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RESEARCH ARTICLE

Overcoming collinearity in path analysis of soybean [*Glycine max* (L.) Merr.] grain oil content

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Abstract

Path analysis allows understanding the direct and indirect effects among traits. Multicollinearity in correlation matrices may cause a bias in path analysis estimates. This study aimed to: a) understand the correlation among soybean traits and estimate their direct and indirect effects on gain oil content; b) verify the efficiency of ridge path analysis and trait culling to overcome colinearity. Three different matrices with different levels of collinearity were obtained by trait culling. Ridge path analysis was performed on matrices with strong collinearity; otherwise, a traditional path analysis was performed. The same analyses were run on a simulated dataset. Trait culling was applied to matrix R originating the matrices R_1 and R_2 . Path analysis for matrices R_1 and R_2 presented a high determination coefficient (0.856 and 0.832, respectively) and low effect of the residual variable (0.379 and 0.410 respectively). Ridge path analysis presented low determination coefficient (0.657) and no direct effects greater than the effects of the residual variable (0.585). Trait culling was more effective to overcome collinearity. Mass of grains, number of nodes, and number of pods are promising for indirect selection for oil content.

Introduction

Soybean has been very important to Brazilian agribusiness. The oleaginous crop comprises approximately 20% oil by weight [1] and this product has been used in several applications, especially as kitchen oil and biodiesel. The crop is an important source of oil and ranks first in oil production in the world scenario [2] and the global demand for soybean oil is arising [3].

Quantitative traits, such as soybean oil, protein content, and grain yield are quantitative traits [4, 5], and are therefore controlled by a large number of genes, undergo a strong environmental influence, and usually show low heritabilities. Thus, the indirect selection is

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recommended to achieve genetic gains [6]. The viability of indirect selection is commonly investigated by correlations studies [7]. However, [8] emphasize that correlations among traits are purely a measure of association and do not investigate the cause-and-effect relationships and according to [9] the association between two traits can be influenced by a third trait or a group of them. [10] proposed path coefficient analysis, which is a multiple regression expanded and allows unfolding the correlation coefficients of direct and indirect effects on a basic variable [11]. Path Analysis was first used in plants by [12] and it was later used in other crops.

The estimation of the path analysis coefficients can be adversely affected by the multicollinearity effects among traits. Multicollinearity occurs when the observation of the independent variables, or their combinations, are correlated [11, 14]. When two variables are highly correlated it becomes hard to individually estimate the relations of each explanatory variable no the main variable because they contribute colletively to explain linear relations [14]. [13] advise performing the multicollinearity diagnostic on the correlation matrix of explanatory variables, since in presence of multicollinearity the variances associated to the path coefficients estimators may become too large and leads to unreliable estimates [14, 15]. In presence of multicollinearity, [12] recommend using ridge regression, which can obtain the solution to the vector parameters, even with significant multicollinearity, by the sum of small constant values (k) in the diagonal of the correlation matrix, as proposed by [16, 17]; or by excluding variables highly associated as presented by [18].

Many studies regarding path analysis on soybean that aimed to study the effects of traits on grain yield are available in the literature. [15, 19–24]. These studies show that it is usual to find high correlations among traits and, consequently, multicollinearity. Thus, methods to overcome multicollinearity must be studied, since most of these papers were not concerned to compare different methods for overcome collinearity.

Furthermore, despite of the large number of scientific papers about path analysis in soybean, no papers concerned grain oil content is found. Thus, the current study aimed to: a) understand the correlation among soybean traits as well as estimate its direct and indirect effects on oil content; b) verify the efficiency of ridge path analysis and trait culling to overcome colinearity problems.

Material and methods

Experimental conduction

The experiment was installed in 2016, February 22, in the Soybean Genetic Breeding Program's greenhouse of the Plant Science Department of the University Federal of Viçosa, Vicosa, Minas Gerais.

The experimental design was a randomized block with 5 replicates. Since the pots filled all tables of the greenhouse, we arranged the treatments in blocks in order to account possible environmental gradient. The treatments consisted of 34 F_1 soybean populations originated from a partial diallel and 15 commercial presented in S1 Appendix according to National Cultivar Registration [25]. Those cultivars used as genitors comprised between 5.5 and 9.0 [26]. The data of commercial cultivars were the ones used for statistical analysis. The plots consisted of two pots containing two plants of the same treatment (genotype). Sowing was performed in a substrate composed of three parts soil, two parts sand, and one part of organic substrate. The greenhouse was adapted using lamps in order to adjust the photoperiod and provide an appropriate environmental since the photoperiod is critical as the flowering inductor in soybean [27].

CODE	DESCRIPTION	ASSESSMENT METHODOLOGY
DF	Days to Flowering	from emergency stage (VE stage) until beginning bloom (R1 stage)
HF	Height of plants on Flowering	from the base until the last apical meristem at R_1 stage
DM	Days to Maturity	measured from the VE stage until full maturity (R ₈ stage)
HD	Hypocotyl Diameter	measured 2 centimeters below the cotyledon node at R ₈ stage
HM	Height of plants on Maturity	measured from the base to the last apical meristem at R ₈ stage
NP	Number of Pods per plant	number of pods per plant measured at R_8 stage
NN	Number of Nodes per plant	number of nodes in the main stem per plant measured at R_8 stage
NB	Number of Branches per plant	number of lateral branches per plant measured at $ m R_8$ stage
NG	Number of Grains per plant	measured at R ₈ stage
MG	Mass of Grains per plant	corrected to 13% of humidity
OC	Oil content of grains	measured in percentage of soybean intact grains at R8 stage and corrected to 13% of humidity
PC	Portein content of grains	

Table 1. Code, description and assessment methodology of traits evaluated in soybean plants.

The experiment was carried in accordance with the recommendations of [28]. Twelve vegetative, productive, and grain quality traits were evaluated in each plant as they reached the appropriate phenological stage, according to [29] as described in Table 1.

Genetic-statistical analysis

The genetic parameters were estimated and the genetic values predicted by analyzing only the genitors' information. The statistical analyses were performed with the genetics and statistics software SELEGEN-REML/BLUP as described by [30, 31], using the following statistical model:

$$y = Xr + Zg + e$$
, where: (1)

Where **y** is an (gr) x 1 vector of response variable, being g and r the number of genotypes and replicates, respectively; **r** is an (r) x 1 vector of unknown fixed effects fixed effects of replicates; **g** is an (g) x 1 vector of random effects of genetic values; and **e** is an (gr) x vector of random effects of error. **X** $(g \times r)$ and **Z** $(r \times g)$ are the incidence matrix.

