

Applied Stats II - Problem Set 1

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Code in PS3_ImeldaFinn.R

Question 1

We are interested in how governments' management of public resources impacts economic prosperity. Our data come from Alvarez, Cheibub, Limongi, and Przeworski (1996) and is labelled `gdpChange.csv` on GitHub. The dataset covers 135 countries observed between 1950 or the year of independence or the first year for which data on economic growth are available ("entry year"), and 1990 or the last year for which data on economic growth are available ("exit year"). The unit of analysis is a particular country during a particular year, for a total $> 3,500$ observations.

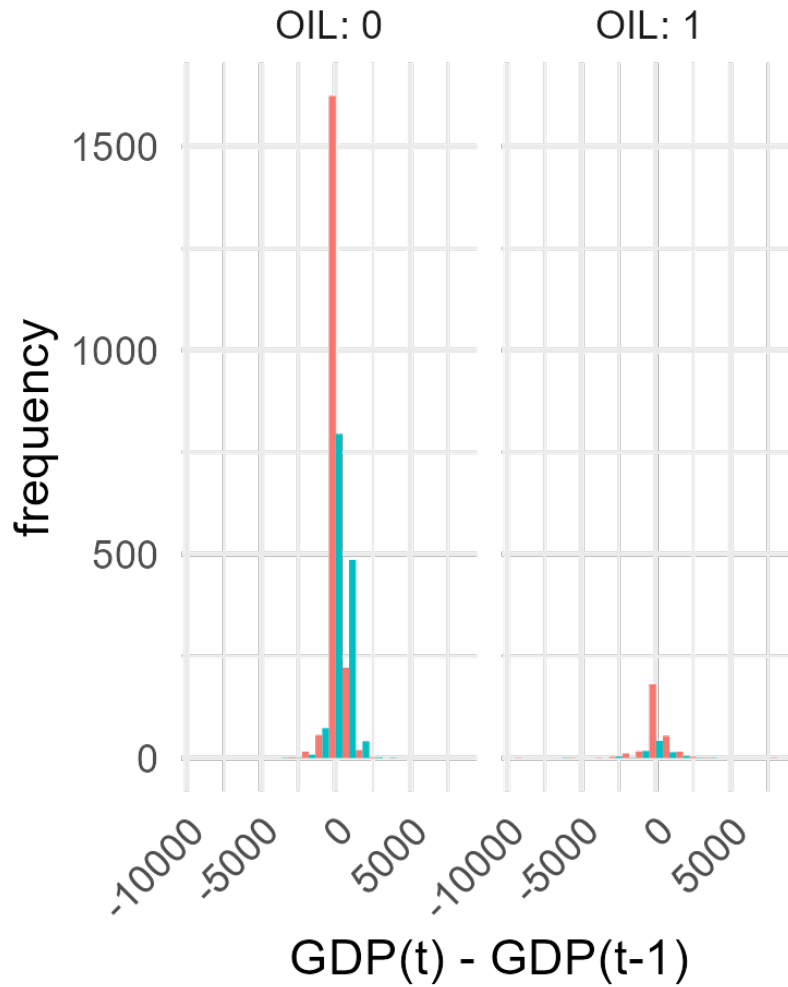
- Response variable:
 - `GDPWdiff`: Difference in GDP between year t and $t-1$. Possible categories include: "positive", "negative", or "no change"
- Explanatory variables:
 - `REG`: 1=Democracy; 0=Non-Democracy
 - `OIL`: 1=if the average ratio of fuel exports to total exports in 1984-86 exceeded 50%; 0= otherwise

The data was read in and `GDPWDiff` was factored. The cutoff point was 0, ie values less than 0 were categorised as *negative*, values equal to 0 were categorised as *no change* and values above 0 were categorised as *positive*. *no change* was set as the reference category.

```
1 gdp <- read_csv("../data/gdpChange.csv")
2 gdp <- rename(gdp, indx = '...1')
3 gdp$diff <- ifelse(gdp$GDPWdiff==0, "no change",
4                   ifelse(gdp$GDPWdiff>0, "positive", "negative"))
5 gdp$diff2 <- relevel(factor(gdp$diff, ordered = FALSE), ref="no change")
```

