## 2.5 Independent Random Variables

1. Consider the following joint probability distribution p(x, y) of two random variables X and Y:

$x \setminus y$	0	1	2	
1	0.15	0.10	0	0.25
2	0.25	0.30	0.20	0.75
	0.40	0.40	0.20	

Recall: A and B are independent if and only if  $P(A \cap B) = P(A) \cdot P(B)$ .

a) Are events  $\{X = 1\}$  and  $\{Y = 1\}$  independent?

$$P(X = 1 \cap Y = 1) = p(1, 1) = 0.10 = 0.25 \times 0.40 = P(X = 1) \times P(Y = 1).$$
 {X = 1} and {Y = 1} are **independent**.

**Def** Random variables X and Y are **independent** if and only if

discrete  $p(x, y) = p_X(x) \cdot p_Y(y)$  for all x, y.

continuous  $f(x, y) = f_{\mathbf{v}}(x) \cdot f_{\mathbf{v}}(y)$  for all x, y.

$$F(x,y) = P(X \le x, Y \le y). \qquad f(x,y) = \frac{\partial^2 F(x,y)}{\partial x \partial y}.$$

**Def** Random variables X and Y are **independent** if and only if

$$F(x, y) = F_X(x) \cdot F_Y(y)$$
 for all  $x, y$ .

b) Are random variables X and Y independent?

$$p(1,0) = 0.15 \neq 0.25 \times 0.40 = p_X(1) \times p_Y(0).$$

X and Y are **NOT independent**.

2. Let the joint probability density function for (X, Y) be

$$f(x,y) = \begin{cases} 60 x^2 y & 0 \le x \le 1, \ 0 \le y \le 1, \ x+y \le 1 \\ 0 & \text{otherwise} \end{cases}$$

Recall:

$$f_X(x) = 30 x^2 (1-x)^2, \quad 0 < x < 1,$$
  
 $f_Y(y) = 20 y (1-y)^3, \quad 0 < y < 1.$ 

Are random variables X and Y independent?

The support of (X, Y) is not a rectangle.

X and Y are **NOT independent**.

OR

Since 
$$f(x, y) \neq f_X(x) \cdot f_Y(y)$$
,

X and Y are **NOT independent**.

3. Let the joint probability density function for (X, Y) be

$$f(x,y) = \begin{cases} x+y & 0 \le x \le 1, 0 \le y \le 1 \\ 0 & \text{otherwise} \end{cases}$$

Are X and Y independent?

$$f_{1}(x) = \int_{0}^{1} (x+y) dy$$

$$= \left[ xy + \frac{1}{2}y^{2} \right]_{0}^{1} = x + \frac{1}{2}, \quad 0 \le x \le 1 ;$$

$$f_{2}(y) = \int_{0}^{1} (x+y) dx = y + \frac{1}{2}, \quad 0 \le y \le 1;$$

$$f(x,y) = x + y \ne \left( x + \frac{1}{2} \right) \left( y + \frac{1}{2} \right) = f_{1}(x) f_{2}(y).$$

X and Y are **NOT independent**.

**4.** Let the joint probability density function for (X, Y) be

$$f(x,y) = \begin{cases} 12 x (1-x) e^{-2y} & 0 \le x \le 1, y \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

Are X and Y independent?

$$f_X(x) = \int_0^\infty 12 x(1-x)e^{-2y} dy = 6x(1-x), \quad 0 < x < 1.$$

$$f_{Y}(y) = \int_{0}^{1} 12 x (1-x) e^{-2y} dx = 2 e^{-2y}, y > 0.$$

Since  $f(x, y) = f_X(x) \cdot f_Y(y)$  for all x, y, X and Y are **independent**.

