

Augmented Inverse Propensity Weighting

(Doubly Robust Estimation)

From IP Weighting part, we saw how to unbiasedly estimate the average treatment effect through inverse probability weighting. Our goal is still to estimate the average causal treatment effect from observational data, adjusting for confounding. This time, we are introducing the Augmented inverse Propensity Weighting (AIPW) method to estimate the average treatment effect. This method is also known as doubly robust estimation. This is because AIPW estimator remains consistent for the average treatment effect if either the propensity score model or the outcome regression is misspecified but the other is properly specified.

Regression Model & Adjustment by Regression Modeling

Let Y is the outcome, A is the treatment, and L is the set of all measured covariates (confounders).

Regression of Y on A and L :

We can identify the regression $E[Y | A, L]$, as this depends on the observed data.

For continuous outcome, $E[Y | A, L] = \beta_0 + \beta_1 A + L^T \beta_2$

$E[Y | A = 1, L]$ is the regression among treated and $E[Y | A = 0, L]$ among control.

If we average over all possible values of L , we get the following result:

$$\begin{aligned} & E(E[Y | A = 1, L]) \\ &= E(E[Y^1 | A = 1, L]) \quad (\text{by consistency}) \\ &= E(E[Y^1 | L]) \quad (\text{by conditional exchangeability under no unmeasured confounders, } Y^a \perp\!\!\!\perp A | L) \\ &= E[Y^1] \end{aligned}$$

Similarly, $E(E[Y | A = 0, L]) = E[Y^0]$

$$\begin{aligned} \text{Thus, } \Delta &= E[Y^1] - E[Y^0] \\ &= E(E[Y^1 | A = 1, L]) - E(E[Y^1 | A = 0, L]) \\ &= \beta_0 + \beta_1 + L^T \beta_2 - \beta_0 + L^T \beta_2 \\ &= \beta_1 \end{aligned}$$

$$\therefore \hat{\Delta} = \hat{\beta}_1$$

Propensity Score Model & Adjustment by Inverse Weighting

Propensity score: $P(A = 1 \mid L) = E[I(A = 1) \mid L] = E[A \mid L]$

Under no unmeasured confounders, $Y^a \perp\!\!\!\perp A \mid E[A \mid L]$

Fitting a logistic regression for propensity score:

$$P(A = 1 \mid L) = \frac{\exp(\alpha_0 + L^T \alpha_1)}{1 + \exp(\alpha_0 + L^T \alpha_1)}$$

Rather than use the difference of simple averages $\bar{Y}^{(1)} - \bar{Y}^{(0)}$, estimate Δ by the difference of inverse propensity score weighted averages:

$$\hat{\Delta}_{IPW} = \frac{1}{n} \sum_{i=1}^n \frac{A_i Y_i}{\hat{P}(A=1|L_i)} - \frac{1}{n} \sum_{i=1}^n \frac{(1-A_i) Y_i}{1-\hat{P}(A=1|L_i)}$$

If $\hat{P}(A = 1 \mid L) = P(A = 1 \mid L)$, the true propensity score:

$$\begin{aligned} & E \left[\frac{AY}{P(A=1|L)} \right] \\ &= E \left[\frac{AY^1}{P(A=1|L)} \right] \quad (\text{by } AY = A(Y^1 A + Y^0(1-A)) = AY^1 \text{ for } A^2 = A) \\ &= E \left(E \left[\frac{AY^a}{P(A=1|L)} \mid Y^a, L \right] \right) \\ &= E \left(\frac{Y^1}{P(A=1|L)} E[A \mid Y^a, L] \right) \\ &= E \left(\frac{Y^1}{P(A=1|L)} E[A \mid L] \right) \quad (\text{under no unmeasured confounders}) \\ &= E \left(\frac{Y^1}{E[A|L]} E[A \mid L] \right) \\ &= E[Y^1] \end{aligned}$$

$$\text{Similarly, } E \left[\frac{(1-A)Y}{1-P(A=1|L)} \right] = E[Y^0]$$

$\therefore \hat{\Delta}_{IPW}$ estimates $\Delta (E[Y^1] - E[Y^0])$ when the propensity model is identical to the true propensity score.