A

Change in GDP year-on-year



REG:1=democracy ■ 0 ■

OIL=1 : average ratio of fuel exports to total exports in 1984-86 exceeded 50%; REG=1 : democracy

variable	OIL=0		OIL=1	
	REG=0	REG=1	REG=0	REG=1
GDPWdiff				
Min / Max	-2506.0 / 2821.0	-3741.0 / 3722.0	-9257.0 / 7867.0	-5997.0 / 3555.0
Med [IQR]	50.0 [-30.0;215.0]	293.0 [13.5;644.8]	140.5 [-50.0;463.0]	39.0 [-527.2;421.8]
Mean (std)	106.0 (395.3)	319.4 (561.6)	141.2 (1142.3)	-46.5 (1228.4)
N (NA)	1939 (0)	1408 (0)	288 (0)	86 (0)

1. An unordered multinomial logit with `GDPwdiff` as the output was constructed as follows:

```
1 multinom_model <- multinom(diff2 ~ REG + OIL, data = gdp )
```

The results of the model are shown in Table 1. The predicted probabilities are shown in Table 2, Figure 2. All of the predicted classes are *positive*. This isn't unexpected as Figure 1 shows that the data is skewed towards positive values.

Table 1: Multinomial, unordered

	<i>Dependent variable:</i>	
	negative	positive
	(1)	(2)
REG	1.379 t = 1.794 p = 0.073*	1.769 t = 2.306 p = 0.022**
OIL	4.784 t = 0.695 p = 0.488	4.576 t = 0.665 p = 0.507
Constant	3.805 t = 14.058 p = 0.000***	4.534 t = 16.842 p = 0.000***
Akaike Inf. Crit.	4,690.770	4,690.770
<i>Note:</i>	*p<0.1; **p<0.05; ***p<0.01	

The baseline category is regime `REG = 0` (non-democracy) and `OIL = 0` (not a significant oil exporter). The predicted probability of having no change in GDP when in the baseline category is 0.7%.

```
1 round(predict(multinom_model, newdata = data.frame(REG=0, OIL=0),
2         type = "probs"), 2)
3
4 predict_data <- data.frame(REG = rep(c(0,1), each = 2),
5                               OIL= rep(c(0,1), 2))
6 cbind(predict_data, predict(multinom_model,
7                             newdata = predict_data, type = "class"))
8
```

```
1 no change  negative  positive
2      0.01      0.32      0.67
```

```

3 REG OIL predict(multinom_model, newdata = predict_data, type = "class")
4
5 1 0 0 positive
6 2 0 1 positive
7 3 1 0 positive
8 4 1 1 positive
9
10      no change negative positive Sum
11 no change      0      0      16   16
12 negative      0      0     1105 1105
13 positive      0      0     2600 2600
14 Sum          0      0     3721 3721
15

```

Table 2: Predicted results from unordered Multinomial

	REG	OIL	level	probability
1	0	0	no change	0.007
2	0	1	no change	0.0001
3	1	0	no change	0.001
4	1	1	no change	0.00001
5	0	0	negative	0.323
6	0	1	negative	0.373
7	1	0	negative	0.246
8	1	1	negative	0.287
9	0	0	positive	0.670
10	0	1	positive	0.627
11	1	0	positive	0.753
12	1	1	positive	0.713

Holding OIL constant:

- a change in REG from 0 to 1 increases the log-odds of `diff= positive` vs. `diff= no change` by 1.769
- a change in REG from 0 to 1 multiplies the odds of `diff= positive` vs. `diff= no change` by a factor of $e^{1.769} = 5.87$.

