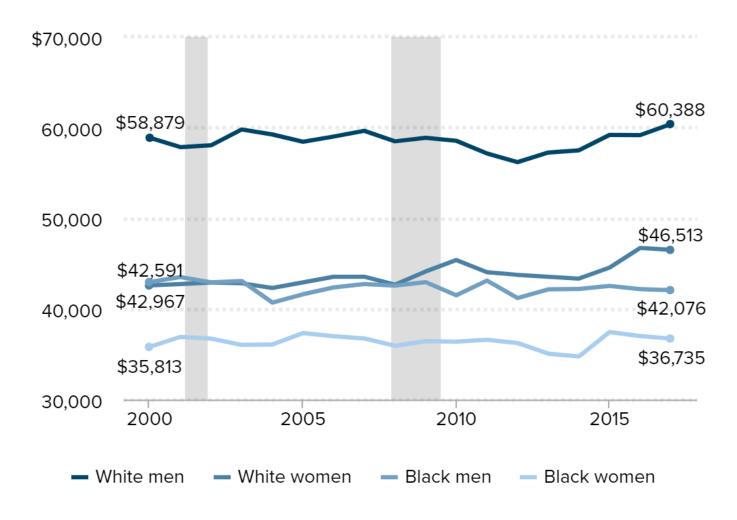
# Hypothesis Test for Two Proportions

# Real median earnings of full-time, full-year black workers and white workers, by gender, 2000–2017

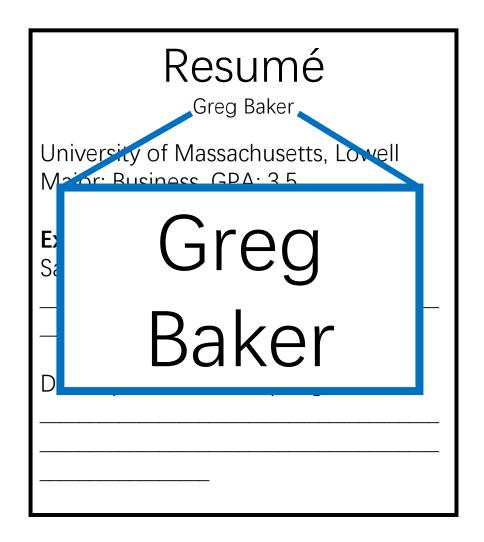


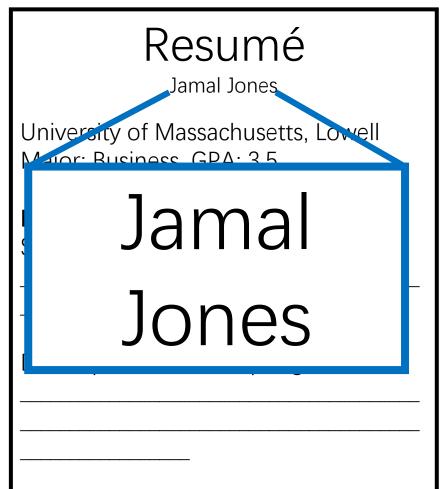
Hiring discrimination

Researchers wanted to test if hiring discrimination was a factor in labor markets

*Economic Policy Institute, 2018:* <a href="https://www.epi.org/blog/black-workers-have-made-no-progress-in-closing-earnings-gaps-with-white-men-since-2000/">https://www.epi.org/blog/black-workers-have-made-no-progress-in-closing-earnings-gaps-with-white-men-since-2000/</a>

