Nonparametric dobuly robust estimation of continuous treatment

This work is to verify the proof of theorems introduced in the paper (Kennedy et al, 2016).

Guide to notation

$Z = (L, A, Y) =$ observed data arising from distribution P with density $p(z) = p(y \mid l, a)p(a \mid l)p(l)$ and support $\text{supp}(Z) = \mathcal{Z} = \mathcal{L} \times \mathcal{A} \times \mathcal{Y}$

$\mathbb{P}_n = \frac{1}{n} \sum_i \delta_{z_i} =$ empirical measure so that $\mathbb{P}_n f(Z) = \frac{1}{n} \sum_i f(z_i)$

$\mathbb{P}(\mho) = \mathbb{P}\{\mho(\mathbb{Z})\} = \int_{\mathcal{Z}} f(z)dP(z) =$ expectation for new Z treating f as fixed (so $\mathbb{P}(\hat{\mho})$ is random if $\hat{f}$ depends on sample, in which case $\mathbb{P}(\hat{\mho}) \neq E(\hat{f})$)

$\pi(a \mid l) = p(a \mid l) = \frac{\partial}{\partial a} P(A \leq a \mid l) =$ conditional density of treatment A

$\hat{\pi}(a \mid l) =$ user-specified estimator of $\pi(a \mid l)$, which converges to limit $\overline{\pi}(a \mid l)$ that may not equal true π

$\omega(a) = p(a) = \frac{\partial}{\partial a} P(A \leq a) = E[\pi(a \mid L)] = \int_{\mathcal{L}} \pi(a \mid l)dP(l) =$ density of A

$\hat{\omega(a)} = \mathbb{P}_{\ltimes}\{\hat{\pi}(a \mid l)\} = \int_{\mathcal{L}} \pi(a \mid l)dP(l) =$ estimator of ω, which converges to limit $\overline{\omega}(a)$ that may not equal true ω

$\mu(l \mid a) = E(Y \mid L = l, A = a) = \int_{\mathcal{Y}} ydP(y \mid l, a) =$ conditional mean outcome

$\hat{\mu}(l \mid a) =$ user-specified estimator of $\mu(l \mid a)$, which converges to limit $\overline{\mu}(l, a)$ that may not equal true μ

$\hat{m}(a) = \mathbb{P}_{\ltimes}\{\hat{\mu}(L, a)\} = \int_{\mathcal{L}} \hat{\mu}(l, a)d\mathbb{P}_{\ltimes}(l) = \frac{1}{n} \sum_i \hat{\mu}(l_i, a)$

$\psi = \int_{\mathcal{A}} \int_{\mathcal{L}} \mu(l, a)\omega(a)dP(l)da$

Theorem 1.

Under a nonparametric model, the efficient influence function for ψ is
$\xi(Z; \pi, u) - \psi + \int_{\mathcal{A}}\{\mu(L, a) - \int_{\mathcal{L}} \mu(l, a)dP(l)\}\omega(a)da$ where

$$\xi(Z; \pi, u) = \frac{Y - \mu(L, A)}{\pi(A \mid L)} \int_{\mathcal{L}} \pi(A \mid l)dP(l) + \int_{\mathcal{L}} \mu(l, a)dP(l)$$

Importantly, the function $\xi(Z; \pi, u)$ satisfies its desired double robustness property, i.e., that $E[\xi(Z; \pi, u) \mid A = a] = \theta(a)$ if either $\overline{\pi} = \pi$ or $\overline{\mu} = \mu$ where $\theta(a) = E(Y^a)$

This motivates estimating the effect curve $\theta(a)$ by estimating the nuisance functions (π, μ), and then regressing the estimated psuedo-outcome

$$\hat{\xi}(Z; \hat{\pi}, \hat{\mu}) = \frac{Y - \hat{\mu}(L, A)}{\hat{\pi}(A, L)} \int_{\mathcal{L}} \hat{\pi}(A, L)d\mathbb{P}_n(l) + \int_{\mathcal{L}} \hat{\mu}(l, a)d\mathbb{P}_n(l)$$

on treatment A using off-the-shelf nonparametric regression or machine learning methods.