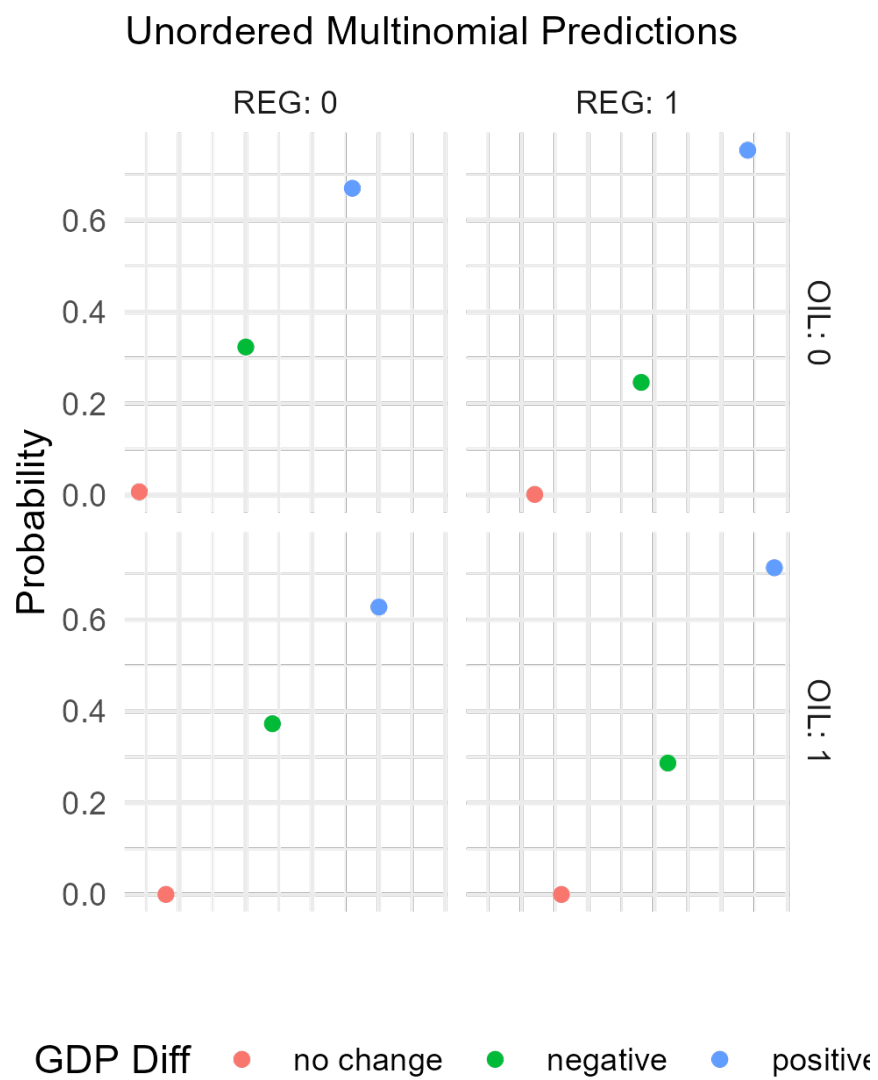


Figure 2: GDP Diff - unordered multinomial predictions

2. The factored `GDPdiff` response variable (from 1) was ordered: (*negative* < *no change* < *positive*), i.e. the cutoff (0) was unchanged. An ordered multinomial logit model was created as follows:

```
1 # get an ordered factor for GDP difference
2 gdp$ordered_diff <- ordered(gdp$diff,
3                             labels=c("negative", "no change", "positive"))
```

The results of the model are shown in Table 3. The baseline category is regime `REG` = 0 (non-democracy) and `OIL` = 0 (not a significant oil exporter). The predicted probability of having no change in GDP when in the baseline category is 0.5% (Table 4).

Table 3: Multinomial Logit, ordered

	<i>Dependent variable:</i>
	ordered_diff
REG	0.398 (0.075) t = 5.300 p = 0.00000***
OIL	-0.199 (0.116) t = -1.717 p = 0.086*
negative no change	-0.731 (0.048) t = -15.360 p = 0.000***
no change positive	-0.710 (0.048) t = -14.955 p = 0.000***
Observations	3,721
<i>Note:</i>	*p<0.1; **p<0.05; ***p<0.01

The coefficients and their confidence intervals are:

```
1 cbind(logOdds = coef(ord_model), confint(ord_model))
2 # logOdds 2.5 % 97.5 %
```

```

3 #REG 0.3984834 0.2516548 0.54643410
4 #OIL -0.1987177 -0.4237548 0.03019571
5

```

Holding REG constant:

- a change in OIL from 0 to 1 changes the log-odds of `diff= no change` vs. `diff = negative` by -0.199
- a change in OIL from 0 to 1 multiplies the odds of `diff= no change` vs. `diff = negative` by a factor of $e^{-0.199} = 0.82$.