# The Race/Resumé Study

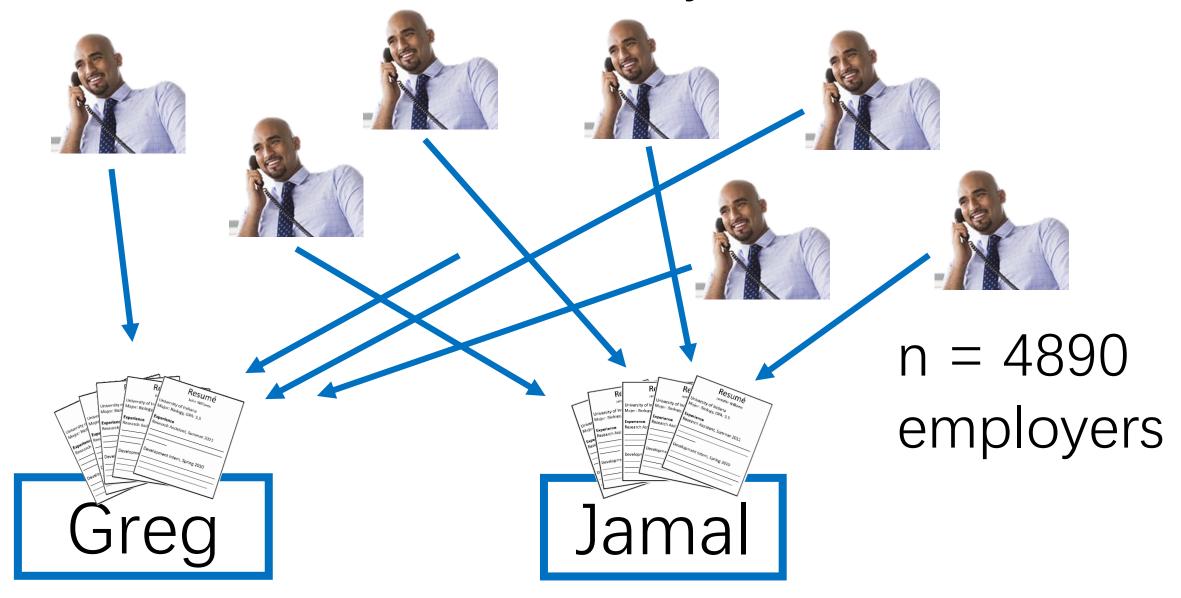




# The jobs

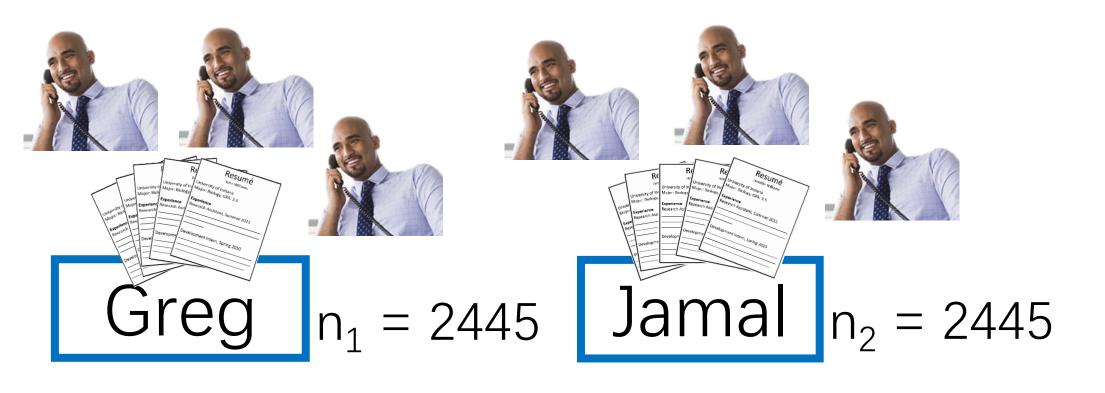
- Wide swath of jobs in the following industries: sales, administrative support, clerical services, and customer services
- Large range of **positions**, from "cashier work at retail establishments and clerical work in a mailroom to office and sales management positions."

