

**Homework #11****(due Friday, December 4, by 5:00 p.m. CST)***No credit will be given without supporting work.*

11. Suppose that students arrive to a certain on-campus COVID-19 testing location according to a Poisson process with average rate  $\lambda$  per minute. We plan to observe this location for the next 4 minutes in order to test  $H_0: \lambda = 3$  vs.  $H_1: \lambda < 3$ .

That is, let  $X_1, X_2, X_3, X_4$  be a random sample of size  $n = 4$  from a Poisson distribution with mean  $\lambda$ . Consider the test  $H_0: \lambda = 3$  vs.  $H_1: \lambda < 3$ .

- a) Find the best rejection region with the significance level closest to 0.05.

Reject  $H_0$  if  $X_1 + X_2 + X_3 + X_4 \leq c$ .

$X_1 + X_2 + X_3 + X_4$  has Poisson distribution with mean  $4\lambda$ .

$$\alpha = P(\text{Reject } H_0 \mid H_0 \text{ is true}) = P(X_1 + X_2 + X_3 + X_4 \leq c \mid \lambda = 3)$$

$$= P(\text{Poisson}(4 \cdot 3) \leq c) = P(\text{Poisson}(12) \leq c).$$

$$P(\text{Poisson}(12) \leq 6) = 0.046 \approx 0.05.$$

Reject  $H_0$  if  $\mathbf{X_1 + X_2 + X_3 + X_4 \leq 6}$ .

- b) What is the power of the Rejection Region obtained in part (a) if  $\lambda = 2$ ? if  $\lambda = 1.2$ ?

$$\text{Power}(\lambda = 2) = P(X_1 + X_2 + X_3 + X_4 \leq 6 \mid \lambda = 2) = P(\text{Poisson}(4\lambda) \leq 6 \mid \lambda = 2)$$

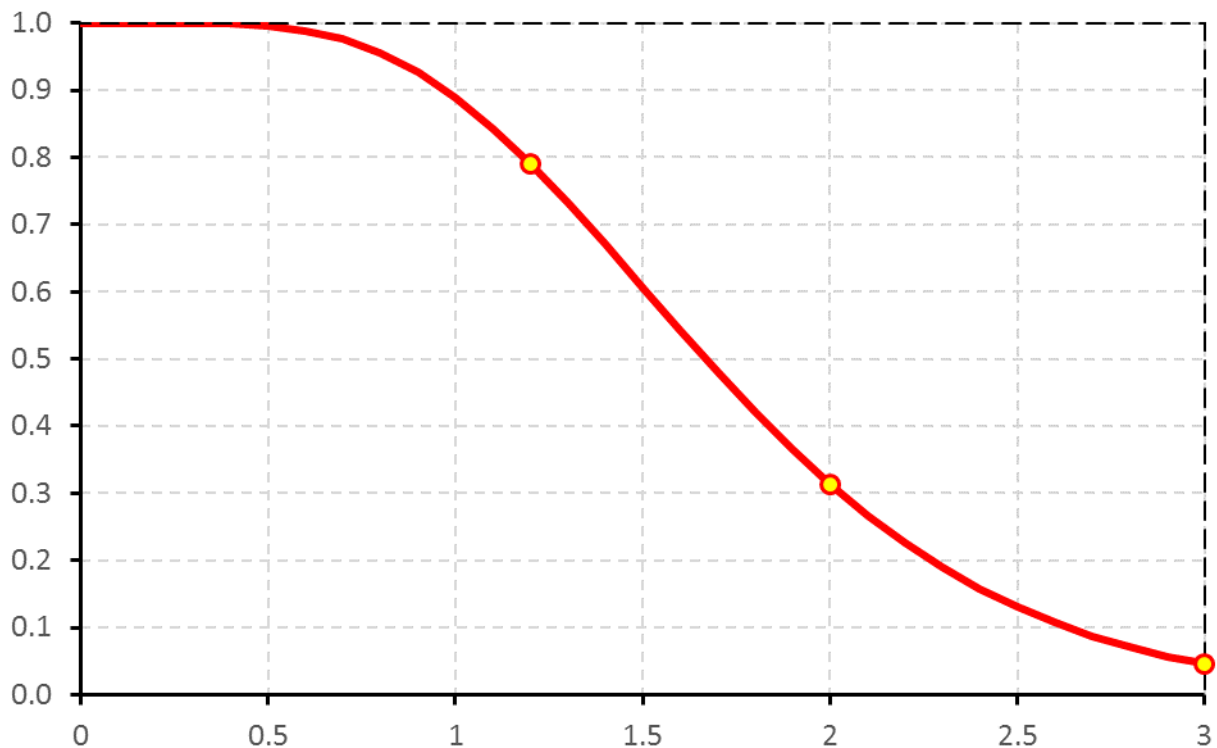
$$= P(\text{Poisson}(8) \leq 6) = \mathbf{0.313}.$$

