

EMPIRICAL BANKING AND FINANCE: TUTORIAL № 4

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Set WD, Install Packages and Load Data

```
1 cd "Z:\Downloads"
2
3 ssc install estout, replace
4 ssc install reghdfe, replace
5 ssc install ftools, replace
6
7 use dataEmpBF_Tutorial4.dta
```

1 Data & Descriptives

1.1 Create descriptive statistics and comment briefly.

```
1 estpost summarize // should delete observations with NAs
2 esttab using sumr.tex, cells("count mean sd min max") noobs
```

Table 1: Summary Statistics for all Variables

	count	mean	sd	min	max
CountryCode	785	445.0994	250.5633	36	842
c	0
year	785	1996.469	11.55728	1964	2012
logGDP	785	28.32083	2.400927	24.64919	35.50195
logGDP_future	695	8.401251	6.564033	-18.8221	32.32899
private_credit	785	128.2249	54.83109	17.02012	339.2
household_credit	785	49.92093	28.04941	.1457143	144.7189
firm_credit	785	78.80823	33.25036	7.479578	227.4264
private_credit_past	785	8.677822	16.09333	-60.47037	114.603
household_credit_past	785	4.60116	6.189944	-11.19909	34.95463
firm_credit_past	785	4.016322	12.37292	-48.23518	86.50771

- CountryCode/c: We see that the Country has no missing values. Looking on other statistics makes no sense since the variable is just an arbitrary ID.
- Year: We have no missing values for any observations and observations are between the years 1964 and 2012.
- LogGDP: The log for the real GDP in constant prices has also no missing observations. We see values between 24.6 and 35.5, the interpretation is not really possible since the values are measured in local currency units.
- LogGDP_future: There are 90 missing values. For some countries the change in GDP from t to t+3 is negative and for other it is positive.
- Private_Credit: Private non-financial debt to GDP varies a lot, the minimum value is 17 while the maximum value is 339.
- household_credit: Household debt to GDP also varies a lot, from nearly zero up to 144.
- firm_credit: Non-financial firm debt to GDP also varies a lot, from 7 to 227.
- private_credit_past¹: The change in the ratio of private credit to GDP from t-4 to t-1 is in mean about 8 and has a range from -60 to 114.
- household_credit_past: The change in the ratio of private household to GDP from t-1 to t-4 is in mean about 4 and has a range from -11 to 35.
- Firm_Credit_Past: The change in the ratio of firm credit to GDP from t-1 to t-4 is in mean about 4 and has a range from -48 to 87.

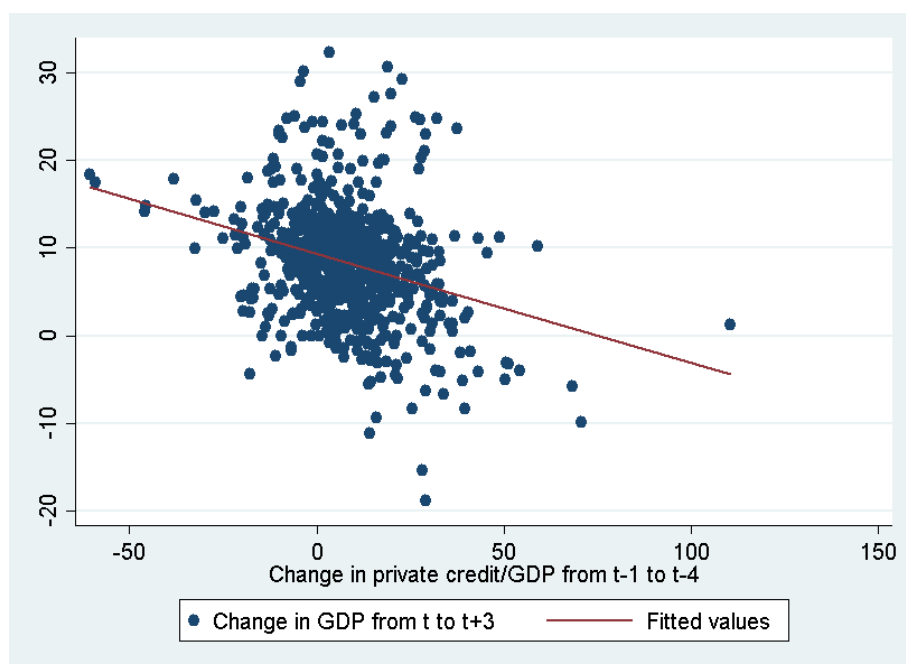
1.2 Correlation of past change in private credit and future change in GDP

```

1 twoway scatter logGDP_future private_credit_past ///
2 || lfit logGDP_future private_credit_past
3 corr logGDP_future private_credit_past

```

¹according to the label of variable in the dataset, the change is measured from t-1 to t-4. But since it's counterintuitive and also different in the Milan (2017)



Correlation coefficient = -0.2896

- `logGDP_future`: For each point in time the difference of the GDP three years later and the GDP of that year is calculated. The real GDP is used.
- `private_credit_past`: For each point in time the difference of the private credit to GDP of the previous year and the same measure from four years earlier is calculated.