# The Race/Resumé Study



# The Race/Resumé Study

# Measured which group got more callbacks from potential employers



### The results

#### **Treatment 1** Treatment 2

	Commonly- White Names	Commonly- Black Names	Total
Called back	246	164	410
Not called back	2199	2281	4480
Total	2445	2445	4890

Comparing the proportion who received callbacks from both treatments.

$$n_1 = 2445$$
  $n_2 = 2445$ 

$$n_2 = 2445$$

$$\hat{p}_1 = \frac{246}{2445} = 0.101$$
  $\hat{p}_2 = \frac{164}{2445} = 0.067$ 

# Two-Sample Situation

If there's hiring discrimination,  $\hat{p}_1 > \hat{p}_2$ 

### **Group 1: White**

 $\hat{p}_1$  = proportion of commonly-white name apps that got callback.

$$\hat{p}_1 = \frac{246}{2445} = 0.101$$



 $\hat{p}_2$  = proportion of commonly-black name apps that got callback.

$$\hat{p}_2 = \frac{164}{2445} = 0.067$$

Are these proportions different enough to show discrimination, or could this difference have been a result of chance alone?

# Hypotheses

 $H_0: p_1 = p_2$ 

 $H_A: p_1 > p_2$ 

There is no discrimination, so the callback rate is the same in both groups. You're seeing if there's evidence to reject this default claim.

There is discrimination, in which case the commonlywhite named applications received a higher rate of callbacks.

#### Where:

 $p_1$  is the proportion of all applicants with commonly-white names who'd receive callbacks when applying to jobs like the ones in this study.  $p_2$  is the proportion of all applicants with commonly-black names who'd receive callbacks when applying to jobs like the ones in this study.

# Setting up the Hypotheses

$$H_0: p_1 = p_2$$
  
 $H_A: p_1 > p_2$  OR  $H_0: p_1 - p_2 = 0$   
 $H_A: p_1 - p_2 > 0$ 

Preferred

#### Where:

 $p_1$  is the proportion of **all** applicants with commonly-**white** names who'd receive callbacks when applying to jobs like the ones in this study.  $p_2$  is the proportion of **all** applicants with commonly-**black** names who'd receive callbacks when applying to jobs like the ones in this study.

### Calculations

Since null assumes  $p_1 = p_2$ , so we can **combine** the proportion who got callbacks into one estimate:  $\hat{p}_c$ 

#### Under certain conditions:

$$\hat{p}_1 - \hat{p}_2 \sim N(\mu = 0, \sigma = \sqrt{\frac{\hat{p}_c (1 - \hat{p}_c)}{n_1} + \frac{\hat{p}_c (1 - \hat{p}_c)}{n_2}})$$

Centered at zero (since null assumes **no difference** between callback rates)

$$\widehat{p}_1 = \frac{246}{2445} = 0.101$$

$$\widehat{p}_2 = \frac{164}{2445} = 0.067$$

Combined proportion 
$$\widehat{p}_c = \frac{246+164}{2445+2445} = 0.084$$

### Calculations

Since null assumes  $p_1 = p_2$ , so we can **combine** the proportion who got callbacks into one estimate:  $\hat{p}_c$ 

#### Under certain conditions:

$$\hat{p}_1 - \hat{p}_2 \sim N(\mu = 0, \sigma = 0.0079)$$

Centered at zero (since null assumes **no difference** between callback rates)

#### The Data:

The actual difference in callback rates from the experiment  $\hat{p}_1 - \hat{p}_2 = \mathbf{0.034}$ 

How unlikely was our data?

Check the p-value! 
$$_{1} = \frac{246}{2445} = 0.101$$

$$\widehat{p}_2 = \frac{164}{2445} = 0.067$$

Combined proportion 
$$\hat{p}_c = \frac{246+164}{2445+2445} = 0.084$$

### Conclusion

Under my assumption that there is no difference in callback rates, the actually observed data (a 3.4% difference in callback rates among 4890 employers) is highly unlikely (p-value = 0.00001 < alpha level of 0.05). So, I reject my earlier assumption. There's convincing evidence that commonly-white named resumés receive a higher callback rate.

State: State the hypotheses, significance level, and define your parameters

$$H_0: p_1 - p_2 = 0$$
  
 $H_A: p_1 - p_2 > 0$   $\alpha = 0.05$ 

#### Where:

 $p_1$  is the proportion of **all** applicants with commonly-white names who'd receive callbacks when applying to jobs like the ones in this study.  $p_2$  is the proportion of **all** applicants with commonly-black names who'd receive callbacks when applying to jobs like the ones in this study.

