

# **Regularised B-splines projected Gaussian Process priors to estimate time-trends in age-specific COVID-19 deaths**

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ISBA World Meeting  
29.06.2022

- **The regularised B-splines projected Gaussian Process prior**
- Theoretical properties: **smoothness** and **computational efficiency**
- **Benchmark** against standard GP, standard B-splines, Bayesian P-splines, ...
- **Applications**
  1. estimating time-trends in age-specific COVID-19 deaths
  2. estimating sex and age-specific flows of HIV-1 transmission
  3. estimating sex and age-specific contact patterns

**Regularised B-splines projected  
Gaussian Process priors**

## Problem set-up

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Let our support be a 2D grid  $\mathcal{G}$  with points  $(x_1, \dots, x_n) = \mathcal{X} \subset \mathbb{R}$  on the first dimension and points  $(y_1, \dots, y_m) = \mathcal{Y} \subset \mathbb{R}$  on the second dimension. The grid  $\mathcal{G} = \mathcal{X} \times \mathcal{Y}$  is the set of  $N = n \times m$  coordinates points.

A multivariate random variable  $\mathbf{z} = \{z(x, y)\}_{(x, y) \in \mathcal{G}}$  is defined on the 2D grid. Its likelihood takes the form,

$$z(x, y) \mid f(x, y), \theta \sim F(f(x, y), \theta),$$

where  $F$  is a distribution parametrised by a 2D random function  $f(x, y)$  and some parameters  $\theta$ .

# Two-dimensional Gaussian Process

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A natural starting point for modelling the random function  $\mathbf{f}$  is a two-dimension Gaussian Process,

$$\mathbf{f} \mid \phi \sim \mathcal{GP}(0, \mathbf{K}),$$

where  $\mathbf{K} = \mathbf{K}(\phi)$  with entries defined by some kernel function  $k(.,.)$ .

Because our inputs are on a Cartesian product grid, we can decompose the covariance matrix (Saatçi, 2011, Gonen et al., 2011),

$$\mathbf{K} = \mathbf{K}_2 \otimes \mathbf{K}_1.$$

**Advantage:** Inherits all the properties of Gaussian Processes

**Disadvantage:** The time complexity of this approach scales to  $\mathcal{O}(2N^{3/2})$

# Two-dimensional B-splines surface

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B-splines basis functions are constructed from polynomial pieces that are joined at certain values over the input space, called knots. We use cubic B-splines.

$\mathbf{f}$  can be modelled with a tensor product of B-splines given by,

$$f(x, y) = \sum_{i=1}^I \sum_{j=1}^J \beta_{i,j} B_i^1(x) B_j^2(y).$$

Where  $\mathbf{B}^1$  of size  $I \times n$  and  $\mathbf{B}^2$  of size  $J \times m$  are matrices of B-splines basis functions defined over  $\mathcal{X}$  and over  $\mathcal{Y}$ .

**Advantage:** Smooth (piecewise infinitely differentiable between the knots, and of continuity  $\mathcal{C}^2$  on the knots).

**Disadvantage:** Need to choose the number of knots. Too few knots will not capture complex signals and too many will overfit the surface.

# Regularised B-splines

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**Idea:** Use many knots and apply a regularisation on the B-splines parameters to impose smoothness.

- Frequentist: smoothing splines (O'Sullivan, 1986,9) and P-splines (Eilers and Marx, 1996; Eilers et al., 2006).
  - Penalty applied on the second derivative of the fitted curve or on finite differences of adjacent B-splines coefficients.
- Bayesian: Bayesian P-splines (Land and Brezger (2004, 2006).
  - Conditional Autoregressive prior on the B-splines parameters.

# B-splines projected two-dimensional Gaussian Process prior

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Regularised two-dimensional B-splines view:

$$f(x, y) = \sum_{i=1}^I \sum_{j=1}^J \beta_{i,j} B_i^1(x) B_j^2(y).$$
$$\boldsymbol{\beta} \mid \boldsymbol{\phi} \sim \mathcal{GP}(0, \mathbf{K}_{\boldsymbol{\beta}}).$$

where  $\mathbf{K}_{\boldsymbol{\beta}} = \mathbf{K}_{\boldsymbol{\beta}}(\boldsymbol{\phi})$  with entries defined by a kernel function  $k_{\boldsymbol{\beta}}$

