) \log p(Y|X, W) dW - \mathsf{KL}(q_{\theta}(W)||p(W))$$

- Instead, all we need is
  - prior: p(W)
  - variational distribution:  $q_{\theta}(W)$
  - likelihood: p(Y|X, W)

• How can we define a likelihood function?

- Let's use a Gaussian process.
- We need a kernel function.

We will use following kernel function:

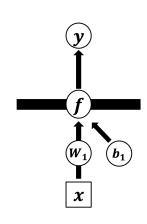
$$K(\mathbf{x}, \mathbf{y}) = \int p(\mathbf{w})p(b)\sigma(\mathbf{w}^T\mathbf{x} + b)\sigma(\mathbf{w}^T\mathbf{y} + b)d\mathbf{w}db.$$

• Then we use Monte Carlo integration with *K* terms:

$$\hat{K}(\mathbf{x}, \mathbf{y}) = \frac{1}{K} \sum_{k=1}^{K} \sigma(\mathbf{w}_{k}^{T} \mathbf{x} + b_{k}) \sigma(\mathbf{w}_{k}^{T} \mathbf{y} + b_{k})$$

where  $\mathbf{w}_k \sim p(\mathbf{w})$  and  $b_k \sim p(b)$ .

• Note that K will be the number hidden units in the hidden layer.



• Using  $\hat{K}$  as the covariance function of the Gaussian process, we have the followings:

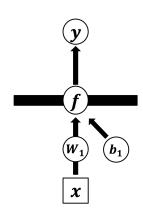
$$\mathbf{w}_{k} \sim p(\mathbf{w}), b_{k} \sim p(b)$$

$$W_{1} = [\mathbf{W}_{k}]_{k=1}^{K}, \mathbf{b} = [b_{k}]_{k=1}^{K}$$

$$\hat{K}(\mathbf{x}, \mathbf{y}) = \frac{1}{K} \sum_{k=1}^{K} \sigma(\mathbf{w}_{k}^{T} \mathbf{x} + b_{k}) \sigma(\mathbf{w}_{k}^{T} \mathbf{y} + b_{k})$$

$$F|X, W_{1}, b \sim \mathcal{N}(0, \hat{K}(X, X))$$

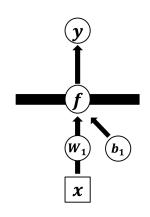
$$Y|F \sim \mathcal{N}(F, \tau^{-1}I_{N})$$
where  $W_{1} \in \mathbb{R}^{Q \times K}$ .



• The marginal likelihood becomes:

$$p(Y|X) = \int p(Y|F)p(F|W_1, b, X)p(W_1)p(b)$$
 where the integration is w.rt.  $F$ ,  $W_1$ , and  $b$ .

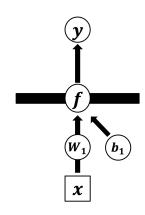
- What we have so far is a marginal likelihood where the W<sub>1</sub> and b are auxiliary random variables.
- Where do they came from?
  - $\hat{K}(\mathbf{x}, \mathbf{y}) = \frac{1}{K} \sum_{k=1}^{K} \sigma(\mathbf{w}_{k}^{T} \mathbf{x} + b_{k}) \sigma(\mathbf{w}_{k}^{T} \mathbf{y} + b_{k}).$



• Suppose  $\phi(\mathbf{x}) \in \mathbb{R}^K$  is a hidden layer corresponding to  $\mathbf{x} \in \mathbb{R}^Q$ :

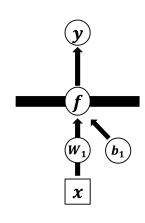
$$\phi(\mathbf{x}, W_1, b) = \sqrt{\frac{1}{\kappa}} \sigma(W_1^T \mathbf{x} + b).$$

- Let  $\Phi = [\phi(\mathbf{x}_n, W_1, b)]_{n=1}^N \in \mathbb{R}^{N \times K}$  be a concatenated feature matrix.
- Then we have  $\hat{K}(X,X) = \Phi \Phi^T$ .