$$\begin{aligned}
 \text{Power}(\lambda = 1.2) &= P(X_1 + X_2 + X_3 + X_4 \leq 6 \mid \lambda = 1.2) \\
 &= P(\text{Poisson}(4\lambda) \leq 6 \mid \lambda = 1.2) \\
 &= P(\text{Poisson}(4.8) \leq 6) = \mathbf{0.791}.
 \end{aligned}$$

c) Suppose  $x_1 = 1$ ,  $x_2 = 3$ ,  $x_3 = 1$ ,  $x_4 = 2$ . Find the p-value of the test.

$$x_1 + x_2 + x_3 + x_4 = 7.$$

$$\begin{aligned}
 \text{P-value} &= P(\text{value of } \sum_{i=1}^{n=4} X_i \text{ as extreme or more extreme than } 7 \mid H_0 \text{ true}) \\
 &= P(X_1 + X_2 + X_3 + X_4 \leq 7 \mid \lambda = 3) = P(\text{Poisson}(12) \leq 7) = \mathbf{0.090}.
 \end{aligned}$$



7. Let  $\psi > 0$  be a population parameter, and let  $X_1, X_2, \dots, X_n$  be a random sample from a distribution with probability density function

$$f(x; \psi) = \frac{2}{\sqrt{\pi\psi}} e^{-x^2/\psi}, \quad x > 0, \quad \text{zero otherwise.}$$

Recall:  $W = X^2$  has  $\text{Gamma}(\alpha = \frac{1}{2}, \theta = \psi)$  distribution.

$\sum_{i=1}^n X_i^2$  is a sufficient statistic for  $\psi$ .

Suppose  $n = 4$ . We wish to test  $H_0: \psi = 1.5$  vs.  $H_1: \psi > 1.5$ .

p) Consider rejection region Reject  $H_0$  if  $\sum_{i=1}^{n=4} x_i^2 \geq 7.2$ . Find ...

- i) ... the significance level  $\alpha$ ;
- ii) ... the power if  $\psi = 2$  and if  $\psi = 4$ .

$\sum_{i=1}^{n=4} X_i^2$  has  $\text{Gamma}(\alpha = \frac{n}{2} = 2, \theta = \psi)$  distribution.

(i)  $\alpha = \text{significance level} = P(\text{Reject } H_0 \mid H_0 \text{ is true}) = P(\sum_{i=1}^{n=4} X_i^2 \geq 7.2 \mid \psi = 1.5)$   
 $= P(\text{Gamma}(\alpha = 2, \theta = 1.5) \geq 7.2)$

```
> 1-pgamma(7.2, 2, 1/1.5)
[1] 0.04773253
```