We see a negative correlation of the two variables, implying a negative relation between the change in private credit to GDP and subsequent output growth. The relation is not perfect ($|r| \neq 1$) and medium strong.

A theoretical explanation could be that banks are most willing to lend to households when unemployment is low and the ratings of their customers are good and vice versa, when we assume the demand function does not change over time this implies a pro-cyclical amount of debt over time. When the business cycle has a length of around six years this theory fits to the data and is also consistent with the longrun findings (trend) in the papers of the previous Problem Sets.

1.3 How does the result fit to results we have seen in previous Problems

Porta et al (1998) and Levine et al (2000) both find a positive relation between private credit to GDP and subsequent growth. The findings are not necessary inconsistent with those of Milan (2017) since Milan (2017) considers the short run while the other two papers look on a longer time horizon. So it might be the case that the relation is pro-cyclical and the trend of growth is positively influenced by private credit at the same time.

2 Regression set 1

2.1 Regression of future LogGDP on past private Credit with country fixed effects

```

1 reghdfe logGDP_future private_credit_past, a(c)
2 est sto m1
3 esttab m1 using reg1.tex, se obslast scalar(F) r2 ///
4 title("Regression 1") replace

```

We observe a negative relation, when `private_credit_past` increases by one standard deviation in average `logGDP_future` is 0.2918 standard deviations lower:

Table 2: Regression 1

Model	$y_{i,t} = \alpha_i + \beta x_{i,t} + \varepsilon_{i,t}$
	logGDP_future
private_credit_past	-0.119*** (0.0149)
_cons	9.312*** (0.230)
R^2	0.387
F	63.20
N	695

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

2.2 Regression 1 with reg command

```

1 reg logGDP_future private_credit_past i.CountryCode, robust
2 est sto m2
3 esttab m2 using reg2.tex, se obslast scalar(F) r2 ///
4 title("Regression 2") replace

```

The results for the coefficient as well as for the R^2 are similar as expected, due to the fact that the methods of fixed effects and first differences are similar in terms of the coefficients of interest:

Table 3: Regression 2

Model	$y_{i,t} = \alpha_i + \beta x_{i,t} + \varepsilon_{i,t}$
	logGDP_future
private_credit_past	-0.119*** (0.0163)
_cons	11.25*** (0.472)
R^2	0.387
F	34.08
N	695

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

2.3 Regression 1 using demeaned variables

```

1 by c, sort : summarize logGDP_future private_credit_past
2 foreach i in logGDP_future private_credit_past{
3   egen mean_`i`=mean(`i'), by(c)
4   ge demeaned_`i`= `i' - mean_`i'
5 }
6 reg demeaned_logGDP_future demeaned_private_credit_past, robust
7 est sto m3
8 esttab m3 using reg3.tex, se aic obslast scalar(F) bic r2 ///
9 title("Regression 3") replace

```

Again we observe the same² coefficient, because the methods are same³ here. The R^2 differs as we see:

²the slight difference is due to computation errors/rounding in the demeaning process

³in terms of the estimation of the coefficient of the independent variable

Table 4: Regression 3

Model	$y_{i,t} - \bar{y}_i = \beta(x_{i,t} - \bar{x}_i) + \varepsilon_{i,t}$
	demeaned_logGDP_future
demeaned_private_credit_past	-0.117*** (0.0156)
_cons	-0.0891 (0.197)
R^2	0.086
F	56.82
N	695

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

2.4 Which R^2 should we report

We should report the R^2 from the demeaned ols regression, because it gives us the most adequate measure of the goodness of fit imposed by the variable of interest. In the other specifications the R^2 is much higher and dependent on fixed effects variables due to the explanatory power⁴ of them.

2.5 Assumption allowing causal interpretation

We have to assume that there is no omitted variable⁵, which is not constant overtime for each given country. Then we would have the normal OVB problem, if the omitted variable is constant overtime for each given country we would already control for it by the country fixed effects.

2.6 Fixed effects for countries legal origin, a good idea?

When clustering over country the clustering over legal origin is already controlled for - the legal origin of a country is constant over time.

⁴not necessary casual power, but residual variance is reduced

⁵which is correlated with the residual and the independent variable

3 Regression set 2

3.1 Regression of logGDP_future on private_credit_past with country fixed effects

- (1) Using the default standard errors
- (2) Using standard errors robust to heteroskedasticity
- (3) Clustering standard errors at the country level
- (4) Clustering standard errors at the country level and at the year level

```
1 reghdfe logGDP_future private_credit_past, a(c) //1
2 est sto d1
3 reghdfe logGDP_future private_credit_past, a(c) vce(r) //2
4 est sto d2
5 reghdfe logGDP_future private_credit_past, a(c) vce(cluster c) //3
6 est sto d3
7 reghdfe logGDP_future private_credit_past, a(c) vce(cluster c year) //4
8 est sto d4
9 esttab d1 d2 d3 d4 using reg4.tex, se obslast scalar(F) r2 ///
10 title("Regression 4") replace
```