• Finally we have the following likelihood:

$$p(Y|X) = \int \mathcal{N}(Y; 0, \Phi\Phi^T + \tau^{-1}) p(W_1) p(b)$$
 where the integration is w.r.t.  $W_1$  and  $b$ .



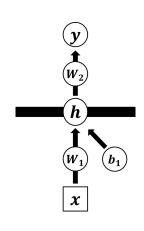
- Here comes a tricky part.
- The probability distribution of  $Y \in \mathbb{R}^{N \times D}$  can be written as joint distribution of column vectors  $\mathbf{y}_d \in \mathbb{R}^N$ .
- Then for each y<sub>d</sub> has the following distribution:

$$\mathcal{N}(\mathbf{y}_d; 0, \Phi\Phi^T + \tau I)$$

$$= \int \mathcal{N}(\mathbf{y}_d; \Phi\mathbf{w}_d, \tau^{-1}I) \mathbf{N}(\mathbf{w}_d; 0, I) d\mathbf{w}_d$$

where  $\mathbf{w}_d \sim \mathcal{N}(0, I)$  and  $\Phi$  is a concatenated feature matrix of

$$\phi(\mathbf{x}, W_1, b) = \sqrt{\frac{1}{K}} \sigma(W_1^T \mathbf{x} + b).$$



 Now we know that y<sub>d</sub> has the following distribution:

$$\mathcal{N}(\mathbf{y}_d; 0, \Phi\Phi^T + \tau I)$$

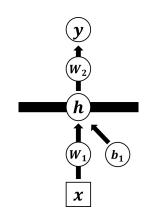
$$= \int \mathcal{N}(\mathbf{y}_d; \Phi\mathbf{w}_d, \tau^{-1}I) \mathbf{N}(\mathbf{w}_d; 0, I) d\mathbf{w}_d$$

where  $\mathbf{w}_d \sim \mathcal{N}(0, I)$  and  $\Phi$  is a concatenated feature matrix of

$$\phi(\mathbf{x}, W_1, b) = \sqrt{\frac{1}{\kappa}} \sigma(W_1^T \mathbf{x} + b).$$

• Writing  $W_2 = [\mathbf{w}_d]_{d=1}^D \in \mathbb{R}^{K \times D}$  then we have the following likelihood:

$$p(Y|X) = \int p(Y|X, W_1, W_2, b) p(W_1) p(W_2) p(b)$$



• Finally, we have:

$$\begin{split} p(Y|X) &= \int p(Y|X, W_1, W_2, b) p(W_1) p(W_2) p(b) \\ \text{where } X \in \mathbf{R}^{N \times Q}, \ \Phi \in \mathbf{R}^{N \times K}, \ Y \in \mathbf{R}^{N \times D}, \\ W_1 \in \mathbf{R}^{Q \times K}, \ \text{and} \ W_2 \in \mathbf{R}^{K \times D}. \end{split}$$

Recall ELBO:

$$\mathsf{ELBO} = \int q_{\theta}(W) \log p(Y|X,W) dW - \mathsf{KL}(q_{\theta}(W)||p(W))$$
 where  $q_{\theta}(W) = q(W_1,W_2,b_1)$  and  $\log p(Y|X,W) = p(Y|X,W_1,W_2,b_1)$ .

• To sum up, we have

$$\int_{W_1W_2b_1} q_{\theta}(W_1, W_2, b_1) \log p(Y|X, W_1, W_2, b_1) dW_1 dW_2 db_1$$

$$- \mathsf{KL}(q_{\theta}(W_1, W_2, b_1) || p(W_1, W_2, b_1)).$$

• However, integration with respect to  $\{W_1, W_2, b_1\}$  is still problematic.

$$\begin{split} \int_{W_1W_2b_1} & q_{\theta}(W_1, W_2, b_1) \log p(Y|X, W_1, W_2, b_1) dW_1 dW_2 db_1 \\ & - \mathsf{KL}(q_{\theta}(W_1, W_2, b_1) || p(W_1, W_2, b_1)). \end{split}$$

