

**Summary:**

Let  $X_1, X_2$  be two random variables with mean  $\mu_1, \mu_2$  and variance  $\sigma_1^2, \sigma_2^2$  respectively.

**Covariance:**  $Cov(X_1, X_2) = \mathbb{E}([X_1 - E(X_1)][X_2 - E(X_2)]) = \mathbb{E}(X_1 X_2) - \mu_1 \mu_2$ .

**Correlation:**  $\rho_{X_1, X_2} = \frac{Cov(X_1, X_2)}{\sqrt{Var(X_1)Var(X_2)}} = \frac{Cov(X_1, X_2)}{\sigma_1 \sigma_2}$ .

**Covariance matrix:**

$$\Sigma = \begin{pmatrix} Cov(X_1, X_1) & Cov(X_1, X_2) \\ Cov(X_2, X_1) & Cov(X_2, X_2) \end{pmatrix} = \begin{pmatrix} \sigma_1^2 & \rho_{X_1, X_2} \sigma_1 \sigma_2 \\ \rho_{X_1, X_2} \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}.$$

The determinant  $Det(\Sigma) = \sigma_1^2 \sigma_2^2 (1 - \rho_{X_1, X_2}^2)$ . Thus covariance matrix of  $X_1, X_2$  is invertible when the correlation is not  $\pm 1$ . A bi-variate normal density can be describe in term of inverse of covariance matrix.

**Markov's Inequality :** Let  $X$  be a non negative random variable with expectation  $E(X) = \mu$ . Then for any  $a > 0, P(X \geq a) \leq \frac{\mu}{a}$ .

**Chebyshev's Inequality :** Let  $X$  be a random variable with mean  $\mu$  and variance  $\sigma^2$ . Then for any  $a > 0, P(|X - \mu| \geq a) \leq \frac{\sigma^2}{a^2}$ .

**Limit Theorems**

Let  $\{X_k\}$  be a sequence of independent random variables and identical probability distributions with mean  $\mu$  and variance  $\sigma^2$ . Let  $S_n = X_1 + X_2 + \dots + X_n$ . Then  $\mathbb{E}(S_n) = n\mu$  and  $Var(S_n) = n\sigma^2$ .

**Weak Law of Large Number:** Let  $\{X_k\}$  be as above. Then for any  $\epsilon > 0, \lim_{n \rightarrow \infty} P(|\frac{S_n}{n} - \mu| \geq \epsilon) = 0$ .

**Strong Law of Large Number:** Let  $\{X_k\}$  be as above. Then  $P(\lim_{n \rightarrow \infty} \frac{S_n}{n} = \mu) = 1$ .

**Central Limit Theorem:** Let  $\{X_k\}$  be as above. Then for any  $t \in \mathbb{R}$ ,

$$\lim_{n \rightarrow \infty} P(\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-\frac{x^2}{2}} dx = \Phi(t).$$

Under the assumption that the moment generating function  $\phi_{X_k}$  exists, you saw a proof of Central Limit Theorem using of the Continuity Theorem. Recall that Moment generating function

$$\phi_X(t) = \mathbb{E}(e^{tX}) = 1 + \sum_{k=1}^{\infty} \frac{t^k}{k!} E(X^k).$$

For standard normal random variable  $X, \phi_X(t) = e^{\frac{t^2}{2}}$ . Note that the  $n$ -th moment  $E(X^n) = \frac{d^n \phi_X(t)}{dt^n} |_{t=0}$ .

**Continuity Theorem:** Let  $Z, Z_1, Z_2, \dots$  be a family of random variable such that  $\lim_{n \rightarrow \infty} \phi_{Z_n}(t) = \phi_Z(t)$  on an interval  $(-\alpha, \alpha)$ . Then  $\lim_{n \rightarrow \infty} F_{Z_n}(t) = F_Z(t)$  for all  $t$  where  $F_Z$  is continuous.

**Question**

1. Let  $(X, Y)$  be a bi-variate random variable with joint density  $f(x, y) = \frac{2}{x} e^{-2x}$  for  $0 < y \leq x < \infty$ . Find the Covariance matrix and Correlation.

2. Find the moment generating function of exponential random variable  $X$  with parameter  $\lambda$  and compute the third moment.
3. Let  $\{X_k\}$  be a sequence of independent Poisson random variables with parameter 1. Let  $S_n = X_1 + X_2 + \cdots + X_n$ . Estimate the probability  $P(S_{20} > 15)$  using Markov's inequality and Chebyshev's inequality. See if you can obtain a better estimate using Central limit theorem.
4. Let  $f$  be a continuous function on the interval  $I = [0, 1]$ . Consider the associated Bernstein polynomials  $B_n(x) = \sum_{k=0}^n f\left(\frac{k}{n}\right) \frac{n!}{k!(n-k)!} x^k (1-x)^{n-k}$ . Using Strong Law of large number show that  $\{B_n(x)\}$  converges to  $f(x)$  for  $0 < x < 1$ .