

Discrete Time Stochastic Processes:

Consider system in state $X(t)$ at time t . $X(t)$ is a random variable.

First limit $X(t)$ to outcomes from a countable sample space:

$$\Omega = \{0, 1, 2, \dots\}.$$

Also limit time index t to integer values.

Sometimes $X(t+1)$ depends on entire history of system,
e.g. @ $t, t-1, t-2, \dots$

Defn: A Markov Chain is a system whose next state depends only on its current state:

$$\Pr[X(t+1)=j \mid X(t)=i_t, X(t-1)=i_{t-1}, \dots; X(0)=i_0] = \Pr[X(t+1)=j \mid X(t)=i_t].$$

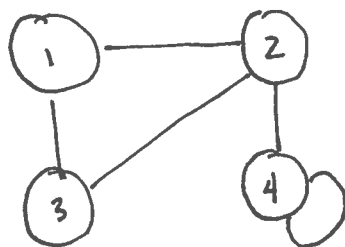
Defn: The transition matrix for a Markov chain is an arrangement of the individual $\Pr[X(t+1)=j \mid X(t)=i_t]$ values such that:

$$P(t) = \{P_{ij}(t)\} = \Pr[X(t+1)=j \mid X(t)=i_t].$$

Often P does not depend on time so $P(t) = P$.

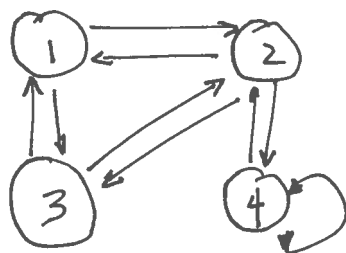
Ex: Cave system.

Consider 4 caves connected by tunnels.



Start in cave 1. Select (randomly) a tunnel and move to the next cave. Suppose each tunnel takes 1 hr to traverse.

Represent the caves as a state transition diagram:



then

$$P = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} & 0 \\ \frac{1}{3} & 0 & \frac{1}{3} & \frac{1}{3} \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ 0 & \frac{1}{3} & 0 & \frac{2}{3} \end{bmatrix}$$

Defn: A stochastic matrix is a matrix whose rows sum to 1.

Notice that P is a stochastic matrix.

Ex (cont)

Find \Pr [Cave K after 3 transitions]

Define $\pi(t) = \begin{bmatrix} \pi_1(t) & \pi_2(t) & \pi_3(t) & \pi_4(t) \end{bmatrix}$.

↓

$$\pi_i(t) = \text{Prob}[\text{in cave } i \text{ at time } t].$$

Start at: $\pi(0) = \begin{bmatrix} 1 & 0 & 0 & 0 \end{bmatrix}$.

Then $\pi_i(t+1) = \sum_{j=1}^4 \pi_j(t) \cdot p_{ji}$ [matrix form: $\pi(t+1) = \pi(t) P$].

that is: $\pi(1) = \pi(0) \cdot P = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix}$.

$$\pi(2) = \pi(1) \cdot P = \begin{bmatrix} \frac{5}{12} & \frac{3}{12} & \frac{2}{12} & \frac{2}{12} \end{bmatrix}.$$

$$\pi(3) = \pi(2) \cdot P = \begin{bmatrix} \frac{12}{72} & \frac{25}{72} & \frac{21}{72} & \frac{14}{72} \end{bmatrix}$$

In this case: as $t \rightarrow \infty$, $\pi(t) = \pi$ a limiting (steady-state)
↑
 constant.

value that does not depend on $\pi(0)$.

Defn: A Markov chain is aperiodic iff $d = \gcd \{ n : P(X_n = i | X_0 = i) > 0 \}$
= 1 for all states.

Notice a Markov chain with transition matrix P is

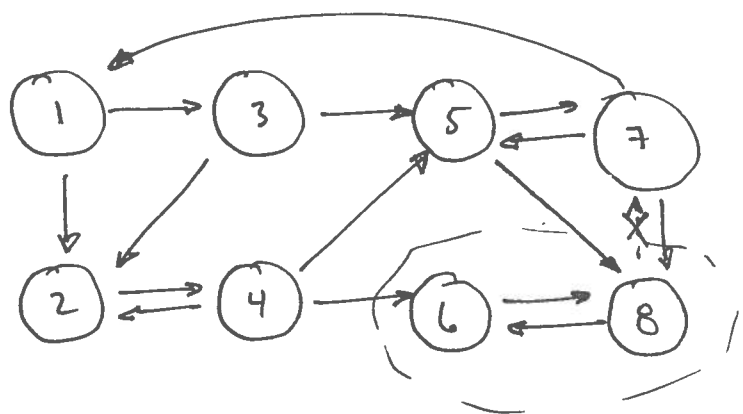
periodic if $|\lambda_1| = |\lambda_2| = 1$ for 2 distinct eigenvalues.

$$\left(\lambda \cdot P = \lambda \cdot x \right).$$

↑ eigenvalue ↑ eigenvector

Defn: A Markov chain is irreducible iff $P_{ij}^{(n)} > 0$ for all i and j and some $n \geq 1$.

i.e. every state "communicates" with every other state.



not irreducible.
states 6 and 8 are attracting.