Plan: Name your inference method and check conditions

We will conduct a two-sample z-test for  $p_1 - p_2$ , if all conditions are met.

#### **Conditions**

# Recall: Why we check conditions

$$\hat{p} \sim \text{Normal}\left(\mu = 0, \sigma = \sqrt{\frac{\hat{p}_c (1 - \hat{p}_c)}{n_1} + \frac{\hat{p}_c (1 - \hat{p}_c)}{n_2}}\right)$$

- 3) Large counts
  - → approx. normal shape

- 2) 10% condition
  - → calculable spread

- 1) Random condition→ unbiased center

Plan: Name your inference method and check conditions

We will conduct a two-sample z-test for  $p_1 - p_2$ , if all conditions are met.

#### **Conditions**

#### 1. Random:

Employers were randomly assigned either a commonly-white or commonly-black named resumé

#### 3. Large Counts:

$$n_1 \hat{p}_c \ge 10$$

$$n_1(1-\hat{p}_c) \ge 10$$

$$n_2 \hat{p}_c \ge 10$$

$$n_2(1-\hat{p}_c) \ge 10$$

Only have to do **10**% when sampling. However, this is an experiment. We don't have to check this condition!

Plan: Name your inference method and check conditions

We will conduct a two-sample z-test for  $p_1 - p_2$ , if all conditions are met.

#### **Conditions**

1. **Random:** Employers were randomly **assigned** either a commonly-white or commonly-black named resumé

#### 2. Large Counts:

$$n_1 \hat{p}_c \ge 10$$

$$(2445)(.084) \ge 10$$

$$n_1(1 - \hat{p}_c) \ge 10$$

$$(2445)(1 - .084) \ge 10$$

$$n_2 \hat{p}_c \ge 10$$
(2445)(.084)  $\ge 10$ 

$$n_2(1 - \hat{p}_c) \ge 10$$
 $(2445)(1 - .084) \ge 10$ 

Plan: Name your inference method and check conditions

We will conduct a two-sample z-test for  $p_1 - p_2$ , if all conditions are met.

#### **Conditions**

1. **Random:** Employers were randomly **assigned** either a commonly-white or commonly-black named resumé

#### 2. Large Counts:

$$n_1 \hat{p}_c \ge 10$$
 $205.4 \ge 10$ 
 $n_1(1 - \hat{p}_c) \ge 10$ 
 $2239.6 \ge 10$ 

$$n_2 \hat{p}_c \ge 10$$
 $205.4 \ge 10$ 
 $n_2(1 - \hat{p}_c) \ge 10$ 
 $2239.6 \ge 10$ 

<u>Do:</u> Perform calculations (if conditions met), report the test statistic and the p-value

$$z = 4.231$$
  
p-value = 0.00001

Conclude: Reject or fail to reject  $H_0$  and justify

$$H_0: p_1 - p_2 = 0$$
  
 $H_A: p_1 - p_2 > 0$   $\alpha = 0.05$   $z = 4.231$   
p-value = 0.00001

Conclusions template: Because our p-value (\_\_\_\_) is less/greater than our alpha level (\_\_\_\_), we reject/fail to reject  $H_0$ . We do/don't have convincing evidence that ( $H_A$  in context).

Conclude: Reject or fail to reject  $H_0$  and justify

$$H_0: p_1 - p_2 = 0$$
  
 $H_A: p_1 - p_2 > 0$   $\alpha = 0.05$   $z = 4.231$   
p-value = 0.00001

Because our p-value (0.00001) is **less** than our alpha level (0.05), we **reject**  $H_0$ . We **do** have convincing evidence that commonly-white name resumés get a higher callback rate for jobs similar to the ones in this study.

# Hypothesis Test for Two Means

# Standard deviations known

When two random samples are independently selected and when  $n_1$  and  $n_2$  are both large or the population distributions are (at least approximately) normal, the distribution of

$$z = \frac{\bar{x}_1 - \bar{x}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

is described (at least approximately) by the standard normal (z) distribution.