Low-rank two-dimensional Gaussian Process view:

$$\mathbf{f} \mid \boldsymbol{\phi} \sim \mathcal{GP}\left(0, (\mathbf{B}^2 \otimes \mathbf{B}^1)^T \mathbf{K}_{\boldsymbol{\beta}} (\mathbf{B}^2 \otimes \mathbf{B}^1)\right).$$



# B-splines projected two-dimensional Gaussian Process prior

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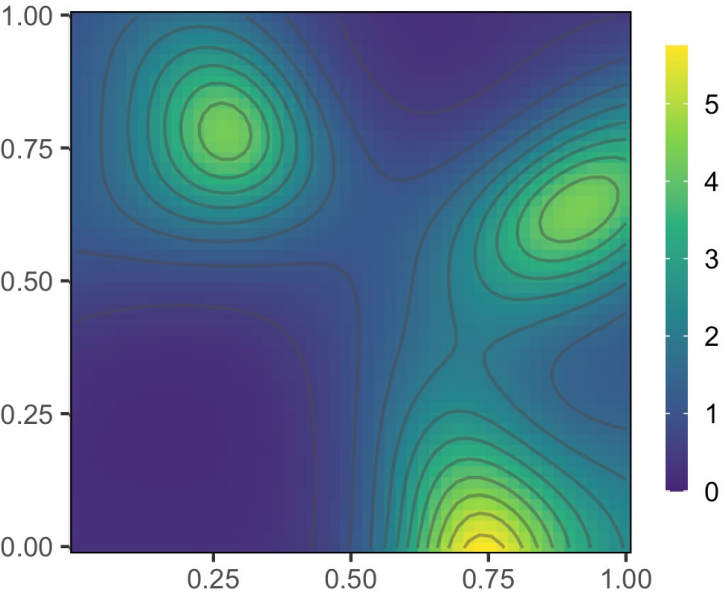
## Advantages:

1. The 2D random function  $\mathbf{f}$  inherits smoothness properties from the B-splines. We show that the kernel function obtained by projecting a base kernel function with cubic B-splines is  $C^2$ .
2. The B-splines parameters  $\boldsymbol{\beta}$  are regularised as the negative determinant of the covariance matrix in the log-likelihood plays the same role of a complexity penalty.
3. The time complexity compared to a full rank two-dimensional Gaussian Process reduces from  $\mathcal{O}(2N^{3/2})$  to  $\mathcal{O}(2(I \times J)^{3/2})$

