

Examples of PIFs

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1 Degenerated

Let $P(X = k) = 1$ for some $k > 0$. Let $R(x; \theta) \geq 1$ y $f(x) > 0$ arbitrary. The Population Impact Fraction is:

$$\text{PIF} = 1 - \frac{E[R(f(x); \theta)]}{E[R(x; \theta)]} = 1 - \frac{R(f(k); \theta)}{R(k; \theta)} \quad (1)$$

2 Multivariate normal

Let $X \sim N_k(\mu, \Sigma)$ normal multivariate of dimension k . Let $R(X; \theta) = e^{\theta^T X}$ and $f(X) = AX + b$. The Population Impact Fraction is:

$$\text{PIF} = 1 - \frac{E[R(f(x); \theta)]}{E[R(x; \theta)]} = 1 - \frac{\overbrace{\frac{1}{(2\pi)^k \sqrt{|\Sigma|}} \int_{\mathcal{X}} R(f(x); \theta) e^{-\frac{(x-\mu)^T \Sigma^{-1} (x-\mu)}{2}} dx}^{\text{N}}}{\underbrace{\frac{1}{(2\pi)^k \sqrt{|\Sigma|}} \int_{\mathcal{X}} R(x; \theta) e^{-\frac{(x-\mu)^T \Sigma^{-1} (x-\mu)}{2}} dx}_{\text{D}}} \quad (2)$$

where

$$D = \frac{1}{(2\pi)^k \sqrt{|\Sigma|}} \int_{\mathcal{X}} R(x; \theta) e^{-\frac{(x-\mu)^T \Sigma^{-1} (x-\mu)}{2}} dx = e^{\mu^T \theta + \frac{1}{2} \theta^T \Sigma \theta} \quad (3)$$

on the other hand:

$$\begin{aligned}
N &= \frac{1}{(2\pi)^k \sqrt{|\Sigma|}} \int_{\mathcal{X}} e^{\theta^T f(x)} e^{-\frac{(x-\mu)^T \Sigma^{-1} (x-\mu)}{2}} dx \\
&= \frac{e^{\theta^T b}}{(2\pi)^k \sqrt{|\Sigma|}} \int_{\mathcal{X}} e^{\theta^T A x + b} e^{-\frac{(x-\mu)^T \Sigma^{-1} (x-\mu)}{2}} dx \\
&= e^{(A\mu)^T \theta + \frac{1}{2} \theta^T A \Sigma A^T \theta + \theta^T b}
\end{aligned} \tag{4}$$

Substituting in (2) we obtain:

$$PIF = 1 - \frac{e^{\mu^T \theta + \frac{1}{2} \theta^T \Sigma \theta}}{e^{(A\mu)^T \theta + \frac{1}{2} \theta^T A \Sigma A^T \theta + \theta^T b}} \tag{5}$$

3 Exponencial

Let $X \sim \text{Exp}(\lambda)$ y $R(X; \theta) = 1 + \sum_{i=1}^n \theta_i x^i$. Let $f(x) = k_1 x + b$. The PIF is:

$$\begin{aligned}
PIF &= 1 - \frac{\int_0^\infty \frac{1}{\lambda} e^{-x/\lambda} \left[1 + \sum_{i=1}^n \theta_i f(x)^i \right] dx}{\int_0^\infty \frac{1}{\lambda} e^{-x/\lambda} \left[1 + \sum_{i=1}^n \theta_i x^i \right] dx} \\
&= 1 - \frac{1 + \int_0^\infty e^{-x/\lambda} \left[\sum_{i=1}^n \theta_i f(x)^i \right] dx}{1 + \int_0^\infty e^{-x/\lambda} \left[\sum_{i=1}^n \theta_i x^i \right] dx} \\
&= 1 - \frac{1 + \sum_{i=1}^n \theta_i \int_0^\infty [e^{-x/\lambda} f(x)^i] dx}{1 + \sum_{i=1}^n \theta_i \int_0^\infty [e^{-x/\lambda} x^i] dx}
\end{aligned} \tag{6}$$

where:

$$\int_0^\infty [e^{-x/\lambda} x^i] dx = (-\lambda)^{i+1} \Gamma(i+1, \frac{1}{\lambda} x) \Big|_{x=0}^{x \rightarrow \infty} = (-\lambda)^{i+1} \Gamma(i+1) \tag{7}$$

on the other hand:

$$\begin{aligned}
\int_0^\infty \left[e^{-x/\lambda} f(x)^i \right] dx &= \int_0^\infty \underbrace{\left[e^{-x/\lambda} (k_1 x + b)^i \right]}_{(k_1 x + b)^i = \sum_{j=0}^i \binom{i}{j} (k_1 x)^{i-j} b^j} dx \\
&= \sum_{j=0}^i \left[\binom{i}{j} b^j k_1^{i-j} \int_0^\infty \left(x^{i-j} e^{-x/\lambda} \right) dx \right] \\
&= \sum_{j=0}^i \left[\binom{i}{j} b^j k_1^{i-j} (-\lambda)^{i-j+1} \Gamma(i-j+1, \frac{1}{\lambda} x) \right] \Bigg|_{x=0}^{x \rightarrow \infty} \\
&= \sum_{j=0}^i \left[\binom{i}{j} b^j k_1^{i-j} (-\lambda)^{i-j+1} \Gamma(i-j+1) \right]
\end{aligned} \tag{8}$$