Doubly Robust Estimator

$$\begin{aligned} \hat{\Delta}_{DR} &= \frac{1}{n} \sum_{i=1}^n \left[\frac{A_i Y_i}{\hat{P}(A=1|L_i)} - \frac{A_i - \hat{P}(A=1|L_i)}{\hat{P}(A=1|L_i)} \hat{E}(Y_i \mid A_i = 1, L_i) \right] - \frac{1}{n} \sum_{i=1}^n \left[\frac{(1-A_i) Y_i}{1-\hat{P}(A=1|L_i)} + \frac{A_i - \hat{P}(A=1|L_i)}{1-\hat{P}(A=1|L_i)} \hat{E}(Y_i \mid A_i = 0, L_i) \right] \\ &= \hat{\mu}_{1,DR} - \hat{\mu}_{0,DR} \end{aligned}$$

What does $\hat{\mu}_{1,DR}$ estimate?

$$\begin{aligned} &= E \left[\frac{AY}{P(A=1|L)} - \frac{A-P(A=1|L)}{P(A=1|L)} E(Y \mid A = 1, L) \right] \\ &= E \left[\frac{AY^1}{P(A=1|L)} - \frac{A-P(A=1|L)}{P(A=1|L)} E(Y \mid A = 1, L) \right] \\ &= E \left[Y^1 + \frac{A-P(A=1|L)}{P(A=1|L)} \{Y^1 - E(Y \mid A = 1, L)\} \right] \quad (\text{by algebra}) \\ &= E[Y^1] + E \left[\frac{A-P(A=1|L)}{P(A=1|L)} \{Y^1 - E(Y \mid A = 1, L)\} \right] \end{aligned}$$

Does $\hat{\Delta}_{DR}$ estimate Δ ?

$\hat{\mu}_{1,DR}$ estimates

$$E[Y^1] + E \left[\frac{A-P(A=1|L)}{P(A=1|L)} \{Y^1 - E(Y | A = 1, L)\} \right] \quad (*)$$

In order for $\hat{\mu}_{1,DR}$ to estimate $E[Y_1]$, the second term in $(*)$ must be 0.

It is 0 when either **postulated propensity score model** or **postulated regression model** is correct.

Scenario 1: Propensity score model $P(A = 1 | L)$ is correct

$$\begin{aligned} & E \left[\frac{A-P(A=1|L)}{P(A=1|L)} \{Y^1 - E(Y | A = 1, L)\} \right] \\ &= E \left(E \left[\frac{A-P(A=1|L)}{P(A=1|L)} \{Y^1 - E(Y | A = 1, L)\} \mid Y^1, L \right] \right) \\ &= E \left(\{Y^1 - E(Y | A = 1, L)\} E \left[\frac{A-P(A=1|L)}{P(A=1|L)} \mid Y^1, L \right] \right) \\ &= E \left(\{Y^1 - E(Y | A = 1, L)\} E \left[\frac{A-E(A|L)}{E(A|L)} \mid Y^1, L \right] \right) \\ &= E \left(\{Y^1 - E(Y | A = 1, L)\} E \left[\frac{E[A|Y^1, L] - E(A|L)}{E(A|L)} \right] \right) \\ &= E \left(\{Y^1 - E(Y | A = 1, L)\} E \left[\frac{E[A|L] - E(A|L)}{E(A|L)} \right] \right) \quad (\text{under no unmeasured confounders}) \\ &= 0 \end{aligned}$$

$\therefore \hat{\mu}_{1,DR}$ estimates $E[Y^1]$ and similarly $\hat{\mu}_{0,DR}$ estimates $E[Y^0]$ when propensity score model is correct