The significance of the random effect was obtained by the likelihood ratio test by the analysis of deviance as recommended by 30. The analysis of deviance is derived from the difference between the deviance of the complete model and reduced model missing the effect which must be tested using chi-squared test with one degree of freedom [30, 32].

After obtaining the genetic values, we estimated the correlations between the traits. Then, a multicollinearity diagnosis in the original correlation matrix (matrix R) composed of 11 explanatory traits was performed. The multicollinearity diagnosis was performed considering the condition number as stated by, therefore, the matrices were classified as with weak, moderate, or strong collinearity for condition numbers lower than 100, between 100 and 1000, and greater than 1000, respectively [32].

Two procedures were used to address the problems caused by collinearity. Firstly, the path analysis through the ridge regression onto matrix R (Fig 1) as proposed by [16, 17] and secondly, traditional path analysis after eliminating traits which showed high correlation with another (trait culling). For the path analysis with trait culling, two new correlation matrices were obtained: R_1 and R_2 by the elimination of different traits in each matrix, and the traditional path analysis were run as shown in the diagrams (Figs 2 and 3, respectively). The traits eliminated were chosen considering its correlation with other traits kept in the matrix and how often they are evaluated in soybean breeding programs, for example: HF and HM are

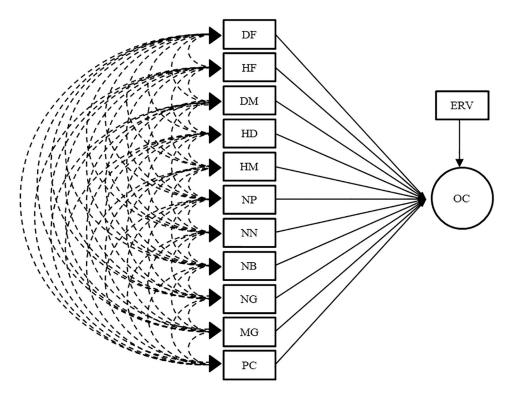


Fig 1. Path analysis diagram for matrices R and Rs's. Oil content was considered the main trait and DF—days to flowering; HF—height of plants on flowering; DM—days to maturity; HD—hypocotyl diameter; HM—height of plants on maturity; NP—number of pods per plant; NN—number of nodes in the main stem per plant; NB—number of branches per plant; NG—number of grains per plant; MG—mass of grains per plant; PC—protein content in grains; OC—oil content in grains and ERV—effect of residual variable, the explanatory variables.

highly correlated (0.92) and HF is less frequently measured in soybean breeding. Herewith, the reduction of collinearity was expected.

To establish the constant value K, the direct effects were plotted in a graph as a function of the values of K. The minimum value of K that provided stable estimates was chosen to run the analysis, as recommended by [12].

Simulation of a new dataset and validation of the path analysis procedures

To validate the results obtained with the experimental data, one hundred of new populations composed of 100 individuals were simulated using the means and covariance matrix of the original dataset in order to keep biological sense in the variables. The large number of simulated populations of few individuals avoids the reproduction of results obtained with original dataset.

Then, one correlation matrix was estimated for each population. After obtained the correlation matrices for each population the same procedures of path analysis were adopted to the simulated dataset. The same set of variables and procedures performed for R, R_1 and R_2 were also performed respectively for simulated matrices Rs's, R_1s 's and R_2s 's.

The best linear unbiased predictions (BLUPs) of the genotypes and variance components were estimated using the software SELEGEN-REML/BLUP [31]; and the correlation coefficients, path analysis, and multicollinearity diagnosis were performed for real data using the software GENES [33] and for simulated data using R software 3.5.3 [34].

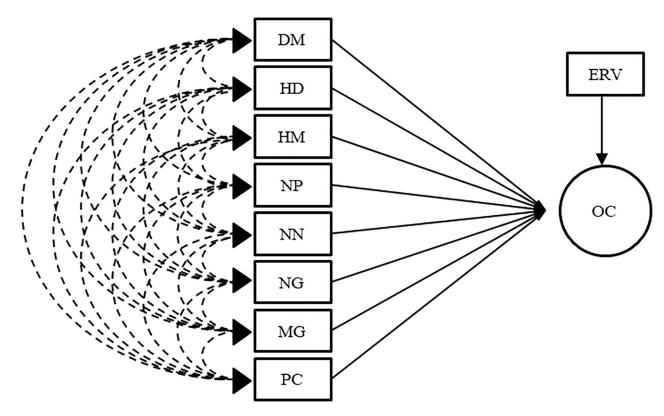


Fig 2. Path analysis diagram for matrices R₁ and R₁s'. Oil content was considered the main trait and DM—days to maturity; HD—hypocotyl diameter; HM—height of plants on maturity; NP—number of pods per plant; NN—number of nodes in the main stem per plant; NG—number of grains per plant; MG—mass of grains per plant; PC—protein content in grains; OC—oil content in grains and ERV-effect of residual variable the explanatory variables.

Results

Significant genetic variability was found for all vegetative, reproductive and grain quality traits. The lowest and the greater heritability (h^2) estimate were found for OC (0.368) and DF (0.992) respectively (Table 2). Grain quality traits (OC and PC) and MG showed lower heritability estimated than most of traits. Accuracy ($r_{\hat{a}a}$) values varied from 0.854 for OC to 0.992 for DF. The experimental coefficients of variation (CVe%) varied from 2.81 for PC to 20.3 for NS. Genotypic coefficient of variation (CVg%) ranged from 32.1 for HF to 3.00 for PC and CVg% values were greater than CVe% for all traits. The relative coefficient of variation, which is the ratio between the CVg% and CVe% ranged from 20.3 for HF and 2.81 for NP.

