

Examples for 10/19/2020 (2) and 10/23/2020 (2)
and 10/30/2020 (3) and 11/04/2020 (2) (continued)

1. Let $\beta > 0$ and let X_1, X_2, \dots, X_n be a random sample from the distribution with probability density function

$$f(x; \beta) = \frac{\beta^2 \ln x}{x^{\beta+1}}, \quad x > 1, \quad \text{zero otherwise.}$$

Recall: Examples for 11/04/2020 (2):

$$Y = \sum_{i=1}^n \ln X_i \text{ is a sufficient statistic for } \beta.$$

Examples for 10/19/2020 (2):

$$W = \ln X \text{ has a Gamma}(\alpha = 2, \theta = \frac{1}{\beta}) \text{ distribution.}$$

$$\Rightarrow Y = \sum_{i=1}^n \ln X_i = \sum_{i=1}^n W_i \text{ has a Gamma}(\alpha = 2n, \theta = \frac{1}{\beta}) \text{ distribution.}$$

We wish to test $H_0: \beta = 1.4$ vs. $H_1: \beta > 1.4$.

- a) Find the form of the uniformly most powerful rejection region.

Hint: Let $\beta > 1.4$. Start with

$$\frac{L(H_0; x_1, x_2, \dots, x_n)}{L(H_1; x_1, x_2, \dots, x_n)} = \frac{L(1.4; x_1, x_2, \dots, x_n)}{L(\beta; x_1, x_2, \dots, x_n)} = \frac{\prod_{i=1}^n f(x_i; 1.4)}{\prod_{i=1}^n f(x_i; \beta)} \leq k.$$

Simplify this. Since $Y = \sum_{i=1}^n \ln X_i$ is a sufficient statistic for β ,

and the final form of the “best” rejection region should look like this:

$$\text{“Reject } H_0 \text{ if } \sum_{i=1}^n \ln x_i \left[\leq \text{ or } \geq \right] c \text{”}.$$

The direction of the inequality sign is what you are trying to determine.

Let $\beta > 1.4$.

$$\frac{L(1.4)}{L(\beta)} = \frac{\prod_{i=1}^n \frac{1.4^2 \ln x_i}{x_i^{1.4+1}}}{\prod_{i=1}^n \frac{\beta^2 \ln x_i}{x_i^{\beta+1}}} = \left(\frac{1.4}{\beta}\right)^{2n} \left(\prod_{i=1}^n x_i\right)^{\beta-1.4} \leq k.$$

$$\Leftrightarrow \left(\prod_{i=1}^n x_i\right)^{\beta-1.4} \leq k_1 = k \left(\frac{\beta}{1.4}\right)^{2n}.$$

$$\Leftrightarrow (\beta - 1.4) \sum_{i=1}^n \ln x_i \leq k_2 = \ln k_1.$$

$$\Leftrightarrow \sum_{i=1}^n \ln x_i \leq c = \frac{k_2}{\beta - 1.4}, \quad \text{since } \beta > 1.4.$$

Intuition: β is “ λ ”.

$$E(W) = \alpha \theta = \frac{2}{\beta}.$$

Large $\beta \Rightarrow$ small $\ln x$.

The sign is opposite from the sign in H_1 .

$$\text{Reject } H_0 \text{ if } \sum_{i=1}^n \ln x_i \leq c.$$

b) Suppose $n = 5$. Find the uniformly most powerful rejection region with $\alpha = 0.05$.

Hint 1: $Y = \sum_{i=1}^n \ln X_i = \sum_{i=1}^n W_i$ has a Gamma($\alpha = 2n$, $\theta = \frac{1}{\beta}$) distribution.

Hint 2: Want c such that

$$0.05 = \alpha = P(\text{Reject } H_0 \mid H_0 \text{ is true}) = P\left(\sum_{i=1}^5 \ln X_i \geq c \mid \beta = 1.4\right).$$

Hint 3: If T has a Gamma(α , $\theta = 1/\lambda$) distribution, then

$2T/\theta = 2\lambda T$ has a $\chi^2(2\alpha)$ distribution (a chi-square distribution with 2α degrees of freedom).

$\sum_{i=1}^n \ln X_i = \sum_{i=1}^n W_i$ has a Gamma($\alpha = 2n = 10$, $\theta = \frac{1}{\beta}$) distribution.

