

STAT 425

Collinearity

Collinearity

Consider a MLR model with a design matrix $\mathbf{X}_{n \times p}$ including the intercept. If the columns of \mathbf{X} are **orthogonal** to each other (i.e., the sample correlation of any two predictors is equal to 0), then the LS problem is greatly simplified:

$$\hat{\beta}_j = [(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}]_j = \frac{\mathbf{X}_{\cdot j}^\top \mathbf{y}}{\|\mathbf{X}_{\cdot j}\|^2}$$

where $\mathbf{X}_{\cdot j}$ denotes the j -th column of \mathbf{X} .

In other words, in this case (only) the LS regression coefficient for the j -th predictor does not depend on whether other predictors are included in the model or not.

Collinearity

- In practice, we often encounter problems in which many of the predictors are highly correlated.
- In such cases, the values and sampling variance of regression coefficients can be highly dependent on the particular predictors chosen for the model.

Exact Collinearity

- If there exists a set of constants c_1, c_2, \dots, c_p (at least one of them is non-zero), such that the corresponding linear combination of the columns of \mathbf{X} is zero, i.e.:

$$\sum_{j=1}^p c_j \mathbf{X}_{.j} = \mathbf{0}$$

then the columns of \mathbf{X} are called **linearly dependent** and there is **exact collinearity**. That is, at least one column in the design matrix \mathbf{X} can be expressed as a linear combination of other columns.

What happens when the columns of \mathbf{X} are collinear?

- ① $(\mathbf{X}^\top \mathbf{X})^{-1}$ does not exist,
- ② The LS estimate $\hat{\beta}$ is not unique, and
- ③ The coefficients of the linear model are not identifiable.

Example: Suppose the 1st column of \mathbf{X} is the intercept, and the 2nd column of \mathbf{X} is the vector $(2, 2, \dots, 2)^\top$. Then if $(\hat{\beta}_1, \hat{\beta}_2, \hat{\beta}_3, \dots)^\top$ is one LS estimate of β , the vector $(\hat{\beta}_1 - c, \hat{\beta}_2 + c/2, \hat{\beta}_3, \dots)^\top$ is also an estimate of β , where c is any real number.

Note: In case of exact collinearity the column space of \mathbf{X} has dimension $< p$. In this case we can often fit an equivalent model by eliminating one or more redundant variables.

Approximate Collinearity

- We generally do not need to worry about exact collinearity¹, but **approximate collinearity**. That is, at least one column $\mathbf{X}_{.j}$ can be approximated by the others:

$$\mathbf{X}_{.k} \approx - \sum_{j \neq k} c_j \mathbf{X}_{.j} / c_k$$

A simple diagnostic for this is to obtain the regression of $\mathbf{X}_{.k}$ on the remaining predictors, and if the corresponding R_k^2 is close to 1, we would diagnose approximate collinearity.

¹ **R** can detect it and fix it automatically

Why approximate collinearity is a problem?

- In a multiple regression $Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + e$, the LS estimate $\hat{\beta}_k$ is unbiased with variance:

$$\text{var}(\hat{\beta}_k) = \sigma^2 \left(\frac{1}{1 - R_k^2} \right) \left(\frac{1}{\sum_{i=1}^n (x_{ik} - \bar{x}_{.k})^2} \right)$$

where R_k^2 is the R-square from the regression of $\mathbf{X}_{.k}$ on the remaining predictors. When R_k^2 is close to 1, the variance of $\hat{\beta}_k$ is large. Consequently we will have:

- ① large Mean Square Error
 - ② large (inflated) p-value to the corresponding t-test, i.e, we could **miss** a significant predictor.
- The quantity $\left(\frac{1}{1 - R_k^2} \right)$ is the **variance inflation factor** (VIF) for the k -th coefficient of the model

Example: Car position data

Data on 38 drivers:

- Age: Drivers age in years
- Weight: Drivers weight in lbs
- HtShoes: height with shoes in cm
- Ht: height without shoes in cm
- Seated: seated height in cm
- Arm: lower arm length in cm
- Thigh: thigh length in cm
- Leg: lower leg length in cm
- hipcenter: horizontal distance of the midpoint of the hips from a fixed location in the car in mm

```
library(faraway)
data(seatpos)
attach(seatpos)
g=lm(hipcenter ~ ., seatpos)
summary(g)
```


Example: Car position data

Collinearity Symptoms: None of the individual variables is significant.
Large standard errors. High correlation among variables

```
##  
## Call:  
## lm(formula = hipcenter ~ ., data = seatpos)  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max   
## -73.827 -22.833  -3.678   25.017   62.337   
##  
## Coefficients:  
##              Estimate Std. Error t value Pr(>|t|)      
## (Intercept)  436.43213   166.57162    2.620   0.0138 *      
## Age           0.77572     0.57033    1.360   0.1843      
## Weight        0.02631     0.33097    0.080   0.9372      
## HtShoes       -2.69241     9.75304   -0.276   0.7845      
## Ht            0.60134    10.12987    0.059   0.9531      
## Seated        0.53375     3.76189    0.142   0.8882      
## Arm          -1.32807     3.90020   -0.341   0.7359      
## Thigh         -1.14312     2.66002   -0.430   0.6706      
## Leg          -6.43905     4.71386   -1.366   0.1824      
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##  
## Residual standard error: 37.72 on 29 degrees of freedom  
## Multiple R-squared:  0.6866, Adjusted R-squared:  0.6001  
## F-statistic: 7.94 on 8 and 29 DF, p-value: 1.306e-05
```

Calculate Variance Inflation Factor of model matrix X (after removing the first column) using function **vif(.)**.

```
# Variance Inflation Factor (VIF)
round(vif(x), dig=2)
```

```
##      Age  Weight HtShoes      Ht  Seated      Arm  Thigh      Leg
##      2.00    3.65  307.43  333.14    8.95    4.50    2.76    6.69
```

```
sqrt(307.43)
```

```
## [1] 17.53368
```

Standard error of the estimated predictor $\hat{\beta}_{HtShoes}$ is approximately 17 times larger than it would have been without collinearity.

A global measure of collinearity

- A global measure of collinearity is given by examining the eigenvalues of $\mathbf{X}^\top \mathbf{X}$. A popular measure is the **condition number** of $\mathbf{X}^\top \mathbf{X}$, denoted by:

$$\kappa = (\text{largest eigenvalue} / \text{smallest eigenvalue})^{1/2}$$

An empirical rule for declaring collinearity is $\kappa \geq 30$

- Note that κ is not scale-invariant, so we should standardize each column of \mathbf{X} (i.e. each column should have zero mean and sample variance equal to 1, before calculating the condition number).

Example: Car Seat Position data

```
# Standardize matrix
x = model.matrix(g)[-1]
x = x - matrix(apply(x,2, mean), 38,8, byrow=TRUE)
x = x / matrix(apply(x, 2, sd), 38,8, byrow=TRUE)
apply(x,2,mean)
```

```
##           Age           Weight           HtShoes           Ht           Seated
## -2.193512e-17  2.810252e-16  9.566280e-16  1.941574e-16 -1.073010e-15
##           Arm           Thigh           Leg
## -1.070022e-16  8.909895e-17 -9.114182e-17
```

```
apply(x,2,var)
```

```
##      Age  Weight HtShoes      Ht  Seated      Arm  Thigh      Leg
##      1      1      1      1      1      1      1      1      1
```

```
e = eigen(t(x) %*% x)
sqrt(e$val[1]/e$val)
```

```
## [1] 1.000000  2.141737  3.497636  4.852243  5.404643  6.384606 10.615424
## [8] 59.766197
```

Symptoms and Remedies of Collinearity

- Possible symptoms of collinearity:
 - ① high pair-wise (sample) correlation between predictors
 - ② high VIF
 - ③ high condition number
 - ④ R^2 is relatively large but none of the predictor is significant.
- What to do with collinearity?

Remove some predictors from highly correlated groups of predictors.

Another method we study later: regularize the model using penalized Least Squares estimation