Genotypic values prediction for each genotype is presented in <u>Table 3</u> and correlation coefficients estimated between genetic values are presented in <u>Table 4</u>. The correlation coefficients indicates that MG is high influenced by DF, DM, NN, and NP. Besides this, MG was more correlated to NP (0,79). With the respect to the oil content, NS showed the greater correlation with OC (0.75). The oil content was also strongly correlated with the yield components NN (0.64), NP (0.65), NG (0.67) and also with MG (0.67). High and significant correlations were found for number of pods (NP) and NN (0.88), and for NN and DM (0.86). Regarding the traits of grain quality, a high and negative correlation were found for PC × NS (-0.56), PC × NG (-0.53), and PC × OC (-0.60). The means of the correlation coefficients from the 100

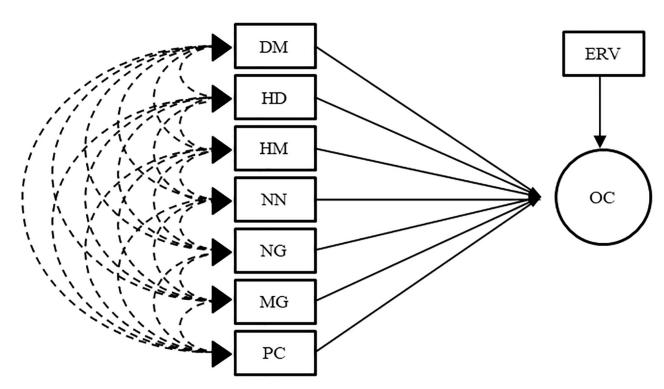


Fig 3. Path analysis diagram for matrices R_2 and R_2 s's. Oil content was considered the main trait DM—days to maturity; HD—hypocotyl diameter; HM—height of plants on maturity; NN—number of nodes in the main stem per plant; NG—number of grains per plant; MG—mass of grains per plant; PC—protein content in grains; OC—oil content in grains and ERV-effect of residual variable the explanatory variables.

Table 2. Genetic and environmental parameters estimated for days to flowering (DF), height of plants on flowering (HF), days to maturity (DM), hypocotyl diameter (HD), height of plants on maturity (HM), number of pods per plant (NP), number of nodes in the main stem per plant (NN), number of branches per plant (NB), number of grains per plant (NG), mass of grains per plant (MG), oil content in grains (OC), and protein content in grains (PC) evaluated in 15 soybean genotypes.

Parameters	DF	HF	DM	HD	НМ	NP	NN	NB	NG	MG	ос	PC
σ_{g}^{2}	34.3 *	201 *	90.2 *	0.333 *	174 *	52.5 *	1.84 *	0.486 *	151 *	2.85 *	0.667 *	1.38 *
$\sigma_{\rm e}^2$	2.81	24.4	28.5	0.228	71.9	10.6	0.544	0.832	65.4	1.79	1.24	1.21
$\sigma_{\rm f}^2$	37.1	224	119	0.561	246	63.2	2.39	1.32	217	4.64	1.91	2.58
h_{g}^{2}	0.924	0.891	0.76	0.593	0.708	0.832	0.772	0.368	0.698	0.613	0.35	0.532
$r_{\hat{a}a}$	0.992	0.988	0.97	0.938	0.961	0.98	0.972	0.863	0.959	0.942	0.854	0.922
CVg%	14.09	32.1	8.07	11.5	19.5	23.9	11.206	15.5	18.4	14.8	4.28	3
CVe%	4.03	11.2	4.53	9.54	12.5	10.8	6.091	20.3	12.1	11.7	5.84	2.81
CVr	3.49	2.86	1.78	1.21	1.56	2.22	1.84	0.764	1.52	1.26	0.733	1.07
PEV	0.552	4.76	5.36	0.04	13.3	2.04	0.103	0.124	12	0.319	0.181	0.206
SEP	0.743	2.18	2.31	0.2	3.64	1.43	0.321	0.352	3.47	0.565	0.425	0.453
Mean	41.5	44.1	118	5.01	67.8	30.3	12.1	4.49	66.8	11.4	19.1	39.1

^{*} Significant at the 0.01 probability level by the LRT test (Likelihood Ratio Test) using chi-squared test with one degree of freedom;

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 $^{^{\}dagger}\sigma_{g}^{2}$ = genotypic variance; σ_{e}^{2} = residual variance; σ_{f}^{2} = individual phenotypic variance; h_{g}^{2} = wide hereditability of individual plots; r_{fa} = selection accuracy of genotypes, assuming no plots lost; CVg% = genotypic coefficient of variation; CVe% = residual coefficient of variation; CVr = CVg/CVe = relative coefficient of variation; PEV = prediction error variance, assuming no plots lost; SEP = standard deviation of predicted genetic values, assuming no plots lost and Mean = general mean of three experiment.

Table 3. Genotypic values (u + g) of the genotypes evaluated for days to flowering (DF), height of plants on flowering (HF), days to maturity (DM), hypocotyl diameter (HD), height of plants on maturity (HM), number of nodes in the main stem per plant (NN), number of pods per plant (NP), number of branches per plant (NB), number of grains per plant (NG), mass of grains per plant (MG), oil content in grains (OC), and protein content in grains (PC) evaluated in 15 soybean genotypes.

Variety						Tra	its					
	DF	HF	DM	HD	HM	NN	NP	NB	NG	MG	ОС	PC
						u +	g					
BMX Apolo RR	36.0	21.4	106.0	4.5	45.4	20.9	10.1	3.5	54.2	9.0	17.9	41.2
DM 5958 IPRO	34.4	28.2	106.2	4.0	51.2	23.4	10.9	4.4	54.1	9.1	19.4	38.4
NA 5909 RR	35.2	42.7	107.4	4.6	72.4	26.1	10.5	4.9	60.0	10.7	18.8	38.6
TMG 7062 IPRO	36.7	35.6	103.5	4.7	59.5	22.1	10.9	4.1	53.0	10.9	18.9	38.8
DM 6563 IPRO	36.7	30.6	114.1	4.8	61.2	26.6	11.9	3.7	64.5	10.9	19.0	38.0
BMX Potência RR	37.1	29.7	111.3	4.7	59.9	24.0	11.3	4.2	58.6	9.7	18.3	39.6
BMX Desafio RR	34.3	30.8	117.3	5.6	62.7	23.8	11.7	3.9	58.2	10.8	18.1	40.2
M 7739 IPRO	45.4	50.5	119.4	5.5	66.5	33.1	12.3	5.0	76.4	12.5	19.5	37.6
P98Y30 RR	44.8	56.7	123.3	5.0	79.9	37.1	13.6	5.5	80.1	12.9	20.6	38.8
M8349 IPRO	45.4	46.0	124.3	5.1	63.7	35.5	12.6	5.0	82.6	13.9	19.8	37.8
TMG 132 RR	47.4	52.0	128.8	4.8	70.4	36.5	13.3	5.1	80.3	11.2	19.1	38.8
M 9056 RR	51.2	65.1	127.3	5.8	94.1	42.6	15.0	5.0	88.4	13.4	19.5	38.2
M 8221 RR	46.2	50.7	119.7	5.8	69.2	40.8	12.5	4.5	66.8	12.3	19.0	40.3
P98N71	45.8	57.1	131.2	4.9	76.3	31.1	13.1	4.5	61.0	10.9	19.3	39.8
Conquista	46.3	64.0	125.7	5.5	84.6	31.0	12.1	4.1	63.6	13.4	19.3	40.1

u + g = general mean summed to genotypic effects for each genotype.