$= P(\text{Poisson}(\frac{7.2}{1.5}) \leq 2 - 1) = P(\text{Poisson}(4.8) \leq 1) = \mathbf{0.048}.$

```
> ppois(2-1, 7.2/1.5)
[1] 0.04773253
```

$$= P\left(\frac{2}{1.5} T_2 \geq \frac{2}{1.5} \cdot 7.2 \mid \theta = 1.5\right) = P(\chi^2(4) \geq 9.6).$$

```
> 1-pchisq(9.6,4)
[1] 0.04773253
```

$$\int_{7.2}^{\infty} \frac{1}{1.5^2} x^{2-1} e^{-\frac{x}{1.5}} dx = 0.04773\dots$$

$$\begin{aligned} \text{(ii)} \quad \text{Power}(\psi = 2) &= P(\text{Reject } H_0 \mid \psi = 2) = P\left(\sum_{i=1}^{n=4} X_i^2 \geq 7.2 \mid \psi = 2\right) \\ &= P(\text{Gamma}(\alpha = 2, \theta = 2) \geq 7.2) \end{aligned}$$

```
> 1-pgamma(7.2,2,1/2)
[1] 0.1256891
```

$$= P\left(\text{Poisson}\left(\frac{7.2}{2}\right) \leq 2 - 1\right) = P(\text{Poisson}(3.6) \leq 1) = \mathbf{0.126}.$$

```
> ppois(2-1,7.2/2)
[1] 0.1256891
```

$$= P\left(\frac{2}{2} T_2 \geq \frac{2}{2} \cdot 7.2 \mid \theta = 2\right) = P(\chi^2(4) \geq 7.2).$$

```
> 1-pchisq(7.2,4)
[1] 0.1256891
```

$$\int_{7.2}^{\infty} \frac{1}{2^2} x^{2-1} e^{-\frac{x}{2}} dx = 0.12568\dots$$

$$\begin{aligned} \text{Power}(\psi = 4) &= P(\text{Reject } H_0 \mid \psi = 4) = P\left(\sum_{i=1}^{n=4} X_i^2 \geq 7.2 \mid \psi = 4\right) \\ &= P(\text{Gamma}(\alpha = 2, \theta = 4) \geq 7.2) \end{aligned}$$

```
> 1-pgamma(7.2,2,1/4)
[1] 0.4628369
```

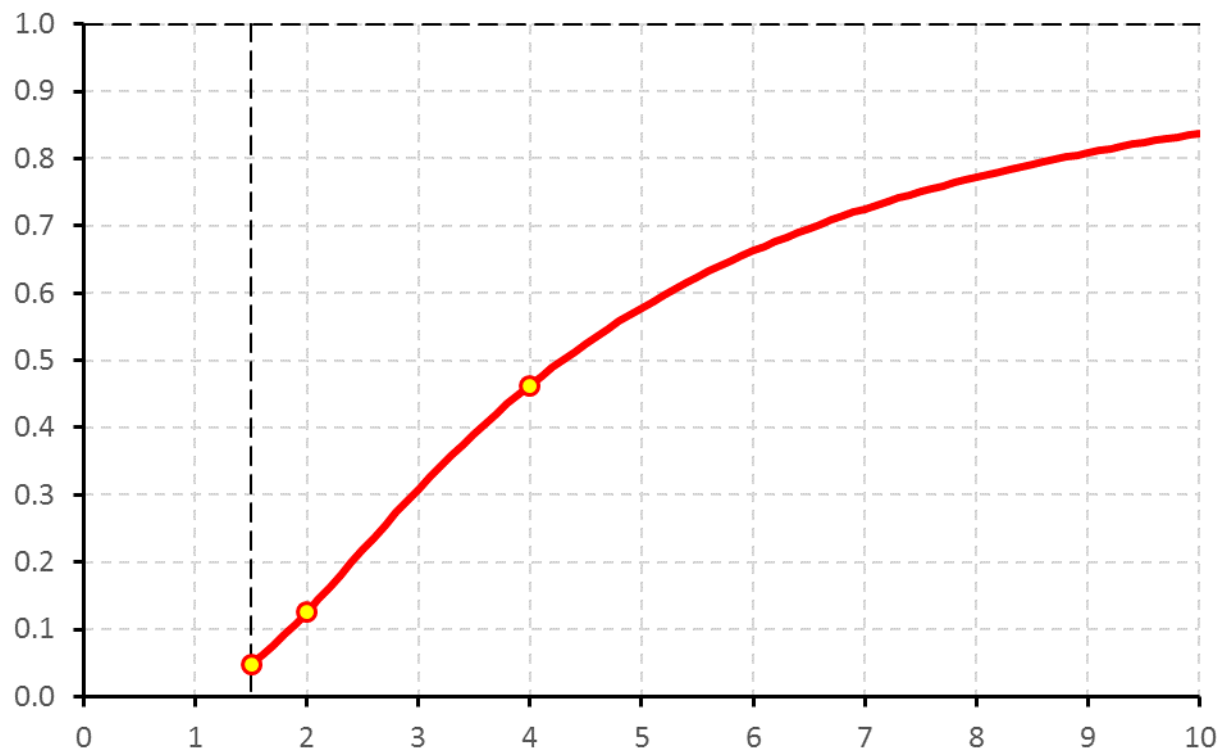
$$= P(\text{Poisson}(\frac{7.2}{4}) \leq 2 - 1) = P(\text{Poisson}(1.8) \leq 1) = \mathbf{0.463}.$$

```
> ppois(2-1,7.2/4)
[1] 0.4628369
```