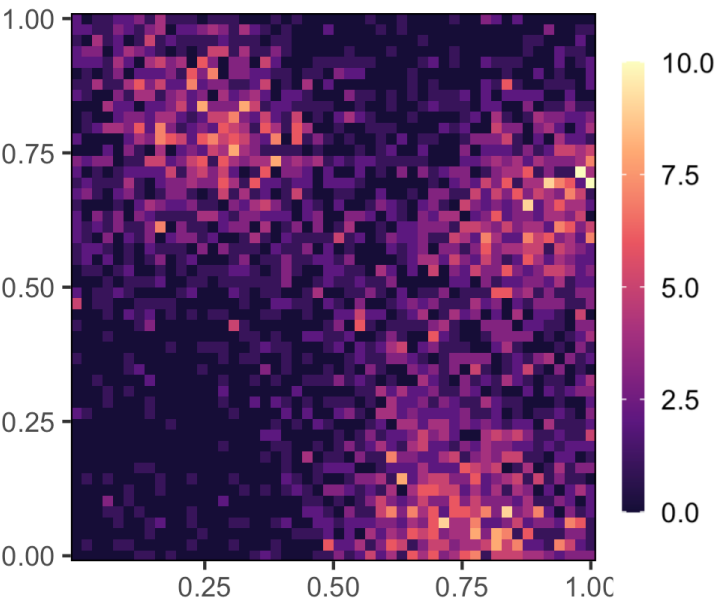
Simulation analysis

# Simulated data

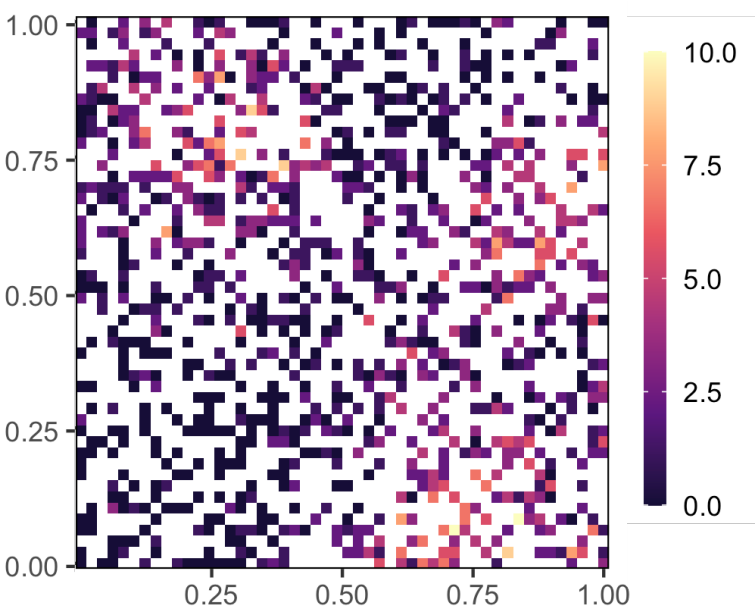
Simulated  
mean surface



Simulated  
count observations



Simulated  
count observations  
included in the training set

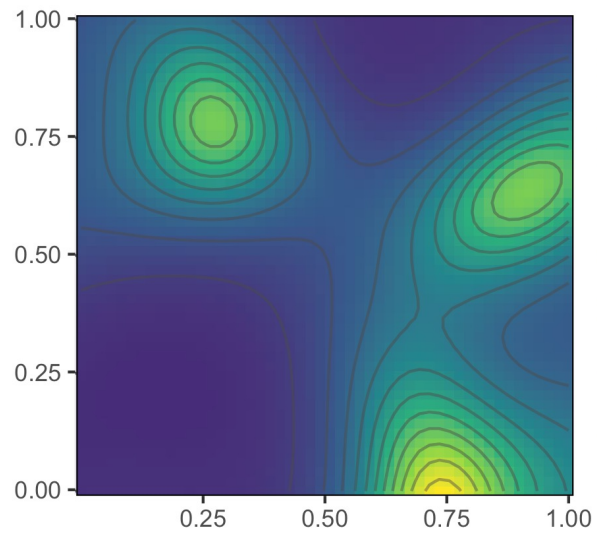


# Results

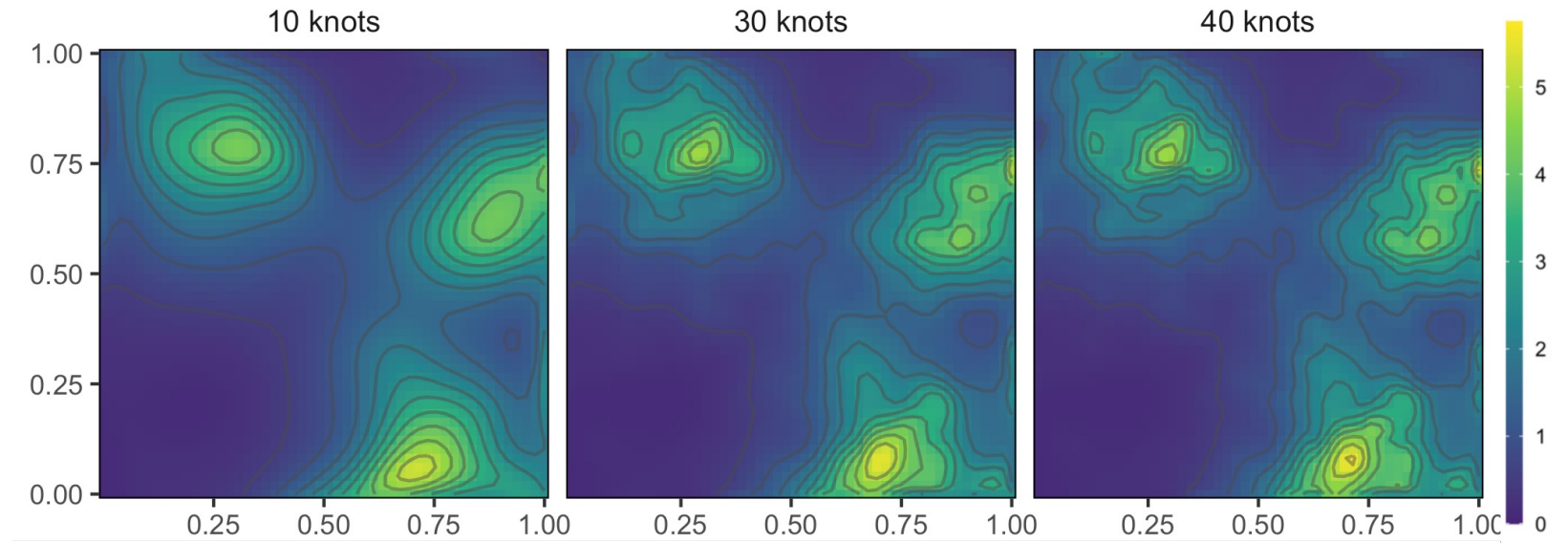
Method	MSE	Runtime in minutes (%reduction runtime)
Standard 2D GP	0.04	47 (0.00%)
Bayesian P-splines		
10 knots	0.09	7 (85.66%)
30 knots	0.13	5 (90.14%)
40 knots	0.13	5 (90.29%)
Regularised B-splines projected 2D GP		
10 knots	0.06	6 (87.50%)
30 knots	0.05	6 (87.07%)
40 knots	0.05	10 (78.28%)