# Standard deviations unknown

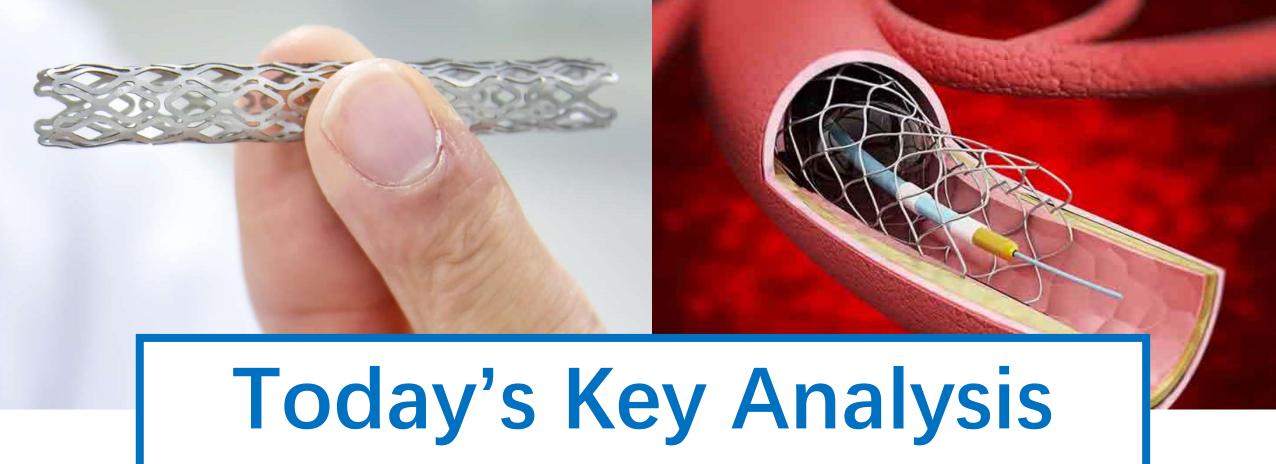
When two random samples are independently selected and when  $n_1$  and  $n_2$  are both large or when the population distributions are normal, the standardized variable

$$t = \frac{\bar{x}_1 - \bar{x}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$$

has approximately a t distribution with

df = 
$$\frac{(V_1 + V_2)^2}{\frac{V_1^2}{n_1 - 1} + \frac{V_2^2}{n_2 - 1}}$$
 where  $V_1 = \frac{s_1^2}{n_1}$  and  $V_2 = \frac{s_2^2}{n_2}$ 

The computed value of df should be truncated (rounded down) to obtain an integer value of df.



Did the stent treatment work?

