

## Examples for 10/19/2020 (4) (continued)

1. The Pareto probability distribution has many applications in economics, biology, and physics. Let  $\beta > 0$  and  $\delta > 0$  be the population parameters, and let  $X_1, X_2, \dots, X_n$  be a random sample from the distribution with probability density function

$$f(x; \beta, \delta) = \frac{\delta \cdot \beta^\delta}{x^{\delta+1}}, \quad x > \beta, \quad \text{zero otherwise.}$$

Suppose  $\delta$  is known.

- l) Obtain the maximum likelihood estimator of  $\beta$ ,  $\hat{\beta}$ .
- m) Suppose  $n = 5$ ,  $\delta = 1.5$ , and  $x_1 = 3.9$ ,  $x_2 = 4.2$ ,  $x_3 = 6$ ,  $x_4 = 9$ ,  $x_5 = 15$ . Find the maximum likelihood estimate of  $\beta$ .
- n) Assume  $n > \frac{1}{\delta}$ . Is the maximum likelihood estimator  $\hat{\beta}$  an unbiased estimator of  $\beta$ ? If  $\hat{\beta}$  is not an unbiased estimator of  $\beta$ , construct an unbiased estimator of  $\beta$  based on  $\hat{\beta}$ .
- o) Assume  $n > \frac{2}{\delta}$ . Find  $\text{MSE}(\hat{\beta}) = (\text{bias}(\hat{\beta}))^2 + \text{Var}(\hat{\beta})$ .

- p) Assume  $\delta > 1$ . Obtain a method of moments estimator of  $\beta$ ,  $\tilde{\beta}$ .
- q) Suppose  $n = 5$ ,  $\delta = 1.5$ , and  $x_1 = 3.9$ ,  $x_2 = 4.2$ ,  $x_3 = 6$ ,  $x_4 = 9$ ,  $x_5 = 15$ . Find a method of moments estimate of  $\beta$ .
- r) Assume  $\delta > 1$ . Is the method of moments estimator of  $\beta$ ,  $\tilde{\beta}$ , an unbiased estimator of  $\beta$ ? If  $\tilde{\beta}$  is not an unbiased estimator of  $\beta$ , construct an unbiased estimator of  $\beta$  based on  $\tilde{\beta}$ .
- s) Assume  $\delta > 2$ . Find  $\text{MSE}(\tilde{\beta}) = (\text{bias}(\tilde{\beta}))^2 + \text{Var}(\tilde{\beta})$ .
- t) Which estimator is “better”,  $\hat{\beta}$  or  $\tilde{\beta}$ ? *Justify your decision.*

“Hint”:  $E(\bar{V}) = \mu_V = E(V)$ .  $\text{Var}(\bar{V}) = \frac{\sigma_V^2}{n} = \frac{\text{Var}(V)}{n}$ .

$$\text{Var}(V) = E(V^2) - [E(V)]^2.$$

$$E(a \odot) = a E(\odot).$$

$$\text{Var}(a \odot) = a^2 \text{Var}(\odot).$$

$$F_{\max X_i}(x) = (F(x))^n.$$

$$f_{\max X_i}(x) = n \cdot (F(x))^{n-1} \cdot f(x).$$

$$F_{\min X_i}(x) = 1 - (1 - F(x))^n.$$

$$f_{\min X_i}(x) = n \cdot (1 - F(x))^{n-1} \cdot f(x).$$

## Answers:

1. The Pareto probability distribution has many applications in economics, biology, and physics. Let  $\beta > 0$  and  $\delta > 0$  be the population parameters, and let  $X_1, X_2, \dots, X_n$  be a random sample from the distribution with probability density function

$$f(x; \beta, \delta) = \frac{\delta \cdot \beta^\delta}{x^{\delta+1}}, \quad x > \beta, \quad \text{zero otherwise.}$$

Suppose  $\delta$  is known.

- 1) Obtain the maximum likelihood estimator of  $\beta$ ,  $\hat{\beta}$ .