Scenario 2: Regression model $E(Y | A = 1, L)$ is correct

$$\begin{aligned} & E \left[\frac{A-P(A=1|L)}{P(A=1|L)} \{Y^1 - E(Y | A = 1, L)\} \right] \\ &= E \left(E \left[\frac{A-P(A=1|L)}{P(A=1|L)} \{Y^1 - E(Y | A = 1, L)\} \mid A, L \right] \right) \\ &= E \left(\frac{A-P(A=1|L)}{P(A=1|L)} E \left[\{Y^1 - E(Y | A = 1, L)\} \mid A, L \right] \right) \\ &= E \left(\frac{A-P(A=1|L)}{P(A=1|L)} E \left[\{Y^1 - E(Y | A = 1, L)\} \mid A, L \right] \right) \\ &= E \left(\frac{A-P(A=1|L)}{P(A=1|L)} \{E[Y^1 | A, L] - E(Y | A = 1, L)\} \right) \\ &= E \left(\frac{A-P(A=1|L)}{P(A=1|L)} \{E[Y^1 | A, L] - E(Y^1 | A = 1, L)\} \right) \\ &= E \left(\frac{A-P(A=1|L)}{P(A=1|L)} \{E[Y^1 | L] - E(Y^1 | L)\} \right) \quad (\text{under no unmeasured confounders}) \\ &= 0 \end{aligned}$$

$\therefore \hat{\mu}_{1,DR}$ estimates $E[Y^1]$ and similarly $\hat{\mu}_{0,DR}$ estimates $E[Y^0]$ when regression model is correct

Therefore, $\hat{\Delta}_{DR}$ estimates Δ !

Generating data with true ATE = 0.5

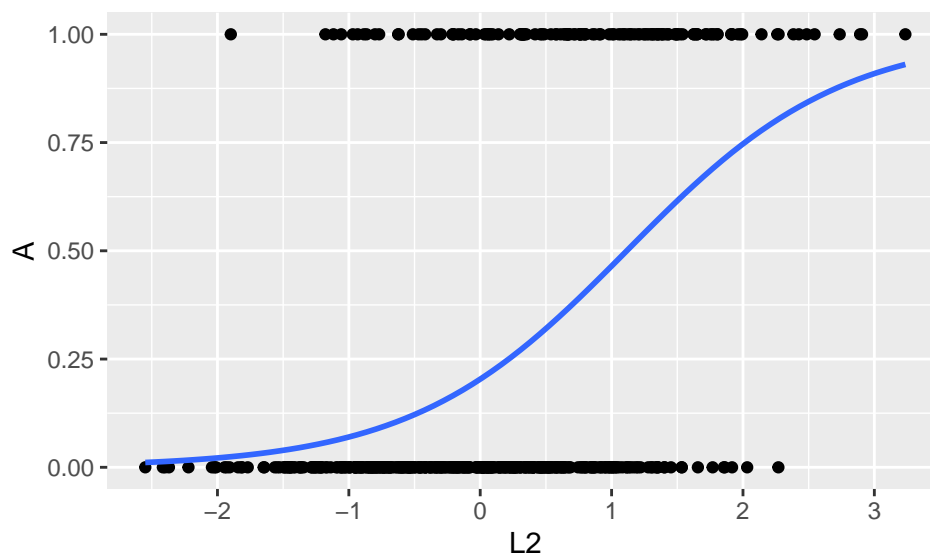
```
pre_data <- defData(varname = "L1", formula = "0", variance = 1,
  dist = "normal")
pre_data <- defData(pre_data, varname = "L2", formula = "0", variance = 1,
  dist = "normal")
pre_data <- defData(pre_data, varname = "L3", formula = "0;1",
  dist = "uniform")
pre_data <- defData(pre_data, varname = "A", formula = "-1.3 + 1.2*L2",
  dist = "binary", link = "logit")
pre_data <- defData(pre_data, varname = "Y", formula = "L1 + 0.5*L2 + 0.5*A",
  variance = 0.5, dist = "normal")
set.seed(7777)
df <- genData(500, pre_data)
```

