

# Ho-Lee model calibration

IntQuant 2021



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## Agenda

- Short rate models
- Fitting the curve
- Pricing swaptions
- Calibration
- Pricing Bermudans
- Implementation
- Q&A

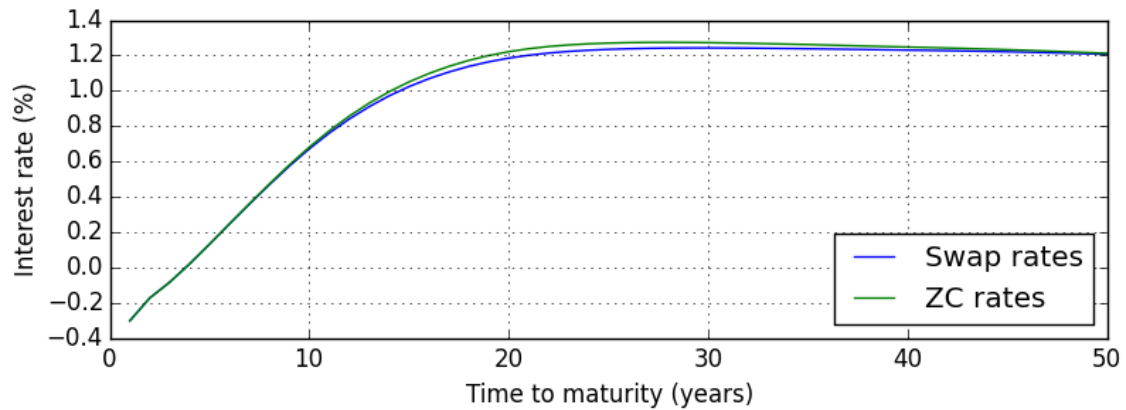


Figure 1: EURIBOR yield curve. Swap rates are the numbers quoted on the market and zero-coupon rates are obtained by bootstrapping and used for discounting.

**Interest rate curve** Discounting with annual rate  $r$ :  $B(0, T) = \left(1 + \frac{1}{r}\right)^{-T}$

With continuous compounding:  $B(0, T) = e^{-rT}$

In general  $r$  is different for different  $T$ :

Which rate should we model?

**Instantaneous rate** Instantaneous rate is the annualized rate used for discounting for an infinitesimally short maturity:

$$r(t) = \lim_{T \rightarrow t^+} -\frac{\ln B(t, T)}{T - t}$$

Hence:

$$B(0, T) = \mathbb{E} \left( e^{-\int_0^T r(t) dt} \right)$$

- Not used directly for discounting any product
- Not quoted on the market
- Not corresponding to any “real” object
- Convenient theoretical concept
- Single parameter for entire curve
- Closed-form solutions in many models

## Short-rate models

one-factor models without jumps:  $dr_t = a(t, r_t) dt + b(t, r_t) dW_t$

Model	Dynamics (SDE)	$r > 0$	$r \sim$	AB	AO
Vařicek	$dr_t = (\theta - \alpha r_t)dt + \sigma dW_t$	N	$\mathcal{N}$	Y	Y
CIR	$dr_t = (\theta - \alpha r_t)dt + \sigma\sqrt{r_t}dW_t$	Y*	$NC\chi^2$	Y	Y
RB	$dr_t = \theta r_t dt + \sigma r_t dW_t$	Y	$LN$	Y	N
Ho-Lee	$dr_t = \theta_t dt + \sigma dW_t$	N	$\mathcal{N}$	Y	Y
Hull-White	$dr_t = (\theta_t - \alpha r_t)dt + \sigma_t dW_t$	N	$\mathcal{N}$	Y	Y
BDT	$d \ln r_t = (\theta_t + \frac{\sigma'_t}{\sigma_t} \ln r_t)dt + \sigma_t dW_t$	Y	$LN$	N	N
BK	$d \ln r_t = (\theta_t + \phi_t \ln r_t)dt + \sigma_t dW_t$	Y	$LN$	N	N

AB = analytical formula for zero-coupon bond pricing

AO = analytical formula for bond option pricing

For more details & examples see e.g.: D. Brigo, F. Mercurio, Interest Rate Models - Theory and Practice [1].