$$L(\beta) = \prod_{i=1}^n f(x_i; \beta, \delta) = \dots$$

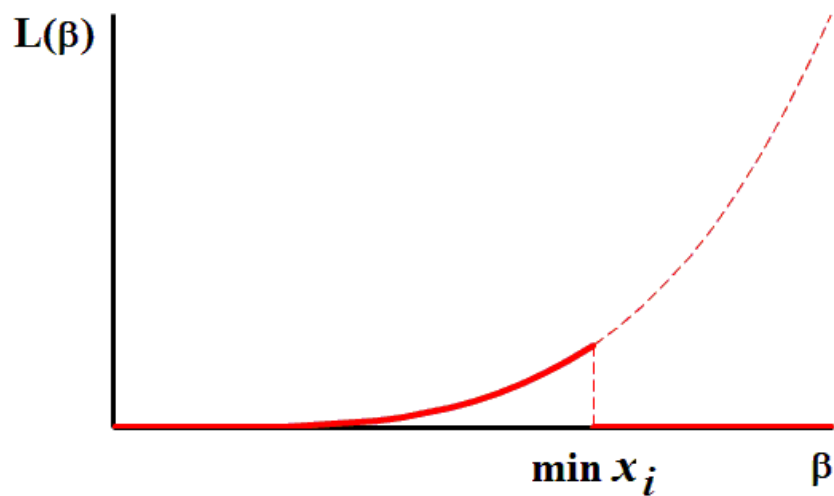
$$\dots = \delta^n \cdot \beta^{n\delta} \cdot \left( \prod_{i=1}^n x_i \right)^{-(\delta+1)}, \quad x_1 > \beta, x_2 > \beta, \dots, x_n > \beta,$$

$$\dots = 0, \quad \text{otherwise.}$$

$$\ln L(\beta) = n \cdot \ln \delta + n \delta \cdot \ln \beta - (\delta + 1) \cdot \sum_{i=1}^n \ln x_i.$$

$$\frac{d}{d\beta} \ln L(\beta) = \frac{n\delta}{\beta} = 0 \quad ???$$

$$x_1 > \beta, x_2 > \beta, \dots, x_n > \beta \quad \Leftrightarrow \quad \beta < \min x_i.$$



$$\Rightarrow \quad \hat{\beta} = \min X_i.$$

- m) Suppose  $n = 5$ ,  $\delta = 1.5$ , and  $x_1 = 3.9$ ,  $x_2 = 4.2$ ,  $x_3 = 6$ ,  $x_4 = 9$ ,  $x_5 = 15$ .  
Find the maximum likelihood estimate of  $\beta$ .

$$x_1 = 3.9, x_2 = 4.2, x_3 = 6, x_4 = 9, x_5 = 15. \quad \hat{\beta} = \min x_i = \mathbf{3.9}.$$

- n) Assume  $n > \frac{1}{\delta}$ . Is the maximum likelihood estimator  $\hat{\beta}$  an unbiased estimator of  $\beta$ ?  
 If  $\hat{\beta}$  is not an unbiased estimator of  $\beta$ , construct an unbiased estimator of  $\beta$  based on  $\hat{\beta}$ .

$$F_X(x) = P(X \leq x) = \int_{\beta}^x \frac{\delta \cdot \beta^{\delta}}{u^{\delta+1}} du = -\frac{\beta^{\delta}}{u^{\delta}} \Big|_{\beta}^x = 1 - \frac{\beta^{\delta}}{x^{\delta}}, \quad x > \beta.$$

$$F_{\min X_i}(x) = 1 - (1 - F_X(x))^n = 1 - \left(\frac{\beta^{\delta}}{x^{\delta}}\right)^n = 1 - \frac{\beta^{\delta n}}{x^{\delta n}}, \quad x > \beta.$$

$$f_{\min X_i}(x) = \frac{\delta n \beta^{\delta n}}{x^{\delta n+1}}, \quad x > \beta.$$

$$\begin{aligned} E(\hat{\beta}) &= E(\min X_i) = \int_{\beta}^{\infty} x \cdot \frac{\delta n \beta^{\delta n}}{x^{\delta n+1}} dx = \delta n \beta^{\delta n} \cdot \int_{\beta}^{\infty} x^{-\delta n} dx \\ &= \frac{\delta n}{\delta n - 1} \beta \neq \beta. \end{aligned}$$

$\hat{\beta}$  is NOT an unbiased estimator of  $\beta$ .

$$\text{Consider } \hat{\hat{\beta}} = \frac{\delta n - 1}{\delta n} \hat{\beta} = \frac{\delta n - 1}{\delta n} \min X_i.$$

$$E(\hat{\hat{\beta}}) = \frac{\delta n - 1}{\delta n} E(\hat{\beta}) = \beta.$$

$\hat{\hat{\beta}}$  is an unbiased estimator of  $\beta$ .

o) Assume  $n > \frac{2}{\delta}$ . Find  $\text{MSE}(\hat{\beta}) = (\text{bias}(\hat{\beta}))^2 + \text{Var}(\hat{\beta})$ .