Proof of Theorem 1

By definition the efficient influence function for ψ is the unique function $\phi(Z)$ that satisfies $\psi'_\epsilon(0) = E[\phi(Z)\ell'_\epsilon(Z;0)]$, where $\psi(\epsilon)$ represents the parameter of interest as a functional on the parametric submodel and $\ell(w \mid \overline{w}; \epsilon) = log\ p(w \mid \overline{w}; \epsilon)$ for any partition $(W, \overline{W}) \subseteq Z$. Therefore

$$\ell'_\epsilon(z; \epsilon) = \ell'_\epsilon(y \mid l, a; \epsilon) + \ell'_\epsilon(a \mid l; \epsilon) + \ell'_\epsilon(l; \epsilon)$$

The authors give two important properties of such score functions of such score functions $\ell'_\epsilon(w \mid \overline{w}; \epsilon)$ that will be used throughout this proof. First note that since $\ell_\epsilon(w \mid \overline{w}; \epsilon)$ is a log transformation of $p(w \mid \overline{w}; \epsilon)$, it follows that $\ell'_\epsilon(w \mid \overline{w}; \epsilon) = p'(w \mid \overline{w}; \epsilon)/p(w \mid \overline{w}; \epsilon)/$. Similarly, as with any score function, note that $E[\ell'_\epsilon(W \mid \overline{W}; \epsilon) \mid \overline{W}] = 0$ since

$$\int_\mathcal{W} \ell'_\epsilon(w \mid \overline{w}; 0)dP(w \mid \overline{w}) = \int_\mathcal{W} dP'(w \mid \overline{w}) = \tfrac{\partial}{\partial \epsilon} \int_\mathcal{W} P(w \mid \overline{w}) = 0$$

The goal in this proof is to show that $\psi'_\epsilon(0) = E[\phi(Z)]\ell'_\epsilon(Z;0)$ for the proposed influence function $\phi(Z) = \xi(Z; \pi, u) - \psi + \int_\mathcal{A}\{\mu(L, a) - \int_\mathcal{L}\{\mu(l, a)dP(l)\}\omega(a)da$. First we will give an expression for $\psi'_\epsilon(0)$. By definition $\psi(\epsilon) = \int_\mathcal{A} \theta(a; \epsilon)\omega(a; \epsilon)da$ because

$$\int_\mathcal{L} \mu(l, a)dP(l) = E[\mu(l, a)] = E[E(Y \mid L = l, A = a)] = E[Y^a].$$

Hence, $\psi'_\epsilon(0) = \int_\mathcal{A}\{\theta'(a; 0)\omega(a) + \theta(a)\omega'(a; 0)\}da = E[\theta'_\epsilon(A; 0) + \theta(A)\ell'_\epsilon(A; 0)]$ because

$$\omega(a) = p(a) \text{ and } \ell'_\epsilon(A; 0) = \tfrac{P'_\epsilon(A;0)}{P(A;0)} \iff \omega'_\epsilon(A; 0) = \ell'_\epsilon(A; 0)P(A; 0)$$

Also since $\theta(a; \epsilon) = \int_\mathcal{L} \int_\mathcal{Y} y\ p(y \mid l, a; \epsilon)p(l; \epsilon)\ d\eta(y)\ dv(l)$, we have

$\theta'(a; 0) = \int_\mathcal{L} \int_\mathcal{Y} y\ \{p'_\epsilon(y \mid l, a; 0)p(l) + p(y \mid l, a)p'(l; 0)\}\ d\eta(y)\ dv(l)$

$\quad = \theta'(a; 0) = \int_\mathcal{L} \int_\mathcal{Y} y\ \{\ell'_\epsilon(y \mid l, a; 0)p(y \mid l, a)p(l) + p(y \mid l, a)\ell'_\epsilon(l; 0)p(l)\}\ d\eta(y)\ dv(l)$

$\quad = E\big[E\{Y\ell'_\epsilon(Y \mid L, A; 0) \mid L, A = a\}\big] + E\big[\mu(L, a)\ell'_\epsilon(L; 0)\big]$