\* Assuming the Feller condition is satisfied, i.e.  $\theta \geq \frac{1}{2}\sigma^2$

**Model parameters** Vařicek model:

$$dr(t) = (\theta - \alpha r(t))dt + \sigma dW(t)$$

$\theta$  - drift parameter. Calculated from swap rates. A term structure  $\theta(t)$  needed to fit the full interest rate curve.

$\sigma$  - volatility. Calibrated to swaptions. A term structure  $\sigma(t)$  needed to fit all ATM swaptions or all co-terminal swaptions for a Bermudan.

$\alpha$  - mean reversion rate. Determines the short rate's autocorrelation:

$$\text{Corr}(r(s), r(t)) = \sqrt{\frac{e^{2\alpha s} - 1}{e^{2\alpha t} - 1}}, \quad t > s > 0$$

higher  $\alpha$  - lower autocorrelation.

**Distribution of  $r(t)$  in Ho-Lee [3]** Dynamics:

$$dr(t) = \theta(t)dt + \sigma dW(t)$$

By integrating from  $s$  to  $t$ , we get:

$$r(t) = r(s) + \int_s^t \theta(u)du + \sigma(W(t) - W(s))$$

In particular, for  $s = 0$ , using the fact that  $W(t) \sim \mathcal{N}(0, t)$ , we get

$$\mathbb{E}(r(t)) = r(0) + \int_0^t \theta(u)du$$

$$\text{Var}(r(t)) = \sigma^2 t$$

Finally,  $r(t) \sim \mathcal{N}\left(r(0) + \int_0^t \theta(u)du, \sigma^2 t\right)$

**Distribution of  $r(t)$  in Hull-White [5] and Vařiček [11]** Dynamics:

$$dr(t) = (\theta(t) - \alpha r(t))dt + \sigma(t)dW(t) \quad (1)$$

This cannot be integrated directly because of the  $\alpha r(t)dt$  term on the RHS. Note that by Itô's lemma

$$d(e^{\alpha t}r(t)) = e^{\alpha t}(dr(t) + \alpha r(t)dt)$$

Hence, using (1), we get

$$d(e^{\alpha t}r(t)) = e^{\alpha t}\theta(t)dt + e^{\alpha t}\sigma(t)dW(t)$$

Integrating from  $s$  to  $t$  and multiplying by  $e^{-\alpha t}$ , we obtain

$$r(t) = r(s)e^{-\alpha(t-s)} + \int_s^t \theta(u)e^{-\alpha(t-u)}du + \int_s^t \sigma(u)e^{-\alpha(t-u)}dW(u)$$

Hull-White:

$$\begin{aligned} \mathbb{E}(r(t)) &= r(0)e^{-\alpha t} + \int_0^t \theta(u)e^{-\alpha(t-u)}du \\ \text{Var}(r(t)) &= \int_0^t \sigma(u)^2 e^{-2\alpha(t-u)}du \end{aligned} \quad (2)$$

In Vařiček this becomes (assuming  $\alpha \neq 0$ ):

$$\begin{aligned} \mathbb{E}(r(t)) &= r(0)e^{-\alpha t} + \theta \cdot \frac{1 - e^{-\alpha t}}{\alpha} \\ \text{Var}(r(t)) &= \sigma^2 \cdot \frac{1 - e^{-2\alpha t}}{2\alpha} \end{aligned}$$

As before,  $r(t) \sim \mathcal{N}(\mathbb{E}(r(t)), \text{Var}(r(t)))$ .

**Zero-coupon bond pricing** We want to calculate

$$B(t, T) = \mathbb{E} \left( e^{-\int_t^T r(s)ds} | \mathcal{F}_t \right) \quad (3)$$

For  $s > t$ :

$$r(s) = r(t)e^{-\alpha(s-t)} + \int_t^s \theta(u)e^{-\alpha(s-u)}du + \int_t^s \sigma(u)e^{-\alpha(s-u)}dW(u)$$

We integrate from  $t$  to  $T$ . Note that  $t \leq u \leq s \leq T$ .