$$= P(\frac{2}{4} T_2 \geq \frac{2}{4} \cdot 7.2 \mid \theta = 4) = P(\chi^2(4) \geq 3.6).$$

```
> 1-pchisq(3.6,4)
[1] 0.4628369
```

$$\int_{7.2}^{\infty} \frac{1}{4} x^{2-1} e^{-\frac{x}{4}} dx = 0.46283\dots$$



- q) Suppose  $x_1 = 0.2, x_2 = 0.6, x_3 = 1.1, x_4 = 1.7$ .  
Find the p-value of the test.

$$\sum_{i=1}^{n=4} x_i^2 = 0.2^2 + 0.6^2 + 1.1^2 + 1.7^2 = 4.5.$$

$$\begin{aligned} \text{p-value} &= P\left(\sum_{i=1}^{n=4} X_i^2 \text{ as extreme or more extreme than } \left(\sum_{i=1}^{n=4} x_i^2\right)_{\text{observed}} \mid H_0 \text{ true}\right) \\ &= P\left(\sum_{i=1}^{n=4} X_i^2 \geq 4.5 \mid \psi = 1.5\right) \\ &= P(\text{Gamma}(\alpha = 2, \theta = 1.5) \geq 4.5) \end{aligned}$$

```
> 1-pgamma(4.5, 2, 1/1.5)
[1] 0.1991483
```

$$= P\left(\text{Poisson}\left(\frac{4.5}{1.5}\right) \leq 2 - 1\right) = P(\text{Poisson}(3.0) \leq 1) = \mathbf{0.199}.$$

```
> ppois(2-1, 4.5/1.5)
[1] 0.1991483
```

$$= P\left(\frac{2}{1.5} T_2 \geq \frac{2}{1.5} \cdot 4.5 \mid \theta = 1.5\right) = P(\chi^2(4) \geq 6.0).$$

```
> 1-pchisq(6.0, 4)
[1] 0.1991483
```

$$\int_{4.5}^{\infty} \frac{1}{1.5^2} x^{2-1} e^{-\frac{x}{1.5}} dx = 0.19914...$$

For fun:

- r) Find a uniformly most powerful rejection region with the significance level  $\alpha = 0.05$ .

$$H_0: \psi = 1.5 \text{ vs. } H_1: \psi > 1.5. \quad n = 4.$$

Let  $\psi > 1.5$ .