Therefore

$$\psi'_\epsilon(0) = \int_\mathcal{A} \Big(E\big[E\{Y\ell'_\epsilon(Y \mid L, A; 0) \mid L, A = a\}\big] + E\big[\mu(L, a)\ell'_\epsilon(L; 0)\big] + \theta(a)\ell'_\epsilon(a; 0)\Big)\omega(a)da$$

Now we will consider the covariance

$E[\phi(Z)\ell'_\epsilon(Z; 0)] = E[\phi(Z)\{\ell'_\epsilon(Y \mid L, A; 0) + \ell'_\epsilon(A, L; 0)\}]$ because

$$\ell'_\epsilon(z; \epsilon) = \ell'_\epsilon(y \mid l, a; \epsilon) + \ell'_\epsilon(a \mid l; \epsilon) + \ell'_\epsilon(l; \epsilon) = \ell'_\epsilon(y \mid l, a; 0) + \ell'_\epsilon(a, l; 0)$$

which we need to show equals the earlier expression for $\psi'_\epsilon(0)$.

The proposed efficient influence function given in the theorem 1 is

$$\xi(Z; \pi, u) = \tfrac{Y - \mu(L, A)}{\pi(A \mid L)} \int_\mathcal{L} \pi(A \mid l)dP(l) + \int_\mathcal{L} \mu(l, a)dP(l)$$

where we define $m(a) = \int_\mathcal{L} \mu(l, a)dP(l)$ as the marginalized version of the regression function μ, so that $m(a) = \theta(a)$ if μ is the true regression function.

$$m(a) = \int_\mathcal{L} \mu(l, a)dP(l) = \int_\mathcal{L} \int_\mathcal{Y} y\ dP(y \mid l, a)dP(l) = \int_\mathcal{L} \int_\mathcal{Y} y\ p(y \mid l, a)p(l)\ d\eta(y)\ dv(l) = \theta(a)$$

Thus, $E\big[\phi(Z)\ell'_\epsilon(Y\mid L,A;0)\big]$

$= E\Big(\Big[\frac{Y-\mu(L,A)}{\pi(A\mid L)/\omega(A)} + \int_{\mathcal{A}}\{\mu(L,a-\theta(a)\}\omega(a)\ da + \theta(A) - \psi\Big]\ell'_\epsilon(Y\mid L,A;0)\Big)$

$= E\Big[\frac{Y\ell'_\epsilon(Y\mid L,A;0)}{\pi(A\mid L)/\omega(A)}\Big]$

$= E\Big[\frac{E[Y\ell'_\epsilon(Y\mid L,A;0)\mid L,A]}{\pi(A\mid L)/\omega(A)}\Big]$

$= \int_{\mathcal{A}} E\big[E\{Y\ell'_\epsilon(Y\mid L,A;0)\mid L,A=a\}\big]\omega(a)\ da$

where the first equality follows since

$$\phi(Z) = \frac{Y-\mu(L,A)}{\pi(A\mid L)}\int_{\mathcal{L}}\pi(A\mid l)dP(l) + \int_{\mathcal{L}}\mu(l,a)dP(l) - \psi + \int_{\mathcal{A}}\{\mu(L,a)-\int_{\mathcal{L}}\mu(l,a)dP(l)\}\omega(a)da$$

$$= \frac{Y-\mu(L,A)}{\pi(A\mid L)/\omega(A)} + \theta(A) - \psi + \int_{\mathcal{A}}\{\mu(L,a)-\theta(A)\}\omega(a)da$$

, the second since $E\big[\ell'_\epsilon(Y\mid L,A;0)\mid L,A\big]=0$, the third by iterated expectation conditioning on L and A, the last by iterated expectation conditioning on L gives:

$E\Big(E\Big[\frac{E[Y\ell'_\epsilon(Y\mid L,A;0)\mid L,A]}{\pi(A\mid L)/\omega(A)}\mid L\Big]\Big)$

$= E\Big(E\Big[\frac{E[Y\ell'_\epsilon(Y\mid L,A;0)\mid L,A]}{\pi(A\mid L)/E[\pi(A\mid L)]}\mid L\Big]\Big)$

$= E\Big[\frac{E[Y\ell'_\epsilon(Y\mid L,A;0)\mid L,A]}{\pi(A\mid L)}E[\pi(A\mid L)]\Big]$