Thm: (Perron Frobenius): Let P be a regular stochastic matrix.
Suppose P is irreducible and aperiodic. Then P has a unique positive eigenvector π with $\pi_i > 0$ for $1 \leq i \leq N$,

$$\pi P = \pi.$$

Thm: (Markov Chain Ergodic Theorem): For every non negative vector $x \in \mathbb{R}^N$ then if P satisfies the requirements for Perron-Frobenius

$$\lim_{n \rightarrow \infty} x \cdot P^n = \pi.$$

think $\left[\left[\left[x(0) \cdot P \right] \cdot P \right] \cdot P \right] \dots \rightarrow \pi.$

Analytic solution: (intuitive w/o proof):

If it exists then: $\underline{\pi P = \pi}$. \leftarrow Linear equations.

it exists then "flow of probability" into each state must balance flow leaving.

$$\text{i.e.} \quad \sum_j \pi_i P_{ij} = \sum_j \pi_j P_{ji}$$

\nearrow leaving j ($i \rightarrow j$) \nwarrow entering j ($j \leftarrow i$).

Also remember that $\sum_i \pi_i = 1$.

Solving gives: $\pi = \left[\frac{2}{10} \quad \frac{3}{10} \quad \frac{2}{10} \quad \frac{3}{10} \right]$.

Ex: (In honor of Valentines Day) Picking hearts.

Pick card from a deck of 52 cards (4 suits, 13 each). If heart drawn keep it. Else randomly return it to the deck. How long to pick all 13 hearts?

Note: After picking h hearts $\text{prob}[\text{next card is heart}] = \frac{13-h}{52-h}$.

Using a Markov chain: $H(k)$ # of hearts picked up after k draws:

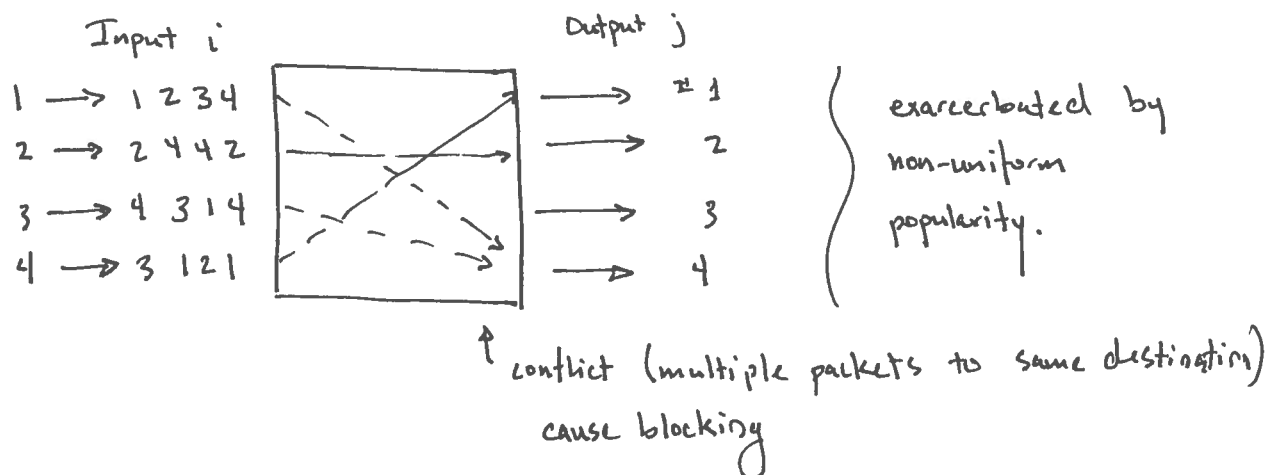
$$P_{ij}(k) = P_{ij} = \begin{cases} 0 & j \neq i, i+1 \\ \frac{39}{52-h} & j = i \\ \frac{13-h}{52-h} & j = i+1. \end{cases}$$

instead interested in the transient behavior.

Note: here limiting distribution not very interesting: $\lim_{k \rightarrow \infty} \{ \pi_{13}(k) \} = 1$.

Ex: Head of Line (HOL) Blocking switch

Consider $N \times N$ blocking switch ... only one packet delivered at a time to a particular output.



Start w/ 2×2 HOL switch.

Consider heavily loaded switch. \rightarrow every input always has a packet for transmission

Then state (x_1, x_2) w/ $x_i \in \{0, 1\}$ specifies output port is requested by each input.

There are 4 possible states: $(0, 0)$; $(0, 1)$; $(1, 0)$; $(1, 1)$

\rightarrow $(0, 0)$ and $(1, 1)$ allow only a single packet to transmit
 $(0, 1)$ and $(1, 0)$ allow both packets to transmit.

So for instance: current state $(0, 0)$. One packet delivered, one packet held.
 replaced on next interval: equally likely 0 or 1.
 "uniform popularity".