

When can a Bayesian divide and conquer?

General optimality guarantees for distributed nonparametric Bayesian inference

Lasse Vuursteen



This talk is based on joint works with Botond Szabó and Harry van Zanten.



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- Popular methods like Gaussian process regression typically have computational complexity $O(n^3)$ and memory complexity $O(n^2)$ even in conjugate models.

Divide and conquer

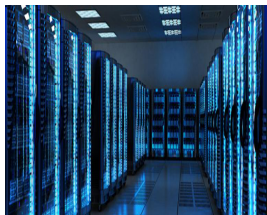
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Divide and conquer - formally

Formally:

- 1 Observations X_1, \dots, X_n are divided over n machines, machine $j = 1, \dots, n$ receiving n/m observations $X^{(j)} = (X_1^{(j)}, \dots, X_{n/m}^{(j)})$

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The distribution of the average of the draws of the local posteriors give us a generalized “posterior”: the m -times convolution of the local posteriors rescaled with m :

$$\tilde{\Pi}(A | X^{(1:m)}) := \bigotimes_{j=1}^m \Pi(m \cdot | X^{(j)})(A).$$

We perform inference on θ using $\tilde{\Pi}(\cdot | X^{(1:m)})$.

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- “*Consensus Monte Carlo*” in Steven L Scott et al.(2016), “Bayes and big data: The consensus Monte Carlo algorithm”

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$$O(n^3) \text{ becomes } \begin{cases} O\left(\frac{n^3}{m^3}\right) & \text{if ran in parallel,} \\ O\left(\frac{n^3}{m^2}\right) & \text{otherwise.} \end{cases}$$

So we can even obtain linear or constant complexity in n if $m \simeq n$.

- It is super simple to implement (even parallelly): no need to rewrite code for existing Bayesian methods (it is “plug-and-play”).

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Short answer: *(almost) always, but it is not always good. Inference might become highly suboptimal.*

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In this work, we provide general sufficiency conditions that guarantee optimal frequentist convergence rates for the divide and conquer strategy.

- We demonstrate that the guarantees hold for a variety of nonparametric models and priors.

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Short answer: *(almost) always, but it is not always good. Inference might become highly suboptimal.*

In this work, we provide general sufficiency conditions that guarantee optimal frequentist convergence rates for the divide and conquer strategy.

- We demonstrate that the guarantees hold for a variety of nonparametric models and priors.
- Similar guarantees are provided for methods that are based on general loss functions instead of likelihoods for G-Bayes and M-estimation.

Setting and notation

In this talk, I will assume iid $X_i \sim P_{\theta^\circ}$ for some $\theta^\circ \in \Theta$ and that the model $\{P_\theta : \theta \in \Theta\}$ is dominated.

As shorthand notation, let

$$\ell(\theta_1, \theta_2; X_i) := \log \left(\frac{p_{\theta_1}}{p_{\theta_2}}(X_i) \right)$$

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$$\bar{\theta} = \frac{1}{m} \sum_{j=1}^m \theta^{(j)}.$$

Recall:

$$\bar{\theta} \sim \tilde{\Pi}(\cdot | \mathbf{X}^{(1:m)}) := \bigotimes_{j=1}^m \Pi(m \cdot | \mathbf{X}^{(j)}).$$

Divide and conquer convergence rate

Definition: The posterior distribution $\Pi(\cdot | X^{(n)})$ is said to *contract at rate* $\varepsilon_n \rightarrow 0$ in d at θ° if

$$\Pi\left(\theta \in \Theta : d(\theta, \theta^\circ) \leq M_n \varepsilon_n | X^{(n)}\right) \rightarrow 1 \text{ in } P_{\theta^\circ}\text{-probability,}$$

for any $M_n \rightarrow \infty$ as $n \rightarrow \infty$.

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Suppose the full data posterior contracts at rate ε_n , does

$$\tilde{\Pi}\left(\bar{\theta} : d(\bar{\theta}, \theta^\circ) \leq M_n \varepsilon_n | X^{(1:m)}\right) \xrightarrow{P_{\theta^\circ}} 1$$

also for all $M_n \rightarrow \infty$?

Divide and conquer convergence rate Theorem

Theorem ((Szabó, V., van Zanten))

Let ε_n be a sequence of positive numbers for which the **likelihood geometry condition** and **testing condition** hold. Assume furthermore that $\Pi^{(j)}$'s are such that the **prior mass** and **entropy conditions** are satisfied.

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Let ε_n be a sequence of positive numbers for which the **likelihood geometry condition** and **testing condition** hold. Assume furthermore that $\Pi^{(j)}$'s are such that the **prior mass** and **entropy conditions** are satisfied.

Then, ε_n is a contraction rate for the generalized posterior. That is, it holds that

$$\tilde{\Pi} \left(\bar{\theta} : d(\bar{\theta}, \theta^\circ) \leq M_n \varepsilon_n \mid X^{(1:m)} \right) \xrightarrow{P_{\theta^\circ}} 1 \quad (1)$$

for any $M_n \rightarrow \infty$.