$$\begin{aligned} \int_t^T \int_t^s \theta(u)e^{-\alpha(s-u)}duds &= \int_t^T \theta(u) \left( \int_u^T e^{-\alpha(s-u)}ds \right) du \\ \int_t^T \int_t^s \sigma(u)e^{-\alpha(s-u)}dW(u)ds &= \int_t^T \sigma(u) \left( \int_u^T e^{-\alpha(s-u)}ds \right) dW(u) \end{aligned}$$

Let's denote

$$D(u, T) = \int_u^T e^{-\alpha(s-u)} ds \quad (4)$$

After integration:

$$\begin{aligned} \text{Hull-White:} \quad D(u, T) &= \frac{1 - e^{-\alpha(T-u)}}{\alpha} \\ \text{Ho-Lee:} \quad D(u, T) &= T - u \end{aligned}$$

$$\begin{aligned} \int_t^T r(s)ds &= r(t)D(t, T) + \int_t^T \theta(u)D(u, T)du + \int_t^T \sigma(u)D(u, T)dW(u) \\ \int_t^T r(s)ds - r(t)D(t, T) &\sim \mathcal{N}\left(\int_t^T \theta(u)D(u, T)du, \int_t^T \sigma(u)^2 D(u, T)^2 du\right) \end{aligned}$$

$$X \sim \mathcal{N}(m, \sigma^2) \Rightarrow \mathbb{E}(e^{-X}) = e^{-m + \frac{1}{2}\sigma^2}$$

Hence, using (3):

$$B(t, T) = \exp\left(-r(t)D(t, T) - \int_t^T \theta(u)D(u, T)du + \frac{1}{2} \int_t^T \sigma(u)^2 D(u, T)^2 du\right) \quad (5)$$

Ho-Lee:

$$B(t, T) = \exp\left(-r(t)(T-t) - \int_t^T \theta(u)(T-u)du + \frac{1}{6}\sigma^2(T-t)^3\right)$$

Vašiček:

$$B(t, T) = \exp\left(-r(t)D(t, T) - \theta \int_t^T D(u, T)du + \frac{1}{2}\sigma^2 \int_t^T D(u, T)^2 du\right)$$

### Fitting the curve

$B(0, T)$  - implied from the yield curve

$\int_0^T \theta(u)D(u, T)du$  - chosen to fit  $B(0, T)$

In Vašiček we only can get approximate fit by manipulating  $\theta$  and  $\alpha$ .

In Hull-White and Ho-Lee we calculate from (5)

$$\int_0^T \theta(u)D(u, T)du = -\ln B(0, T) - r(0)D(0, T) + \frac{1}{2} \int_0^T \sigma(u)^2 D(u, T)^2 du \quad (6)$$

Inputs in the formula:

- $B(0, T), r(0)$  - implied from yield curve
- $\sigma(u)$  - model parameter, calibrated to swaptions
- $D(u, T)$  - depends on choice of  $\alpha$

$\theta$  not directly used for pricing - only  $\int_t^T \theta(u)D(u, T)du$  is needed

**Fitting the curve - Ho-Lee** In the particular case of the Ho-Lee model (i.e.  $\alpha = 0$  and  $\sigma(u) = \sigma$ ), (6) simplifies to:

$$\int_0^T \theta(u)(T-u)du = -\ln B(0, T) - r(0)T + \frac{1}{6}\sigma^2 T^3$$

Differentiating twice wrt  $T$ , we get

$$\theta(T) = \frac{\partial}{\partial T} f(0, T) + \sigma^2 T \quad (7)$$

where  $f(t, T)$  is the instantaneous forward rate for  $T$  at time  $t$ :

$$f(t, T) = -\frac{\partial \ln B(t, T)}{\partial T}$$

$f(0, T)$  - obtained directly from the yield curve

$$f(t, t) = r(t)$$

**Bond option pricing** Differentiating (5) we get:

$$dB(t, T) = r(t)B(t, T)dt - \sigma(t)D(t, T)B(t, T)dW(t)$$

Solving analogically to Black-Scholes model, we get option prices:

$$\begin{aligned} \mathbf{BC}(0; S, T, K) &= B(0, T)N(d_+) - KB(0, S)N(d_-) \\ \mathbf{BP}(0; S, T, K) &= KB(0, S)N(-d_-) - B(0, T)N(-d_+) \end{aligned} \quad (8)$$

where

$$d_{\pm} = \frac{\ln \frac{B(0, T)}{B(0, S)K} \pm \frac{1}{2}\nu(0, S)}{\sqrt{\nu(0, S)}},$$

$$\begin{aligned} \nu(0, S) &= \text{Var}(\ln B(S, T)) \stackrel{(5)}{=} D(S, T)^2 \text{Var}(r(S)) \\ &\stackrel{(2)}{=} D(S, T)^2 \int_0^S \sigma(u)^2 e^{-2\alpha(S-u)} du \end{aligned} \quad (9)$$

