

**0. 2.1.6** (7th and 6th edition)

Let  $f(x, y) = e^{-x-y}$ ,  $0 < x < \infty$ ,  $0 < y < \infty$ , zero elsewhere, be the pdf of  $X$  and  $Y$ . Then if  $Z = X + Y$ , compute  $P(Z \leq 0)$ ,  $P(Z \leq 6)$ , and, more generally,  $P(Z \leq z)$ , for  $0 < z < \infty$ . What is the pdf of  $Z$ ?

$$F_Z(z) = P(Z \leq z) = P(Y \leq z - X)$$

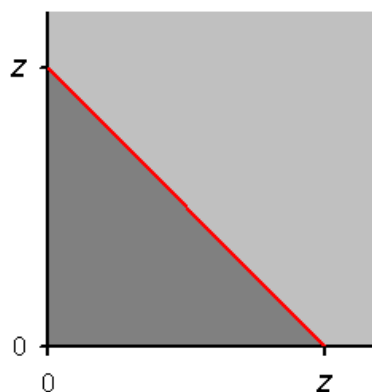
$$= \int_0^z \left( \int_0^{z-x} e^{-x-y} dy \right) dx$$

$$= \int_0^z e^{-x} \left( \int_0^{z-x} e^{-y} dy \right) dx$$

$$= \int_0^z e^{-x} (1 - e^{-z+x}) dx$$

$$= \int_0^z e^{-x} dx - \int_0^z e^{-z} dx = 1 - e^{-z} - z e^{-z}, \quad z > 0.$$

$$f_Z(z) = F'_Z(z) = e^{-z} - e^{-z} + z e^{-z} = z e^{-z}, \quad z > 0.$$



Another approach:

$X$  and  $Y$  are two independent Exponential random variables with mean 1.

$$M_X(t) = \frac{1}{1-t}, \quad t < 1. \quad M_Y(t) = \frac{1}{1-t}, \quad t < 1.$$

$$\Rightarrow M_{X+Y}(t) = M_X(t) \cdot M_Y(t) = \left( \frac{1}{1-t} \right)^2, \quad t < 1.$$

$$\Rightarrow Z = X + Y \text{ has a Gamma distribution with } \alpha = 2, \theta = 1.$$

$$\Rightarrow f_Z(z) = \frac{1}{\Gamma(2) \cdot 1^2} \cdot z^{2-1} \cdot e^{-z/1} = z e^{-z}, \quad z > 0.$$

b)  $W = 2X + Y;$

$$M_{2X+Y}(t) = E(e^{2Xt+Yt}) = M_X(2t) \cdot M_Y(t) = \left(\frac{1}{1-2t}\right) \cdot \left(\frac{1}{1-t}\right), \quad t < 1/2.$$

$$f_{2X+Y}(w) = ???$$

$$\begin{aligned} F_W(w) &= P(W \leq w) = P(Y \leq w - 2X) = \int_0^{w/2} \left( \int_0^{w-2x} e^{-x-y} dy \right) dx \\ &= \int_0^{w/2} e^{-x} \left( -e^{-y} \right) \Big|_0^{w-2x} dx = \int_0^{w/2} (e^{-x} - e^{-w+x}) dx \\ &= \left( -e^{-x} - e^{-w} e^x \right) \Big|_0^{w/2} = 1 + e^{-w} - 2e^{-w/2}, \quad w > 0. \end{aligned}$$

$$f_W(w) = F'_W(w) = e^{-w/2} - e^{-w}, \quad w > 0.$$

c)  $V = \frac{Y}{X};$

$$\begin{aligned} F_V(v) &= P(V \leq v) = P(Y \leq vX) = \int_0^\infty \left( \int_{y/v}^\infty (e^{-x-y}) dx \right) dy \\ &= \int_0^\infty e^{-y} \left( \int_{y/v}^\infty (e^{-x}) dx \right) dy = \int_0^\infty e^{-y} (e^{-y/v}) dy = \int_0^\infty e^{-(1+1/v)y} dy \\ &= \frac{v}{1+v} = 1 - \frac{1}{1+v}, \quad v > 0. \end{aligned}$$