$$\begin{aligned} \frac{L(1.5)}{L(\psi)} &= \frac{L(1.5; x_1, x_2, \dots, x_n)}{L(\psi; x_1, x_2, \dots, x_n)} = \frac{\prod_{i=1}^n \frac{2}{\sqrt{\pi 1.5}} e^{-x_i^2/1.5}}{\prod_{i=1}^n \frac{2}{\sqrt{\pi \psi}} e^{-x_i^2/\psi}} \\ &= \left( \frac{\psi}{1.5} \right)^{n/2} \exp \left\{ \left( \frac{1}{\psi} - \frac{1}{1.5} \right) \sum_{i=1}^n x_i^2 \right\} \leq k. \end{aligned}$$

$$\Leftrightarrow \exp \left\{ \left( \frac{1}{\psi} - \frac{1}{1.5} \right) \sum_{i=1}^n x_i^2 \right\} \leq k_1.$$

$$\Leftrightarrow \left( \frac{1}{\psi} - \frac{1}{1.5} \right) \sum_{i=1}^n x_i^2 \leq k_2.$$

$$\psi > 1.5 \quad \Rightarrow \quad \frac{1}{\psi} - \frac{1}{1.5} < 0$$

$$\Leftrightarrow \sum_{i=1}^n x_i^2 \geq c.$$

$$\text{Intuition: } \psi \text{ is “}\theta\text{”}. \quad E(W) = \alpha \theta = \frac{1}{2} \psi.$$

$$\text{Large } \psi \quad \Rightarrow \quad \text{large } w = x^2.$$

The sign is the same as the sign in  $H_1$ .

$$\sum_{i=1}^{n=4} X_i^2 \text{ has Gamma}(\alpha = \frac{n}{2} = 2, \theta = \psi) \text{ distribution.} \quad \sum_{i=1}^{n=4} X_i^2 = T_2.$$

$$\begin{aligned} 0.05 &= \alpha = P(\text{Reject } H_0 \mid H_0 \text{ is true}) = P\left(\sum_{i=1}^{n=4} X_i^2 \geq c \mid \psi = 1.5\right) \\ &= P(T_2 \geq c \mid \theta = 1.5) \end{aligned}$$

If  $T_\alpha$  has a  $\text{Gamma}(\alpha, \theta)$  distribution, then  $\frac{2}{\theta} T_\alpha$  has a  $\chi^2(2\alpha)$  distribution.

$$\begin{aligned} &= P\left(\frac{2}{\theta} T_2 \geq \frac{2}{\theta} c \mid \theta = 1.5\right) = P\left(\frac{2}{1.5} T_2 \geq \frac{2}{1.5} c \mid \theta = 1.5\right) \\ &= P(\chi^2(4) \geq \frac{2}{1.5} c). \end{aligned}$$

$$\Rightarrow \quad \frac{2}{1.5} c = \chi_{0.05}^2(4) = 9.488. \quad \Rightarrow \quad c = \mathbf{7.116}.$$

$$\text{Reject } H_0 \text{ if } \sum_{i=1}^{n=4} x_i^2 \geq \mathbf{7.116}.$$

```
> qgamma(0.95, 2, 1/1.5)
[1] 7.115797
>
> qchisq(0.95, 2*2)
[1] 9.487729
> qchisq(0.95, 2*2)*(1.5/2)
[1] 7.115797
```



8. Let  $\beta > 0$  be a population parameter, and let  $X_1, X_2, \dots, X_n$  be a random sample from a distribution with probability density function

$$f(x; \beta) = \beta (1-x)^{\beta-1}, \quad 0 < x < 1, \quad \text{zero otherwise.}$$

Recall:  $W = -\ln(1-X)$  has an Exponential( $\theta = \frac{1}{\beta}$ )  
 $=$  Gamma( $\alpha = 1, \theta = \frac{1}{\beta}$ ) distribution.

$\sum_{i=1}^n (-\ln(1-X_i))$  is a sufficient statistic for  $\beta$ .

Suppose  $n = 3$ . We wish to test  $H_0: \beta = 0.2$  vs.  $H_1: \beta > 0.2$ .

- p) Find the uniformly most powerful rejection region with significance level  $\alpha = 0.10$ .