Then $2\beta \sum_{i=1}^5 \ln X_i$ has a $\chi^2(2\alpha = 4n = 20 \text{ degrees of freedom})$ distribution.

$$\begin{aligned} 0.05 &= P\left(\sum_{i=1}^5 \ln X_i \leq c \mid \beta = 1.4\right) = P\left(2\beta \sum_{i=1}^5 \ln X_i \leq 2\beta c \mid \beta = 1.4\right) \\ &= P(\chi^2(20) \leq 2.8c). \end{aligned}$$

$$\Rightarrow 2.8c = \chi_{0.95}^2(20) = 10.85. \quad \Rightarrow c = \mathbf{3.875}.$$

The uniformly most powerful rejection region is “Reject H_0 if $\sum_{i=1}^5 \ln x_i \leq 3.875$.”

So $\alpha = 0.05\ldots$ Yes.

And $\alpha = 10\ldots$ Yes.



c) Find the power of the rejection region from (b) if $\beta = 4$.

Hint: $\text{Power}(\beta) = P(\text{Reject } H_0 \mid H_0 \text{ is NOT true}) = P\left(\sum_{i=1}^5 \ln X_i \geq c \mid \beta = 4\right)$.

Suggestion: If T has a $\text{Gamma}(\alpha, \theta = 1/\lambda)$ distribution, where α is an integer, then $F_T(t) = P(T \leq t) = P(X_t \geq \alpha)$ and $P(T > t) = P(X_t \leq \alpha - 1)$, where X_t has a $\text{Poisson}(\lambda t = t/\theta)$ distribution.

$$\begin{aligned}\text{Power}(\beta = 4) &= P\left(\sum_{i=1}^5 \ln X_i \leq 3.875 \mid \beta = 4\right) \\ &= P\left(\text{Gamma}(\alpha = 10, \theta = \frac{1}{4}) \leq 3.875\right) \\ &= P(\text{Poisson}(3.875 \cdot 4) \geq 10) = 1 - P(\text{Poisson}(15.5) \leq 9) \\ &= 1 - 0.055 = \mathbf{0.945}.\end{aligned}$$

```
> 1-ppois(9,15.5)
[1] 0.9448095
```

$$\begin{aligned}\text{Power}(\beta = 4) &= P\left(\sum_{i=1}^5 \ln X_i \leq 3.875 \mid \beta = 4\right) \\ &= P\left(\text{Gamma}(\alpha = 10, \theta = \frac{1}{4}) \leq 3.875\right) \\ &= P(\chi^2(20) \leq 8 \cdot 3.875) = P(\chi^2(20) \leq 31).\end{aligned}$$

```
> pchisq(31,20)
[1] 0.9448095
```

```
> pgamma(3.875,10,4)
[1] 0.9448095
```

d) Find the power of the rejection region from (b) if $\beta = 2$.

Hint 1: $\text{Power}(\beta) = P(\text{Reject } H_0 \mid H_0 \text{ is NOT true}) = P\left(\sum_{i=1}^5 \ln X_i \geq c \mid \beta = 2\right)$.

Hint 2: Excel: =GAMMA.DIST(x , α , θ , 1) $P(\text{Gamma}(\alpha, \theta) \leq x)$,

R: pgamma(x , α , λ) $P(\text{Gamma}(\alpha, 1/\lambda) \leq x)$.

OR

If T has a $\text{Gamma}(\alpha, \theta = 1/\lambda)$ distribution, where α is an integer,

then $F_T(t) = P(T \leq t) = P(X_t \geq \alpha)$ and $P(T > t) = P(X_t \leq \alpha - 1)$,

where X_t has a $\text{Poisson}(\lambda t = t/\theta)$ distribution.

Excel: =POISSON.DIST(x , λ , 1) $P(\text{Poisson}(\lambda) \leq x)$,

=POISSON.DIST(x , λ , 0) $P(\text{Poisson}(\lambda) = x)$.

R: ppois(x , λ) $P(\text{Poisson}(\lambda) \leq x)$,

dpois(x , λ) $P(\text{Poisson}(\lambda) = x)$.

OR

If T has a $\text{Gamma}(\alpha, \theta = 1/\lambda)$ distribution, then

$2T/\theta = 2\lambda T$ has a $\chi^2(2\alpha)$ distribution (a chi-square distribution with 2α degrees of freedom).