• Use re-parametrization trick  $\tilde{W} = g(W, \epsilon_W)$  where  $\epsilon_W \sim p(\epsilon)$  is a random variable.

$$\begin{split} \int_{\epsilon_{W_1},\epsilon_{W_2},\epsilon_{b_1}} & q_{\theta}(\epsilon_{W_1},\epsilon_{W_2},\epsilon_{b_1}) log \, p(Y|X,\tilde{W}_1,\tilde{W}_2,\tilde{b}_1) d\epsilon_{W_1} d\epsilon_{W_2} d\epsilon_{b_1} \\ & - \mathsf{KL}(q_{\theta}(W_1,W_2,b_1) || p(W_1,W_2,b_1)). \end{split}$$

Approximate the integral with Mote Carlo integration:

$$\mathcal{L}_{GP-MC} \triangleq -\sum_{n=1}^{N} \log p(y_n | x_n, \hat{W}_1^n, \hat{W}_2^n, \hat{b}_1^n) + \text{KL}(q(W_1, W_2, b_1) || p(W_1, W_2, b_1))$$

with realizations  $\hat{W}_1^n$ ,  $\hat{W}_2^n$ ,  $\hat{b}_1^n$  from dropout.

• With Gaussian prior on the parameters  $W_1$ ,  $W_2$ , and  $b_1$ :

$$\mathcal{L}_{GP-MC} = -\sum_{n=1}^{N} \log p(y_n | x_n, \hat{W}_1^n, \hat{W}_2^n, \hat{b}_1^n)$$

$$+ \lambda_1 ||W_1||^2 + \lambda_2 ||W_2||^2 + \lambda_3 ||b_1||^2$$

with realizations  $\hat{W}_1^n$ ,  $\hat{W}_2^n$ ,  $\hat{b}_1^n$  from dropout.

• To sum up, we have:

$$\mathcal{L}_{GP-MC} = -\sum_{n=1}^{N} \log p(y_n | x_n, \hat{W}_1^n, \hat{W}_2^n, \hat{b}_1^n)$$

$$+ \lambda_1 ||W_1||^2 + \lambda_2 ||W_2||^2 + \lambda_3 ||b_1||^2$$

with realizations  $\hat{W}_1^n, \hat{W}_2^n, \hat{b}_1^n$  from dropout.

• Once we have  $W_1$ ,  $W_2$ , and  $b_1$ , predictive mean and variance can be approximated as:

$$\mathbb{E}_{q}(y_{*}) = \int y_{*} q(y_{*}|x_{*}) dy_{*}$$

$$\approx \frac{1}{T} \sum_{t=1}^{T} \hat{y}_{*}(x_{*}, \hat{W}_{1,t}, \hat{W}_{w,t}, \hat{b}_{1,t})$$

$$\mathbb{E}_{q}(y_{*}y_{*}) \approx \tau^{-1} + \frac{1}{T} \sum_{t=1}^{T} \hat{y}_{*}(x_{*}, \hat{W}_{1,t}, \hat{W}_{w,t}, \hat{b}_{1,t}) \hat{y}_{*}(x_{*}, \hat{W}_{1,t}, \hat{W}_{w,t}, \hat{b}_{1,t})$$

It can be naturally extended to multi-layer neural networks.

# Descent

Stein Variational Gradient

- Proposed a general purpose variational inference algorithm by
  - iteratively transports a set of particles to match the target distribution
  - by applying a form of functional gradient descent that minimizes the KL divergence.