Table 5: Regression 4

	(1)	(2)	(3)	(4)
	logGDP_future	logGDP_future	logGDP_future	logGDP_future
private_credit_past	-0.119*** (0.0149)	-0.119*** (0.0163)	-0.119*** (0.0297)	-0.119*** (0.0319)
_cons	9.312*** (0.230)	9.312*** (0.219)	9.312*** (0.228)	9.312*** (0.526)
R^2	0.387	0.387	0.387	0.387
F	63.20	52.99	15.96	13.85
N	695	695	695	695

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

3.2 Compare the standard errors across the four types of adjustments and comment

We see that the standard error assuming homoscedasticity is lowest and the robust standard error is higher when we cluster over countries. Clustering additionally over years increases the standard error even more.

As we already know from the very beginning of the lecture robust standard errors are always higher than those calculated assuming homoscedasticity. We also know that due to the additional covariance terms usually clustered standard errors are even higher than simple robust standard errors. The clustered standard errors for country and year are higher than those clustered only over countries, since the number of observations available for the estimation of each element of the covariance matrix is lower with more clusters.

3.3 What are the “theoretical” arguments in favor of and against clustering standard errors at the country-level? What could be the reason for the additional clustering at the year level?

- Main argument for clustering: If we do not cluster standard errors / assume that observations are iid while this is not the case we underestimate the standard errors by a large extent and maybe wrongly report significance of our coefficients.
- (Argument) against clustering: If we cluster standard errors while the iid assumption holds we overestimate the standard errors, but this is hard to argue.
- Argument for clustering over Years: Clustering over years makes much sense since its commonly known and empirical proven that there is often autocorrelation to a large extent for GDP variables.

3.4 Does the coefficient change?

No, when we look on their calculation than we see that the covariance matrix is not part of the equation.

3.5 Clustering standard errors at the country-year level, a good idea?

The result would be the same covariance matrix as with normal robust estimation. So since we want to check for the covariances and not underestimate our standard errors, it is a bad idea.

4 Regression set 4

4.1 Replication of columns (2), (3) and (4) of Table III in [Mian et al., 2017]

```
1 reghdfe logGDP_future household_credit_past, ///  
2 a(c) vce(cluster c year) // HH
```



```

3  est  sto  f1
4  reghdfe logGDP_future firm_credit_past, ///
5  a(c) vce(cluster c year) // Firm
6  est  sto  f2
7  reghdfe logGDP_future household_credit_past firm_credit_past, ///
8  a(c) vce(cluster c year) // Firm and HH
9  est  sto  f3
10 reghdfe logGDP_future household_credit_past firm_credit_past, ///
11 a(c year) vce(cluster c year) // Firm and HH and additional year FE
12 est  sto  f4
13 esttab f1 f2 f3 f4 using reg5.tex, se obslast scalar(F) r2 ///
14 title("Regression 5") replace

```

Table 6: Regression 5

	(2)	(3)	(4)	(4 add year FE)
	logGDP_future	logGDP_future	logGDP_future	logGDP_future
household_credit_past	-0.366*** (0.0786)		-0.337*** (0.0794)	-0.221** (0.0674)
firm_credit_past		-0.0978* (0.0398)	-0.0411 (0.0355)	-0.0379 (0.0324)
_cons	10.00*** (0.463)	8.718*** (0.560)	10.01*** (0.452)	9.495*** (0.319)
R^2	0.411	0.353	0.415	0.659
F	21.64	6.038	12.30	5.605
N	695	695	695	695

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

4.2 F-Test: Are the coefficients on household credit and firm credit in the regression of column (4) are equal

```

1  reghdfe logGDP_future household_credit_past firm_credit_past, ///
2  a(c) vce(cluster c year) // Firm and HH
3  test firm_credit_past==household_credit_past

```

- H_0 : The coefficients are equal $\iff \beta_{HHCredPast} - \beta_{FirmCredPast} = 0$
- H_1 : Negation of H_0
- The Teststatistic is F Distributed: $F(1, 29) = 9.45$ $Prob > F = 0.0046$

- We reject the H_0 , that the coefficients are equal, at the 1% level

4.3 Re-run the regression of column (4) and add year fixed effects

```
1 reghdfe logGDP_future household_credit_past firm_credit_past, ///
2 a(c year) vce(cluster c year) // Firm and HH and additional year FE
```

The results are already shown in table 6 column (4 add year FE).

We see that the R^2 went up by a large extend, but this was expected and does not tell us much. Standard errors where not effected by large extend. Also all relevant coefficients did not change by significant extent except for the coefficient of household_credit_past. This coefficient rised from -0.337 to -0.221 which implies that there probably was an omitted variable correlated with the error of the regression and household_credit_past, which is constant accross households but changes over time. This caused a downward biased coefficient.

Example: a nonpersistent shock in t-1 hits all households and varies over time, it is negatively correlated with household_credit_past and has an positive impact on the change of GDP from t to t+3 by reducing t but not t+3.