Hint 1: Let  $\beta > 0.2$ . Start with

$$\frac{L(H_0; x_1, x_2, \dots, x_n)}{L(H_1; x_1, x_2, \dots, x_n)} = \frac{L(0.2; x_1, x_2, \dots, x_n)}{L(\beta; x_1, x_2, \dots, x_n)} = \frac{\prod_{i=1}^n f(x_i; 0.2)}{\prod_{i=1}^n f(x_i; \beta)} \leq k.$$

Simplify this. Since  $Y = \sum_{i=1}^n (-\ln(1-X_i))$  is a sufficient statistic for  $\beta$ ,

and the final form of the “best” rejection region should look like this:

$$\text{“Reject } H_0 \text{ if } \sum_{i=1}^n (-\ln(1-x_i)) [\leq \text{ or } \geq] c \text{”}.$$

The direction of the inequality sign is what you are trying to determine.

Hint 2:  $Y = \sum_{i=1}^n (-\ln(1-X_i)) = \sum_{i=1}^n W_i$  has a Gamma( $\alpha = n, \theta = \frac{1}{\beta}$ ) distribution.

Hint 3: Want  $c$  such that

$$0.10 = \alpha = P(\text{Reject } H_0 \mid H_0 \text{ is true}) = P\left(\sum_{i=1}^{n=3} (-\ln(1-X_i)) \geq c \mid \beta = 0.2\right).$$

Hint 4: If  $T_\alpha$  has a  $\text{Gamma}(\alpha, \theta = 1/\lambda)$  distribution,

then  $\frac{2}{\theta} T_\alpha = 2\lambda T_\alpha$  has a  $\chi^2(2\alpha)$  distribution.

Let  $\beta > 0.2$ .

$$\begin{aligned} \frac{L(0.2)}{L(\beta)} &= \frac{L(0.2; x_1, x_2, \dots, x_n)}{L(\beta; x_1, x_2, \dots, x_n)} = \frac{\prod_{i=1}^n 0.2 (1-x_i)^{-0.8}}{\prod_{i=1}^n \beta (1-x_i)^{\beta-1}} \\ &= \left( \frac{0.2}{\beta} \right)^n \left( \prod_{i=1}^n (1-x_i) \right)^{0.2-\beta} \leq k. \end{aligned}$$

$$\Leftrightarrow \left( \prod_{i=1}^n (1-x_i) \right)^{0.2-\beta} \leq k_1.$$

$$\Leftrightarrow (\beta - 0.2) \sum_{i=1}^n (-\ln(1-x_i)) \leq k_2.$$

$$\beta > 0.2 \quad \Rightarrow \quad \beta - 0.2 > 0$$

$$\Leftrightarrow \sum_{i=1}^n (-\ln(1-x_i)) \leq c.$$

Intuition:  $\beta$  is “ $\lambda$ ”.

$$E(W) = \alpha \theta = \frac{1}{\beta}.$$

Large  $\beta \Rightarrow$  small  $w = -\ln(1-x)$ .

The sign is opposite from the sign in  $H_1$ .

$\sum_{i=1}^{n=3} (-\ln(1-X_i))$  has  $\text{Gamma}(\alpha = n = 3, \theta = \frac{1}{\beta})$  distribution.

$$\sum_{i=1}^{n=3} \left( -\ln(1-X_i) \right) = T_3.$$

$$\begin{aligned} 0.10 = \alpha &= P(\text{Reject } H_0 \mid H_0 \text{ is true}) = P\left(\sum_{i=1}^{n=3} \left( -\ln(1-X_i) \right) \leq c \mid \beta = 0.2\right) \\ &= P(T_3 \leq c \mid \lambda = 0.2) = P(T_3 \leq c \mid \theta = 5) \end{aligned}$$

If  $T_\alpha$  has a  $\text{Gamma}(\alpha, \theta)$  distribution,

then  $\frac{2}{\theta} T_\alpha = 2\lambda T_\alpha$  has a  $\chi^2(2\alpha)$  distribution.