#### Background

- Let x be a continuous random variable of interest in  $\mathcal{X} \subset \mathbb{R}^d$  and  $\{D_k\}$  is a set of i.i.d. observation
- $p_0(x)$  is prior,  $p(x) = \bar{p}(x)/Z$  is posterior where  $\bar{p}(x) \triangleq p_0(x) \prod_{i=1}^N p(D_k|x)$  where the conditioning on data D is omitted.
- Let  $k(x,x'): \mathcal{X} \times \mathcal{X} \to \mathbb{R}$  is a kernel and  $\mathcal{H}$  is the reproducing kernel Hilbert space corresponding to  $k(\cdot,\cdot)$ , i.e.,  $\mathcal{H} = \{f: f(x) = \sum_{i=1}^m a_i k(x,x_i), a_i \in \mathbb{R}, m \in \mathbb{N}, x_i \in \mathcal{X}\}$ , equipped with an inner product  $\langle f,g \rangle_{\mathcal{H}} = \sum_{ij} a_i b_j k(x_i,x_j)$ .
- $\phi(x) = [\phi_1(x), \dots, \phi_d(x)]^T$  is a smooth vector function.

• Stein's identity (for 1-D case)

$$\mathbb{E}_{x \sim p}[\mathcal{A}_p \phi(x)] = \int_{x \in \mathcal{X}} \nabla_x p(x) \phi(x) + \nabla_x \phi(x) p(x) dx$$
$$= \int_{x \in \mathcal{X}} \nabla_x (p(x) \phi(x)) dx$$
$$= 0$$

- This holds when  $p(x)\phi(x) = 0$ ,  $\forall x \in \partial \mathcal{X}$  for compact  $\mathcal{X}$ .
- For  $\mathcal{X} = \mathbb{R}^d$ , it is sufficient to show that  $p(x)\phi(x)$  goes to zero for  $x = \infty$  using Green's identity.
  - Suppose  $g(x) = p(x)\phi(x)$ . Then, very loosely speaking,

$$\int_{x \in \mathcal{X}} \nabla_x g(x) dx = \oint g(x) dx.$$

#### Stein discrepancy

• Intuitively speaking, it is the maximum violation of Stein's identity for some  $\phi$  in a set  $\mathcal{F}$ :

$$\mathbb{D}(q,p) = \max_{\phi \in \mathcal{F}} \{ \mathbb{E}_{x \sim q} [ \operatorname{trace}(\mathcal{A}_p \phi(x)) ] \}.$$

 $\bullet$  Here, the choice of the function set  ${\mathcal F}$  is critical as it decides the discriminative power.

- Kernelized Stein discrepancy (KSD)
  - Kernelized Stein discrepancy (KSD) assumes that the set is in the unit ball of RKHS:

$$\mathbb{D}(q,p) = \max_{\phi in \mathcal{F}} \{ \mathbb{E}_{\mathbf{x} \sim q} [ \operatorname{trace}(\mathcal{A}_p \phi(\mathbf{x})) ], \text{ s.t. } \|\phi\|_{\mathcal{H}^{d} \leq 1} \}.$$

• The optimal solution of KSD is:

$$\phi(x) = \phi_{q,p}^*(x) / \|\phi_{q,p}^*\|_{\mathcal{G}^d}$$

where

$$\phi_{p,q}^*(\cdot) = \mathbb{E}_{x \sim q}[\mathcal{A}_p k(x,\cdot)]$$

and

$$\mathbb{D}(q,p) = \|\phi_{q,p}^*\|_{\mathcal{H}^d}.$$

• If we rewrite the KSD, we have:

$$\mathbb{D}(q,p) = \left\| \int_{x \in \mathcal{X}} \left[ \nabla_x \log p(x) k(x,\cdot) + \nabla_x k(x,\cdot) \right] q(x) dx \right\|_{\mathcal{H}^d}$$

and since p(x) is considered through the score function  $\nabla_x \log p(x) = \nabla_x \log \bar{p}(x)$ , unnormalized posterior  $\bar{p}(x)$  can be used without loss of generality.