### Measurement of Outcome

No Symptoms

Death

0 1 2 3 4 5 6

If stent works, it will move patients down this scale

# Measurement of Outcome

No Symptoms

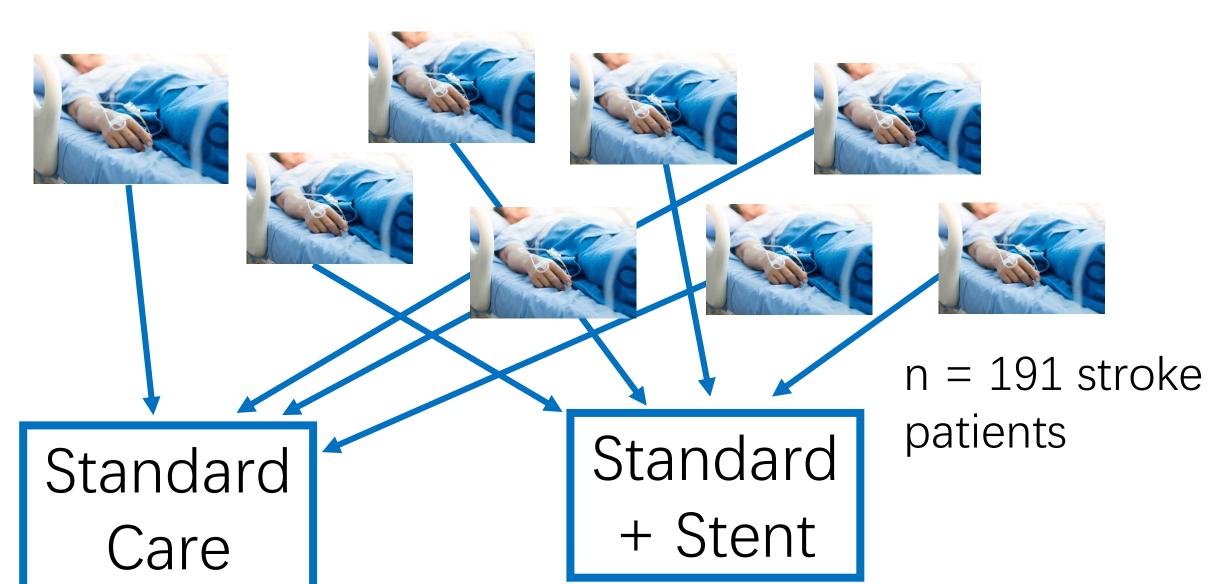
Death

0 1 2 3 4 5 6

From here on: referred to as "disability score"

# Our Study

#### \*Random Assignment\*



# Our Study





Standard Care

 $n_1 = 93$ 

Standard + Stent

 $n_2 = 98$ 

# Our Study

Measured which group had lower mean disability score





Standard Care

 $n_1 = 93$ 

Standard + Stent

 $n_2 = 98$ 

# **Topics**

- 1. Two-sample t-test for a difference of means
- 2. Four step process

# Setting up the Hypotheses

$$H_0$$
:  $\mu_S = \mu_C$ 

$$H_A: \mu_S < \mu_C$$

$$H_0: \mu_S - \mu_C = 0$$

$$H_A: \mu_S - \mu_C < 0$$

#### Where:

 $\mu_{\mathcal{S}}$  is the mean disability score of all patients who'd receive stents.

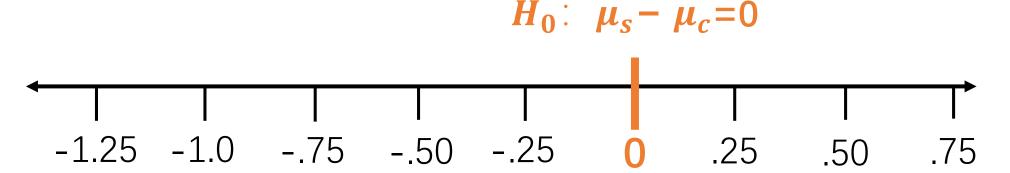
 $\mu_{c}$  is the mean disability score of all patients who'd receive current standard of care.