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simulated populations showed to be very similar to those estimated from the original dataset as also shown in Table 4.

Multicollinearity diagnosis analysis showed severe multicollinearity for matrix R; moderate, but very close to the limit to be weak for matrix R_1 (not contemplating the traits DF, HF and NB); and weak for matrix R_2 (not contemplating the traits DF, HF, NN and NB) (Table 5).

There was no direct or indirect effect higher than the residual effect (0.585) and the determination coefficient was not satisfactory (0.657) for matrix R (Table 6). For the R₁ matrix, path analysis showed high determination coefficient (0.856) and the effect of the residual variable was outdone by direct effects of traits on OC (Table 7). The traits DM, HD, HM, and NG presented direct effects with sign opposite to the correlation coefficient in analysis of the matrix R₁. A weak association of DM and HM with OC was observed. On the other hand, NP, NN, and MG were noticed to be highly associated with OC by the direct effects 0.427, 0.632 and 0.967.

Strong direct effects of NN (0.804) and MG (1.05) on OC were found (Table 8). The ERV (0.410) and R^2 (0.832) found for path analysis on matrix R_2 fitted better than the analysis of the matrix R_1 , but not as well as the analysis of the matrix R_1 . Analysis of the matrix R_2 revealed results very similar to those of matrix R_1 .

The path analyses on Rs's matrices revealed a lower average k-value (0.334) and a higher determination coefficient (0.701) as presented in Table 9. The same promising traits (NP, NN and MG) for indirect selection in matrices R_1 were also appointed by path analysis on R_1 s's matrices whose average estimates are shown in Table 9. The average condition number estimated for simulated populations were slightly greater than those estimated for the original dataset.

Traits	DF	HF	DM	HD	HM	NP	NN	NB	NG	MG	OC	PC
DF	1	0.9	0.86	0.61	0.75	0.91	0.86	0.61	0.81	0.78	0.59	-0.19
		0.02	0.03	0.06	0.04	0.02	0.03	0.06	0.04	0.04	0.07	0.1
HF	0.9**	1	0.82	0.6	0.92	0.83	0.8	0.65	0.68	0.8	0.65	-0.2
			0.03	0.06	0.01	0.03	0.03	0.06	0.05	0.04	0.07	0.1
DM	0.86**	0.82**	1	0.59	0.72	0.79	0.86	0.51	0.71	0.69	0.51	-0.1
				0.07	0.05	0.05	0.03	0.08	0.05	0.06	0.09	0.11
HD	0.62*	0.61*	0.59*	1	0.62	0.63	0.58	0.22	0.5	0.73	0.14	0.06
					0.06	0.06	0.08	0.1	0.07	0.05	0.11	0.1
HM	0.75**	0.92**	0.73**	0.63*	1	0.73	0.79	0.56	0.62	0.73	0.53	-0.19
						0.04	0.03	0.07	0.05	0.04	0.08	0.1
NP	0.91**	0.83**	0.78**	0.65**	0.73**	1	0.88	0.72	0.87	0.78	0.65	-0.29
							0.03	0.05	0.02	0.04	0.06	0.09
NN	0.86**	0.80**	0.86**	0.59*	0.79**	0.88**	1	0.64	0.85	0.7	0.64	-0.33
								0.06	0.03	0.06	0.07	0.09
NB	0.62*	0.66*	0.51*	0.24	0.58*	0.72**	0.65**	1	0.77	0.56	0.76	-0.57
									0.04	0.06	0.04	0.07
NG	0.81**	0.68*	0.72**	0.52*	0.62*	0.87**	0.85**	0.77**	1	0.77	0.66	-0.54
										0.04	0.06	0.07
MG	0.78**	0.81**	0.69**	0.73**	0.73**	0.79**	0.70**	0.57**	0.77**	1	0.67	-0.36
											0.06	0.09
OC	0.60*	0.66*	0.51	0.16	0.54*	0.65**	0.64**	0.75**	0.67**	0.67**	1	-0.6
												0.06
PC	-0.19	-0.2	-0.09	0.05	-0.19	-0.28	-0.33	-0.56*	-0.53*	-0.36	-0.60*	1

Table 4. Estimates of genotypic correlation between the traits DF, HF, DM, HD, HM, NP, NN, NB, NG, MG, OC and PC for fifteen soybean genotypes from the original data set below the diagonal and for the simulated populations above the diagonal, including the standard deviation bellow.

Discussion

Genetic control and variance components

The genetic variability noticed between soybean cultivars shows that they can be used as genitor in breeding programs. The variability between genitors is fundamental to generate segregating populations that provide genetic gains [7].

Low heritability and accuracy were expected for mass of grains and quality traits, since they are controlled by a larger number of genes and are more affected by environmental factors [5, 6]. The values of accuracy found in this research are considered high according to [35]. These results indicate good experimental quality; thus, the data set and data analysis can be used in trustworthy inferences by researchers.

According a classification proposed by [36] the residual coefficients of variation (CVe%) were classified as low which indicates good experimental quality. The greatest value of CVe% was found for number of branches. According to [37] soybean plants have the ability to adjust the number of branches to compensate for the distance from other plants and intercept the maximum amount of light. The authors named this phenotypic plasticity. Despite equal spacing between pots, randomizing the treatments inside the blocks induced different levels of competition to the same genotype in different blocks. This effect was accounted for the environmental effects, therefore, decreases the coefficient of variance.

^{*} Significant at the 0.05 probability level;

^{**} Significant at the 0.01 probability level according to Mantel's test.

Table 5. Eigenvalues, condition number (CN) and its standard deviations between parenthesis of the correlation matrices R, R_1 , R_2 , R_3 , R_1 s and R_2 s between explanatory variables.