**Pricing swaptions** Fixed leg:

$$N \cdot \sum_{i=1}^n K \cdot \tau_i \cdot B(t, T_i)$$

Floating leg:

$$N \cdot (1 - B(t, T_n))$$

- $N$  - notional
- $K$  - strike (fixed coupon)

- $T_0$  - exercise of the option
- $T_1, \dots, T_n$  - fixed coupon payments
- $\tau_i = T_{i-1} - T_i$

Value of Payer Swaption at time  $T_0$ :

$$\mathbf{PS}(T_0; N, K, \mathbf{T}) = N \cdot \left( 1 - K \cdot \sum_{i=1}^n \tau_i \cdot B(T_0, T_i) - B(T_0, T_n) \right)^+ \quad (10)$$

**Pricing swaptions - Jamshidian's trick [6]** Express (10) as sum of options. Note that all  $B(T_0, T_i)$  are decreasing functions of a single factor:  $r(T_0)$ .

Denote:

$$B(t, T) = F(t, r(t); T)$$

Find the break-even rate  $\tilde{r}$ , for which:

$$1 - K \cdot \sum_{i=1}^n \tau_i \cdot F(T_0, \tilde{r}; T_i) - F(T_0, \tilde{r}; T_n) = 0 \quad (11)$$

Option is exercised if and only if  $r(T_0) \geq \tilde{r}$ . In such case

$$B(T_0, T_i) \leq K_i := F(T_0, \tilde{r}; T_i), \quad i = 1, 2, \dots, n$$

Finally,

$$\mathbf{PS}(T_0; N, K, \mathbf{T}) = K \cdot \sum_{i=1}^n \tau_i \cdot (K_i - B(T_0, T_i))^+ + (K_n - B(T_0, T_n))^+$$

Payer swaption:

$$\mathbf{PS}(0; N, K, \mathbf{T}) = K \cdot \sum_{i=1}^n \tau_i \cdot \mathbf{BP}(0; T_0, T_i, K_i) + \mathbf{BP}(0; T_0, T_n, K_n) \quad (12)$$

Receiver swaption:

$$\mathbf{RS}(0; N, K, \mathbf{T}) = K \cdot \sum_{i=1}^n \tau_i \cdot \mathbf{BC}(0; T_0, T_i, K_i) + \mathbf{BC}(0; T_0, T_n, K_n) \quad (13)$$

where bond option values are given by (8)

For more details on all the calculations, see:

D. McInerney, T. Zastawniak, Stochastic Interest Rates [9].

## Calibration of volatility

One needs to consider the following:

- Parametrization of  $\sigma(t)$
- Calibration strategy
- Choice of instruments
- What about  $\alpha$ ?

A very good overview of these points is provided in:

- S. Gurrieri, M. Nakabayashi, T. Wong, Calibration Methods of Hull-White Model [2]

**Parametrization of  $\sigma(t)$**  Recall ((5) and (2)) that  $\sigma(t)$  is used in the below expressions which we would like to calculate analytically:

$$\int_t^T \sigma(u)^2 D(u, T)^2 du \quad \text{and} \quad \int_t^T \sigma(u)^2 e^{-2\alpha(T-u)} du$$

Idea 1:  $\sigma(t) = \sigma$  (constant - this is the case in Ho-Lee and Vašíček) - easy for calculation but cannot replicate the market swaption prices

Idea 2: Piecewise constant - if we're only interested in options expiring on dates  $T_0, T_1, \dots, T_n$ , then shape of  $\sigma(t)$  between these dates doesn't matter, as long as the total levels are calibrated

$$\sigma(t) = \begin{cases} \sigma_0, & t \leq T_0 \\ \sigma_i, & t \in (T_{i-1}, T_i], i = 1, \dots, n \\ \sigma_n, & t > T_n \end{cases}$$

