

# <proof of Sample Correlation limiting distribution>

$$(X_i, Y_i) \stackrel{i.i.d}{\sim} (\mu, \sigma^2)$$

for the sake of simplicity we assume that

$$E(X_i) = 0, E(Y_i) = 0.$$

we let

$$\begin{pmatrix} m_x \\ m_y \\ m_{xx} \\ m_{yy} \\ m_{xy} \end{pmatrix} = \frac{1}{n} \begin{pmatrix} \sum X_i \\ \sum Y_i \\ \sum X_i^2 \\ \sum Y_i^2 \\ \sum X_i Y_i \end{pmatrix}$$

$$s_x^2 = m_{xx} - m_x^2, \quad s_y^2 = m_{yy} - m_y^2, \quad s_{xy} = m_{xy} - m_x m_y$$

$$\text{then } r = s_{xy} / s_x s_y$$

By CLT.

$$\sqrt{n} \left\{ \begin{pmatrix} m_x \\ m_y \\ m_{xx} \\ m_{yy} \\ m_{xy} \end{pmatrix} - \begin{pmatrix} 0 \\ 0 \\ \sigma_x^2 \\ \sigma_y^2 \\ \sigma_{xy} \end{pmatrix} \right\} \xrightarrow{d} N_5 \left( \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \text{Cov}(X_1, X_1) & \dots & \text{Cov}(X_1, X_5) \\ \text{Cov}(Y_1, X_1) & & \vdots \\ \vdots & & \ddots \\ \text{Cov}(X_5, X_1) & & \text{Cov}(X_5, X_5) \end{pmatrix} \right)$$

We want to asymptotic distribution of  $r$ .

So define function  $g: \mathbb{R}^5 \rightarrow \mathbb{R}^3$  which yields  $(s_x^2, s_y^2, s_{xy})$ .

Then

$$\sqrt{n} \left( g \begin{pmatrix} m_x \\ m_y \\ \vdots \\ m_{xy} \end{pmatrix} - g \begin{pmatrix} 0 \\ 0 \\ \sigma_x^2 \\ \sigma_y^2 \\ \sigma_{xy} \end{pmatrix} \right) \xrightarrow{d} N_3 \left( \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \nabla g^T \Sigma \nabla g \right)$$

which satisfy

$$\nabla g \begin{pmatrix} a \\ b \\ c \\ d \\ e \end{pmatrix} = \begin{pmatrix} -2a & 0 & -b \\ 0 & -2b & -a \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{pmatrix}.$$

$$\begin{aligned} g(a, b, c, d, e) \\ = (c - a^2, d - b^2, e - ab) \end{aligned}$$

So

$$\sqrt{n} \left( g \begin{pmatrix} \bar{m}_x \\ \bar{m}_y \\ \bar{m}_{xx} \\ \bar{m}_{yy} \\ \bar{m}_{xy} \end{pmatrix} - g \begin{pmatrix} 0 \\ 0 \\ \sigma_x^2 \\ \sigma_y^2 \\ \sigma_{xy} \end{pmatrix} \right) \rightarrow N_3 \left( \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \Sigma^* \right)$$

where

$$\Sigma^* = \begin{pmatrix} \text{Cov}(X_1^2, X_1^2) & \text{Cov}(X_1^2, Y_1^2) & \text{Cov}(X_1^2, X_1 Y_1) \\ \text{Cov}(Y_1^2, X_1^2) & \text{Cov}(Y_1^2, Y_1^2) & \text{Cov}(Y_1^2, X_1 Y_1) \\ \text{Cov}(X_1 Y_1, X_1^2) & \text{Cov}(X_1 Y_1, Y_1^2) & \text{Cov}(X_1 Y_1, X_1 Y_1) \end{pmatrix}$$

precisely

$$\sqrt{n} \left( \begin{pmatrix} \bar{s}_x^2 \\ \bar{s}_y^2 \\ \bar{s}_{xy} \end{pmatrix} - \begin{pmatrix} \sigma_x^2 \\ \sigma_y^2 \\ \sigma_{xy} \end{pmatrix} \right) \rightarrow N_3(0, \Sigma^*)$$

As an aside,  $\sqrt{n}(\bar{s}_x^2 - \sigma_x^2) \xrightarrow{d} N(0, \frac{\text{Var}(X_1^2)}{\parallel})$   
 $E((X_1 - \mu)^4) - \sigma^4$

Next, we use Delta method one more time.

which  $h: \mathbb{R}^3 \rightarrow \mathbb{R}$ ,  $h(a, b, c) = \frac{c}{\sqrt{a}\sqrt{b}}$

$$\nabla h(\cdot) = \begin{pmatrix} \frac{c}{\sqrt{b}} \cdot -\frac{1}{2} a^{-\frac{3}{2}} \\ \frac{c}{\sqrt{a}} \cdot -\frac{1}{2} b^{-\frac{3}{2}} \\ \frac{1}{\sqrt{ab}} \end{pmatrix} \Rightarrow \nabla h \begin{pmatrix} \bar{s}_x^2 \\ \bar{s}_y^2 \\ \bar{s}_{xy} \end{pmatrix} = \begin{pmatrix} \frac{\bar{s}_{xy}}{\bar{s}_y} \cdot -\frac{1}{2} \cdot \frac{1}{\bar{s}_x^3} \\ \frac{\bar{s}_{xy}}{\bar{s}_x} \cdot -\frac{1}{2} \cdot \frac{1}{\bar{s}_y^3} \\ \frac{1}{\bar{s}_x \bar{s}_y} \end{pmatrix}$$

So that

$$\sqrt{n}(\bar{r} - \rho) \xrightarrow{d} N(0, \nabla h^T \Sigma^* \nabla h)$$

When  $E(X) = E(Y) = 0$ ,  $\text{Var}(X) = \text{Var}(Y) = 1$ ,  $\sqrt{n}(\bar{r} - \rho) \xrightarrow{d} N(0, (1 - \rho^2)^2)$