If random variables X and Y are independent, then

$$E(g(X) \cdot h(Y)) = E(g(X)) \cdot E(h(Y)).$$

$$\Rightarrow M_{X,Y}(t_1, t_2) = M_{X,Y}(t_1, 0) \cdot M_{X,Y}(0, t_2) = M_X(t_1) \cdot M_Y(t_2).$$

5. Suppose the probability density functions of  $T_1$  and  $T_2$  are

$$f_{T_1}(x) = \alpha e^{-\alpha x}, \quad x > 0, \qquad f_{T_2}(y) = \beta e^{-\beta y}, \quad y > 0,$$

respectively. Suppose  $T_1$  and  $T_2$  are independent. Find  $P(2T_1 > T_2)$ .

$$P(2T_1 > T_2) = \int_0^\infty \left( \int_{y/2}^\infty \left( \alpha \beta e^{-\alpha x - \beta y} \right) dx \right) dy = \int_0^\infty \beta e^{-\beta y} \left( \int_{y/2}^\infty \left( \alpha e^{-\alpha x} \right) dx \right) dy$$
$$= \int_0^\infty \beta e^{-\beta y} \left( e^{-\alpha y/2} \right) dy = \int_0^\infty \beta e^{-(\alpha/2 + \beta)y} dy = \frac{2\beta}{\alpha + 2\beta}.$$

6. Let X and Y be two independent random variables, X has a Geometric distribution with the probability of "success" p = 1/3, Y has a Poisson distribution with mean 3. That is,

$$p_{X}(x) = \left(\frac{1}{3}\right) \cdot \left(\frac{2}{3}\right)^{x-1}, \quad x = 1, 2, 3, \dots,$$

$$p_{Y}(y) = \frac{3^{y} e^{-3}}{y!}, \quad y = 0, 1, 2, 3, \dots.$$

a) Find P(X = Y).

$$P(X = Y) = \sum_{k=1}^{\infty} p_X(k) \cdot p_Y(k) = \sum_{k=1}^{\infty} \left(\frac{1}{3}\right) \cdot \left(\frac{2}{3}\right)^{k-1} \cdot \frac{3^k e^{-3}}{k!}$$

$$= e^{-3} \cdot \sum_{k=1}^{\infty} \frac{2^{k-1}}{k!} = \frac{e^{-3}}{2} \cdot \left[\sum_{k=0}^{\infty} \frac{2^k}{k!} - 1\right] = \frac{e^{-3}}{2} \cdot \left[e^2 - 1\right]$$

$$= \frac{e^{-1} - e^{-3}}{2} \approx 0.159.$$

b) Find P(X = 2Y).

$$P(X = 2Y) = \sum_{k=1}^{\infty} p_X(2k) \cdot p_Y(k) = \sum_{k=1}^{\infty} \left(\frac{1}{3}\right) \cdot \left(\frac{2}{3}\right)^{2k-1} \cdot \frac{3^k e^{-3}}{k!}$$
$$= \frac{e^{-3}}{2} \cdot \sum_{k=1}^{\infty} \left(\frac{4}{3}\right)^k \cdot \frac{1}{k!} = \frac{e^{-3}}{2} \cdot \left[e^{4/3} - 1\right] \approx 0.069544.$$

c) Find P(X > Y).

$$P(X > Y) = \sum_{y=0}^{\infty} \sum_{x=y+1}^{\infty} \left(\frac{1}{3}\right) \cdot \left(\frac{2}{3}\right)^{x-1} \cdot \frac{3^{y} e^{-3}}{y!} = \sum_{y=0}^{\infty} \left(\frac{2}{3}\right)^{y} \cdot \frac{3^{y} e^{-3}}{y!}$$
$$= e^{-3} \cdot \sum_{y=0}^{\infty} \frac{2^{y}}{y!} = e^{-1} \approx 0.368.$$

## 2.4 Covariance and Correlation Coefficient

Covariance of X and Y

$$\sigma_{XY} = \text{Cov}(X,Y) = \text{E}[(X - \mu_X)(Y - \mu_Y)] = \text{E}(XY) - \mu_X \mu_Y$$

- (a) Cov(X,X) = Var(X);
- (b) Cov(X,Y) = Cov(Y,X);
- (c)  $\operatorname{Cov}(aX+b,Y) = a\operatorname{Cov}(X,Y);$
- (d) Cov(X+Y,W) = Cov(X,W) + Cov(Y,W).