Propensity Model

```
logistic_model <- glm(A ~ L2, family = binomial, data = df)
logistic_model %>% broom::tidy()
```

```
## # A tibble: 2 x 5
##   term      estimate std.error statistic  p.value
##   <chr>      <dbl>    <dbl>    <dbl>   <dbl>
## 1 (Intercept) -1.36    0.132   -10.4 3.37e-25
## 2 L2          1.22    0.138    8.87 7.06e-19
```

```
ggplot(df, aes(L2, A)) + geom_point() +
  geom_smooth(method = "glm", se = FALSE, method.args = list(family = "binomial"))
```



Regression Model

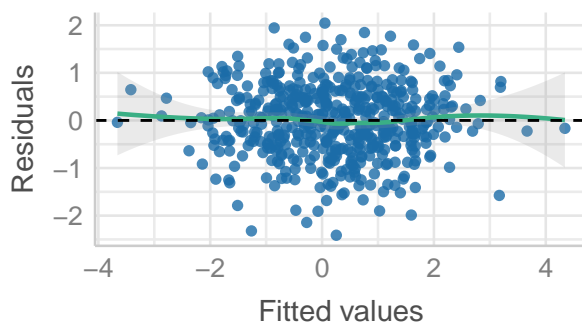
```
linear_model <- lm(Y ~ L1 + L2, data = df)
linear_model %>% broom::tidy()
```

```
## # A tibble: 3 x 5
##   term      estimate std.error statistic  p.value
##   <chr>      <dbl>    <dbl>    <dbl>    <dbl>
## 1 (Intercept)  0.0867    0.0343      2.52 1.19e- 2
## 2 L1          1.04     0.0335     31.0 3.44e-118
## 3 L2          0.572    0.0338     16.9 4.71e- 51
```

```
check_model(linear_model,
             check = c("linearity", "qq", "normality", "outliers"))
```

