## Estimating Marginal Returns to Education

Pedro Carneiro, James J. Heckman, and Edward J. Vytlacil

This paper estimates marginal returns to college for individuals induced to enroll in college by different marginal policy changes. The recent instrumental variables literature seeks to estimate this parameter, but in general it does so only under strong assumptions that are tested and found wanting. We show how to utilize economic theory and local instrumental variables estimators to estimate the effect of marginal policy changes. Our empirical analysis shows that returns are higher for individuals with values of unobservables that make them more likely to attend college. We contrast our estimates with IV estimates of the return to schooling.

TABLE 2—DEFINITIONS OF THE VARIABLES USED IN THE EMPIRICAL ANALYSIS

| Variable                             | Definition  |
|--------------------------------------|---|
| Y                                    | Log wage in 1991 (average of all nonmissing wages between 1989 and 1993)  |
| S = 1                                | If ever enrolled in college by 1991; zero otherwise   |
| X                                    | AFQT, a mother's education, number of siblings, average log earnings 1979–2000 in county of residence at 17, average unemployment 1979–2000 in state of residence at 17, urban residence at 14, cobort dummies, years of experience in 1991, average local log earnings in 1991, local unemployment in 1991           |
| $\mathbb{Z}\backslash\mathbb{X}^{b}$ | Presence of a college at age 14 (Card 1995; Stephen V. Cameron and Christopher Taber 2004), local earnings at 17 (Cameron and Heckman 1998; Cameron and Taber 2004), local unemployment at 17 (Cameron and Heckman 1998), local tuition in public four-year colleges at 17 (Thomas J. Kane and Cecilia E. Rouse 1995) |

TABLE 3—COLLEGE DECISION MODEL: AVERAGE MARGINAL DERIVATIVES

|   | Average derivative |
|---|--------------------|
| Controls (X)  |                    |
| Corrected AFQT                                      | 0.2826             |
|   | (0.0114)***        |
| Mother's years of schooling                         | 0.0441             |
|   | (0.0059)***        |
| Number of siblings                                  | -0.0233            |
|   | (0.0068)***        |
| Urban residence at 14                               | 0.0340             |
|   | (0.0274)           |
| "Permanent" local log earnings at 17                | 0.1820             |
|   | (0.0941)**         |
| "Permanent" state unemployment rate at 17           | 0.0058             |
|   | (0.0165)           |
| Instruments (Z)                                     |                    |
| Presence of a college at 14                         | 0.0529             |
|   | (0.0273)**         |
| Local log eamings at 17                             | -0.2687            |
|   | (0.1008)***        |
| Local unemployment rate at 17                       | 0.0149             |
| (in percent)  | (0.0100)           |
| Tuition in 4 year public colleges at 17             | -0.0027            |
| (in \$100)  | (0.0017)*          |
| Test for joint significance of instruments: p-value | 0.0001             |

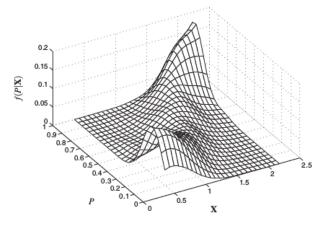


Figure 2. Support of P Conditional on  $\mathbf X$ 

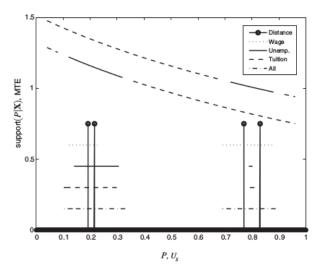
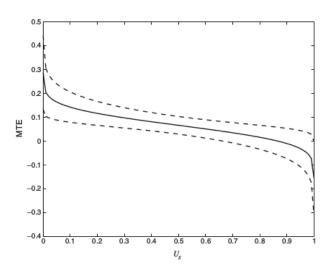


Figure 7. Support of P for Fixed X



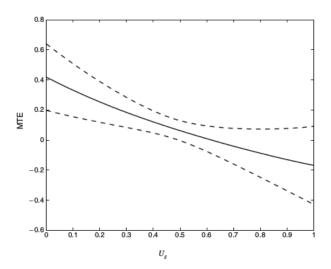


Figure 4.  $E(Y_1 - Y_0 | \mathbf{X}, U_S)$  with 90 Percent Confidence Interval—Locally Quadratic Regression Estimates

