

Homework #12

(due Friday, December 11, by 5:00 p.m. CST)

No credit will be given without supporting work.

12. Consider two (discrete) probability mass functions:

x	1	2	3	4	5	6	7	8	9
$p_0(x)$	0	0.05	0.05	0.10	0.20	0.40	0.10	0.05	0.05
$p_1(x)$	0.08	0.08	0.10	0.25	0.05	0.08	0.16	0.20	0

You have only a single observation of X on which to base your choice between

$$H_0: X \text{ has p.m.f. } p_0(x) \quad \text{vs.} \quad H_1: X \text{ has p.m.f. } p_1(x).$$

- a) (Warm-up) Consider the rejection region Reject H_0 if $x \in \{1, 2, 3\}$. Find ...
- (i) ... the significance level α , (ii) ... the power.

$$\begin{aligned} \alpha &= P(\text{Reject } H_0 \mid H_0 \text{ is true}) = P(X \in \{1, 2, 3\} \mid H_0 \text{ is true}) \\ &= p_0(1) + p_0(2) + p_0(3) = 0 + 0.05 + 0.05 = \mathbf{0.10}. \end{aligned}$$

$$\begin{aligned} \text{Power} &= P(\text{Reject } H_0 \mid H_0 \text{ is NOT true}) = P(X \in \{1, 2, 3\} \mid H_0 \text{ is NOT true}) \\ &= p_1(1) + p_1(2) + p_1(3) = 0.08 + 0.08 + 0.10 = \mathbf{0.26}. \end{aligned}$$

b) Find $\lambda(x) = \frac{L(H_0; x)}{L(H_1; x)} = \frac{p_0(x)}{p_1(x)}$ for $x = 1, 2, 3, \dots, 9$.

x	1	2	3	4	5	6	7	8	9
$p_0(x)$	0	0.05	0.05	0.10	0.20	0.40	0.10	0.05	0.05
$p_1(x)$	0.08	0.08	0.10	0.25	0.05	0.08	0.16	0.20	0
$\lambda(x)$	0	0.625	0.5	0.4	4	5	0.625	0.25	∞

most extreme					least extreme			
$\lambda(x)$	0	0.25	0.4	0.5	0.625	4	5	∞
x	1	8	4	3	2 & 7	5	6	9

Intuition on $\lambda(1) = 0$:

If I observe $x = 0$, then I would know with 100% certainty that H_1 is true.

(since $x = 0$ cannot be observed if H_0 is true).

$x = 0$ is the possible value of X that has the most evidence against for H_0 .

A rejection region “Reject H_0 if $x = 0$ ” would have had 0% significance level, yet positive power ($p_1(0) = 0.08$).

Intuition on $\lambda(9) = \infty$:

If I observe $x = 9$, then I immediately know with 100% certainty that H_0 is true

(since $x = 9$ cannot be observed if H_1 is true).

$x = 9$ is the possible value of X that has the most support for H_0 .

Since $p_1(9) = 0$ and $p_0(9) > 0$, we are “infinitely more likely” to observe $x = 9$ if H_0 is true than if H_1 is true.

c) Find the most powerful rejection region with significance level $\alpha = 0.20$.

“Hint”: Reject H_0 if $x \in \{ ??? \}$.

“Hint”: The Neyman-Pearson lemma “strongly suggests” that the values with the smallest $\lambda(x)$ are added to the rejection region first.

	most extreme					least extreme		
$\lambda(x)$	0	0.25	0.4	0.5	0.625	4	5	∞
x	1	8	4	3	2 & 7	5	6	9
$p_0(x)$	0	0.05	0.10	0.05

$$P_0(\{1, 8, 4, 3\}) = 0 + 0.05 + 0.10 + 0.05 = 0.20.$$

Reject H_0 if $x \in \{ \mathbf{1, 8, 4, 3} \}$.

d) Find the power of the rejection region from part (c).

$$P_1(\{1, 8, 4, 3\}) = 0.08 + 0.20 + 0.25 + 0.10 = \mathbf{0.63}.$$

IF $n = 2$, then you have $9 \cdot 9 = 81$ possible samples (x_1, x_2) to consider.

IF $n = 3$, then you have $9 \cdot 9 \cdot 9 = 729$ possible samples (x_1, x_2, x_3) to consider. ...

$n = 5$ on Final Exam? (obviously, with different numbers, so you do not have time to prepare)

8. Let $\beta > 0$ be a population parameter, and let X_1, X_2, \dots, X_n be a random sample from a distribution with probability density function

$$f(x|\beta) = \beta (1-x)^{\beta-1}, \quad 0 < x < 1, \quad \beta > 0, \quad \text{zero otherwise.}$$

Recall that the maximum likelihood estimator for β is $\hat{\beta} = \frac{n}{\sum_{i=1}^n (-\ln(1-X_i))}$.

Let the prior p.d.f. of β be $\text{Gamma}(\alpha, \theta)$. That is,

$$\pi(\beta) = \frac{1}{\Gamma(\alpha) \theta^\alpha} \beta^{\alpha-1} e^{-\beta/\theta}, \quad \beta > 0.$$

- s) Find the posterior distribution of β , given $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$.