# The results

	Stent	Control
Mean Disability	2.26	3.23
Stdev. Disability	1.78	1.78
n	98	93

### The Calculations

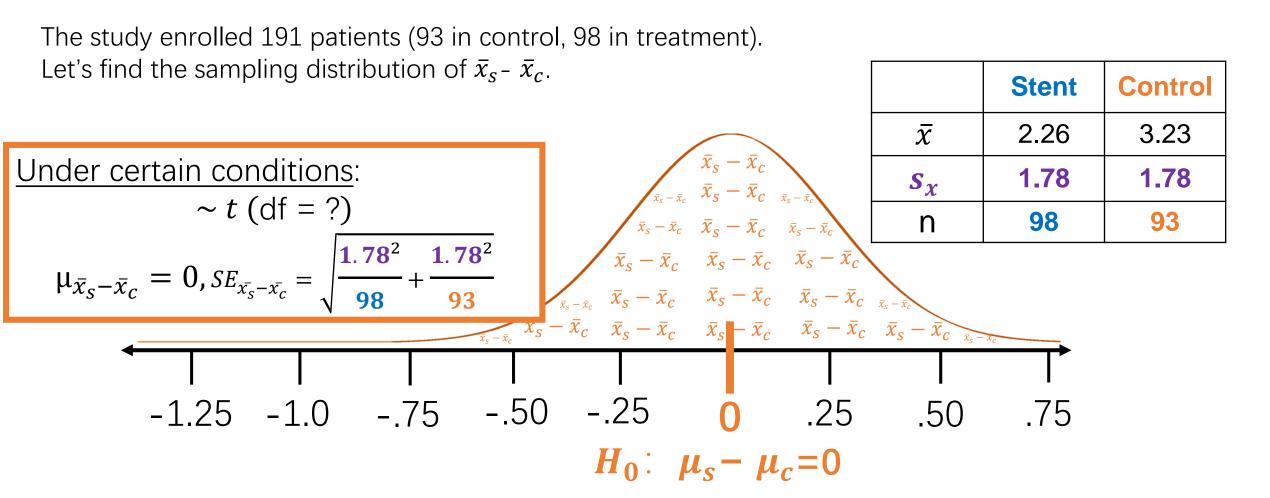
1. Assume the null is true



We assume there is no difference in average disability scores among all patients who would receive stents or the current standard of care:  $\mu_s$ -  $\mu_c$ =0

# The Calculations

2. If the null is in fact true, how likely is the data that you've gathered?



The Data: The actual difference in mean disability score from the experiment  $\bar{x}_s - \bar{x}_c = -0.97$ 

	Stent	Control
$\bar{x}$	2.26	3.23
$S_{\chi}$	1.78	1.78
n	98	93

How unlikely was our data?

Probability of getting this result or

 $\bar{x}_{S} - \bar{x}_{C} = -0.97$ more extreme? p-value=0.0001

Is there convincing statistical evidence that the stent lowered the average disability from stroke?

$$H_0$$
:  $\mu_S - \mu_C = 0$   $\mu_S = 2.26$ ,  $\mu_C = 3.23$   $H_A$ :  $\mu_S - \mu_C < 0$   $\bar{x}_S - \bar{x}_C = -0.97$ 

Is there convincing statistical evidence that the stent lowered the average disability from stroke?

$$H_0: \mu_s - \mu_c = 0$$

$$H_A: \mu_s - \mu_c < 0$$

$$\mu_s = 2.26, \, \mu_c = 3.23$$
 $\bar{x}_s - \bar{x}_c = -0.97$ 

1. Assume null is true

Is there **convincing statistical evidence** that the stent lowered the average disability from stroke?

$$H_0: \mu_s - \mu_c = 0$$
  
 $H_A: \mu_s - \mu_c < 0$ 

1. Assume null is true

$$\mu_s = 2.26, \, \mu_c = 3.23$$
 $\overline{x}_s - \overline{x}_c = -0.97$ 

2. How unlikely was our sampled data? p-value: 0.0001

Is there **convincing statistical evidence** that the stent lowered the average disability from stroke?

$$H_0$$
:  $\mu_s - \mu_c = 0$ 

$$H_A: \mu_S - \mu_C < 0$$

$$\mu_s = 2.26, \, \mu_c = 3.23$$
 $\overline{x}_s - \overline{x}_c = -0.97$ 

Our sample data was very unlikely, if we assume the null is true.

2. How unlikely was our sampled data?

p-value: 0.0001

## Conclude

Note: more concise conclusion template provided later

#### 3. Draw a conclusion:

Under my assumption that stents provide no added benefit, the actually observed data (0.97 point decline in disability in stent group) is highly unlikely (p-value = 0.0001). So, I reject my earlier assumption. There's convincing evidence that the stent lowers disability from stroke.

$$H_0: \mu_s - \mu_c = 0$$
 $H_A: \mu_s - \mu_c < 0$ 

## **Topics**

- 1. Two-sample t-test for a difference of means
- 2. Four step process

## The Four Steps for Inference

A suggested way to **organize** your work so that you get full credit on FRQ's!