Rank				Eigenvalues		
	\mathbf{R}^{\dagger}	R ₁	R_2	Rs	R ₁ s	R ₂ s
1	7.67	5.13	4.72	7.65(0.28)	5.46(0.21)	4.61(0.20)
2	1.40	1.36	1.11	1.44(0.16)	1.19(0.11)	1.19(0.11)
3	0.590	0.550	0.388	0.618(0.09)	0.534(0.09)	0.523(0.09)
4	0.454	0.338	0.336	0.444(0.06)	0.336(0.05)	0.317(0.05)
5	0.314	0.282	0.240	0.302(0.05)	0.203(0.03)	0.185(0.03)
6	0.226	0.183	0.128	0.215(0.03)	0.150(0.03)	0.121(0.02)
7	0.147	0.111	0.077	0.139(0.02)	0.075(0.01)	0.050(0.01)
8	0.104	0.048		0.096(0.01)	0.047(0.01)	
9	0.058			0.054(0.01)		
10	0.037			0.033(0.01)		
11	0.004			0.003(0.00)		
CN	2081	105	86.5	2337(559.91)	119(22.53)	95.3(18.22)
Multicollinearity level	Severe	Moderate	Weak	Severe	Moderate	Weak

 $^{^{\}dagger}$ R and Rs = original set of the correlation matrix between explanatory variables from original and simulated dataset, respectively; R₁ and R₁s = correlation matrix with trait culling to three explanatory variables from original and simulated dataset, respectively; R₂ and R₂s = correlation matrix with trait culling to four explanatory variables from original and simulated dataset, respectively.

Table 6. Estimates of direct and indirect effects for: DF, HF, DM, HD, HM, NP, NN, NB, NG, MG and PC on OC evaluated in 15 soybean genotypes onto the matrix R.

					Expl	anatory vari	ables				
Effect	DF	HF	DM	HD	НМ	NP	NN	NB	NG	MG	PC
Direct on OC	0.035	0.141	0.019	-0.207	0.020	0.086	0.105	0.196	0.022	0.216	-0.205
Indirect via DF	-	0.032	0.031	0.022	0.027	0.032	0.031	0.022	0.029	0.028	-0.007
Indirect via HF	0.127	-	0.115	0.086	0.131	0.117	0.113	0.094	0.096	0.114	-0.029
Indirect via DM	0.017	0.016	-	0.011	0.014	0.015	0.016	0.010	0.014	0.013	-0.002
Indirect via HD	-0.128	-0.126	-0.122	-	-0.129	-0.134	-0.121	-0.049	-0.106	-0.151	-0.010
Indirect via HM	0.015	0.018	0.014	0.012	-	0.014	0.016	0.011	0.012	0.014	-0.004
Indirect via NP	0.079	0.071	0.068	0.056	0.063	-	0.076	0.063	0.075	0.068	-0.024
Indirect via NN	0.090	0.084	0.090	0.061	0.083	0.092	-	0.068	0.089	0.073	-0.034
Indirect via NB	0.121	0.130	0.101	0.046	0.113	0.142	0.126	-	0.151	0.111	-0.109
Indirect via NG	0.018	0.015	0.016	0.012	0.014	0.019	0.019	0.017	-	0.017	-0.012
Indirect via MG	0.169	0.174	0.148	0.158	0.158	0.170	0.152	0.122	0.167	-	-0.077
Indirect via PC	0.039	0.041	0.018	-0.010	0.040	0.057	0.067	0.114	0.108	0.073	-
Total [†]	0.60	0.66	0.51	0.16	0.54	0.65	0.64	0.76	0.67	0.67	-0.60
R ²						0.657					
K						0.445					
ERV						0.585					

[†]Total = correlation coefficient between the trait and OC; R² = Determination coefficient; K = constant used in the analysis and; ERV = effect of the residual variable.

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Table 7. Estimates of direct and indirect effects for: DM, HD, HM, NP, NN, NG, MG and PC on OC evaluated in 15 soybean genotypes onto the matrix R₁.

	Explanatory variables ^{††}												
Effect	DM	HD	НМ	NP	NN	NG	MG	PC					
Direct on OC	-0.126	-0.758	-0.110	0.427	0.632	-0.556	0.967	-0.223					
Indirect via DM	-	-0.075	-0.092	-0.099	-0.109	-0.091	-0.087	0.011					
Indirect via HD	-0.449	-	-0.474	-0.491	-0.445	-0.391	-0.556	-0.037					
Indirect via HM	-0.080	-0.069	-	-0.081	-0.087	-0.069	-0.081	0.021					
Indirect via NP	0.333	0.276	0.312	-	0.374	0.371	0.335	-0.119					
Indirect via NN	0.543	0.371	0.500	0.554	-	0.536	0.443	-0.207					
Indirect via NG	-0.398	-0.286	-0.347	-0.483	-0.472	-	-0.429	0.294					
Indirect via MG	0.663	0.709	0.709	0.760	0.679	0.747	-	-0.344					
Indirect via PC	0.020	-0.011	0.043	0.062	0.073	0.118	0.079	-					
Total [†]	0.51	0.16	0.54	0.65	0.64	0.67	0.67	-0.60					
R ²				0.0	856								
ERV		0.379											

 $^{^{\}dagger}$ Total = correlation coefficient between the trait and OC; R^2 = determination coefficient; and ERV = effect of the residual variable.

According to [38] an statistic usually employed on genetic evaluations is the ratio between the genetic coefficient of variation and the experimental coefficient of variation which is called relative coefficient of variation (CVr) and is related to experimental accuracy. Estimates of relative CVr higher than 1 indicate more accurate and precise inferences about the rank of the genotypes, which was found for 10 traits with the exception of NB and OC.

Genetic correlation among traits and multicollinearity diagnosis

The traits days to flowering, days to maturing, number of nodes and number of pods were found strongly correlated to mass of grains. These results corroborate with results for genetic correlation found by [39].

Table 8. Estimates of direct and indirect effects for: DM, HD, HM, NN, NG, MG, PC on OC evaluated in 15 soybean genotypes onto the matrix R2.