Excel: =CHISQ.DIST(x , degrees of freedom , 1) $P(\chi^2(\text{d.f.}) \leq x)$,

=CHISQ.DIST.RT(x , degrees of freedom) $P(\chi^2(\text{d.f.}) > x)$.

R: pchisq(x , degrees of freedom) $P(\chi^2(\text{d.f.}) \leq x)$,

pchisq(x , degrees of freedom , lower.tail=FALSE).

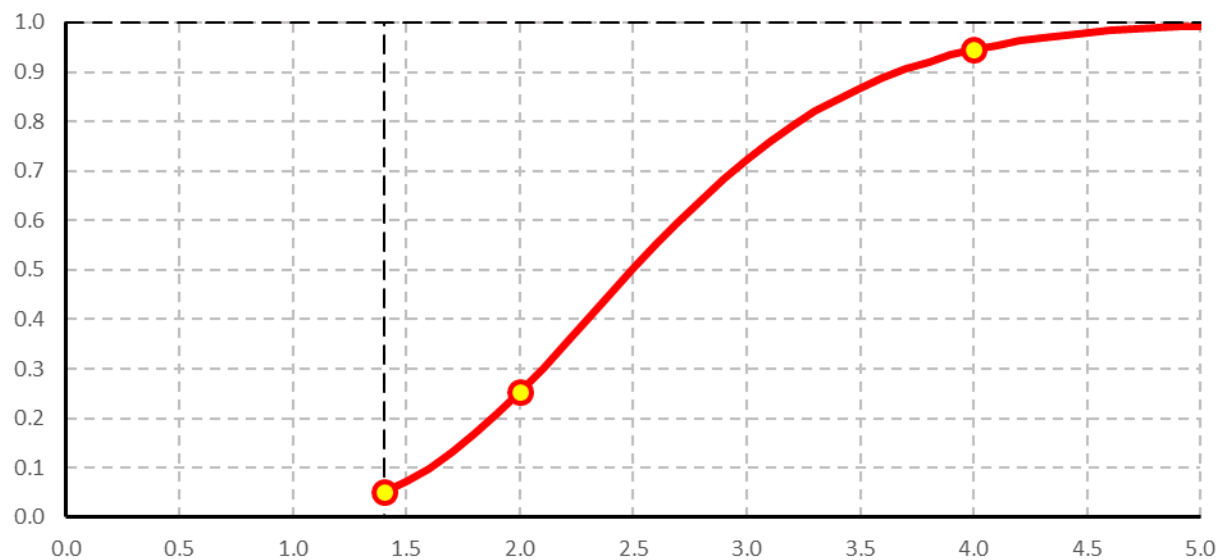
$$\begin{aligned}
 \text{Power}(\beta = 2) &= P\left(\sum_{i=1}^5 \ln X_i \leq 3.875 \mid \beta = 2\right) \\
 &= P\left(\text{Gamma}\left(\alpha = 10, \theta = \frac{1}{2}\right) \leq 3.875\right) \\
 &= P\left(\text{Poisson}(3.875 \cdot 2) \geq 10\right) = 1 - P\left(\text{Poisson}(7.75) \leq 9\right)
 \end{aligned}$$

```
> 1-ppois(9,7.75)
[1] 0.2528812
```

$$\begin{aligned}
 \text{Power}(\beta = 4) &= P\left(\sum_{i=1}^5 \ln X_i \leq 3.875 \mid \beta = 2\right) \\
 &= P\left(\text{Gamma}\left(\alpha = 10, \theta = \frac{1}{2}\right) \leq 3.875\right) \\
 &= P\left(\chi^2(20) \leq 4 \cdot 3.875\right) = P\left(\chi^2(20) \leq 15.5\right).
 \end{aligned}$$

```
> pchisq(15.5,20)
[1] 0.2528812
```

```
> pgamma(3.875,10,2)
[1] 0.2528812
```



1. (continued)

Suppose $n = 5$, and $x_1 = 1.3, x_2 = 1.4, x_3 = 2.0, x_4 = 3.0, x_5 = 5.0$.

e) Find the p-value for the test.

Hint 1: $\dots \sum_{i=1}^5 \ln X_i$ as extreme or more extreme than the observed $\sum_{i=1}^n \ln x_i \dots$

Hint 2: For the p-value, go in the same direction as the “best” rejection region.