The odds of having no change in GDP growth for a country that has oil, are .18% lower compared to a country that doesn't have oil, holding regime status constant.

```

1 round(predict(ord_model, newdata = data.frame(REG=0, OIL=0), type = "
  probs"), 2)
2 predict(ord_model, newdata = data.frame(REG=0, OIL=0), type = "class")
3 cbind(predict_data, predict(ord_model, predict_data, type="class"))
4

```

```

1 negative no change positive
2      0.32      0.00      0.67
3
4 [1] positive
5 Levels: negative no change positive
6
7 REG OIL predict(ord_model, predict_data, type = "class")
8 1   0   0 positive
9 2   0   1 positive
10 3   1   0 positive
11 4   1   1 positive
12
13      negative no change positive Sum
14 negative      0      0     1105 1105
15 no change      0      0      16   16
16 positive      0      0     2600 2600
17 Sum           0      0     3721 3721

```

The predicted probabilities are shown in Table 4, Figure 3.

```

1 pred_ord <- melt(cbind(predict_data,
2                          predict(ord_model, predict_data, type="probs")),
3                  id.vars=c("REG", "OIL"),
4                  variable.name="level", value.name="probability")
5

```

The model predictions for the individual GDP difference categories are give in Table 5
The proportional-odds assumption does not appear to hold for this regression i.e. the coefficients are not consistent (coef for OIL goes from -0.04 to 0.05 to -0.01).

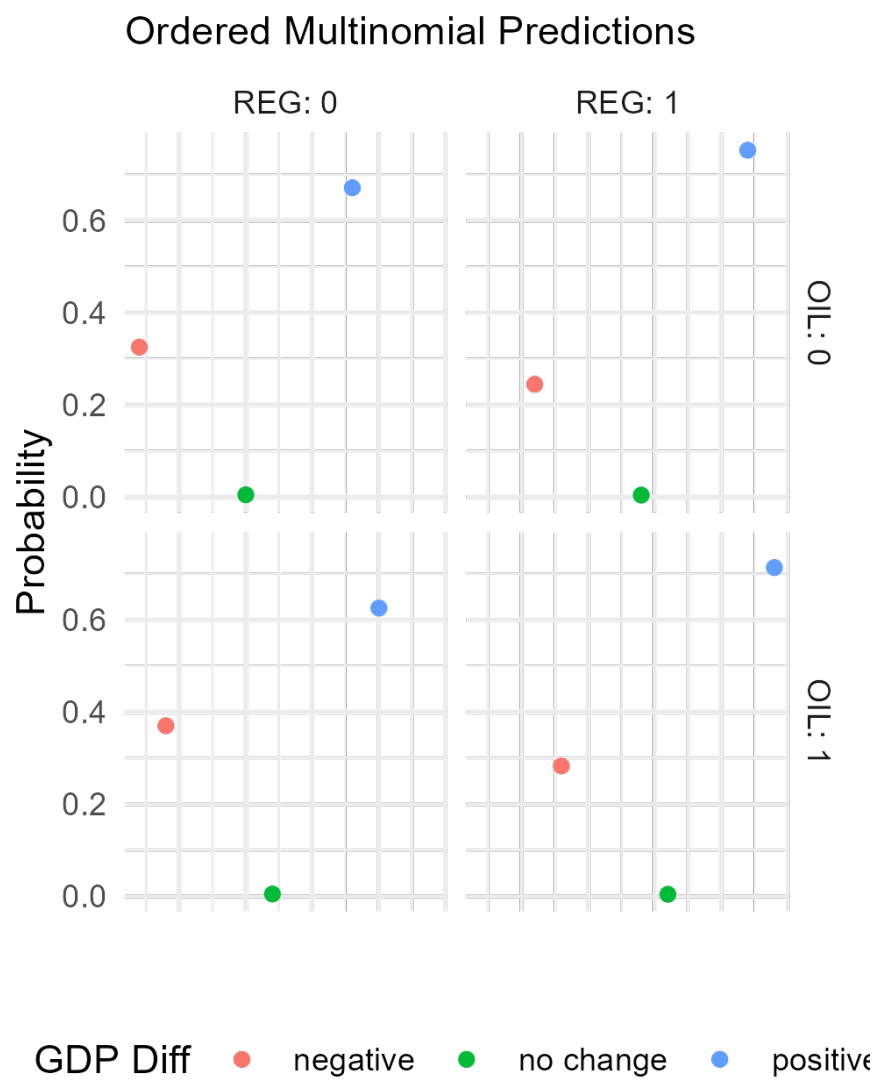


Figure 3: GDP Diff - ordered multinomial predictions

Table 4: Predicted results from Ordered Model

	REG	OIL	level	probability
1	0	0	negative	0.325
2	0	1	negative	0.370
3	1	0	negative	0.244
4	1	1	negative	0.283
5	0	0	no change	0.005
6	0	1	no change	0.005
7	1	0	no change	0.004
8	1	1	no change	0.004
9	0	0	positive	0.671
10	0	1	positive	0.625
11	1	0	positive	0.752
12	1	1	positive	0.713

```

1 id.vars=c("REG", "OIL")
2 for ( i in 1:length(unique(gdp$ordered_diff))) {
3   assign(paste("logit_model", i, sep=""),
4     glm(ifelse(ordered_diff==unique(gdp$ordered_diff)[i],
5       1, 0) ~ REG + OIL, data = gdp),
6     envir = globalenv())
7