# Results

Simulated  
mean surface

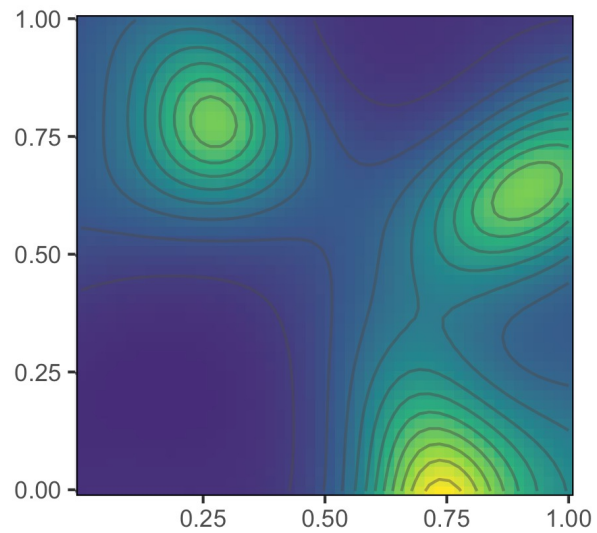


Bayesian P-splines (Lang and Brazer, 2004)

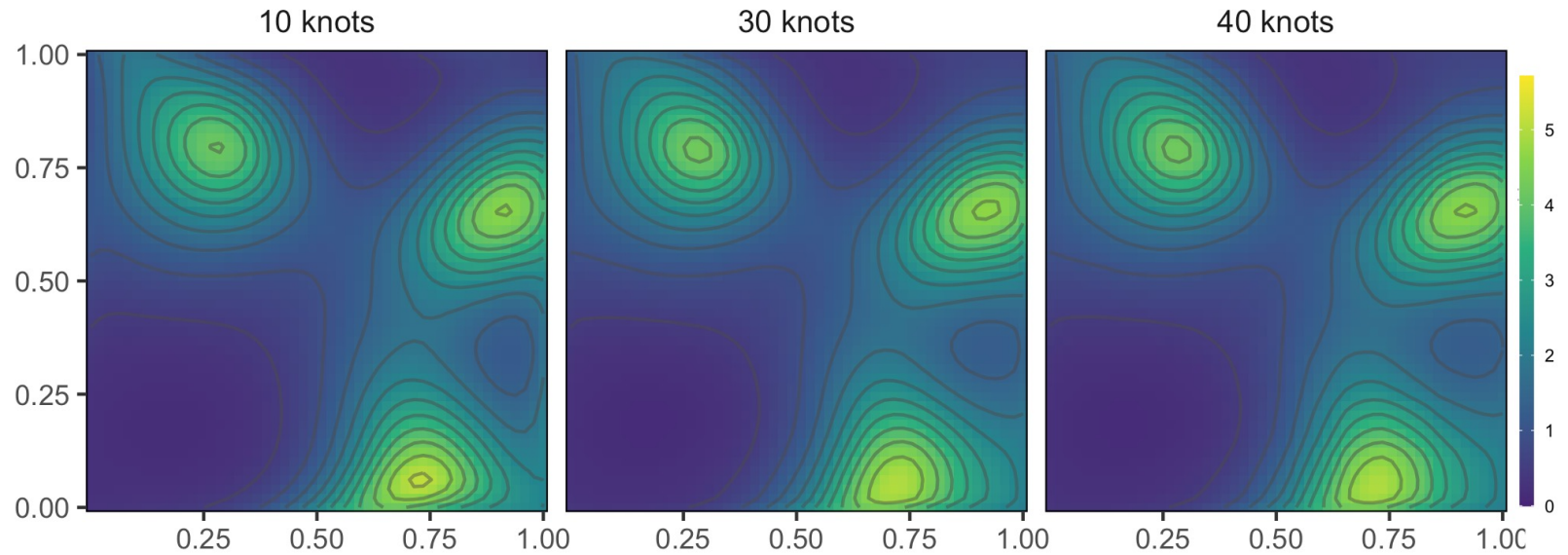


# Results

Simulated  
mean surface



B-splines projected two-dimensional  
Gaussian Process prior



Case study

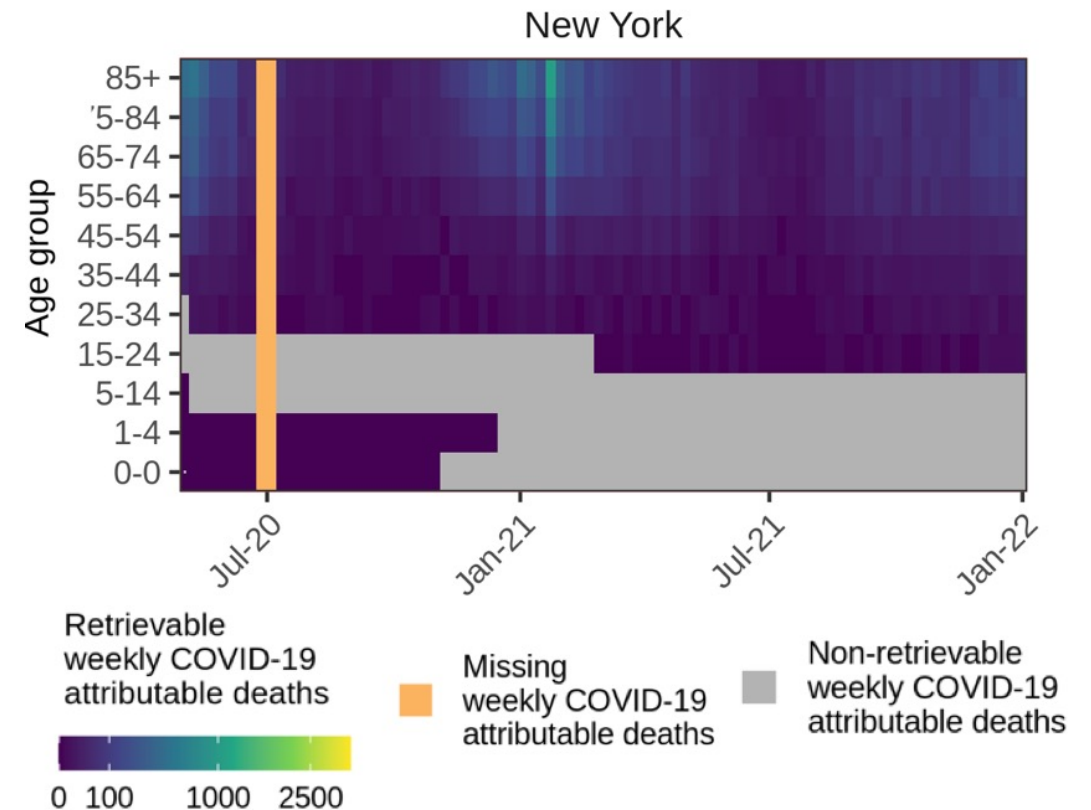
**Time-trends in  
age-specific COVID-19 deaths**

# Problem and data

- Age and state-specific COVID-19 attributable deaths data over time in the United States are reported by the CDC.
- Those data are partially censored, reported with delays, and reported in age bands  $\mathcal{B}$  that may not match those of other data streams.

$$\mathcal{B} = \{0, 1 - 4, 5 - 14, 15 - 24, 25 - 34, 35 - 44, 45 - 54, 55 - 64, 65 - 74, 75 - 84, 85 + \}.$$

→ we provide methods for estimating high resolution age-specific COVID-19 attributable deaths over time without reporting delays.