Table 4— Test of Linearity of  $E(Y|\mathbf{X},P=p)$  Using Polynomials in  $P_i$  and Test of Equality of LATEs Over Different Intervals  $(H_0:LATE^j(U_g^{L_j},U_g^{R_j})-LATE^{j+1}(U_s^{L_{j+1}},U_s^{R_{j+1}})=0)$ 

| Panel A. Test of linearity of $E(Y X,P=p)$ using models with different orders of polynomials in $P^a$ |                           |                                    |                                  |                                       |                 |              |  |  |  |
|---|---------------------------|------------------------------------|----------------------------------|---------------------------------------|-----------------|--------------|--|--|--|
| Panel A. Test of linearity of   | $E(Y \mathbf{X}, P = j$   | p) using model                     | s with differen                  | t orders of poly                      | momials in Pa   |              |  |  |  |
| Degree of polynomial  |                           |                                    |                                  |                                       |                 |              |  |  |  |
| for model   | 2                         | 3                                  | 4                                | 5                                     |                 |              |  |  |  |
| p-value of joint test of  | 0.035                     | 0.049                              | 0.086                            | 0.122                                 |                 |              |  |  |  |
| nonlinear terms   |                           | 0.0                                | 167                              |                                       |                 |              |  |  |  |
| Adjusted critical value   |                           |                                    | -                                |                                       |                 |              |  |  |  |
| Outcome of test   | Outcome of test Reject    |                                    |                                  |                                       |                 |              |  |  |  |
| Panel B. Test of equality of  | LATEs (H <sub>0</sub> : I | $ATE^{j}(U_{S}^{L_{j}},U_{S}^{H})$ | $\frac{1}{2}$ ) - $LATE^{j+1}$ ( | $U_{S}^{L_{j+1}}, U_{S}^{H_{j+1}}) =$ | 0) <sup>b</sup> |              |  |  |  |
| Ranges of Us for LATE   | (0,0.04)                  | (0.08, 0.12)                       | (0.16, 0.20)                     | (0.24, 0.28)                          | (0.32, 0.36)    | (0.40, 0.44) |  |  |  |
| Ranges of $U_S$ for LATE <sup><math>j+1</math></sup>  | (0.08, 0.12)              | (0.16, 0.20)                       | (0.24, 0.28)                     | (0.32, 0.36)                          | (0.40, 0.44)    | (0.48, 0.52) |  |  |  |
| Difference in LATEs   | 0.0689                    | 0.0629                             | 0.0577                           | 0.0531                                | 0.0492          | 0.0459       |  |  |  |
| p-value   | 0.0240                    | 0.0280                             | 0.0280                           | 0.0320                                | 0.0320          | 0.0520       |  |  |  |
| Ranges of Ug for LATE   | (0.48, 0.52)              | (0.56, 0.60)                       | (0.64, 0.68)                     | (0.72, 0.76)                          | (0.80, 0.84)    | (0.88, 0.92) |  |  |  |
| Ranges of $U_S$ for LATE <sup><math>j+1</math></sup>  | (0.56, 0.60)              | (0.64, 0.68)                       | (0.72, 0.76)                     | (0.80, 0.84)                          | (0.88, 0.92)    | (0.96, 1)    |  |  |  |
| Difference in LATEs   | 0.0431                    | 0.0408                             | 0.0385                           | 0.0364                                | 0.0339          | 0.0311       |  |  |  |
| p-value   | 0.0520                    | 0.0760                             | 0.0960                           | 0.1320                                | 0.1800          | 0.2400       |  |  |  |
| Joint p-value   |                           |                                    | 0.0                              | 520                                   |                 |              |  |  |  |

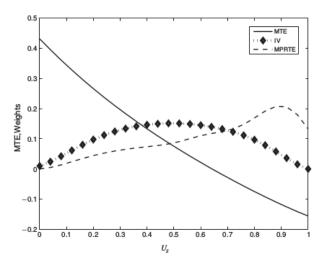


FIGURE 6. WEIGHTS FOR IV AND MPRTE

TABLE 5-RETURNS TO A YEAR OF COLLEGE

| Model                                 |  | Normal              | Semiparametric     |
|---------------------------------------|--|---------------------|--------------------|
| $ATE = E(\beta)$                      |  | 0.0670<br>(0.0378)  | Not identified     |
| $TT = E(\beta   S = 1)$               |  | 0.1433<br>(0.0346)  | Not identified     |
| $TUT = E(\beta   S = 0)$              |  | -0.0066<br>(0.0707) | Not identified     |
| MPR                                   | TE   |                     |                    |
| Policy perturbation                   | Metric   |                     |                    |
| $Z_{\alpha}^{k} = Z^{k} + \alpha$     | $ \mathbf{Z}\gamma - V  < e$   | 0.0662<br>(0.0373)  | 0.0802<br>(0.0424) |
| $P_{\alpha} = P + \alpha$             | P-U  <arepsilon< td=""><td>0.0637<br/>(0.0379)</td><td>0.0865<br/>(0.0455)</td></arepsilon<> | 0.0637<br>(0.0379)  | 0.0865<br>(0.0455) |
| $P_{\alpha} = (1 + \alpha)P$          | $ \frac{P}{U} - 1  < \epsilon$   | 0.0363<br>(0.0569)  | 0.0148<br>(0.0589) |
| Linear IV (Using $P(\mathbf{Z})$ as t |  | 0.0951<br>0.0386)   |                    |
| OLS                                   | -  | 0.0836              |                    |

## **Appendix**

## References

Carneiro, P., Heckman, J. J., & Vytlacil, E. J. (2011). Estimating marginal returns to education. *American Economic Review*, 101(6), 2754–2781.