HINT: $(1-x) = e^{-(\ln(1-x))}$.

$$\begin{aligned} f(x_1, x_2, \dots, x_n, \beta) &= \prod_{i=1}^n \beta (1-x_i)^{\beta-1} \times \frac{1}{\Gamma(\alpha) \theta^\alpha} \beta^{\alpha-1} e^{-\beta/\theta} \\ &= \dots \beta^{n+\alpha-1} e^{-\beta \left(\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta} \right)}. \end{aligned}$$

\Rightarrow the posterior distribution of β , given $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$,

is **Gamma** with New $\alpha = n + \alpha$ and

$$\text{New } \theta = \frac{1}{\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta}}.$$

- t) Find the conditional mean of β , given $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$. Show that it is a weighted average of the maximum likelihood estimate $\hat{\beta}$ and the prior mean $\alpha \theta$. (What are the weights?)

(conditional mean of β given $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$)

$$\begin{aligned}
 &= (\text{New } \alpha) \times (\text{New } \theta) = \frac{n + \alpha}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}} \\
 &= \frac{n}{\sum_{i=1}^n \left(-\ln(1-x_i) \right)} \cdot \frac{\sum_{i=1}^n \left(-\ln(1-x_i) \right)}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}} + \alpha \theta \cdot \frac{\frac{1}{\theta}}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}} \\
 &= (\text{MLE}) \cdot \frac{\sum_{i=1}^n \left(-\ln(1-x_i) \right)}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}} + \left(\frac{\text{prior}}{\text{mean}} \right) \cdot \frac{\frac{1}{\theta}}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}}.
 \end{aligned}$$

- u) Use part (s) to construct a $(1 - \gamma) 100\%$ credible interval for β , given $X_1 = x_1$, $X_2 = x_2, \dots, X_n = x_n$.

$(\beta | x_1, x_2, \dots, x_n)$ has a

Gamma (New $\alpha = n + \alpha$, New $\theta = \frac{1}{\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta}}$) distribution.

$\frac{2}{\text{New } \theta} (\beta | x_1, x_2, \dots, x_n)$ has a $\chi^2(2 \times \text{New } \theta)$ distribution.

$2 \left(\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta} \right) (\beta | x_1, x_2, \dots, x_n)$ has a $\chi^2(2n + 2\alpha)$ distribution.

$$P(\chi_{1-\gamma/2}^2(2n + 2\alpha) < 2 \left(\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta} \right) (\beta | x_1, x_2, \dots, x_n) < \chi_{\gamma/2}^2(2n + 2\alpha)) = 1 - \gamma.$$

$$P\left(\frac{\chi_{1-\gamma/2}^2(2n + 2\alpha)}{2 \left(\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta} \right)} < (\beta | x_1, x_2, \dots, x_n) < \frac{\chi_{\gamma/2}^2(2n + 2\alpha)}{2 \left(\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta} \right)} \right) = 1 - \gamma.$$

$$\left(\frac{\chi_{1-\gamma/2}^2(2n + 2\alpha)}{2 \left(\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta} \right)}, \frac{\chi_{\gamma/2}^2(2n + 2\alpha)}{2 \left(\sum_{i=1}^n (-\ln(1-x_i)) + \frac{1}{\theta} \right)} \right)$$

is a $(1 - \gamma) 100\%$ credible interval for β , given $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$.

- v) Suppose $n = 3$, and $x_1 = 0.31$, $x_2 = 0.77$, $x_3 = 0.93$.
Let $\alpha = 2$, $\theta = 1.20$.

$$n = 3, \quad \sum_{i=1}^n \left(-\ln(1-x_i) \right) = -\ln 0.69 - \ln 0.23 - \ln 0.07 \approx 4.5.$$

- (i) Find the conditional mean of β , given $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$.

$$\frac{n + \alpha}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}} \approx \frac{3 + 2}{4.5 + \frac{1}{1.20}} = \mathbf{0.9375}.$$

Last chance for Alex to be annoying:

$$\hat{\beta} \approx \frac{3}{4.5} = \frac{2}{3}. \quad \text{Prior mean} = \alpha \theta = 2.4.$$

$$\frac{\sum_{i=1}^n \left(-\ln(1-x_i) \right)}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}} \approx \frac{4.5}{4.5 + \frac{1}{1.2}} = 0.84375,$$

$$\frac{\frac{1}{\theta}}{\sum_{i=1}^n \left(-\ln(1-x_i) \right) + \frac{1}{\theta}} \approx \frac{\frac{1}{1.2}}{4.5 + \frac{1}{1.2}} = 0.15625.$$

$$\frac{2}{3} \cdot 0.84375 + 2.4 \cdot 0.15625 = \mathbf{0.9375}.$$

- (ii) Construct a 95% credible interval for β , given $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$.

$$\chi^2_{0.975}(10) = 3.247, \quad \chi^2_{0.025}(10) = 20.48.$$

$$\left(\frac{3.247}{2 \cdot \left(4.5 + \frac{1}{1.2} \right)}, \frac{20.48}{2 \cdot \left(4.5 + \frac{1}{1.2} \right)} \right) \quad (0.3044, 1.92)$$

