

Examples for 10/19/2020 (4) & 10/21/2020 (2) (continued)

1. The Pareto probability distribution has many applications in economics, biology, and physics. Let $\beta > 0$ and $\delta > 0$ be the population parameters, and let X_1, X_2, \dots, X_n be a random sample from the distribution with probability density function

$$f(x; \beta, \delta) = \frac{\delta \cdot \beta^\delta}{x^{\delta+1}}, \quad x > \beta, \quad \text{zero otherwise.}$$

Suppose β is known.

Recall: a method of moments estimator of δ is $\tilde{\delta} = \frac{\bar{X}}{\bar{X} - \beta}$; $E(X) = \frac{\beta \delta}{\delta - 1}$;

the maximum likelihood estimator of δ is $\hat{\delta} = \frac{n}{\sum_{i=1}^n \ln\left(\frac{X_i}{\beta}\right)} = \frac{n}{\sum_{i=1}^n \ln X_i - n \cdot \ln \beta}$;

$W = \ln\left(\frac{X}{\beta}\right) = \ln X - \ln \beta$ has an Exponential($\theta = \frac{1}{\delta}$) distribution.

- u) Show that $\tilde{\delta}$ is a consistent estimator of δ .
(NOT enough to say “because it is a method of moments estimator”)
- v) Show that $\hat{\delta}$ is a consistent estimator of δ .
(NOT enough to say “because it is the maximum likelihood estimator”)

Suppose δ is known.

Recall: $E(X) = \frac{\beta \delta}{\delta - 1}; \quad F_X(x) = 1 - \frac{\beta^\delta}{x^\delta}, \quad x > \beta;$

a method of moments estimator of β is $\tilde{\beta} = \frac{\delta - 1}{\delta} \bar{X};$

the maximum likelihood estimator of β is $\hat{\beta} = \min X_i.$

w) Show that $\tilde{\beta}$ is a consistent estimator of β .

(NOT enough to say “because it is a method of moments estimator”)

x) Show that $\hat{\beta}$ is a consistent estimator of β .

(NOT enough to say “because it is the maximum likelihood estimator”)

Answers:

1. The Pareto probability distribution has many applications in economics, biology, and physics. Let $\beta > 0$ and $\delta > 0$ be the population parameters, and let X_1, X_2, \dots, X_n be a random sample from the distribution with probability density function

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- u) Show that $\tilde{\delta}$ is a consistent estimator of δ .
(NOT enough to say “because it is a method of moments estimator”)

By WLLN, $\bar{X} \xrightarrow{P} \mu = E(X) = \frac{\beta \delta}{\delta - 1}$.

$X_n \xrightarrow{P} a$, g is continuous at $a \Rightarrow g(X_n) \xrightarrow{P} g(a)$

$g(x) = \frac{x}{x - \beta}$ is continuous at μ .

$$\tilde{\delta} = \frac{\bar{X}}{\bar{X} - \beta} = g(\bar{X}) \xrightarrow{P} g(\mu) = \frac{\mu}{\mu - \beta} = \frac{\frac{\beta \delta}{\delta - 1}}{\frac{\beta \delta}{\delta - 1} - \beta} = \delta.$$

- v) Show that $\hat{\delta}$ is a consistent estimator of δ .
 (NOT enough to say “because it is the maximum likelihood estimator”)

$$\text{By WLLN, } \overline{W} \xrightarrow{P} E(W) = \frac{1}{\delta}.$$

$$\hat{\delta} = \frac{n}{\sum_{i=1}^n \ln\left(\frac{X_i}{\beta}\right)} = \frac{1}{\overline{W}} \xrightarrow{P} \frac{1}{E(W)} = \delta.$$

OR

$$\begin{aligned} E(\ln X) &= \int_{\beta}^{\infty} \ln x \cdot \frac{\delta \cdot \beta^{\delta}}{x^{\delta+1}} dx = \delta \cdot \beta^{\delta} \cdot \int_{\beta}^{\infty} \ln x \cdot x^{-\delta-1} dx \\ &= \delta \cdot \beta^{\delta} \cdot \left[\frac{1}{-\delta} \cdot \ln x \cdot x^{-\delta} - \frac{1}{(-\delta)^2} \cdot x^{-\delta} \right] \Bigg|_{\beta}^{\infty} = \ln \beta + \frac{1}{\delta}. \end{aligned}$$

$$\text{By WLLN, } \frac{1}{n} \sum_{i=1}^n \ln X_i = \overline{\ln X} \xrightarrow{P} E(\ln X).$$

$$\hat{\delta} = \frac{n}{\sum_{i=1}^n \ln X_i - n \cdot \ln \beta} = \frac{1}{\overline{\ln X} - \ln \beta} \xrightarrow{P} \frac{1}{E(\ln X) - \ln \beta} = \delta.$$

OR

$$\text{MSE}(\hat{\delta}) = \frac{(n+2) \delta^2}{(n-2)(n-1)} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

$$\Rightarrow \hat{\delta} \xrightarrow{P} \delta.$$

Suppose δ is known.

Recall: $E(X) = \frac{\beta \delta}{\delta - 1}; \quad F_X(x) = 1 - \frac{\beta^\delta}{x^\delta}, \quad x > \beta;$

a method of moments estimator of β is $\tilde{\beta} = \frac{\delta - 1}{\delta} \bar{X};$

the maximum likelihood estimator of β is $\hat{\beta} = \min X_i.$

w) Show that $\tilde{\beta}$ is a consistent estimator of β .

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By WLLN, $\bar{X} \xrightarrow{P} \mu = E(X) = \frac{\beta \delta}{\delta - 1}.$

$$X_n \xrightarrow{P} X, \quad a = \text{const} \Rightarrow a X_n \xrightarrow{P} a X$$

$$\tilde{\beta} = \frac{\delta - 1}{\delta} \bar{X} \xrightarrow{P} \frac{\delta - 1}{\delta} \cdot \frac{\beta \delta}{\delta - 1} = \beta.$$

OR

$$\text{MSE}(\tilde{\beta}) = \frac{\beta^2}{(\delta - 2) \delta n} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

$$\Rightarrow \tilde{\beta} \xrightarrow{P} \beta.$$

- x) Show that $\hat{\beta}$ is a consistent estimator of β .
 (NOT enough to say “because it is the maximum likelihood estimator”)

$$F_{\min X_i}(x) = 1 - (1 - F_X(x))^n = 1 - \left(\frac{\beta^\delta}{x^\delta}\right)^n = 1 - \frac{\beta^{\delta n}}{x^{\delta n}}, \quad x > \beta.$$

Let $\varepsilon > 0$. Then

$$\begin{aligned} P\left(\left|\hat{\beta} - \beta\right| \geq \varepsilon\right) &= P(\min X_i \leq \beta - \varepsilon) + P(\min X_i \geq \beta + \varepsilon) \\ &= 0 + P(\min X_i \geq \beta + \varepsilon) \\ &= 1 - F_{\min X_i}(\beta + \varepsilon) = \left(\frac{\beta}{\beta + \varepsilon}\right)^{\delta n} \rightarrow 0 \quad \text{as } n \rightarrow \infty. \end{aligned}$$

$$\Rightarrow \hat{\beta} \xrightarrow{P} \beta.$$

OR

$$\text{MSE}(\hat{\beta}) = \frac{2\beta^2}{(\delta n - 2)(\delta n - 1)} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

$$\Rightarrow \tilde{\beta} \xrightarrow{P} \beta.$$