

Convergence in distribution:

(1)

Example 1: De Moivre-Laplace Thm. (X_i) i.i.d. with

$$P(X_i = 1) = P(X_i = -1) = 1/2 \Rightarrow \mathbb{E} X_i = 0.$$

$$\bar{X}_n := \frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{P} 0 \quad (LLN)$$

empirical mean

Law of large numbers.

Consider

$$Z_n := \sqrt{n} \bar{X}_n = \frac{1}{\sqrt{n}} \sum_{i=1}^n X_i.$$

$$\Rightarrow \mathbb{E} Z_n = 0 ; \text{Var } Z_n = \text{Var} \left(\frac{1}{\sqrt{n}} \sum_{i=1}^n X_i \right) = \frac{1}{n} \text{Var} \left(\sum_{i=1}^n X_i \right)$$

independence

$$= \frac{1}{n} \sum_{i=1}^n \underbrace{\text{Var}(X_i)}_{=1} = \frac{n}{n} = 1.$$

Thm. (De Moivre-Laplace)

$$P(a \leq Z_n \leq b) \xrightarrow{n \rightarrow \infty} \frac{1}{\sqrt{2\pi}} \int_a^b e^{-x^2/2} dx, \quad a < b. //$$

Rephrased

$$F_{Z_n}(x) \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy = F_Z(x), \quad Z \in \mathcal{N}(0,1)$$

$$F_{Z_n}(x) \rightarrow F_Z(x), \quad Z \in \mathcal{N}(0,1) \quad \text{"convergence in distribution"}$$

CDF of Z_n

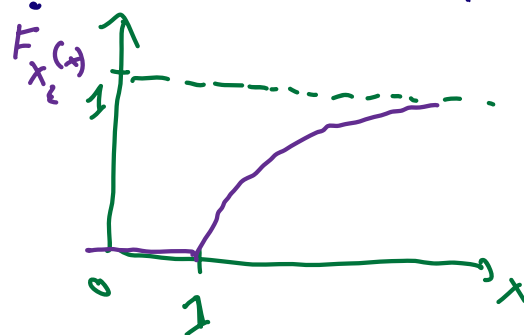
CDF of a
standard Normal r.v.

Example 2: X_1, X_2, \dots iid with CDF

(2)

$$F_{X_k}(x) = \begin{cases} 1 - \frac{1}{x^\alpha} & \text{if } x > 1 \\ 0 & \text{else} \end{cases} \quad \alpha > 0$$

Set $Y_n := n^{1/\alpha} \max_{1 \leq k \leq n} (X_k)$.



$$\Rightarrow F_{Y_n}(x) = P\left(\max_{1 \leq k \leq n} X_k \leq n^{1/\alpha} x\right)$$

$$= P(X_1 \leq n^{1/\alpha} x, X_2 \leq n^{1/\alpha} x, \dots, X_n \leq n^{1/\alpha} x)$$

independent

identically distributed

$$= \prod_{k=1}^n P(X_k \leq n^{1/\alpha} x) = \begin{cases} \left(1 - \frac{1}{n x^\alpha}\right)^n, & \text{if } x > n^{-1/\alpha} \\ 0 & \text{else} \end{cases}$$

$$= \begin{cases} \left(1 - \frac{x^{-\alpha}}{n}\right)^n & \text{if } x > n^{-1/\alpha} \\ 0 & \text{else} \end{cases}$$

$\rightarrow 0$
as $n \rightarrow \infty$

$$\xrightarrow{n \rightarrow \infty} \begin{cases} e^{-x^{-\alpha}} & \text{if } x > 0 \\ 0 & \text{else} \end{cases}$$

Check that this is a CDF: $F_{Y_n}(x) \rightarrow F_Y(x)$ of a random variable.

//

Def: X_n converges in distribution to a random variable X , as $n \rightarrow \infty$, if

$$F_{X_n}(x) \rightarrow F_X(x) \text{ as } n \rightarrow \infty \text{ for all } x \in C(F_X)$$

where $C(F_X) = \{x \in \mathbb{R} : F_X \text{ is continuous at } x\}$.