Idea 3: Piecewise linear or any other interpolation between points  $T_0, T_1, \dots, T_n$  - smoother shape of  $\sigma(t)$  but no improvement on pricing with increased complexity of calibration

## Calibration strategies - bootstrapping

- Choose a swaption (or set of swaptions) expiring at  $T_0$ . From market price obtain  $\text{Var}(r(T_0))$  and hence calculate  $\sigma_0$
- From market price of a swaption expiring at  $T_1$  obtain  $\text{Var}(r(T_1))$ . It depends on  $\sigma_0$  and  $\sigma_1$ . You already know  $\sigma_0$ , so you can imply  $\sigma_1$
- Likewise, having calculated  $\sigma_0, \dots, \sigma_{i-1}$ , you obtain  $\sigma_i$  from prices of swaptions expiring at  $T_i$ .

Pros:



ATM vol	1y	2y	5y	7y	10y	12y	15y	20y	30y
1y	19.9%	22.4%	23.1%	22.1%	21.2%	20.6%	19.9%	19.4%	19.2%
2y	25.4%	26.3%	24.9%	23.6%	22.4%	21.8%	21.1%	20.5%	20.2%
5y	27.7%	27.3%	25.0%	23.9%	22.8%	22.3%	21.5%	20.8%	20.6%
7y	26.0%	25.4%	23.6%	22.7%	21.9%	21.4%	20.8%	20.2%	20.1%
10y	22.7%	22.5%	21.8%	21.2%	20.6%	20.1%	19.5%	18.9%	19.1%
12y	21.6%	21.5%	20.8%	20.4%	19.8%	19.4%	18.8%	18.3%	18.6%
15y	20.4%	20.3%	19.7%	19.3%	18.8%	18.4%	18.0%	17.7%	18.0%
20y	18.7%	18.6%	18.3%	17.9%	17.5%	17.3%	17.1%	17.1%	17.3%
30y	17.9%	17.7%	17.5%	17.6%	17.7%	17.7%	17.6%	17.6%	17.5%

Figure 2: A typical swaption volatility matrix. The presented numbers are log-normal volatilities for ATM strikes

- Computationally efficient - only 1-dim optimization needed
- Potential exact fit for chosen instruments

Cons:

- Results in spiky shape of  $\sigma(t)$
- Tends to overfit short-term instruments at cost of long-term ones

#### Calibration strategies - least squares

- Choose a pool of swaptions  $(\mathbf{S}_i)_{i=1}^m$
- Price each swaption using (12) and (13)
- Calculate the error as a function of  $\bar{\sigma} = (\sigma_0, \dots, \sigma_n)$ :

$$\phi(\bar{\sigma}) = \sum_{i=1}^m |V_{\text{model}}(\mathbf{S}_i; \bar{\sigma}) - V_{\text{market}}(\mathbf{S}_i)|^2$$

- Use a numerical approach to find  $\bar{\sigma}$  which minimizes  $\phi(\bar{\sigma})$ .  
A standard method for this is the Levenberg-Marquardt algorithm [7]. Open-source implementations exist in Python.

Computationally less efficient but generates more realistic results, reasonably fitting all instruments.

**Choice of instruments** Volatilities can be implied from caps or swaptions.

For each expiry we have multiple tenors quoted and only one parameter  $\sigma_i$ , so we cannot fit them all. Which of them to choose?

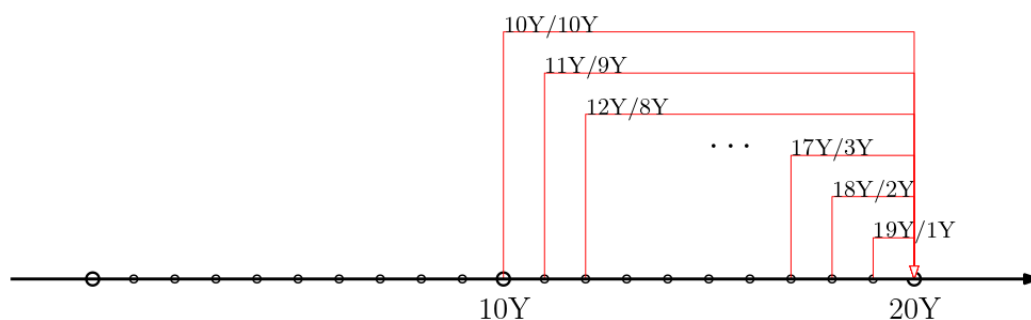


Figure 3: Co-terminal swaptions for a Bermudan

Idea 1: All of them. Put all available products to the least squares algorithm to get somewhere in the middle. You should get reasonable prices for all products but none of them is fit exactly. Good approach when the model needs to be used for many purposes without recalibration