	Explanatory variables ^{††}											
Effect	DM	HD	HM	NN	NG	MG	PC					
Direct on OC	-0.158	-0.744	-0.095	0.804	-0.347	1.05	-0.148					
Indirect via DM	-	-0.094	-0.115	-0.136	-0.113	-0.018	0.014					
Indirect via HD	-0.440	-	-0.466	-0.437	-0.384	-0.546	-0.036					
Indirect via HM	-0.069	-0.059	-	-0.075	-0.059	-0.070	0.018					
Indirect via NN	0.690	0.472	0.636	-	0.682	0.564	-0.260					
Indirect via NG	-0.249	-0.179	-0.217	-0.295	-	-0.268	0.184					
Indirect via MG	0.718	0.768	0.768	0.735	0.809	-	-0.372					
Indirect via PC	0.013	-0.007	0.029	0.048	0.079	0.054	-					
Total [†]	0.51	0.16	0.54	0.64	0.67	0.67	-0.60					
R ²				0.832								
ERV				0.410								

[†]Total = correlation coefficient between the trait and OC; R² = determination coefficient; and ERV = effect of the residual variable.

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^{††} The traits DF, HF and NB were eliminated from the matrix R in originating the matrix R₁.

^{††} The traits DF, HF, NN and NB were eliminated from the matrix R in originating the matrix R₂.

-	-													
	DF	HF	DM	HD	НМ	NP	NN	NB	NG	MG	PC	K	ERV	R ²
Rs's†	0.011	0.169	0.016	-0.265	-0.013	0.099	0.140	0.211	-0.024	0.283	-0.203	0.334	0.544	0.701
	(0.04)	(0.05)	(0.04)	(0.09)	(0.07)	(0.04)	(0.07)	(0.04)	(0.08)	(0.1)	(0.04)	(0.14)	(0.05)	(0.04)
R ₁ s's	-	-	-0.124	-0.772	-0.106	0.416	0.631	-	-0.552	0.969	-0.214	-	0.363	0.866
	-	-	(0.10)	(0.10)	(0.07)	(0.11)	(0.16)	-	(0.12)	(0.11)	(0.07)	-	(0.04)	(0.03)
R ₂ s's	-	-	-0.146	-0.762	-0.091	-	0.803	-	-0.357	1.048	-0.143	-	0.394	0.844
	-	-	(0.10)	(0.10)	(0.08)	-	(0.16)	-	(0.12)	(0.11)	(0.07)	-	(0.04)	(0.03)

Table 9. Mean direct effects and standard deviation between parentheses for explanatory traits, k-value (K), effect of residual variable (ERV) and R-squared (R²) of the path analysis for the simulated populations.

†Rs's = matrices from simulated dataset composed of DF, HF, DM, HD, HM, NP, NN, NB, NG, MG and PC; R₁s's = matrices from simulated dataset composed of DM, HD, HM, NP, NN, NG, MG and PC; R₂s's = matrices from simulated dataset composed of DM, HD, HM, NN, NG, MG and PC.

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The strong correlation between oil content and grain yield found is reinforced by [40]. The high correlation of number of nodes and number of pods with grains yield are reported in the literature as well as found in this research [24, 39]. This is due to the grains yield comes from pods which are developed in axillary buds located on plant stems or branches originated from the nodes. Besides this, the high correlation between number of nodes and days to maturing shows that late maturing genotypes tend to present higher grain yield per plant. These results agree with those of [15].

A strong and negative correlation between oil and protein content were also found by [3, 41, 42] in molecular markers studies. [43] suggested that both traits are controlled by the same gene or group of genes, which were later reported as a major QTL in the crhomossome 20 by [44].

The elimination of traits that are strongly correlated and have biological interpretations such as DF and DM or HF and HM was shown to be effective to overcome multicollinearity. Thus, matrix R_1 or R_2 each contains just one of these traits from these pairs. This discussion is supported by the results of simulated data presented in <u>Table 4</u> that showed very similar results for 100 populations. However the matrices comprised of the same set of variables presented the same level of collinearity problems.

Path analysis results and comparison among procedures

The poor adjustment of the ridge path analysis onto R and Rs's matrices and the greater effects of the residual variable in comparison with the direct effects, leads us to affirm that the estimates are not reliable (Tables 6 and 8). [15] obtained direct effect estimates higher than 1 for two traits in a traditional path analysis on a correlation matrix with strong collinearity problems for soybean. The authors also ran a ridge analysis, which gave estimates between 0 and 1, and these two traits did not show relevant direct effects. The K value used by the authors was of 0.05, whereas the K value needed to overcome collinearity in our original and simulated dataset were much greater. [12] affirm that high K values lead to higher bias in the regression and may provide results without biological mean.

High values for effects of residual variable (ERV) may be found when the number of explanatory variables is not enough to explain variation of the main trait. However, the estimates for matrix R_1 showed that even after reducing the number of explanatory variables, it was possible to explain a large part of the total variance. This result was also observed by [14] when they found greater regression adjustment after trait culling of the explanatory variables. [10] affirm that the exclusion of variables allows estimating path coefficients with biological meaning, what is an advantage of trait culling.

A weak association of HM with OC was observed by the low direct and indirect effects, although HM showed significant correlation with OC. Likewise, a high correlation between PC and OC and weak direct effect was noticed. These results proves that the correlation coefficient does not necessarily give the best idea of association between variables and researches need to be careful at choose traits for indirect selection based just on correlation coefficient. On the other hand, NP, NN and MG are recommended to be used in indirect selection for OC, since the magnitude of direct effects were close to the correlation, which is the sum of direct and indirect effects involving each trait. In addition, these traits are easier, cheaper, and quicker to measure than laboratory procedures to measure OC. Besides this, NP and NN may be evaluated before complete maturation, which allows earlier selection, and MG is routinely evaluated in soybean experiments. In order to achieve genetic gains on OC we also recommend the use of a selection index including the traits that showed high direct effects on OC.

The direct effect of MG on OC from path analysis on R_2 matrix was 1.05, although estimates greater than 1 were not expected, once the matrix R_2 showed weak collinearity problems. The path analysis on R_2 s's matrices also presented means direct effects of MG on OC higher than 1. The strong and positive direct effects of NN and MG on OC showed in R_1 analysis were confirmed by path analysis on R_2 and R_2 s's matrices and encourage researches to use them for indirect selection. The same discussion mentioned above regarding PC in matrix R_1 can be applied to matrix R_2 .