Likelihood geometry condition

The likelihood geometry condition is satisfied if

$$\frac{1}{m} \sum_{j=1}^m \mathbb{E}_{\theta^\circ} \ell \left(\theta^{(j)}, \theta^\circ; \mathbf{x}_i^{(j)} \right) \leq c_\ell \mathbb{E}_{\theta^\circ} \ell \left(m^{-1} \sum_{j=1}^m \theta^{(j)}, \theta^\circ; \mathbf{x}_i^{(j)} \right) \quad (2)$$

whenever $-\mathbb{E}_{\theta^\circ} \ell \left(\bar{\theta}, \theta^\circ; \mathbf{x}_i^{(j)} \right) \leq C_\ell \epsilon_n^2$ for some constants $c_\ell, C_\ell > 0$.

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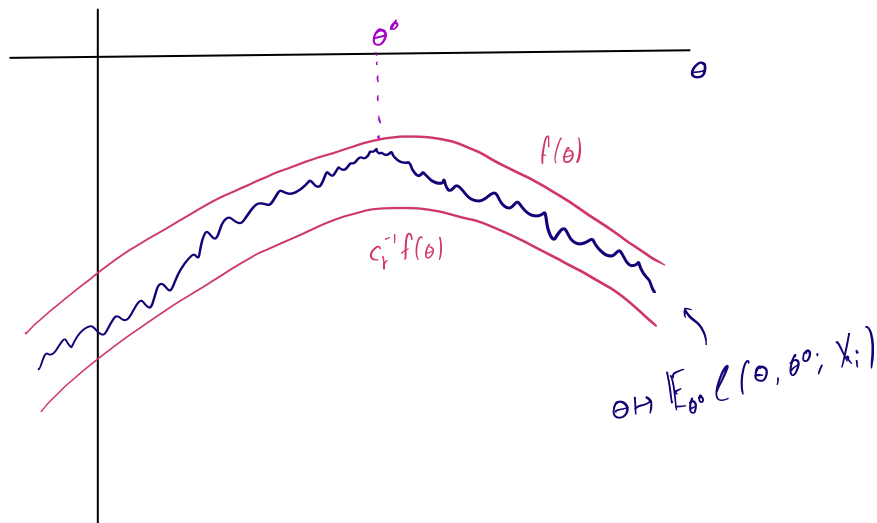
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Likelihood geometry condition – a cartoon



Divide and conquer Ghoshal & van der Vaart conditions

The following **prior mass conditions** and **entropy** and **testing conditions** respectively:

$$\bigotimes_{j=1}^m \Pi^{(j)} \left(\theta^{(1:m)} \in \Theta^m : -\mathbb{E}_{\theta^\circ} \ell \left(\bar{\theta}, \theta^\circ; \mathbf{X}_i^{(j)} \right) \leq \varepsilon_n^2 \right) \geq e^{-C_{KL} n \varepsilon_n^2}, \quad (3)$$

and that there exists a $C_0 > 0$ such that for all $C \geq C_0$ there exists $\mathcal{P}_n \subset \mathcal{P}$ such that

$$\bigotimes_{j=1}^m \Pi^{(j)} (\theta^{(1:m)} \in \Theta^m : \bar{\theta} \in \mathcal{P}_n^c) \leq e^{-C n \varepsilon_n^2} \quad (4)$$

$$\log N(\varepsilon_n, d, \mathcal{P}_n) \leq C n \varepsilon_n^2. \quad (5)$$

Furthermore, writing $\bar{\theta} = m^{-1} \sum_{j=1}^m \theta^{(j)}$, assume that there exists a test T such that

$$P_{\theta^\circ} T \xrightarrow{n \rightarrow \infty} 0 \quad \text{and} \quad \bigotimes_{j=1}^m P_{\theta^{(j)}} (1 - T) \leq e^{-C_{\text{test}} M_n n \varepsilon_n^2} \quad (6)$$

whenever $\bar{\theta} \in \mathcal{P}_n$ and $d(\bar{\theta}, \theta^\circ) \geq M_n \varepsilon_n$.

Divide and conquer - Gaussian processes

Let $\Theta = (\Theta, \|\cdot\|)$ be a separable Banach space and let W be a centered Gaussian process on it.

Suppose that $-\mathbb{E}_{\theta} \ell(\theta', \theta; X_i) \leq \|\theta - \theta'\|^2$ and that W 's *concentration function* satisfies¹

$$\phi_{\theta^\circ}^W(\varepsilon_n) = \inf_{h \in \mathbb{H}: \|\theta^\circ - h\| < \varepsilon_n} \|h\|_{\mathbb{H}}^2 - \log \Pr(\|W\| \leq \varepsilon_n) \lesssim n\varepsilon_n^2.$$

¹If the full data posterior corresponding to W is to contract at rate ε_n , this should be the case; see e.g. [Ismaël Castillo\(2008\)](#), “Lower bounds for posterior rates with Gaussian process priors” and [A. W. van der Vaart and J. H. van Zanten\(June 2008\)](#), “Rates of contraction of posterior distributions based on Gaussian process priors”.