$$\text{E}(\hat{\beta}^2) = \text{E}[(\min X_i)^2] = \int_{\beta}^{\infty} x^2 \cdot \frac{\delta n \beta^{\delta n}}{x^{\delta n + 1}} dx = \frac{\delta n}{\delta n - 2} \beta^2.$$

$$\text{Var}(\hat{\beta}) = \text{Var}(\min X_i) = \frac{\delta n}{\delta n - 2} \beta^2 - \left( \frac{\delta n}{\delta n - 1} \beta \right)^2 = \frac{\beta^2 \delta n}{(\delta n - 2)(\delta n - 1)^2}.$$

$$\text{bias}(\hat{\beta}) = \text{E}(\hat{\beta}) - \beta = \frac{\delta n}{\delta n - 1} \beta - \beta = \frac{\beta}{\delta n - 1}.$$

$$\begin{aligned} \text{MSE}(\hat{\beta}) &= (\text{bias}(\hat{\beta}))^2 + \text{Var}(\hat{\beta}) = \left( \frac{\beta}{\delta n - 1} \right)^2 + \frac{\beta^2 \delta n}{(\delta n - 2)(\delta n - 1)^2} \\ &= \frac{2 \beta^2}{(\delta n - 2)(\delta n - 1)}. \end{aligned}$$

p) Assume  $\delta > 1$ . Obtain a method of moments estimator of  $\beta$ ,  $\tilde{\beta}$ .

$$E(X) = \int_{\beta}^{\infty} x \cdot \frac{\delta \cdot \beta^{\delta}}{x^{\delta+1}} dx = \delta \beta^{\delta} \cdot \int_{\beta}^{\infty} x^{-\delta} dx = \frac{\beta \delta}{\delta - 1}.$$

$$\bar{X} = \frac{\beta \delta}{\delta - 1} \Rightarrow \tilde{\beta} = \frac{\delta - 1}{\delta} \bar{X}.$$

q) Suppose  $n = 5$ ,  $\delta = 1.5$ , and  $x_1 = 3.9$ ,  $x_2 = 4.2$ ,  $x_3 = 6$ ,  $x_4 = 9$ ,  $x_5 = 15$ . Find a method of moments estimate of  $\beta$ .

$$n = 5, \quad \beta = 3, \quad x_1 = 3.9, \quad x_2 = 4.2, \quad x_3 = 6, \quad x_4 = 9, \quad x_5 = 15.$$

$$\sum_{i=1}^n x_i = 38.1. \quad \bar{x} = 7.62. \quad \tilde{\beta} = \frac{1.5-1}{1.5} \cdot 7.62 = \mathbf{2.54}.$$

r) Assume  $\delta > 1$ . Is the method of moments estimator of  $\beta$ ,  $\tilde{\beta}$ , an unbiased estimator of  $\beta$ ? If  $\tilde{\beta}$  is not an unbiased estimator of  $\beta$ , construct an unbiased estimator of  $\beta$  based on  $\tilde{\beta}$ .

$$E(\tilde{\beta}) = \frac{\delta - 1}{\delta} E(\bar{X}) = \frac{\delta - 1}{\delta} \mu = \frac{\delta - 1}{\delta} \cdot \frac{\beta \delta}{\delta - 1} = \beta.$$

$\tilde{\beta}$  is an unbiased estimator of  $\beta$ .

s) Assume  $\delta > 2$ . Find  $\text{MSE}(\tilde{\beta}) = (\text{bias}(\tilde{\beta}))^2 + \text{Var}(\tilde{\beta})$ .

$$E(X^2) = \int_{\beta}^{\infty} x^2 \cdot \frac{\delta \cdot \beta^{\delta}}{x^{\delta+1}} dx = \frac{\beta^2 \delta}{\delta-2}.$$

$$\sigma^2 = \text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{\beta^2 \delta}{\delta-2} - \left(\frac{\beta \delta}{\delta-1}\right)^2 = \frac{\beta^2 \delta}{(\delta-2)(\delta-1)^2}.$$

$$\text{Var}(\tilde{\beta}) = \left(\frac{\delta-1}{\delta}\right)^2 \text{Var}(\bar{X}) = \left(\frac{\delta-1}{\delta}\right)^2 \cdot \frac{\sigma^2}{n} = \frac{\beta^2}{(\delta-2)\delta n}.$$

Since  $\tilde{\beta}$  is an unbiased estimator of  $\beta$ ,  $\text{bias}(\tilde{\beta}) = 0$ ,

$$\text{and } \text{MSE}(\tilde{\beta}) = \text{Var}(\tilde{\beta}) = \frac{\beta^2}{(\delta-2)\delta n}.$$

t) Which estimator is “better”,  $\hat{\beta}$  or  $\tilde{\beta}$ ? *Justify your decision.*

$$\text{MSE}(\hat{\beta}) = \frac{2\beta^2}{(\delta n-2)(\delta n-1)} \sim \frac{\text{const}}{n^2}.$$

$$\text{MSE}(\tilde{\beta}) = \frac{\beta^2}{(\delta-2)\delta n} \sim \frac{\text{const}}{n}.$$

$\text{MSE}(\hat{\beta})$  decreases faster than  $\text{MSE}(\tilde{\beta})$  as  $n$  increases.

$\text{MSE}(\hat{\beta})$  will be smaller than  $\text{MSE}(\tilde{\beta})$  for larger  $n$ .

$\hat{\beta}$  is a better estimator.