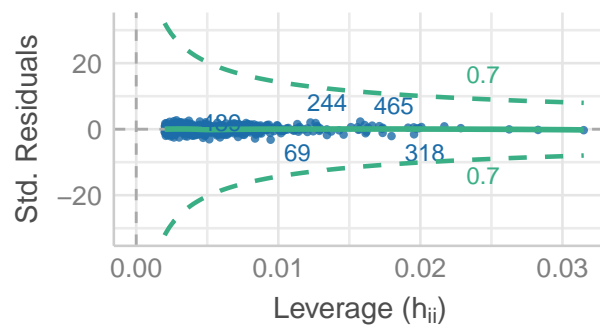
Linearity

Reference line should be flat and horizontal



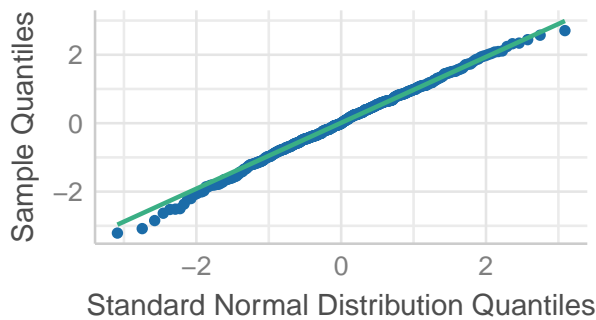
Influential Observations

Points should be inside the contour lines



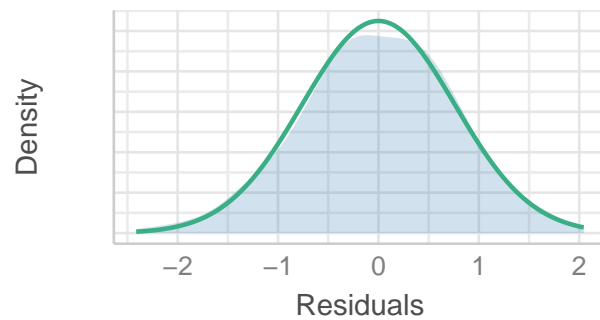
Normality of Residuals

Dots should fall along the line



Normality of Residuals

Distribution should be close to the normal curve



Case when regression model is correct, but propensity model is incorrect

```
ATE.out <- estimate.ATE(pscore.formula = A ~ s(L1),
  pscore.family = binomial,
  outcome.formula.t = Y ~ s(L1) + s(L2),
  outcome.formula.c = Y ~ s(L1) + s(L2),
  outcome.family = gaussian,
  treatment.var = "A", data = df,
  divby0.action = "t", divby0.tol = 0.001,
  var.gam.plot = FALSE, nboot = 50)

print(ATE.out)
```

```
## #####
## AIPW Estimator:
## -----
## Estimated ATE:  0.5043
##               SE          z-statistic      Pr(>|z|)
## Emp. Sandwich  0.0727          6.9339        < 0.0001
## Estim. Asymptotic 0.074          6.8163        < 0.0001
## Bootstrap      0.0978          5.1558        < 0.0001
## #####
## #####
## IPW Estimator:
## -----
## Estimated ATE:  0.9472
##               SE          z-statistic      Pr(>|z|)
## Estim. Asymptotic 0.0766          12.3687        < 0.0001
## Bootstrap      0.0872          10.8584        < 0.0001
## #####
## #####
## Regression Estimator:
## -----
## Estimated ATE:  0.5052
##               SE          z-statistic      Pr(>|z|)
## Estim. Asymptotic 0.074          6.8288        < 0.0001
## Bootstrap      0.0969          5.2143        < 0.0001
## #####
## #####
## General Information
## -----
## Control Value:  A = 0
## Treated Value:  A = 1
## Number of Discarded Units:  0
## Number of Truncated Propensity Scores:  0
## Number of Treated Units Before Discards/Truncations:  139
## Number of Treated Units not Discarded/Truncated:  139
## Number of Control Units Before Discards/Truncations:  361
## Number of Control Units not Discarded/Truncated:  361
## #####
```

Case when propensity model is correct, but regression model is incorrect

```
ATE.out_2 <- estimate.ATE(pscore.formula = A ~ s(L2),
                          pscore.family = binomial,
                          outcome.formula.t = Y ~ s(L1) + s(L3),
                          outcome.formula.c = Y ~ s(L1) + s(L3),
                          outcome.family = gaussian,
                          treatment.var = "A", data = df,
                          divby0.action = "t", divby0.tol = 0.001,
                          var.gam.plot = FALSE, nboot = 50)

print(ATE.out_2)
```

```
## #####
## AIPW Estimator:
## -----
## Estimated ATE: 0.5465
##              SE          z-statistic      Pr(>|z|)
## Emp. Sandwich 0.1184          4.6151      < 0.0001
## Estim. Asymptotic 0.1091          5.0095      < 0.0001
## Bootstrap      0.0642          8.5112      < 0.0001
## #####
## #####
## IPW Estimator:
## -----
## Estimated ATE: 0.2668
##              SE          z-statistic      Pr(>|z|)
## Estim. Asymptotic 0.1118          2.3859      0.017
## Bootstrap      0.1661          1.6067      0.1081
## #####
## #####
## Regression Estimator:
## -----
## Estimated ATE: 0.9469
##              SE          z-statistic      Pr(>|z|)
## Estim. Asymptotic 0.1076          8.7994      < 0.0001
## Bootstrap      0.0746          12.6894      < 0.0001
## #####
## #####
## General Information
## -----
## Control Value: A = 0
## Treated Value: A = 1
## Number of Discarded Units: 0
## Number of Truncated Propensity Scores: 0
## Number of Treated Units Before Discards/Truncations: 139
## Number of Treated Units not Discarded/Truncated: 139
## Number of Control Units Before Discards/Truncations: 361
## Number of Control Units not Discarded/Truncated: 361
## #####
```

Reference

Davidian, M. (n.d.). NC State Department of Statistics. Retrieved January 1, 2022, from <https://www4.stat.ncsu.edu/~davidian/double.pdf>