We estimate the weekly deaths by 1-year age bands  $a \in \mathcal{A} = \{0, 1, \dots, 104, 105\}$  in week  $w \in \mathcal{W} = \{1, \dots, 88\}$  and we denote their expectation by  $\mu_{a,w}$ .

We first decompose  $\mu_{a,w}$  as the product of the weekly deaths for all ages  $\lambda_w$  and the relative contribution  $\pi_{a,w}$  with  $\sum_a \pi_{a,w} = 1$ .

$$\begin{aligned}\mu_{a,w} &= \lambda_w \pi_{a,w} \\ \pi_{a,w} &= \text{softmax}\left([f(a, w)]_{a \in \mathcal{A}}\right) \\ &= \frac{\exp f(a, w)}{\sum_{\tilde{a} \in \mathcal{A}} \exp f(\tilde{a}, w)}.\end{aligned}$$

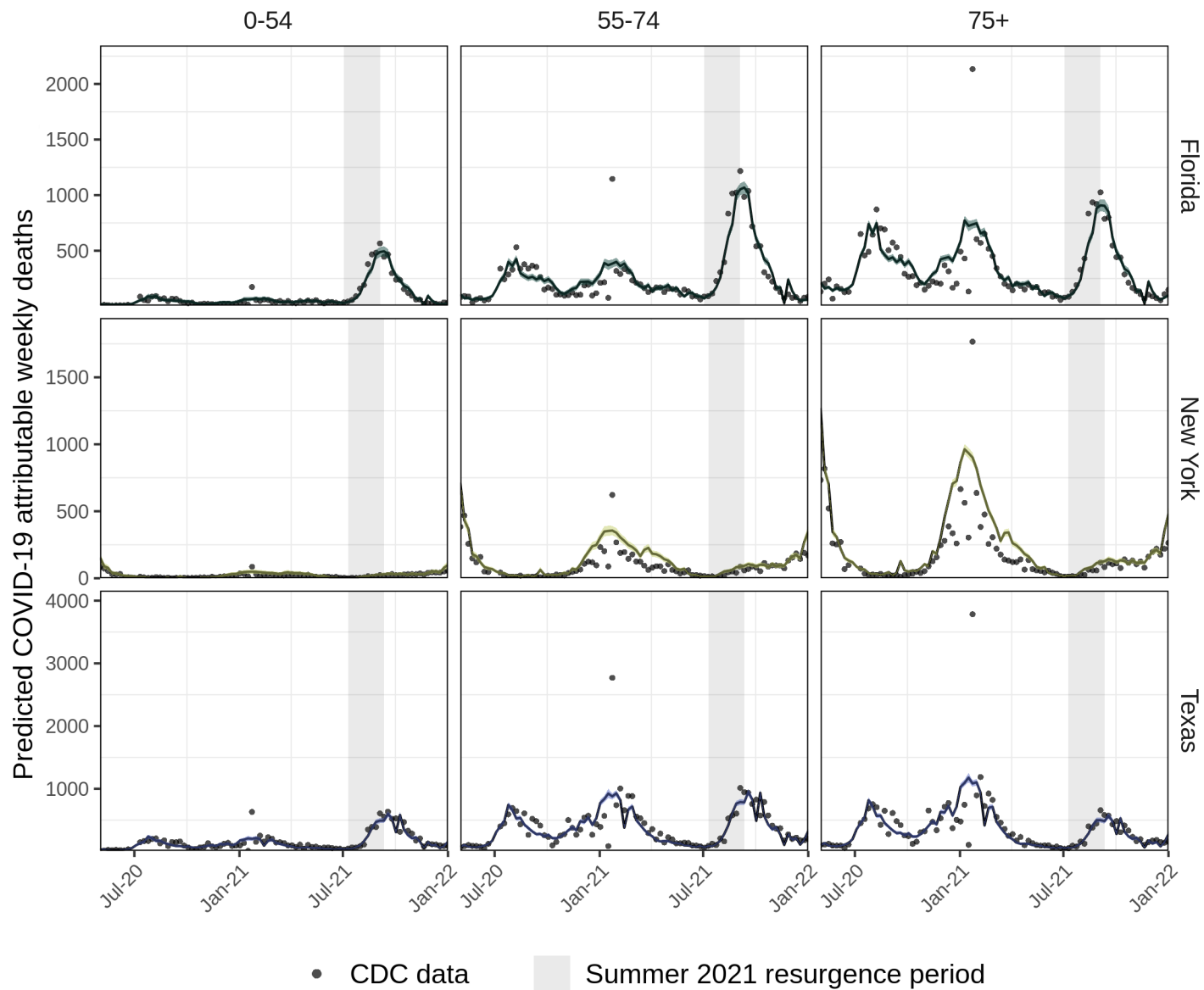
To link the expected weekly deaths by 1-year age band,  $\mu_{a,w}$  to the data, we aggregate them over the age groups specified by the CDC,  $\mu_{b,w} = \sum_{a \in b} \mu_{a,w}$  for all  $b \in \mathcal{B}$ .