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Take each of the $\Pi^{(j)}$ equal to the distribution of $\sqrt{m}W$. Then, the **prior mass** and **entropy** conditions are satisfied.

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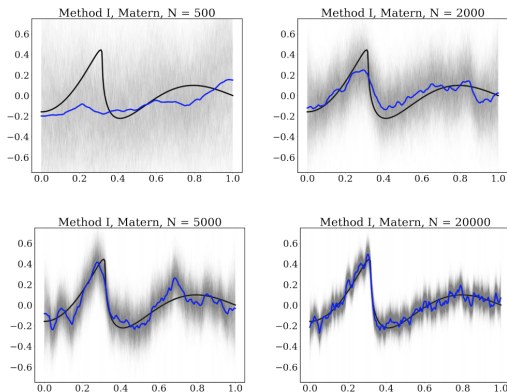
Take each of the $\Pi^{(j)}$ equal to the distribution of $\sqrt{m}W$. Then, the **prior mass** and **entropy** conditions are satisfied.

The prior blow-up of \sqrt{m} turns out to be strictly necessary. We have a theorem that stating that taking $\Pi^{(j)}$ equal to the distribution of W results in a convergence rate that corresponds to throwing away $\frac{n(m-1)}{m}$ observations in typical settings.

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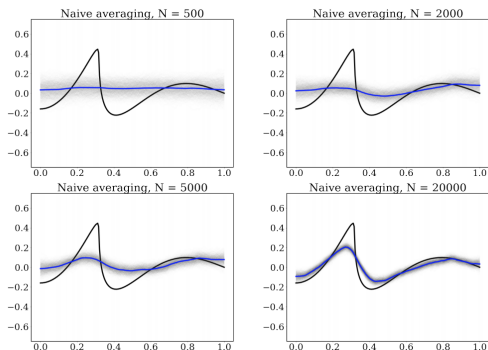
W centered Gaussian with squared exponential kernel. Local prior is with \sqrt{m} blow-up.



Picture is courtesy of [Amine Hadji](#), [Tammo Hesslink](#), and [Botond Szabó\(2022\)](#), “Optimal recovery and uncertainty quantification for distributed Gaussian process regression”.

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Divide and conquer convergence rate Theorem

Theorem ((Szabó, V., van Zanten))

Let ε_n be a sequence of nonnegative numbers and let W be a Gaussian process such that $\phi_{\theta^\circ}^W(\varepsilon_n) \lesssim n\varepsilon_n^2$ holds.

For $j = 1, \dots, m$, let the local priors $\Pi^{(j)}$ be equal to the distribution of $\sqrt{m}W$. Then,

$$\tilde{\Pi}(\bar{\theta} : \|\bar{\theta} - \theta^\circ\| \leq M_n \varepsilon_n | \mathbf{X}^{(1:m)}) \xrightarrow{P_{\theta^\circ}} 1 \quad (7)$$

as long as the **testing condition** and **likelihood geometry condition** are satisfied.

Examples: for which models does all of this hold?

- **Density estimation:** Suppose we observe iid \mathbf{X}_i 's in $[0, 1]^d$ from some Hölder smooth density p° . Take as local priors

$$p_{\sqrt{m}W}(\mathbf{X}_i) = \frac{e^{\sqrt{m}W(\mathbf{X}_i)}}{\int e^{\sqrt{m}W(x)} d\mu(x)}$$

where W is a zero mean Gaussian process in $\ell_\infty([0, 1]^d)$ satisfying (??) with ε_n the optimal contraction rate and suppose that p° lies in the support of W . Then, the generalized posterior contracts at rate ε_n also.

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- **Regression:** A similar statement holds for iid observations $Y_i | P_\theta$ satisfying

$$Y_i = \theta(i/n) + Z_i$$

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with $Z_i \sim N(0, \sigma^2)$ and θ belonging to e.g. a Sobolev space.

- **Classification:** Consider iid observations $(Y_1, \mathbf{X}_1), \dots, (Y_n, \mathbf{X}_n)$ satisfying

$$\Psi(\theta(x)) = \Pr(Y_i = 1 | \mathbf{X}_i = x),$$

conditionally on P_θ for some link function Ψ and some other assumptions, a divide and conquer approach (using e.g. Gaussian process priors) attains the optimal contraction rate for θ° .

Thank you for listening!

References:

- Botond Szabó and Harry Van Zanten(2019), “An asymptotic analysis of distributed nonparametric methods.”
- Amine Hadji, Tammo Hesslink, and Botond Szabó(2022), “Optimal recovery and uncertainty quantification for distributed Gaussian process regression”