```

The cutoff point affects the results. For example, changing the cutoff to split the data into 3 equal-length sections changes the coefficients and the confusion matrices. The resulting models aren't more accurate, because they reduce the number of true positive predictions for *positive* without increasing the true predictions for the other 2 categories enough to compensate. They also have higher deviance. They could be refined to better categorise/predict the data.

```
1 cutoffs: <=14, 14-283, >=283
```

```
2
3 Multinomial model
```

```
4 Dependent variable:
```

```
5
6 negative positive
7 (1) (2)
```

```
8
9 REG 0.184** 1.365***
10 (0.088) (0.087)
11
12 OIL 0.428*** 0.641***
13 (0.139) (0.145)
14
```

Table 5: Comparison of models for GDP Diff categories

	<i>Dependent variable:</i>		
	ordered.diff)[i, 1, 0)		
	negative	no change	positive
REG	0.0832*** (0.0154)	-0.0779*** (0.0153)	-0.0054** (0.0022)
OIL	-0.0416* (0.0251)	0.0474* (0.0250)	-0.0058 (0.0036)
Constant	0.6695*** (0.0102)	0.3235*** (0.0102)	0.0070*** (0.0015)
Observations	3,721	3,721	3,721
Log Likelihood	-2,364.5140	-2,350.4150	4,869.1730
Akaike Inf. Crit.	4,735.0280	4,706.8310	-9,732.3450

Note:

*p<0.1; **p<0.05; ***p<0.01

15 Constant -0.091* -0.649***
 16 (0.051) (0.059)

17
 18 deviance: 7850.215

19
 20 confusion matrix:

21 pred_m_all no change negative positive
 22 no change 799 747 393
 23 negative 82 99 107
 24 positive 353 395 746

25
 26
 27 Ordered Model

28 Dependent variable:

29
 30 odiff

31
 32 REG 0.930***
 33 (0.064)

34
 35 OIL 0.120
 36 (0.105)

37
 38 deviance: 7691.950

39

```
40 confusion matrix:
41 pred_o_all  negative no change positive
42  negative      846      881      500
43  no change       0       0       0
44  positive      395      353      746
```

Question 2

Consider the data set `MexicoMuniData.csv`, which includes municipal-level information from Mexico. The outcome of interest is the number of times the winning PAN presidential candidate in 2006 (`PAN.visits.06`) visited a district leading up to the 2009 federal elections, which is a count. Our main predictor of interest is whether the district was highly contested, or whether it was not (the PAN or their opponents have electoral security) in the previous federal elections during 2000 (`competitive.district`), which is binary (1=close/swing district, 0=“safe seat”). We also include `marginality.06` (a measure of poverty) and `PAN.governor.06` (a dummy for whether the state has a PAN-affiliated governor) as additional control variables.