$$Cov(aX+bY, cX+dY)$$

$$= a c Var(X) + (ad+bc)Cov(X,Y) + b d Var(Y).$$

$$Var(aX + bY) = Cov(aX + bY, aX + bY)$$
$$= a^{2}Var(X) + 2abCov(X,Y) + b^{2}Var(Y).$$

- **0.** Find in terms of  $\sigma_X^2$ ,  $\sigma_Y^2$ , and  $\sigma_{XY}$ :
- a) Cov(2X + 3Y, X 2Y),

$$Cov(2X+3Y,X-2Y) = 2Var(X)-Cov(X,Y)-6Var(Y).$$

b) Var(2X+3Y),

$$Var(2X+3Y) = Cov(2X+3Y,2X+3Y)$$
  
=  $4Var(X) + 12Cov(X,Y) + 9Var(Y)$ .

c) Var(X-2Y).

$$Var(X-2Y) = Cov(X-2Y, X-2Y)$$
$$= Var(X)-4Cov(X,Y)+4Var(Y).$$

$$E(aX + bY) = aE(X) + bE(Y),$$

$$Var(aX + bY) = a^{2}Var(X) + 2abCov(X,Y) + b^{2}Var(Y).$$

If  $X_1, X_2, \dots, X_n$  are n random variables and  $a_0, a_1, a_2, \dots, a_n$  are n+1 constants, then the random variable  $U = a_0 + a_1 X_1 + a_2 X_2 + \dots + a_n X_n$  has mean

$$E(U) = a_0 + a_1 E(X_1) + a_2 E(X_2) + ... + a_n E(X_n)$$

and variance

$$Var(U) = \sum_{i=1}^{n} a_{i}^{2} Var(X_{i}) + \sum_{i=1}^{n} \sum_{j=1}^{n} a_{i} a_{j} Cov(X_{i}, X_{j})$$

$$= \sum_{i=1}^{n} a_{i}^{2} Var(X_{i}) + 2 \sum_{i < j} \sum_{i < j} a_{i} a_{j} Cov(X_{i}, X_{j})$$

Correlation coefficient of X and Y

$$\rho_{XY} = \frac{\sigma_{XY}}{\sigma_{X}\sigma_{Y}} = \frac{Cov(X,Y)}{\sqrt{Var(X)} \cdot \sqrt{Var(Y)}} = E \left[ \left( \frac{X - \mu_{X}}{\sigma_{X}} \right) \cdot \left( \frac{Y - \mu_{Y}}{\sigma_{Y}} \right) \right]$$

- (a)  $-1 \le \rho_{XY} \le 1$ ;
- (b)  $\rho_{XY}$  is either +1 or -1 if and only if X and Y are linear functions of one another.

If random variables X and Y are independent, then

$$E(g(X) \cdot h(Y)) = E(g(X)) \cdot E(h(Y)).$$

$$\Rightarrow \quad Cov(X, Y) = \sigma_{XY} = 0, \quad Corr(X, Y) = \rho_{XY} = 0.$$

If  $X_1, X_2, ..., X_n$  are n independent random variables and  $a_0, a_1, a_2, ..., a_n$  are n + 1 constants, then the random variable  $U = a_0 + a_1 X_1 + a_2 X_2 + ... + a_n X_n$  has variance

$$Var(U) = a_1^2 Var(X_1) + a_2^2 Var(X_2) + ... + a_n^2 Var(X_n)$$

1. Consider the following joint probability distribution p(x, y) of two random variables X and Y:

Find  $Cov(X,Y) = \sigma_{XY}$  and  $Corr(X,Y) = \rho_{XY}$ .

$$Cov(X,Y) = \sigma_{XY} = 1.5 - 1.75 \cdot 0.8 = 0.10.$$

$$E(X^2) = 1 \times 0.25 + 4 \times 0.75 = 3.25.$$

$$Var(X) = E(X^2) - [E(X)]^2 = 3.25 - 1.75^2 = 0.1875.$$

$$E(Y^2) = 0 \times 0.40 + 1 \times 0.40 + 4 \times 0.20 = 1.2.$$

$$Var(Y) = E(Y^2) - [E(Y)]^2 = 1.2 - 0.8^2 = 0.56.$$

$$Corr(X,Y) = \rho_{XY} = \frac{0.10}{\sqrt{0.1875} \cdot \sqrt{0.56}} \approx 0.3086.$$

**2.** Let the joint probability density function for (X, Y) be

$$f(x,y) = \begin{cases} 60 x^2 y & 0 \le x \le 1, \ 0 \le y \le 1, \ x+y \le 1 \\ 0 & \text{otherwise} \end{cases}$$

Find  $Cov(X,Y) = \sigma_{XY}$  and  $Corr(X,Y) = \rho_{XY}$ .