$$f_V(v) = \frac{1}{(1+v)^2}, \quad v > 0.$$

$$d) \quad U = \frac{X}{X+Y};$$

$$F_U(u) = P(U \leq u) = P(X \leq u(X+Y)) = P(Y \geq \frac{1-u}{u} X) = \frac{1}{1 + \frac{1-u}{u}} = u,$$

$$0 < u < 1.$$

$$\left[ P(Y \leq v X) = 1 - \frac{1}{1+v} \quad \Rightarrow \quad P(Y \geq v X) = \frac{1}{1+v}, \quad \text{then } v = \frac{1-u}{u}. \right]$$

$$f_U(u) = 1, \quad 0 < u < 1.$$

U has a Uniform(0, 1) distribution.

OR

$$U = \frac{X}{X+Y} = \frac{1}{1 + \frac{Y}{X}} = \frac{1}{1+V} = g(V).$$

$$V = \frac{1}{U} - 1 = g^{-1}(U) \quad \Rightarrow \quad \frac{dv}{du} = -\frac{1}{u^2}.$$

$$f_U(u) = f_V(g^{-1}(u)) \times \left| \frac{dv}{du} \right| = \frac{1}{\left(1 + \frac{1}{u} - 1\right)^2} \times \left| -\frac{1}{u^2} \right| = 1, \quad 0 < u < 1.$$

U has a Uniform(0, 1) distribution.

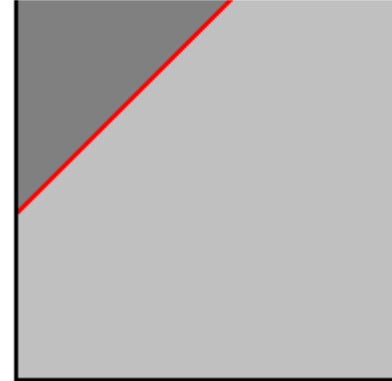
e)  $T = X - Y.$

$$F_T(t) = P(T \leq t) = P(X - Y \leq t) = \dots$$

Case 1:  $t < 0.$

$$\begin{aligned} \dots &= P(Y \geq X - t) \\ &= \int_0^\infty \left( \int_{x-t}^\infty (e^{-x-y}) dy \right) dx \\ &= \int_0^\infty e^{-2x+t} dx \\ &= \frac{1}{2} e^t, \quad t < 0. \end{aligned}$$

$$f_T(t) = \frac{1}{2} e^t, \quad t < 0.$$



Case 2:  $t > 0.$

$$\begin{aligned} \dots &= 1 - P(X \geq Y + t) \\ &= 1 - \int_0^\infty \left( \int_{y+t}^\infty (e^{-x-y}) dx \right) dy \\ &= 1 - \int_0^\infty e^{-2y-t} dy \\ &= 1 - \frac{1}{2} e^{-t}, \quad t > 0. \end{aligned}$$

$$f_T(t) = \frac{1}{2} e^{-t}, \quad t > 0.$$

$$\Rightarrow f_T(t) = \frac{1}{2} e^{-|t|}, \quad -\infty < t < \infty.$$



Another approach:

$X$  and  $Y$  are two independent Exponential random variables with mean 1.

$$M_X(s) = \frac{1}{1-s}, \quad s < 1. \qquad M_Y(s) = \frac{1}{1-s}, \quad s < 1.$$

$$\Rightarrow M_{X-Y}(s) = M_X(s) \cdot M_Y(-s) = \frac{1}{1-s} \cdot \frac{1}{1+s} = \frac{1}{1-s^2}, \quad -1 < s < 1.$$

$\Rightarrow$  ( Recall Example 13 from STAT 400 Review (1) )

$$f_T(t) = \frac{1}{2} e^{-|t|}, \quad -\infty < t < \infty.$$

( double exponential p.d.f. )