A

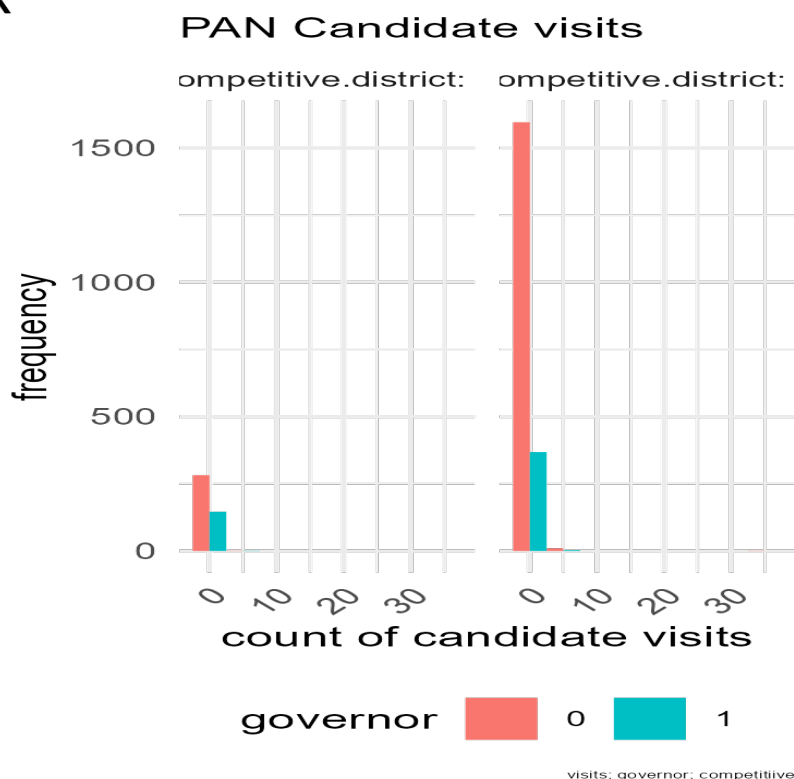


Figure 4: Presidential candidate visits

```
1 mexico <- read_csv("../data/MexicoMuniData.csv")
```

- (a) A Poisson regression model was run because the outcome is a count variable, to consider whether there is evidence that PAN presidential candidates visit swing districts more? The model output is in Table 6. `competitive.district` coefficient is -0.081, but it is not a significant predictor for number of visits.

```

1 mexico_poisson <- glm(PAN.visits.06 ~ competitive.district +
2   marginality.06 + PAN.governor.06, data= mexico, family =
   poisson)

```

A poisson model was run to get the regression coefficients.

Table 6: Poisson Model of candidate visit counts

	<i>Dependent variable:</i>
	PAN.visits.06
	<i>Poisson</i>
competitive.district	−0.081 (0.171)
marginality.06	−2.080*** (0.117)
PAN.governor.06	−0.312* (0.167)
Constant	−3.810*** (0.222)
Observations	2,407
Log Likelihood	−645.606
Akaike Inf. Crit.	1,299.213
<i>Note:</i> *p<0.1; **p<0.05; ***p<0.01	

Hypothesis Test

The summary results for the poisson model for the coefficients are:

1	Coefficients :				
2		Estimate	Std. Error	z value	Pr(> z)
3	(Intercept)	−3.81023	0.22209	−17.156	<2e−16 ***
4	competitive.district	−0.08135	0.17069	−0.477	0.6336
5	marginality.06	−2.08014	0.11734	−17.728	<2e−16 ***
6	PAN.governor.06	−0.31158	0.16673	−1.869	0.0617 .

1. H_0 PAN presidential candidate visits swing districts less than other districts ($E(\lambda|competitive.district = 1) < E(\lambda|competitive.district = 0)$)
2. H_a candidates visit swing districts at least as many times other districts