#### Variational inference using smooth transforms

- The goal of variational inference is to find a simpler distribution
   q\*(x) ∈ Q that minimizes the KL divergence between q(x) and the
   posterior p(x).
- This paper focusses on the sets Q obtained by smooth transforms from a tractable reference distribution z = T(x) where T: X → X is a smooth one-to-one transform and x is drawn from a tractable distribution q<sub>0</sub>(x).
- By the change of variables formula, the density of z is:

$$q_{\mathsf{T}}(z) = q(\mathsf{T}^{-1}(z))|\det(\nabla_z \mathsf{T}^{-1}(z))|.$$

• Now, our problem is to find **T** that maps  $\mathcal{X}$   $\mathcal{X}$  where  $\mathcal{X} \subset \mathbb{R}^d$  and we will perform **steepest descent** on **T** in RKHS.

- Stein operator as the derivative of KL divergence
  - We will consider  $T(x) = x + \epsilon \phi(x)$  where  $\phi(x) \in \mathbb{R}^d$  and  $x \in \mathcal{X}$  (note that x is a continuous random variable of our interest).
  - Now, we will connect the Stein operator and the derivative of KL divergence w.r.t. the perturbation magnitude ε:
  - Let  $T(x) = x + \epsilon \phi(x)$  and  $q_T(z)$  the density of z = T(x) when  $x \sim q(x)$ , we have

$$\nabla_{\epsilon} D_{\mathsf{KL}}(q_{\mathsf{T}} || p)_{\epsilon=0} = -\mathbb{E}_{x \sim q} [\mathsf{trace}(\mathcal{A}_p \phi(x))]$$

where  $\mathcal{A}_p\phi(x)$  is the Stein operator and the optimal direction  $\phi_{q,p}^*$  is

$$\phi_{q,p}^*(\cdot) = \mathbb{E}_{x \sim q}[\nabla_x \log p(x) k(x,\cdot) + \nabla_x k(x,\cdot)]].$$

 φ<sup>\*</sup><sub>q,p</sub>(·) tells us the optimal direction of the random variable of our interest x if we were to minimize the KL divergence between p and q<sub>T</sub>.

#### Stein Variational Gradient Descent

- To implement the iterative procedure, (unnormalized) posterior p(x) is given.
- Then, we draw a set of particles  $\{x_i^0\}_{i=1}^n$  for the initial distribution  $q_0$ .
- Each particle  $x_i$  is updated as follows:

$$x_i^{l+1} \leftarrow x_i^l + \epsilon_l \hat{\phi}^* \big( x_i^l \big)$$

where

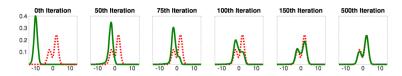
$$\hat{\phi}^*(x) = \frac{1}{n} \sum_{j=1}^n \left[ k(x_j^l, x) \nabla_{x_j^l} \log p(x_j^l) + \nabla_{x_j^l} k(x_j^l, x) \right].$$

The update rule has nice interpretations:

$$\hat{\phi}^*(x) = \frac{1}{n} \sum_{j=1}^n \left[ k(x_j^l, x) \nabla_{x_j^l} \log p(x_j^l) + \nabla_{x_j^l} k(x_j^l, x) \right].$$

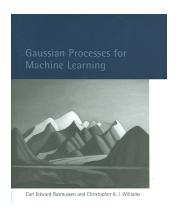
- 1.  $k(x_j^l, x)$ : Similarity between current particle to update x and j-th particle  $x_i^l$ .
- 2.  $\nabla_{x_j'} \log p(x_j')$ : Particle update direction of current particle to update x where it is computed from j-th particle  $x_j'$ .
  - Note that as we are using the score function, i.e., ∇<sub>x</sub> log p(x), unnormalized p(x) can be used!
  - This is also used in policy gradient methods such as REINFORCE.
- 3.  $\nabla_{x_j'}k(x_j',x)$ : This term can be interpreted as running a gradient ascent method on  $k(\cdot,\cdot)$ . As the value of a kernel function usually increases as the distance between two inputs decreases, it can be interpreted as an attractive force between particles.

• How it works: a toy example with 1D Gaussian mixture.





## **Backup slides**



Gaussian process for machine learning [2]

#### References i



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