Hint 3: \dots computed under the assumption that H_0 is true.

Suggestion: If T has a $\text{Gamma}(\alpha, \theta = 1/\lambda)$ distribution, where α is an integer, then $F_T(t) = P(T \leq t) = P(X_t \geq \alpha)$ and $P(T > t) = P(X_t \leq \alpha - 1)$, where X_t has a $\text{Poisson}(\lambda t = t/\theta)$ distribution.

$$\sum_{i=1}^5 \ln x_i = \ln 1.3 + \ln 1.4 + \ln 2.0 + \ln 3.0 + \ln 5.0 \approx 4.$$

$$\begin{aligned} \text{P-value} &= P\left(\sum_{i=1}^5 \ln X_i \leq 4 \mid \beta = 1.4\right) = P\left(\text{Gamma}(\alpha = 10, \theta = \frac{1}{1.4}) \leq 4\right) \\ &= P(\text{Poisson}(4 \cdot 1.4) \geq 10) = 1 - P(\text{Poisson}(5.6) \leq 9) \\ &= 1 - 0.941 = \mathbf{0.059}. \end{aligned}$$

```
> 1-ppois(9,5.6)
[1] 0.05912997
```

```
> pgamma(4,10,1.4)
[1] 0.05912997
```

$$\begin{aligned}\text{P-value} &= P\left(\sum_{i=1}^5 \ln X_i \leq 4 \mid \beta = 1.4\right) = P\left(\text{Gamma}\left(\alpha = 10, \theta = \frac{1}{1.4}\right) \leq 4\right) \\ &= P\left(\chi^2(20) \leq 2.8 \cdot 4\right) = P\left(\chi^2(20) \leq 11.2\right).\end{aligned}$$

```
> pchisq(11.2, 20)
[1] 0.05912997
```

- f) (i) Using the p-value obtained in part (e), state your decision (Reject H_0 or Do NOT Reject H_0) at $\alpha = 0.05$.
- (ii) Does your decision agree with part (b)? That is, does the observed value of $\sum_{i=1}^n \ln x_i$ fall into the rejection region from part (b)?

- i) $\text{p-value} > \alpha \Rightarrow$ Do NOT Reject H_0 .
- $\text{p-value} < \alpha \Rightarrow$ Reject H_0 .

Since $0.059 > 0.05$, **Do NOT Reject H_0 at $\alpha = 0.05$.**

- ii) Part (b): Reject H_0 if $\sum_{i=1}^5 \ln x_i \leq 3.875$.

Observed $\sum_{i=1}^5 \ln x_i \approx 4$ does NOT fall into the rejection region.

Do NOT Reject H_0 at $\alpha = 0.05$.

They agree!!!



g) Find the 90% confidence lower bound for β .

That is, find the 90% confidence interval for β of the form (a, ∞) .

Hint: Recall Examples for 10/30/2020 (3):

A $(1 - \alpha)$ 100 % confidence interval for β :
$$\left(\frac{\chi^2_{1-\alpha/2}(4n)}{2 \sum_{i=1}^n \ln x_i}, \frac{\chi^2_{\alpha/2}(4n)}{2 \sum_{i=1}^n \ln x_i} \right).$$

$$\left(\frac{\chi^2_{1-\alpha}(4n)}{2 \sum_{i=1}^n \ln x_i}, \infty \right)$$
 would also have a $(1 - \alpha)$ 100 % confidence level.

$$\chi^2_{0.90}(20) = 12.44.$$

$$\left(\frac{\chi^2_{1-\alpha}(4n)}{2 \sum_{i=1}^n \ln x_i}, \infty \right) = \left(\frac{12.44}{2 \cdot 4}, \infty \right) \approx (1.555, \infty).$$

- h) (i) Using the p-value obtained in part (e), state your decision (Reject H_0 or Do NOT Reject H_0) at $\alpha = 0.10$.
- (ii) Does your decision agree with part (g)? That is, does the interval in part (g) cover the value $\beta = 1.4$?

i) p-value $> \alpha \Rightarrow$ Do NOT Reject H_0 .

p-value $< \alpha \Rightarrow$ Reject H_0 .

Since $0.059 < 0.10$, **Reject H_0 at $\alpha = 0.10$.**

- ii) $\beta = 1.4$ is NOT covered by the confidence interval in part (g). That is, based on our data set, the lowest “believable” (with 90% confidence) value of β is 1.555. $\beta = 1.4$ is NOT a “believable” (with 90% confidence) value of β .