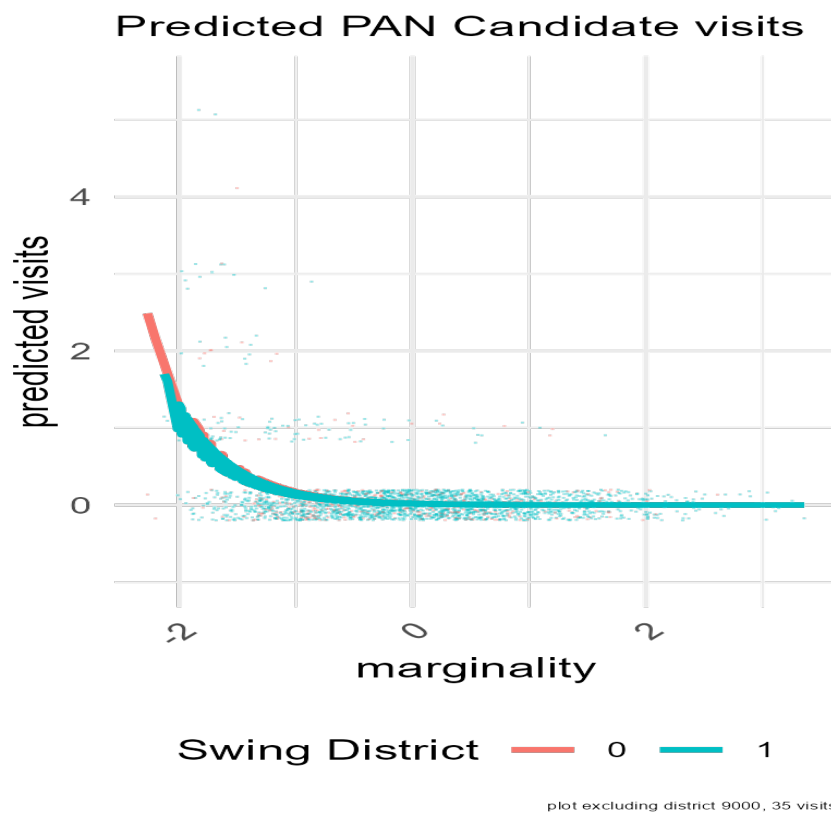


Figure 5: predicted presidential candidate visits

3. the test statistic $= \beta_{competitive} / se_{competitive} = -0.08135 / 0.17069 = -0.477 (\sim N(0, 1))$
4. the α value is 0.05, one-sided, left-tailed z-test
5. the pvalue $p = 0.6833189$ ¹
6. as pvalue is greater than α , we cannot reject the null hypothesis that closely contested districts receive fewer visits.

Using R's `poisson.test`, the p-value is 0.8544, so we also reject the null hypothesis, i.e. there is not evidence to support the theory that presidential candidates visit swing districts more.

```

1  cd1 <-mexico$PAN.visits.06[mexico$competitive.district==1]
2  cd0 <-mexico$PAN.visits.06[mexico$competitive.district==0]
3
4  poisson.test(x=c(sum(cd1), sum(cd0)), T=c(length(cd1), length(cd0)),
5              alternative="greater", conf.level=0.95)
6
7  Comparison of Poisson rates
8
9  data:  c(sum(cd1), sum(cd0)) time base: c(length(cd1), length(cd0))
10 count1 = 176, expected count1 = 181.52, p-value = 0.8544
11 alternative hypothesis: true rate ratio is greater than 1
12 95 percent confidence interval:
13   0.6411102      Inf
14 sample estimates:
15   rate ratio
16   0.8506716
17
```

(b) `marginality.06` and `PAN.governor.06` coefficients.

```

1  > lo_cis<-cbind(logOdds = coef(mexico_poisson), confint(mexico_
2  poisson))
3  Waiting for profiling to be done...
4
5  # Coefficients and confidence intervals
6
7  logOdds      2.5 %      97.5 %
8  (Intercept) -3.81023498 -4.2606981 -3.389583340
9  competitive.district -0.08135181 -0.4063661 0.264275617
10 marginality.06 -2.08014361 -2.3151624 -1.855053854
11 PAN.governor.06 -0.31157887 -0.6484827 0.006518468
12
13 > exp(lo_cis)
14
15 exp(beta) 2.5 % 97.5 %
16 (Intercept) 0.022 0.014 0.034
17 competitive.district 0.922 0.666 1.302
18 marginality.06 0.125 0.099 0.156
19 PAN.governor.06 0.732 0.523 1.007
20
```

¹`pnorm(0.477)`

`marginality.06` is the only coefficient which is significant at $\alpha = 0.01$; `PAN.governor.06` is significant at $\alpha = 0.1$.

The coefficient for `marginality.06` is -2.08 ($CI_{0.05} = -2.315, -1.855$). This means that, keeping all else constant, we expect a decrease of 2.08 in log count for a one-unit increase in `marginality.06`, i.e. if `marginality.06` increases by 1, we expect the estimated mean number of visits to decrease by 87.5% (multiply previous expected count by $e^{-2.08} = 0.125$). Districts with higher marginality receive fewer visits.

The coefficient for `PAN.governor.06` is -0.312, which means that if `PAN.governor.06` switches from 0 to 1, keeping all other variables constant, we expect an decrease in log count of 0.312. If `PAN.governor.06` changes from 0 to 1, we expect the estimated mean number of visits to decrease by 26.8% (ie multiply previous expected count by 0.732). Districts with a PAN governor are expected to receive fewer visits from PAN candidates. Note: $CI_{0.05} = -0.648, 0.007$, which includes 0. The test results suggest that having a PAN governor is not a significant predictor for the number of candidate visits.