$$d_{b,w} \mid \mu_{b,w}, \theta \sim \text{NegBin}(\mu_{b,w}, \theta)$$

with mean  $\mu_{b,w}$  and overdispersion parameter  $\theta > 0$ .

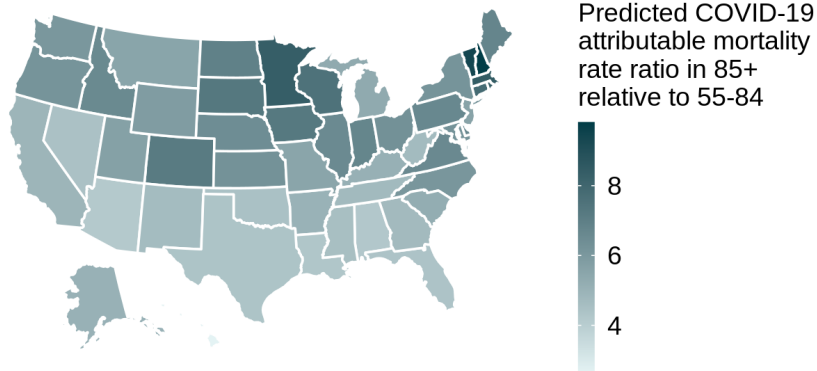
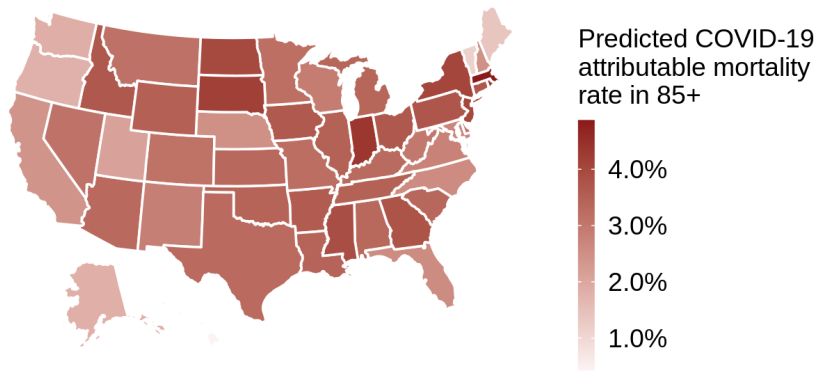
- The JHU data set re-distributes reporting delayed deaths in the CDC data set to earlier dates.
- We require that our posterior predictions of the COVID-19 attributable deaths in age group  $\mathbf{a}$  and week  $\mathbf{w}$ ,  $d_{\mathbf{a},\mathbf{w}}^*$ , sum to  $\sum_{\mathbf{a}} d_{\mathbf{a},\mathbf{w}}^* = d_{\mathbf{w}}^{\text{JHU}}$ .

