HW₃

Question 1

Already known that under logit model, we can have

$$logit\left(Prob(DLT \mid Dose = x)\right) = \frac{1}{\gamma - x_1} \{\gamma \cdot logit(\rho_0) - x_1 \cdot logit(\theta) + [logit(\theta) - logit(\rho_0)] \cdot \textbf{\textit{x}}\}$$

To get the posterior distribution, R is employed. Please see Q1.R for details. Results are shown in the following table.

Prior Distribution	25 percentile after 2 patients	25 percentile after 12 patiens	
$\gamma \sim Unif(50,400)$, $ ho_0 \sim Unif(0,0.2)$	$\gamma=160,\; ho_0=0.0$	$\gamma=250,\; ho_0=0.0$	1
$\gamma \sim Unif(50,650)$, $ ho_0 \sim Unif(0,0.2)$	$\gamma=230,\; ho_0=0.0$	$\gamma=290,\; ho_0=0.0$	2
$\gamma \sim Unif(50,400)$, $ ho_0 \sim Unif(0,0.25)$	$\gamma=150,\;\rho_0=0.1$	$\gamma=240,\;\rho_0=0.0$	3
$\gamma \sim Unif(50,650)$, $ ho_0 \sim Unif(0,0.25)$	$\gamma=240,\;\rho_0=0.0$	$\gamma=280,\;\rho_0=0.1$	4

Thus we can answer that:

The next dose levels for different prior distribution differs.

- o If the prior for the ρ_0 is fixed, the <u>prior distribution for γ will <u>hugely influence the posterior distribution</u> no matter how many patients have been already enrolled. As we can see from above that posterior γ in group 2 is greater than those in group 1 .</u>
- However, when the prior for γ is fiexed, the posterior distribution for γ will not change with the prior for ρ_0 . This may due to the two reasons:
 - The response from patients indicates that the ρ_0 should be small (from the data we can see that first 8 patients does not show AEs).
 - The prior distribution of ρ_0 does not change much (from 0.20 to 0.25) thus has a minor influence on γ .

The next dose levels given different patients information differs.

 As we can see from above form, the next dose level after 12 patients is always larger than the next dose level after 2 patients. This results from the ourcomes from patients since out of 12 patients only 9th patient and 12th patient reflects AEs, which gives more evidence that next dose can be set larger.

Question 2

2.1 Simon's optimal design

Given $p_1=0.6$, and $\alpha=\beta=0.1$,by using https://linus.nci.nih.gov/brb/samplesize/otsd.html we have

Your input Alpha -----> 0.10 Beta ----> 0.10 Response Probability of Poor Drug (P0) ---> 0.40 Response Probability of Good Drug (P1) ---> 0.60

Result

Optimal Two Stage Design	Optimum Design	MinMax Design
First Stage Sample Size (n1)	18	28
Upper Limit For 1st Stage Rejection of Drug (r1)	7	11
Maximum Sample Size (n)	46	41
Upper Limit for 2nd Stage Rejection of Drug (r)	22	20
Expected Sample Size If Response Probability = P0	30.22	33.84
Probability of Early Termination at P0	0.56	0.55

thus

$$N_1=18,\quad a_1=7,\quad N=46,\quad a_2=22$$

Therefore, if 7 or fewer successes in first 18 patients are observed then we should early stop here for futility.

If 22 or fewer successes in total 46 patients are observed then we should stop the trial.

2.2 Bayesian PP

Formulate this question as a hypothesis testing problem:

$$H_0: p \le 0.4, \quad v.s. \quad H_1: p > 0.6$$

where $p_0 = 0.4, p_1 = 0.6$.

1. Prior Distribution for response rate p:

$$p \sim Beta(0.3, 0.7)$$

2. Posterior distribution of the response rate p given X_{10}

$$p \mid X_{10} \sim Beta(0.3 + X_{10}, 0.7 + 10 - X_{10})$$

3. Future Distribution for Y

$$Y \mid X_{10} \sim Beta - Binomial(31, \, 0.3 + X_{10}, \, 10.7 - X_{10})$$

Since $\alpha=\beta=0.1$, set $Q_L=0.1, Q_T=0.9, Q_U=0.9$. Thus, the predictive probability(pp) should be calculated as:

$$PP = \sum_{i=1}^{31} P(Y=i \mid X_{10} = x) imes \mathcal{I}(Pr(p>0.4 \mid Y=i, X_{10} = x) > 0.9)$$

Since we don't know the value for X_{10} , so R is employed to loop over all possible values. Please see Q2.R for details.