$$= P(2\lambda T_3 \leq 2\lambda c \mid \lambda = 0.2) = P\left(\frac{2}{5} T_3 \leq \frac{2}{5} c \mid \theta = 5\right)$$

$$= P(2 \cdot 0.2 \cdot T_3 \leq 2 \cdot 0.2 \cdot c \mid \lambda = 0.2) = P\left(\frac{2}{5} T_3 \leq \frac{2}{5} c \mid \theta = 5\right)$$

$$= P(\chi^2(6) \leq 0.4c).$$

$$\Rightarrow \quad 0.4c = \chi_{0.90}^2(6) = 2.204. \quad \Rightarrow \quad c = \mathbf{5.51}.$$

$$\text{Reject } H_0 \text{ if } \sum_{i=1}^{n=3} \left( -\ln(1-x_i) \right) \leq \mathbf{5.51}.$$

```
> qgamma(0.10,3,0.2)
[1] 5.510327
>
> qchisq(0.10,2*3)
[1] 2.204131
> qchisq(0.10,2*3)/(2*0.2)
[1] 5.510327
```

q) Suppose  $x_1 = 0.31$ ,  $x_2 = 0.77$ ,  $x_3 = 0.93$ . Find the p-value of the test.

Hint 1: Probability that  $\sum_{i=1}^{n=3} (-\ln(1-X_i))$  is as extreme or more extreme than the

observed  $\sum_{i=1}^{n=3} (-\ln(1-x_i)) \dots$

Hint 2: For the p-value, go in the same direction as the “best” rejection region.

Hint 3: ... computed under the assumption that  $H_0$  is true.

$$\sum_{i=1}^{n=3} (-\ln(1-x_i)) = -\ln 0.69 - \ln 0.23 - \ln 0.07 \approx 4.5.$$

$$\begin{aligned} \text{p-value} &= P\left(\sum_{i=1}^{n=3} (-\ln(1-X_i)) \leq 4.5 \mid \beta = 0.2\right) \\ &= P(\text{Gamma}(\alpha = 3, \lambda = 0.2, \theta = 5) \leq 4.5) \end{aligned}$$

```
> pgamma(4.5, 3, 0.2)
[1] 0.06285693
```

$$\begin{aligned} &= P(\text{Poisson}(4.5 \cdot 0.2) \geq 3) = 1 - P(\text{Poisson}(0.9) \leq 2) \\ &= 1 - 0.937 = \mathbf{0.063}. \end{aligned}$$

```
> 1-ppois(3-1, 0.2*4.5)
[1] 0.06285693
```

$$= P(2 \cdot 0.2 \cdot T_3 \leq 2 \cdot 0.2 \cdot 4.5 \mid \lambda = 0.2) = P(\chi^2(6) \leq 1.8).$$

```
> pchisq(2*0.2*4.5, 2*3)
[1] 0.06285693
```

$$\int_0^{4.5} \frac{0.2^3}{2} x^{3-1} e^{-0.2x} dx = 0.06285\dots$$

- r) (i) **Using the p-value obtained in part (q),** state your decision ( Reject  $H_0$  or Do NOT Reject  $H_0$  ) at  $\alpha = 0.05$ .
- (ii) **Using the rejection region obtained in part (p),** state your decision ( Reject  $H_0$  or Do NOT Reject  $H_0$  ) at  $\alpha = 0.10$ .

- (i)  $\text{p-value} > \alpha \Rightarrow \text{Do NOT Reject } H_0.$
- $\text{p-value} < \alpha \Rightarrow \text{Reject } H_0.$

Since  $0.063 > 0.05$ ,

**Do NOT Reject  $H_0$  at  $\alpha = 0.05$ .**

- (ii) Reject  $H_0$  if  $\sum_{i=1}^{n=3} ( - \ln(1-x_i) ) \leq 5.51.$

$$\sum_{i=1}^{n=3} ( - \ln(1-x_i) ) \approx 4.5.$$

$$4.5 \leq 5.51.$$

**Reject  $H_0$  at  $\alpha = 0.10$ .**