Although the matrix R_1 showed moderate to strong problems with collinearity according to [45], the path analysis onto this set of variables was the most suitable for this study, since it provided the greatest determination coefficient of the regression, the lowest effect of the residual variable, and all effects were estimated inside the parameter space, which ranges from -1 to 1 for direct, indirect and residual variable effects. These results corroborate with those found by [14] when the authors tested three methods of path analysis (traditional, ridge path analysis and trait culling) and concluded that the trait culling provided a fit statistics more accurate. The discussions about the quality of the path analysis onto R_1 matrix were validated by path analysis onto R_1 s's matrices according to the great regression adjustment (0.844) and the high effect of the residual variable. Furthermore, the choice of the R_1 matrix allows considering one more variable in the study (NP).

The ridge path analysis was not an appropriated approach for overcome the collinearity problem, since a low R^2 (0.657) and high ERV (0.585) were found. Besides of this, no direct effects greater than ERV were found for analysis onto matrices R and $R_{s's}$ (Tables 6 and 8). Path analysis after trait culling for matrices R_1 and R_2 showed to be the best strategy for overcome collinearity problem since it presented a well-fit of the regression denoted by the values of R^2 (0.856 and, 0.832 respectively) and direct effects greater than the ERV (0.379 and 0.410, respectively). The results found onto simulated matrices supported these findings (Table 9). In addition, trait culling is a simple technique and allows to the researcher exclude the variable that contributes more to the collinearity considering its biological relationship with the main trait.

Conclusions

There is a strong relation of causation of mass of grains per plant, number of nodes, and number of pods with the oil content in grains and they could be used for indirect selection. Trait culling was more suitable and simpler strategy than ridge path analysis to overcome collinearity problems.

Eliminate one between two similar traits biologically was efficient to reduce condition number and, consequently, overcome multicollinearity.

Supporting information

S1 Appendix. Maintainer company, cultivar name and to National Cultivar Registration (RNC) register number of 15 commercial soybean cultivars used for the research. (DOCX)

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Author Contributions

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References

- Lazzarotto JJ, Hirakuri MH. Evolução e perspectivas de desempenho econômico associadas com a produção de soja nos contextos mundial e brasileiro. Londrina Embrapa Soja. 2010; 319: 46.
- Yu J, Zhang Z, Wei J, Ling Y, Xu W, Su Z. SFGD: a comprehensive platform for mining functional information from soybean transcriptome data and its use in identifying acyl-lipid metabolism pathways. BMC Genomics. 2014; 15: 271. https://doi.org/10.1186/1471-2164-15-271 PMID: 24712981
- Liang H, Yu Y, Wang S, Lian Y, Wang T, Wei Y, et al. QTL Mapping of Isoflavone, Oil and Protein Contents in Soybean (Glycine max L. Merr.). Agric Sci China. Elsevier; 2010; 9: 1108–1116. https://doi.org/10.1016/S1671-2927(09)60197-8
- Lee SH, Bailey MA, Mian MAR, Shipe ER, Ashley DA, Parrott WA, et al. Identification of quantitative trait loci for plant height, lodging, and maturity in a soybean population segregating for growth habit. Theor Appl Genet. Springer-Verlag; 1996; 92: 516–523. https://doi.org/10.1007/BF00224553 PMID: 24166318
- da Rodrigues JIS, de Miranda FD, Ferreira A, Borges LL, d S Ferreira MF, Good-God PI V, et al. Mapeamento de QTL para conteúdos de proteína e óleo em soja. Pesqui Agropecu Bras. 2010; 45: 472–480. https://doi.org/10.1590/S0100-204X2010000500006
- **6.** Falconer DS, Mackay TFC. Introduction to quantitative genetics. London: Longman Malaysia; 1996.
- Cruz CD, Regazzi AJ, Carneiro PCS. Modelos biómetricos aplicados ao melhoramento genético. 4th ed. Viçosa: Editora UFV; 2012.
- 8. Vencovsky R, Barriga P. Genética biométrica no fitomelhoramento. São Paulo: Sociedade Brasileira de Genética; 1992.
- Toebe M, Cargnelutti Filho A. Multicollinearity in path analysis of maize (Zea mays L.). J Cereal Sci. Academic Press; 2013; 57: 453–462. https://doi.org/10.1016/j.jcs.2013.01.014