State: hypotheses, significance level, and define your parameters

Plan: Name your inference method and check conditions

Do: Perform calculations (if conditions met), report the test statistic and the p-value

**Conclude:** Reject or fail to reject H<sub>0</sub> and justify

State: State the hypotheses, significance level, and define your parameters

$$H_0: \mu_S - \mu_C = 0$$
  $\alpha = 0.05$   
 $H_A: \mu_S - \mu_C < 0$ 

#### Where:

 $\mu_s$  is the mean disability score of all patients who'd receive stents.  $\mu_c$  is the mean disability score of all patients who'd receive current standard of care.

Plan: Name your inference method and check conditions

We will conduct a **two-sample t-test** for  $\mu_s - \mu_c$ .

$$\frac{(\overline{X_S} - \overline{X_C}) - 0}{\sqrt{\frac{s_S^2}{n_S} + \frac{s_C^2}{n_C}}} \sim t(df)$$
 if the following conditions are met. if all conditions are met.

df=?

<u>Plan:</u> Name your inference method and check conditions

We will conduct a **two-sample t-test** for  $\mu_s - \mu_c$ , if all conditions are met. 2. **Normal/Large Samples:** 

#### Conditions

$$n_s \ge 30$$

$$n_c \ge 30$$

1. Random: Patients were randomly assigned either to receive current standard care or stent treatment.

$$98 \ge 30$$

$$93 \ge 30$$

<u>Do:</u> Perform calculations (if conditions met), report the test statistic and the p-value

```
t = (formula) = -3.76
p-value = (formula) = 0.0001
```

Conclude: Reject or fail to reject  $H_0$  and justify

Conclusions template: Because our p-value (\_\_\_\_) is less/greater than our alpha level (\_\_\_\_), we reject/fail to reject  $H_0$ . We do/don't have convincing evidence that ( $H_A$  in context).

Conclude: Reject or fail to reject  $H_0$  and justify

Because our p-value (0.0001) is **less** than our alpha level (0.05), we **reject**  $H_0$ . We **do** have convincing evidence that stents lower disability from stroke.

*Procedure:* A two-sample *t*-interval for  $\mu_{NRW} - \mu_{RW}$ , the difference in population means of BPA body concentrations in nonretail workers and retail workers

*Checks:* It is given that these are random samples. It is reasonable to assume the samples are independent, both samples sizes ( $528 \ge 30$  and  $197 \ge 30$ ) are large enough so that the CLT applies, and we assume the sample sizes are less than 10% of the populations.

*Mechanics:* Calculator software gives (-0.9521, -0.7479) with df = 332.3.

Conclusion in context: We are 99% confident that the difference in true means of BPA body concentrations in all nonretail and retail workers (nonretail mean minus retail mean) is between -0.75 and  $-0.95 \mu g/L$ .

Because 0 is not in the interval of plausible values for the difference of population means and the entire interval is negative, the interval does support the belief that retail workers carry higher amounts of BPA in their bodies than nonretail workers.

*Parameters:* Let  $\mu_{NFL}$  represent the mean attendance of the population of NFL games. Let  $\mu_{10}$  represent the mean attendance of the population of Big Ten football games.

Hypotheses: 
$$H_0$$
:  $\mu_{NFL} - \mu_{10} = 0$  (or  $\mu_{NFL} = \mu_{10}$ ) and  $H_a$ :  $\mu_{NFL} - \mu_{10} < 0$  (or  $\mu_{NFL} < \mu_{10}$ )

*Procedure:* A two-sample *t*-test for means.

*Checks:* Independent random samples (given); both samples sizes,  $n_{NFL} = 35 \ge 30$  and  $n_{10} = 30 \ge 30$ , are large enough for the CLT to apply; and the sample sizes, 35 and 30, are less than 10% of all NFL and Big Ten football games, respectively.

*Mechanics:* The population SDs are unknown, so we use a t-distribution. Calculator software (such as 2-SampTTest on the TI-84 or 2-Sample tTest on the Casio Prizm) gives t = -0.8301, df = 49.3, and P = 0.2052.

Conclusion in context with linkage to the *P*-value: With a *P*-value this large, 0.2052 > 0.05, there is not sufficient evidence to reject  $H_0$ ; that is, there is not sufficient evidence that the true mean attendance at Big Ten Conference football games is greater than that at NFL games.