Reject $H_0: \beta = 1.4$ at $\alpha = 0.10$.

They agree!!!



Bonus. Let $\theta > 0$ and let X_1, X_2, \dots, X_n be a random sample of size n from the distribution with probability density function

$$f(x; \theta) = \frac{1}{\theta} \cdot x^{\frac{1}{\theta} - 1}, \quad 0 < x < 1.$$

We wish to test $H_0: \theta = 4$ vs. $H_1: \theta < 4$.

Suppose $n = 3$, and $x_1 = 0.2$, $x_2 = 0.3$, $x_3 = 0.5$. Find the p-value.

Recall: Examples for 10/16/2020 (1):

$W = -\ln X$ has Exponential(θ) = Gamma($\alpha = 1, \theta$) distribution.

Examples for 11/02/2020 (1):

$\prod_{i=1}^n X_i$ is a sufficient statistic for θ ,

$\sum_{i=1}^n \ln X_i$ is a sufficient statistic for θ ,

$-\sum_{i=1}^n \ln X_i$ is a sufficient statistic for θ .

$-\sum_{i=1}^3 \ln X_i = \sum_{i=1}^3 W_i$ has a Gamma($\alpha = 3, \theta$) distribution.

θ is “ θ ”. $E(W) = \theta$.

Small $\theta \Rightarrow$ small $w = -\ln x$.

The sign is the same as the sign in H_1 .

Reject H_0 if $-\sum_{i=1}^n \ln x_i \leq c$.

OR

Let $\theta < 4$.

$$\frac{L(4)}{L(\theta)} = \left(\frac{\theta}{4}\right)^n \left(\prod_{i=1}^n x_i\right)^{\frac{1}{4} - \frac{1}{\theta}} \leq k.$$

$$\Leftrightarrow \left(\prod_{i=1}^n x_i\right)^{\frac{1}{4} - \frac{1}{\theta}} \leq k_1$$

$$\frac{1}{4} - \frac{1}{\theta} < 0 \quad \text{since } \theta < 4$$

$$\Leftrightarrow -\sum_{i=1}^n \ln x_i \leq c.$$

Reject H_0 if $-\sum_{i=1}^n \ln x_i \leq c$.

For the p-value, go in the same direction as the “best” rejection region.

$$\text{Observed } -\sum_{i=1}^n \ln x_i = -\ln 0.2 - \ln 0.3 - \ln 0.5 \approx 3.5.$$

$$\text{P-value} = P\left(-\sum_{i=1}^3 \ln X_i \leq 3.5 \mid \theta = 4\right) = P(\text{Gamma}(\alpha = 3, \theta = 4) \leq 3.5).$$

```
> pgamma(3.5, 3, 1/4)
[1] 0.05880372
```

OR

$$\int_0^{3.5} \frac{1}{2 \cdot 4^3} x^{3-1} e^{-\frac{x}{4}} dx = 0.05880 \dots$$

OR

$$\begin{aligned}\text{P-value} &= P\left(-\sum_{i=1}^3 \ln X_i \leq 3.5 \mid \theta = 4\right) = P(\text{Gamma}(\alpha = 3, \theta = 4) \leq 3.5) \\ &= P\left(\text{Poisson}\left(\frac{3.5}{4}\right) \geq 3\right) = 1 - P(\text{Poisson}(0.875) \leq 2) \\ &= 1 - \left(\frac{0.875^0 e^{-0.875}}{0!} + \frac{0.875^1 e^{-0.875}}{1!} + \frac{0.875^2 e^{-0.875}}{2!}\right) \\ &\approx \mathbf{0.0588}.\end{aligned}$$

```
> 1-ppois(2,3.5/4)
[1] 0.05880372
```

OR

$$\begin{aligned}\text{P-value} &= P\left(-\sum_{i=1}^3 \ln X_i \leq 3.5 \mid \theta = 4\right) = P(\text{Gamma}(\alpha = 3, \theta = 4) \leq 3.5) \\ &= P(\chi^2(2 \cdot 3) \leq \frac{2}{4} \cdot 3.5) = P(\chi^2(6) \leq 1.75).\end{aligned}$$

```
> pchisq(2*3.5/4,2*3)
[1] 0.05880372
```