Recall: 
$$f_X(x) = 30x^2 (1-x)^2$$
,  $0 < x < 1$ ,  $E(X) = \frac{1}{2}$ ,  $f_Y(y) = 20y(1-y)^3$ ,  $0 < y < 1$ ,  $E(Y) = \frac{1}{3}$ ,  $E(XY) = \frac{1}{7}$ .

$$Cov(X,Y) = \frac{1}{7} - \frac{1}{2} \cdot \frac{1}{3} = -\frac{1}{42}.$$
  $Var(X) = \frac{9}{252}.$   $Var(Y) = \frac{8}{252}.$ 

$$\rho_{XY} = \frac{-1/42}{\sqrt{9/252} \cdot \sqrt{8/252}} = -\frac{1}{\sqrt{2}} = -\frac{\sqrt{2}}{2} \approx -0.7071.$$

3. Let the joint probability density function for (X, Y) be

$$f(x, y) = \begin{cases} x + y & 0 \le x \le 1, 0 \le y \le 1 \\ 0 & \text{otherwise} \end{cases}$$

Find  $Cov(X,Y) = \sigma_{XY}$  and  $Corr(X,Y) = \rho_{XY}$ .

Recall: 
$$f_X(x) = x + \frac{1}{2}, \quad 0 < x < 1.$$
  $f_Y(y) = y + \frac{1}{2}, \quad 0 < y < 1.$ 

$$\mu_X = \int_0^1 x \left( x + \frac{1}{2} \right) dx = \left[ \frac{1}{3} x^3 + \frac{1}{4} x^2 \right]_0^1 = \frac{7}{12};$$

$$\mu_Y = \int_0^1 y \left( y + \frac{1}{2} \right) dy = \frac{7}{12};$$

$$E(X^2) = \int_0^1 x^2 \left( x + \frac{1}{2} \right) dx = \left[ \frac{1}{4} x^4 + \frac{1}{6} x^3 \right]_0^1 = \frac{5}{12},$$

$$\sigma_X^2 = E(X^2) - \mu_X^2 = \frac{5}{12} - \left( \frac{7}{12} \right)^2 = \frac{11}{144}.$$

Similarly,  $\sigma_Y^2 = \frac{11}{144}$ .

$$E(XY) = \int_{00}^{11} x y(x+y) dx dy = \int_{00}^{11} \left(x^2 y + x y^2\right) dx dy = \int_{0}^{1} \left(\frac{y}{3} + \frac{y^2}{2}\right) dy = \frac{1}{3}.$$

$$Cov(X,Y) = \frac{1}{3} - \frac{7}{12} \cdot \frac{7}{12} = -\frac{1}{144}. \qquad \rho_{XY} = \frac{-1/144}{\sqrt{11/144} \cdot \sqrt{11/144}} = -\frac{1}{11}.$$

**4.** Let the joint probability density function for (X, Y) be

$$f(x,y) = \begin{cases} 12 x (1-x) e^{-2y} & 0 \le x \le 1, y \ge 0 \\ 0 & \text{otherwise} \end{cases}$$

Find  $Cov(X,Y) = \sigma_{XY}$  and  $Corr(X,Y) = \rho_{XY}$ .

Recall: 
$$f_X(x) = 6x(1-x), 0 < x < 1.$$
  $f_Y(y) = 2e^{-2y}, y > 0.$ 

Since X and Y are independent,

$$Cov(X,Y) = \sigma_{XY} = \mathbf{0}, \qquad Corr(X,Y) = \rho_{XY} = \mathbf{0}.$$