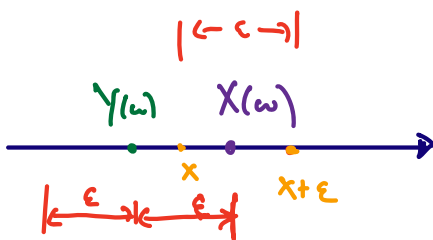
Notation: $X_n \xrightarrow{d} X, n \rightarrow \infty$. Convergence in distribution.

Lemma: $X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{d} X$

Auxiliary claim: X, Y random variables, $x \in \mathbb{R}, \epsilon > 0$. Then

$$P(Y \leq x) \leq P(X \leq x + \epsilon) + P(|Y - X| > \epsilon).$$

Proof of claim: $P(Y \leq x) = \underbrace{P(Y \leq x, |X - Y| \leq \epsilon)}_{\leq P(X \leq x + \epsilon)} + \underbrace{P(Y \leq x, |X - Y| > \epsilon)}_{\leq P(|X - Y| > \epsilon)}$



$$\leq P(X \leq x + \epsilon)$$

$$\leq P(|X - Y| > \epsilon)$$

If $\bar{Y} \leq x$ and $|\bar{X} - \bar{Y}| \leq \epsilon$,
then $\bar{X} \leq x + \epsilon$.

Thus $\{\bar{Y} \leq x\} \cap \{|\bar{X} - \bar{Y}| \leq \epsilon\} \subset \{\bar{X} \leq x + \epsilon\}$.

$$\leq P(X \leq x + \epsilon) + P(|X - Y| > \epsilon).$$

Upside: $F_Y(x) \leq F_{\bar{X}}(x + \epsilon) + P(|\bar{X} - \bar{Y}| > \epsilon)$

Back to the proof: Let $x \in \mathbb{R}$ be a continuity point of F_X . (4.)
 Let $\varepsilon > 0$.

Claim with $Y = X_n$

$$(*) F_{X_n}(x) = \mathbb{P}(X_n \leq x) \leq \mathbb{P}(X \leq x + \varepsilon) + \mathbb{P}(|X_n - X| > \varepsilon) \\ = F_X(x + \varepsilon) + \mathbb{P}(|X_n - X| > \varepsilon).$$

Repeat the argument by switching X and X_n , and changing x to $x - \varepsilon$
 and $x + \varepsilon$ to x .

$$(**) F_X(x - \varepsilon) \leq F_{X_n}(x) + \mathbb{P}(|X_n - X| > \varepsilon).$$

(*) & (**)

$$\Rightarrow F_X(x - \varepsilon) \stackrel{(**)}{\leq} F_{X_n}(x) + \underbrace{\mathbb{P}(|X_n - X| > \varepsilon)}_{\substack{\rightarrow 0, n \rightarrow \infty \\ \text{as } X_n \xrightarrow{P} X}} \stackrel{(*)}{\leq} F_X(x + \varepsilon) + \underbrace{2\mathbb{P}(|X_n - X| > \varepsilon)}_{\substack{\rightarrow 0 \\ n \rightarrow \infty}}$$

We do not know at this point whether $F_{X_n}(x)$ has a limit as $n \rightarrow \infty$. But we must have from the above that

$$(\dagger) F_X(x - \varepsilon) \leq \liminf_{n \rightarrow \infty} F_{X_n}(x) \leq \limsup_{n \rightarrow \infty} F_{X_n}(x) \leq F_X(x + \varepsilon).$$

By assumption x is a continuity point of F_X : $\lim_{\varepsilon \rightarrow 0} F_X(x - \varepsilon) = \lim_{\varepsilon \rightarrow 0} F_X(x + \varepsilon) = F_X(x)$.
 Take now $\varepsilon \rightarrow 0$ in (\dagger) :

$$F_X(x) = \lim_{n \rightarrow \infty} F_{X_n}(x).$$



Used that x is a continuity point of F_X .