- Wright S. Wright_Correlation and Causation.pdf. J Agric Res. 1921; 20: 557–585. https://doi.org/10. 2307/3966855
- Cruz CD, Carneiro PCS, Regazzi AJ. Modelos Biométricos Aplicados Ao Melhoramento Genético. 3rd ed. Viçosa: Editora UFV; 2014.
- Dewey DR, Lu KH. A Correlation and Path-Coefficient Analysis of Components of Crested Wheatgrass Seed Production1. Agron J. Madison, WI: American Society of Agronomy; 1959; 51: 515–518. https://doi.org/10.2134/agronj1959.00021962005100090002x
- Toebe M, Cargnelutti Filho A. Multivariate nonnormality and multicollinearity in path analysis in corn. Pesqui Agropecu Bras. Embrapa Informação Tecnológica; 2013; 48: 466–477. https://doi.org/10.1590/S0100-204X2013000500002
- Olivoto T, de Souza VQ, Nardino M, Carvalho IR, Ferrari M, de Pelegrin AJ, et al. Multicollinearity in Path Analysis: A Simple Method to Reduce Its Effects. Agron J. Madison, WI: The American Society of Agronomy, Inc.; 2017; 109: 131–142. https://doi.org/10.2134/agronj2016.04.0196
- Bizeti HS, Portela De Carvalho CG, Pinto De Souza JR, Destro D. Path analysis under multicollinearity in soybean. Brazilian Arch Biol Technol. Tecpar; 2004; 47: 669–676. https://doi.org/10.1590/S1516-8913200400500001
- Hoerl AE, Kennard RW. Ridge Regression: Biased Estimation for Nonorthogonal Problems. Technometrics. Taylor & Francis; 1970; 12: 55–67. https://doi.org/10.1080/00401706.1970.10488634
- Hoerl AE, Kennard RW. Ridge Regression: Applications to Nonorthogonal Problems. Technometrics. Taylor & Francis; 1970; 12: 69–82. https://doi.org/10.1080/00401706.1970.10488635
- Carvalho SP de. Métodos alternativos de estimação de coeficientes de trilha e índices de seleção, sob multicolínearidade [Internet]. UFV. 1995. http://www.locus.ufv.br/handle/123456789/11048
- Arshad M, Ali N, Ghafoor A. Character correlation and path coefficient in soybean Glycine Max (L.) Merril. Pak J Bot. 2006; 38: 121–130.
- Board JE, Kang MS, Harville BG. Path Analyses Identify Indirect Selection Criteria for Yield of Late-Planted Soybean. Crop Sci. Madison, WI: Crop Science Society of America; 1997; 37: 879–884. https://doi.org/10.2135/cropsci1997.0011183X003700030030x
- 21. Malik MFA, Ashraf M, Qureshi AS, Ghafoor A. Assessment of genetic variability, correlation and path analyses for yield and its components in soybean. Pakistan J Bot. 2007; 39: 405–413.
- 22. Nogueira APO, Sediyama T, de Sousa LB, Hamawaki OT, Cruz CD, Pereira DG, et al. Path Analysis and Correlations Among Traits in Soybean Grown in Two Dates Sowing. Biosci J. Federal University of Uberlandia; 2013; 28: 877–888.
- 23. Pandey JP, Torrie JH. Path Coefficient Analysis of Seed Yield Components in Soybeans (Glycine max (L.) Merr.)1. Crop Sci. Madison, WI: Crop Science Society of America; 1973; 13: 505–507. https://doi.org/10.2135/cropsci1973.0011183X001300050004x
- Teodoro PE, Ribeiro LP, Corrêa CC, Luz Júnior RAA, Zanucio A dos S, Capristo DP, et al. Path analysis in soybean genotypes as function of growth habit. Biosci J. 2015; 31: 794–799. https://doi.org/10.14393/BJ-v31n1a2015-26094
- MAPA. Ministério da Agricultura, Pecuária e Abastecimento. In: Registro Nacional de Cultivares— RNC2 [Internet]. 2019 [cited 27 Apr 2018]. http://sistemas.agricultura.gov.br/snpc/cultivarweb/ cultivares_registradas.php
- Alliprandini LF, Abatti C, Bertagnolli PF, Cavassim JE, Gabe HL, Kurek A, et al. Understanding soybean maturity groups in brazil: Environment, cultivar classifi cation, and stability. Crop Sci. Crop Science Society of America; 2009; 49: 801–808. https://doi.org/10.2135/cropsci2008.07.0390
- 27. de S Câmara GM. Soja: Tecnologia de produção. In: de S Câmara GM, editor. Soja: Tecnologias da produção. Piracicaba: ESALQ; 1998. pp. 256–277.
- 28. T, Silva FL, Borém A. Soja: do plantio à colheita. 1st ed. Viçosa: Editora UFV; 2015.
- 29. Fehr W, Caviness C. Stages of soybean development. Spec Rep. 1977.
- de Resende MDV. Matemática e estatística na análise de experimentos e no melhoramento genético [Internet]. Embrapa Florestas, Colombo. Colombo: Embrapa Floresta; 2007. p. 561.
- de Resende MDV. Software Selegen-REML / BLUP: a useful tool for plant breeding. Crop Breed Appl Biotechnol. 2016; 16: 330–339. http://dx.doi.org/10.1590/1984-70332016v16n4a49.
- 32. de Resende MDV, e Silva FF, Azevedo CF. Estatística Matemática, Biométrica e Computacional: Modelos Mistos, Multivariados, Categóricos e Generalizados (REML/BLUP), Inferência Bayesiana, Regressão Aleatória, Seleção Genômica, QTL-GWAS, Estatística Espacial e Temporal, Competição, Sobrevivência. 1st ed. Viçosa: Suprema; 2014.
- Cruz CD. Genes Software-extended and integrated with the R, Matlab and Selegen. Acta Sci Agron. 2016; 38: 547–552.

- **34.** R Core Team. R: A language and environment for statistical computing [Internet]. Vienna, Austria: R Foundation for Statistical Computing; 2019. https://www.r-project.org/
- de Resende MDV. Genética biométrica e estatística no melhoramento de plantas perenes. Embrapa Informação Tecnológica, Brasília. 1st ed. Brasília, DF: Embrapa Informação Tecnológica; 2002.
- **36.** Pimentel-Gomes F, Garcia CH. Estatística aplicada a experimentos agronômicos e florestais: exposição com exemplos e orientações pra uso de aplicativos. Piracicaba: Fealg; 2002.
- Bradshaw A. D. (1965). Evolutionary Significance of Phenotypic Plasticity in Plants. Advances in Genetics, 13(C), 115–155. https://doi.org/10.1016/S0065-2660(08)60048-6
- **38.** de Resende MDV, Duarte JB. Precisão e controle de qualidade em epxerimentos de avaliação de cultivares. Pesqui Agropecuária Trop. 2007; 37: 182–194.
- Machikowa T, Laosuwan P. Path coefficient analysis for yield of early maturing soybean. Songklanakarin J Sci Technol. 2011; 33: 365–368.
- Wilcox JR, Guodong Z. Relationships between Seed Yield and Seed Protein in Determinate and Indeterminate Soybean Populations. Crop Sci. Madison, WI: Crop Science Society of America; 1997; 37: 361–364. https://doi.org/10.2135/cropsci1997.0011183X003700020009x
- Hwang E-Y, Song Q, Jia G, Specht JE, Hyten DL, Costa J, et al. A genome-wide association study of seed protein and oil content in soybean. BMC Genomics. BioMed Central; 2014; 15: 1. https://doi.org/10.1186/1471-2164-15-1 PMID: 24382143
- 42. Chaudhary J, Patil GB, Sonah H, Deshmukh RK, Vuong TD, Valliyodan B, et al. Expanding Omics Resources for Improvement of Soybean Seed Composition Traits. Front Plant Sci. 2015; 6: 1021. https://doi.org/10.3389/fpls.2015.01021 PMID: 26635846
- Diers BW, Keim P, Fehr WR, Shoemaker RC. RFLP analysis of soybean seed protein and oil content. Theor Appl Genet. 1992; 83: 608–612. https://doi.org/10.1007/BF00226905 PMID: 24202678
- Lestari P, Van K, Lee J, Kang YJ, Lee S-H. Gene divergence of homeologous regions associated with a major seed protein content QTL in soybean. Front Plant Sci. Frontiers Media S.A.; 2013; 4: 176. https:// doi.org/10.3389/fpls.2013.00176 PMID: 23761803
- 45. Montgomey DC, Peck EA. Introduction to linear regression analysis. New York: J. Wiley; 1981.