Substituting in (6) we obtain:

$$\text{PIF} = 1 - \frac{1 + \sum_{i=1}^n (-\lambda)^{i+1} \Gamma(i+1)}{1 + \sum_{i=1}^n \sum_{j=0}^i \left[\binom{i}{j} b^j k_1^{i-j} (-\lambda)^{i+1-j} \Gamma(i+1-j) \right]} \tag{9}$$

4 Lognormal

Let $X \sim \text{Lognormal}(\mu, \sigma)$. Let $f(x) = k_1 x + k_2$ with $k_1 \geq 1$ y sea $R(x; \theta) = e^{\theta x}$. Then:

$$\begin{aligned}
\text{PIF} &= 1 - \frac{\overbrace{E[e^{\theta k_1 x + k_2}]}^{=\infty}}{\underbrace{E[e^{\theta x}]}_{=\infty}} = 1 - e^{k_2} \frac{\int_0^\infty e^{k_1 x} \frac{1}{x} e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} dx}{\int_0^\infty e^x \frac{1}{x} e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} dx} \\
&= 1 - e^{k_2} \lim_{t \rightarrow 0} \frac{\overbrace{\int_0^{\frac{1}{t}} e^{k_1 x} \frac{1}{x} e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} dx}^{f(t)}}{\underbrace{\int_0^{\frac{1}{t}} e^x \frac{1}{x} e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} dx}_{g(t)}} \stackrel{\text{l'Hôpital}}{=} 1 - e^{k_2} \lim_{t \rightarrow 0} \frac{f'(t)}{g'(t)} \quad (10) \\
&= 1 - e^{k_2} \lim_{t \rightarrow 0} \frac{\left(\frac{1}{t^2}\right) e^{k_1 \frac{1}{t}} t e^{-\frac{(\log(1/t) - \mu)^2}{2\sigma^2}}}{\left(\frac{1}{t^2}\right) e^{\frac{1}{t}} t e^{-\frac{(\log(1/t) - \mu)^2}{2\sigma^2}}} \\
&= 1 - \lim_{t \rightarrow 0} e^{(k_1 - 1)/t} = \begin{cases} 1 - e^{k_2} & \text{if } k_1 = 1, \\ -\infty & \text{if } k_1 > 1, \\ 1 & \text{if } k_1 < 1. \end{cases}
\end{aligned}$$

5 Discrete finite

Let X random variable that takes values in the set $\{x_1, x_2, \dots, x_n\}$ with respective probabilities p_1, p_2, \dots, p_n such that $\sum_i p_i = 1$. Let $f(x)$ arbitrary and let $R(x) = \sum_{j=0}^l \theta_j x^j$ with $\theta_0 = 1$. The PIF is:

$$\text{PIF} = 1 - \frac{E[R(f(x); \theta)]}{E[R(x; \theta)]} = 1 - \frac{\sum_{k=1}^n p_k R(f(x_k); \theta)}{\sum_{k=1}^n p_k R(x_k; \theta)} = 1 - \frac{\sum_{k=1}^n \sum_{j=0}^l p_k \theta_j f(x_k)^j}{\sum_{k=1}^n \sum_{j=0}^l p_k \theta_j x_k^j} \quad (11)$$

6 Poisson

Let $X \sim \text{Poisson}(\lambda)$ assume $R(X; \theta) = \sum_{i=0}^n \theta_i x^i$ with $\theta_0 = 1$. Sea $f(x) = ax + b$ arbitrary. Then:

$$\begin{aligned}
\text{PIF} &= 1 - \frac{E[R(f(x); \theta)]}{E[R(x; \theta)]} = 1 - \frac{E\left[\sum_{i=0}^n \theta_i f(x)^i\right]}{E\left[\sum_{i=0}^n \theta_i x^i\right]} \\
&= 1 - \frac{\sum_{i=0}^n \theta_i E[(ax + b)^i]}{\sum_{i=0}^n \theta_i E[x^i]} \\
&= 1 - \frac{\sum_{i=0}^n \theta_i E\left[\sum_{k=0}^i \binom{i}{k} (ax)^{i-k} b^k\right]}{\sum_{i=0}^n \theta_i E[x^i]} \\
&= 1 - \frac{\sum_{i=0}^n \sum_{k=0}^i \overbrace{\theta_i \binom{i}{k} a^{i-k} b^k}^{\pi_{i,k}} E[x^{i-k}]}{\sum_{i=0}^n \theta_i E[x^i]} \\
&= 1 - \frac{\sum_{i=0}^n \sum_{k=0}^i \pi_{i,k} h^{(i-k)}(0)}{\sum_{i=0}^n \theta_i h^{(i)}(0)}
\end{aligned} \tag{12}$$

where $h(t) = e^{\lambda(e^t - 1)}$ y $h^{(m)}$ denotes the m -th derivative of h .