Summary:

⑤

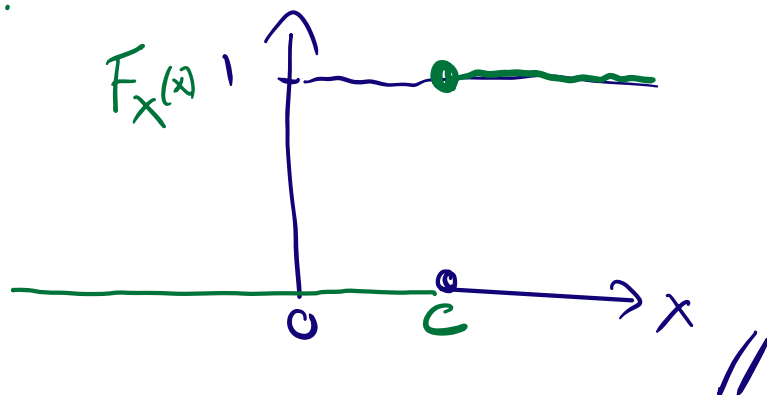
$$X_n \xrightarrow{\text{a.s.}} X \implies X_n \xrightarrow{P} X \implies X_n \xrightarrow{d} X \text{ as } n \rightarrow \infty$$
$$\uparrow$$
$$X_n \xrightarrow{2} X$$

Thm: X_1, X_2, \dots a sequence of random variables. Let $c \in \mathbb{R}$. Then

$$X_n \xrightarrow{P} c \iff X_n \xrightarrow{d} \delta(c)$$

\uparrow
number
not a random variable.

Shorthand for $X \in \delta(c)$.

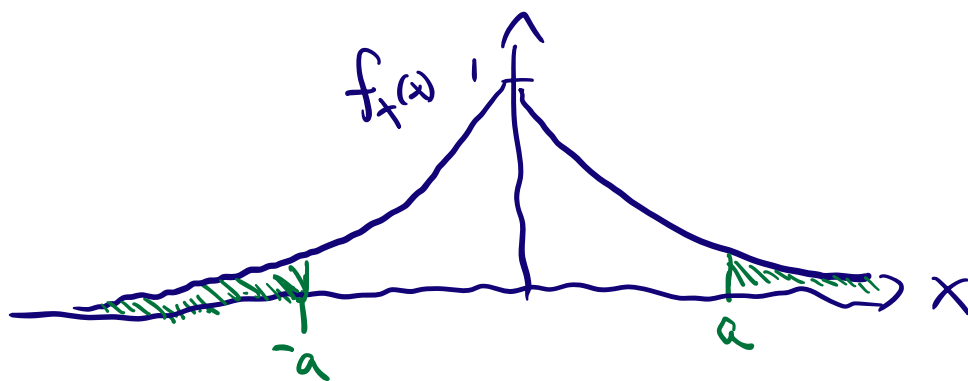


Proof: Exercise, page 157 in [AG].

✓ Laplace distribution with parameter $\lambda=1$.

Example: $X \in \mathcal{L}(1)$, $f_X(x) = \frac{1}{2} e^{-|x|}$, $x \in \mathbb{R}$

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X_1, X_2, \dots a sequence of random variables defined by

$$X_{2n} \stackrel{\text{even}}{=} X, \quad X_{2n-1} \stackrel{\text{odd}}{=} -X, \quad n=1, 2, 3, \dots$$

Note $X_{2n} \stackrel{d}{=} X_{2n-1}$, because f_X is symmetric about zero.

In particular, $X_n \xrightarrow{d} X$, as $n \rightarrow \infty$. ($F_{X_n}(x) = F_X(x)$)

But, $P(|X| > a) > 0$, for any $a > 0$.

Hence, for any $\varepsilon > 0$:

$$P(|X_n - X| > \varepsilon) = \begin{cases} 0 & \text{for } n \text{ even, because } X_n = X. \\ P(|X| > \frac{\varepsilon}{2}) > 0 & \text{for } n \text{ odd.} \end{cases}$$

$$\Rightarrow X_n \not\xrightarrow{P} X$$

X_n does not converge in probability.

and hence $P(|X_n - X| > \varepsilon) = P(|X| > \frac{\varepsilon}{2}) > 0$ for n odd. //