### Choice of instruments - hedging

Idea 2: Hedging instruments. For pricing an exotic one needs to correctly price the hedging instruments. E.g. to price a Bermudan swaption one should calibrate to its co-terminal European swaptions

Example: Bermudan swaption, first exercise: 10Y, then every 1Y, tenor 10Y.

Hedging swaptions: 10Y/10Y, 11Y/9Y, 12Y/8Y, 13Y/7Y, ..., 19Y/1Y

For e.g. 11Y/9Y data may not be available (Figure 2) - time interpolation needed

Alternatively, calibrate directly to the available instruments, e.g.:

10Y/10Y, 12Y/10Y, 12Y/7Y, 15Y/5Y, 20Y/1Y

### Calibration of mean reversion (Vašíček and Hull-White)

Idea 1: Calibrate to correlation-dependent instruments, e.g. Bermudans - the higher auto-correlation, the lower switch value. Market data needed

Idea 2: Try to fit swaptions with different tenors and one expiry - term structure of  $\alpha$  may be needed

Idea 3: Calibrate together with  $\sigma$  to fit the pool of swaptions

Idea 4: Choose a fixed value based on an statistical analysis of historical observations

Whatever  $\alpha$  you choose (including  $\alpha = 0$ ), make sure to justify your choice.

**Pricing Bermudan swaptions** At each exercise date  $T_i$  the option holder has a right to choose: continue or exercise. Choosing to exercise means getting an European swaption with exercise date  $T_i$

and maturity  $T_n$ :

$$\mathbf{Ex}(T_i) = \mathbf{PS}(T_i; N, K, (T_{i+1}, \dots, T_n))$$

Continuation means getting at time  $T_{i+1}$  another Bermudan swaption with one exercise date less:

$$\mathbf{Cont}(T_i) = B(T_i, T_{i+1}) \mathbb{E}(\mathbf{Berm}(T_{i+1}; N, K, (T_{i+1}, \dots, T_n)) | \mathcal{F}_{T_i}) \quad (14)$$

Therefore,

$$\mathbf{Berm}(T_i) = \max\{\mathbf{Ex}(T_i), \mathbf{Cont}(T_i)\}$$

We can calculate  $\mathbf{Ex}(T_i)$  but  $\mathbf{Cont}(T_i)$  looks like a vicious circle...

But for the last exercise  $\mathbf{Cont}(T_{n-1}) = 0$ . Apply the backward propagation.

To calculate the expectation in (14) one needs to simulate  $r(T_{i+1})$  knowing  $r(T_i)$ .

- Monte Carlo? - note that for each value of  $r(T_i)$  we need a big number of paths for  $r(T_{i+1})$  and the simulations get multiply nested. Even for 3-4 exercise dates the number of calculations becomes unmanageable.
- Longstaff-Schwartz method. In the simulations we already have realized payoffs of  $\mathbf{Berm}(T_{i+1})$  for many different values of  $r(T_i)$ . By a clever use of least-squares method one can obtain the expectation of these payoffs as a function of  $r(T_i)$ . For more details see [8] or recall the lecture Introduction to MC Simulation in Finance.
- It is possible to explicitly calculate the conditional distribution of  $r(T_{i+1})$  given  $r(T_i)$ . This however requires a numerical calculation of an improper integral. This is done using a discretization of possible values of  $r$  to a finite set  $r_0, \dots, r_N$  and calculating  $\mathbf{Berm}(T_{n-1}; r_j)$  for  $j = 1, \dots, N$ . Then, by backward propagation, one calculates  $\mathbf{Berm}(T_{n-2}; r_j)$  for  $j = 1, \dots, N$ , etc. Formulas on the next slide.