- (c) The estimated mean number of visits from the winning PAN presidential candidate for a hypothetical district that was competitive (`competitive.district=1`), had an average poverty level (`marginality.06 = 0`), and a PAN governor (`PAN.governor.06=1`).

```

1   mex_pred_data <- data.frame(competitive.district = 1,
2                                   marginality.06=0,
3                                   PAN.governor.06=1)
4   pred_mex <- cbind(predict(mexico_poisson,
5                                   mex_pred_data,
6                                   type= "response", se.fit =TRUE),
7                                   mex_pred_data)
8   # create lower and upper bounds for CIs
9   pred_mex$lowerBound <- pred_mex$fit - 1.96 * pred_mex$se.fit
10  pred_mex$upperBound <- pred_mex$fit + 1.96 * pred_mex$se.fit
11
12  round(pred_mex,3)
13
14
15      fit se.fit residual.scale competitive.district marginality.06
16 1 0.015 0.003                1                1                0
17 PAN.governor.06 lowerBound upperBound
18 1                1        0.009      0.021
19
20
21
```

$$\lambda = e^{\beta_0 + \beta_{competitive} \times competitive + \beta_{marginality} \times marginality + \beta_{governor} \times governor}$$

$$= e^{-3.810 - 0.081 \times 1 + -2.080 \times 0 + -0.312 \times 1} = e^{-4.203} = 0.015 \text{ (The mean visits is 0.092, the median is 0.)}$$

The estimated mean number of visits, in the time frame, by the winning PAN presidential candidate to a district which was a swing state, with average poverty (=0) and a PAN governor is 0.015.

Validation There are 2,272 zero count values in our dataset.

```

1 dispersiontest(mexico_poisson)
2
3   Overdispersion test
4
5 data:  mexico_poisson
6 z = 1.0668, p-value = 0.143
7 alternative hypothesis: true dispersion is greater than 1
8 sample estimates:
9 dispersion
10      2.09834

```

A zero-inflated poisson model was run for comparison, the only coefficient with a significant deviance is the `marginality.06`, where an increase of 1 unit in marginality corresponds to a increase in log-odds of a 0 count value of 0.872(Table 8). The zero-inflated poisson changes the value of the coefficients, but doesn't change their statistical significance.

Table 7: Zero-inflation model

	<i>Dependent variable:</i>
	PAN.visits.06 <i>zero-inflated count data</i>
competitive.district	0.900* (0.511)
marginality.06	0.872*** (0.302)
PAN.governor.06	-0.175 (0.412)
Constant	1.272* (0.675)
Observations	2,407
Log Likelihood	-600.386
<i>Note:</i>	*p<0.1; **p<0.05; ***p<0.01

```

1 anova(mexico_poisson, zeroinfl_poisson, test = "Chi")
2 Analysis of Deviance Table
3

```

Table 8: Zero Infl Poisson vs Poisson Model

	<i>Dependent variable:</i>	
	PAN.visits.06	
	<i>zero-inflated count data</i>	<i>Poisson</i>
competitive.district	0.402 (0.312)	−0.081 (0.171)
marginality.06	−1.240*** (0.261)	−2.080*** (0.117)
PAN.governor.06	−0.470* (0.271)	−0.312* (0.167)
Constant	−1.914*** (0.498)	−3.810*** (0.222)
Observations	2,407	2,407
Log Likelihood	−600.386	−645.606
Akaike Inf. Crit.		1,299.213
<i>Note:</i>	*p<0.1; **p<0.05; ***p<0.01	

4	Model: poisson , link: log					
5						
6	Response: PAN.visits.06					
7						
8	Terms added sequentially (first to last)					
9						
10		Df	Deviance	Resid. Df	Resid. Dev	Pr(>Chi)
11	NULL			2406	1473.87	
12	competitive.district	1	0.91	2405	1472.96	0.34078
13	marginality.06	1	478.03	2404	994.93	< 2e-16 ***
14	PAN.governor.06	1	3.68	2403	991.25	0.05502
15						

There is one outlier with visits = 35; predicted visits based on the model are 0.467. Excluding that district (ie assuming it's bad data) would lead to changes in the model, but would only change the prediction in c) from 0.015 to 0.016.

1	PAN.visits.06		
2	outlier removed	default	
3			
4	competitive.district	−0.261	−0.081
5		(0.176)	(0.171)
6	marginality.06	−1.954***	−2.080***
7		(0.124)	(0.117)
8	PAN.governor.06	−0.129	−0.312*
9		(0.172)	(0.167)
10	Constant	−3.727***	−3.810***
11		(0.227)	(0.222)
12			
13	Observations	2,406	2,407
14	Log Likelihood	−521.651	−645.606
15	Akaike Inf. Crit.	1,051.301	1,299.213
16			