IDS 702: Module 6.3

Unconfoundedness and overlap

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OBSERVATIONAL STUDIES

- We will not focus on randomized experiments since most of the data you will have to analyze in practice are actually based on observational studies.
- In observational studies, we do not control or know the assignment mechanism.
- In addition, the presence of measured and unmeasured confounders can create unbalance between the groups.
- Again, to do causal inference, we have to make some structural (often untestable) assumptions, e.g. on the treatment assignment, for identifying causal effects.
- Once we have those general assumptions, we also usually have to make model assumptions to do the actual estimation.

ESTIMANDS

Once again, we will focus on the following estimands:

The average treatment effect (ATE):

$$au = \mathbb{E}[Y_i(1) - Y_i(0)].$$

The average treatment effect for the treated (ATT):

$$au = \mathbb{E}[Y_i(1) - Y_i(0)|W_i = 1].$$

The average treatment effect for the control (ATC):

$$au = \mathbb{E}[Y_i(1) - Y_i(0)|W_i = 0].$$

For binary outcomes, causal odds ratio (OR) or risk ratio (RR)::

$$au=rac{\mathbb{P}\mathrm{r}[Y_i(1)=1]/\mathbb{P}\mathrm{r}[Y_i(1)=0]}{\mathbb{P}\mathrm{r}[Y_i(0)=1]/\mathbb{P}\mathrm{r}[Y_i(0)=0]}.$$

ESTIMANDS

■ The relationship between ATE, ATT and ATC is given by

$$ext{ATE} = \mathbb{P} ext{r}[W_i = 1] \cdot ext{ATT} + \mathbb{P} ext{r}[W_i = 0] \cdot ext{ATC}$$

- In randomized experiments, ATT is equivalent to ATC because treatment and control groups are similar/comparable.
- ATE is then also equivalent to ATT (and ATC).
- In observational studies, ATE is usually different from ATT and ATC.
- The above relation does not hold for ratio estimands.

ASSUMPTIONS: UNCONFOUNDEDNESS

We will need two major assumptions (in addition to SUTVA). The first, we already talked about, that is,

Assumption 1: Unconfoundedness

$$Y_i(0), Y_i(1) \perp W_i | X_i,$$

or using the equivalent form from last class,

$$\Pr[W_i = 1 | X_i, Y_i(0), Y_i(1)] = \Pr[W_i = 1 | X_i]$$

- Assumes that within subgroups defined by values of observed covariates, the treatment assignment is random.
- Rules out unobserved confounders.
- Randomized experiments satisfy unconfoundedness.
- Untestable in most observational studies, but sensitivity can be checked.

IMPLICATIONS OF UNCONFOUNDEDNESS

Under unconfoundedness, it turns out that

$$\mathbb{P}\mathrm{r}[Y(w)|X] = \mathbb{P}\mathrm{r}[Y^{\mathrm{obs}}|X, W = w] \quad w = 0, 1.$$

■ That is, the observed distribution of Y in treatment arm W=w equals the distribution of the potential outcomes Y(w).

Why does this matter or how does this help us?

- Well, the causal estimands are essentially expectations and probabilities.
- Recall again that ATE is

$$ext{ATE} = \mathbb{E}[Y_i(1) - Y_i(0)].$$

ATE can then be estimated from the observed data using

$$ext{ATE} = \mathbb{E}_X \left(\mathbb{E}[Y^{ ext{obs}}|X,W=1] - \mathbb{E}[Y^{ ext{obs}}|X,W=0]
ight).$$

lacktriangle Note that we need to average out over the distribution of X since the original formula for ATE does not depend on any X.

ASSUMPTIONS: OVERLAP

Assumption 2: Overlap (or positivity)

$$0 < \mathbb{P}r[W_i = 1|X_i] < 1$$
, for all i .

Notice that this is the probabilistic assignment from last class, that is,

$$0 < \Pr[W_i = 1 | X_i, Y_i(0), Y_i(1)] < 1.$$

■ However, we can exclude $\{Y_i(0), Y_i(1)\}$ now because of the unconfoundedness assumption.

$$e(x) = \mathbb{P}\mathrm{r}[W_i = 1 | X_i = x]$$

is usually called the propensity score.

MPLICATIONS OF OVERLAP

- Overlap implies that, in large samples, for all possible values of the covariates, there are both treated and control units.
- This is important within the potential outcomes (or counterfactual) framework both conceptually and operationally (variance inflation).
- Unlike unconfoundedness, overlap can be directly checked from the data often using the estimated propensity scores.
- Unconfoundedness and positivity jointly define the strong ignorability assumption.

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WHAT'S NEXT?

MOVE ON TO THE READINGS FOR THE NEXT MODULE!