$$\mathbb{E}(\mathbf{Berm}(T_{i+1}; r(T_{i+1})) | r(T_i)) = \frac{1}{\sqrt{2\pi\tilde{\sigma}^2}} \int_{-\infty}^{\infty} \mathbf{Berm}(T_{i+1}; u) \exp\left(-\frac{(u - \tilde{\mu})^2}{2\tilde{\sigma}^2}\right) du$$

where, in Ho-Lee:

$$\begin{aligned} \tilde{\mu} &= r(T_i) - f(0, T_i) + f(0, T_{i+1}) + \sigma^2 T_i \tau_{i+1} \\ \tilde{\sigma}^2 &= \tau_{i+1} \sigma^2 \end{aligned} \quad (15)$$

in Hull-White:

$$\begin{aligned} \tilde{\mu} &= (r(T_i) - f(0, T_i)) e^{-\alpha \tau_i} + f(0, T_{i+1}) \\ &\quad + \int_0^{T_i} \frac{\sigma^2(u)}{\alpha} (e^{-\alpha(T_i + T_{i+1} - 2u)} - e^{-2\alpha(T_{i+1} - u)}) du \\ \tilde{\sigma}^2 &= \int_{T_i}^{T_{i+1}} \sigma(u)^2 e^{-2\alpha(T_{i+1} - u)} du \end{aligned} \quad (16)$$

## Implementation

- Curve bootstrapping
- Monte Carlo
- Updating  $\theta$
- Pricing swaptions

### Implementation - curve bootstrapping

Input: swap rates

Output: zero rates, discount factors

- Make sure what is the convention of swap: payment frequencies, day count conventions, etc.
- For  $T_0$  convert the annual (or other appropriate) compounding to the continuous one.
- For  $T_i$  use precomputed zero rates for  $T_0, \dots, T_{i-1}$  and swap rate for  $T_i$  to compute the zero rate for  $T_i$ . E.g. if swap is annual and  $T_i = i + 1$ , then

$$r_{\text{swap}}(T_i) \sum_{k=0}^i B(0, T_k) = 1 - B(0, T_i)$$

- Use some interpolation method for zero rates between  $T_i$  and  $T_{i+1}$

### Implementation - Monte Carlo

Input: model, number of paths, time step

Output: matrix of simulated  $r$ 's for each time and path

- For short time steps you can use directly (1):

$$\begin{aligned} dt &= t[i+1] - t[i] \\ r[i + 1, \text{path}] &= (\text{theta}[i] - \text{alpha} * r[i, \text{path}]) * dt \\ &\quad + \text{sigma}[t] * \text{randn}(\text{mean}=0, \text{stddev}=\text{sqrt}(dt)) \end{aligned}$$

More time steps = longer computations

- If you're only interested in nodes  $T_0, \dots, T_n$  you can skip the intermediate time steps by using the forward distribution of  $r$  given by (15) or (16). The integrals are easily calculated if  $\theta$  and  $\sigma$  are piecewise constant. This approach is more efficient but cannot be used e.g. for American-style options.

To price any product calculate its value for each path using  $r[n, \text{path}]$  and then take average for all paths.

MC can be used for all calculations and is easy to implement. However, it requires huge amount of calculations to get a reasonable precision.

**Implementation – updating  $\theta$**  For Ho-Lee  $\theta$  is given by (7):

$$\theta(T) = \frac{\partial}{\partial T} f(0, T) + \sigma^2 T$$

where the differentiation needs to be performed numerically based on a reasonable curve interpolation.

Note that the simulation needs to be rerun with an updated  $\theta$  whenever either:

- curve moves - then  $B(0, T)$  and  $f(0, T)$  change
- $\sigma$  is changed (also during the calibration)

### Implementation - pricing swaptions

- For any times  $0 < t < T$  and any hypothetical value of  $r(t)$  calculate  $F(t, r(t); T)$ , using (5)
- Use any numerical root-finding algorithm (e.g. Newton or Brent) to find  $\tilde{r}$  solving (11).
- Use formulas (12) and (13) along with (8) to get the price

You can use Monte Carlo instead which is easier to implement but less efficient.

In any case, make sure you always use the correct  $\theta$ , marked to the yield curve and depending on  $\sigma$ .

## References

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