X_{10}	0	1	2	3	4	5	6	7	8	9	10
PP	1e-5	0.0007	0.0094	0.0575	0.2018	0.4567	0.7308	0.9112	0.9824	0.9982	0.9999

Thus, we have to

- ullet Stop for futility if $X_{10} \leq 3$ and stop for efficacy if $X_{10} \geq 7$
- ullet Continue the next state if $4 \le X_{10} \le 6$

So to monitor the trial continously, we have the following rejection region (stop for futility):

- 3/10, 3/11, 4/12, 4/13, 5/14, 5/15, 6/16, 6/17, 7/18, 7/19, 8/20
- 8/21, 9/22, 9/23, 10/24, 10/25, 11/26, 11/27, 12/28, 12/29, 13/30,
- 13/31, 14/32, 14/33, 15/34, 16/35, 16/36, 17/37, 17/38, 18/39, 19/40, 20/41

and the following accept region (stop for efficacy):

- 7/10,8/11, 9/12, 9/13,10/14, 10/15, 11/16 ,11/17,12/18, 12/19, 13/20,
- 13/21, 14/22, 14/23, 15/24, 15/25, 16/26, 16/27, 16/28, 17/29, 17/30,
- 18/31, 18/32, 19/33, 19/34, 19/35, 20/36, 20/37, 21/38, 21/39, 21/40, 21/41

2.3 Table

Predictive probability based designs

$ heta_L$	$ heta_T$	r/N_{max}	$PET(p_0)$	$\mathbb{E}[N \mid p_0]$	α	β
0.001	[0.861, 0.918]	20 / 41	0.8711	31.07	0.0961	0.0977
		NA / 42				
0.001	[0.869, 0.922]	21 / 43	0.8785	32.12	0.0908	0.0926
		NA / 44				
0.001	[0.876, 0.926]	22 / 45	0.8854	33.12	0.0859	0.0880
0.022	[0.847, 0.906]	22 / 46	0.8721	27.45	0.1000	0.0838

2.4 Comparison

Parameter setting : $p_0=0.4, p_1=0.6, \alpha=\beta=0.1$

Your input							
Alpha> 0.10							
Beta> 0.10							
Response Probability of Poor Drug (P0)> 0.40							
Response Probability of Good Drug (P1)> 0.60							
Result							
Result Optimal Two Stage Design	Optimum Design	MinMax Design					
	Optimum Design	MinMax Design					
Optimal Two Stage Design							
Optimal Two Stage Design First Stage Sample Size (n1)	18	28					
Optimal Two Stage Design First Stage Sample Size (n1) Upper Limit For 1st Stage Rejection of Drug (n1)	18	28 11					
Optimal Two Stage Design First Stage Sample Size (n1) Upper Limit For 1st Stage Rejection of Drug (r1) Maximum Sample Size (n)	18 7 46 22	28 11 41					

	Parameter selection: Maximize Power. Theta_L and theta_T ranges:								
Theta_L a		anges: 0010							
Theta_T		9180							
	0.0010	3100							
			Null Case						
Pat.No.	Rej. Reg.	Rej. Reg.	Prob./(p0)	Prob./(p0)	Prob. Cont./(p				
	(Negative)	(Positive)	(Negative)	(Positive)					
10	1	11	0.0464	0.0000	0.9536				
11	1	12	0.0000	0.0000	0.9536				
12	1	13	0.0000	0.0000	0.9536				
13	2	14	0.0261	0.0000	0.9275				
14	2	15	0.0000	0.0000	0.9275				
15	2	16	0.0000	0.0000	0.9275				
16	3	17	0.0213	0.0000	0.9062				
17	3	18	0.0000	0.0000	0.9062				
18	4	19	0.0295	0.0000	0.8767				
19	4	20	0.0000	0.0000	0.8767				
20	5	21	0.0342	0.0000	0.8425				
21	5	22	0.0000	0.0000	0.8425				
22	6	23	0.0369	0.0000	0.8056				
23	6	24	0.0000	0.0000	0.8056				
24	7	25	0.0383	0.0000	0.7673				
25	7	26	0.0000	0.0000	0.7673				
26	8	27	0.0386	0.0000	0.7286				
27	8	28	0.0000	0.0000	0.7286				
28	9	29	0.0383	0.0000	0.6903				
29	10	30	0.0625	0.0000	0.6278				
30	10	31	0.0000	0.0000	0.6278				
31	11	32	0.0456	0.0000	0.5822				
32	12	33	0.0670	0.0000	0.5153				
33	12	34	0.0000	0.0000	0.5153				
34	13	35	0.0449	0.0000	0.4703				
35	14	36	0.0628	0.0000	0.4075				
36	15	37	0.0670	0.0000	0.3405				
37	16	38	0.0640	0.0000	0.2765				
38	17	39	0.0574	0.0000	0.2191				
39	18	40	0.0493	0.0000	0.1698				
40	19	41	0.0408	0.0000	0.1289				
41	20	42	0.0329	0.0961	0.0000				
======= Sum:			0.9039	0.0961					
			3.3033	2.0301					
	Theta_L)/(p0)								
	Theta_U)/(p0)								
PET Tota	L / (pø):	0.8711 31.0686							

Compared with Simon Optimal Design (SOD for short), predictive probabilty (PP) is harder to stop at the early stage of trials. However, PP can provides a more precise and continous instruction about when we should stop the trial and the confidence level of our choice whereas SOD only gives two discrete rejection point. This is because PP includes the uncertatinty of priors while SOD ingores it.