# Results: Predicted age-specific COVID-19 attributable deaths over time

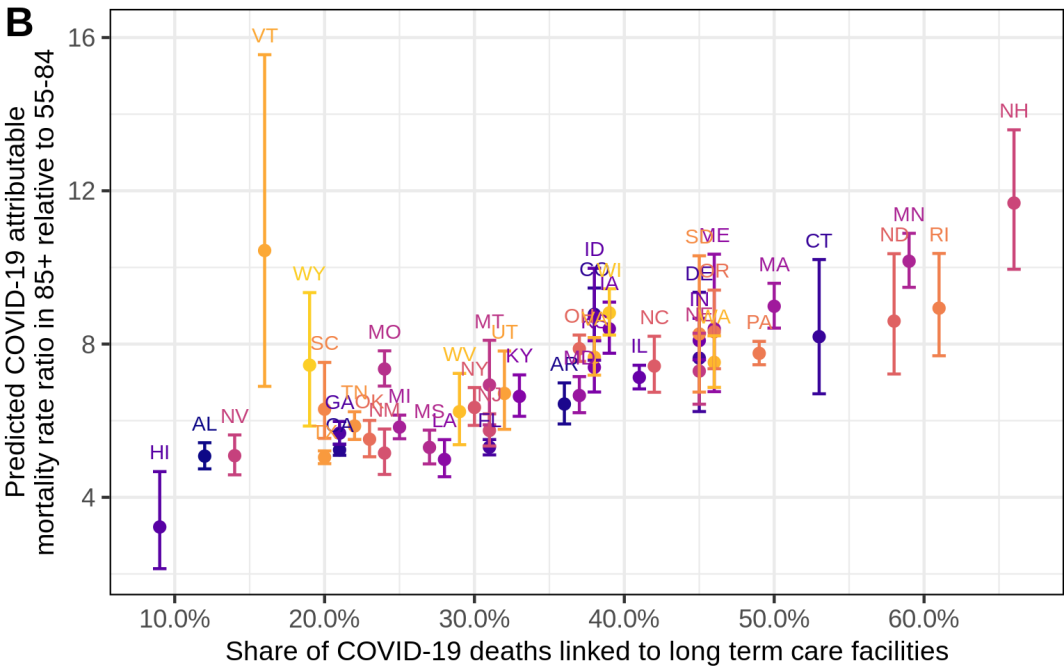


# Results: Posterior predictions of COVID-19 attributable mortality

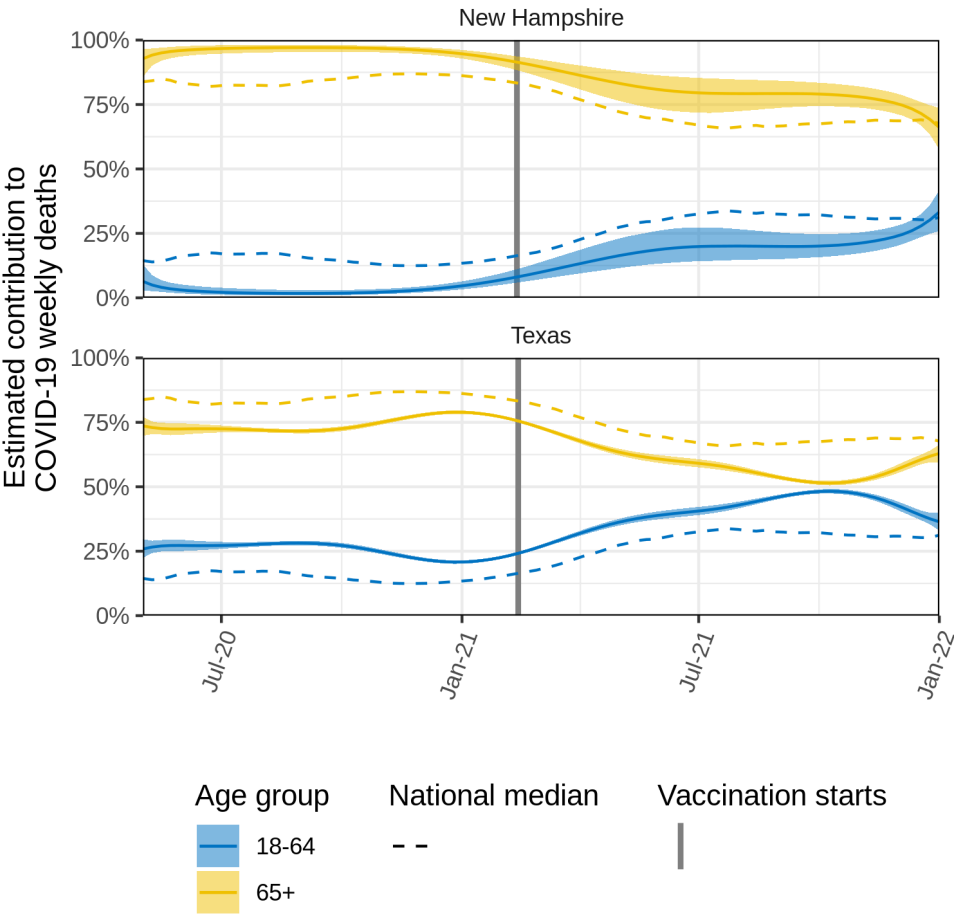
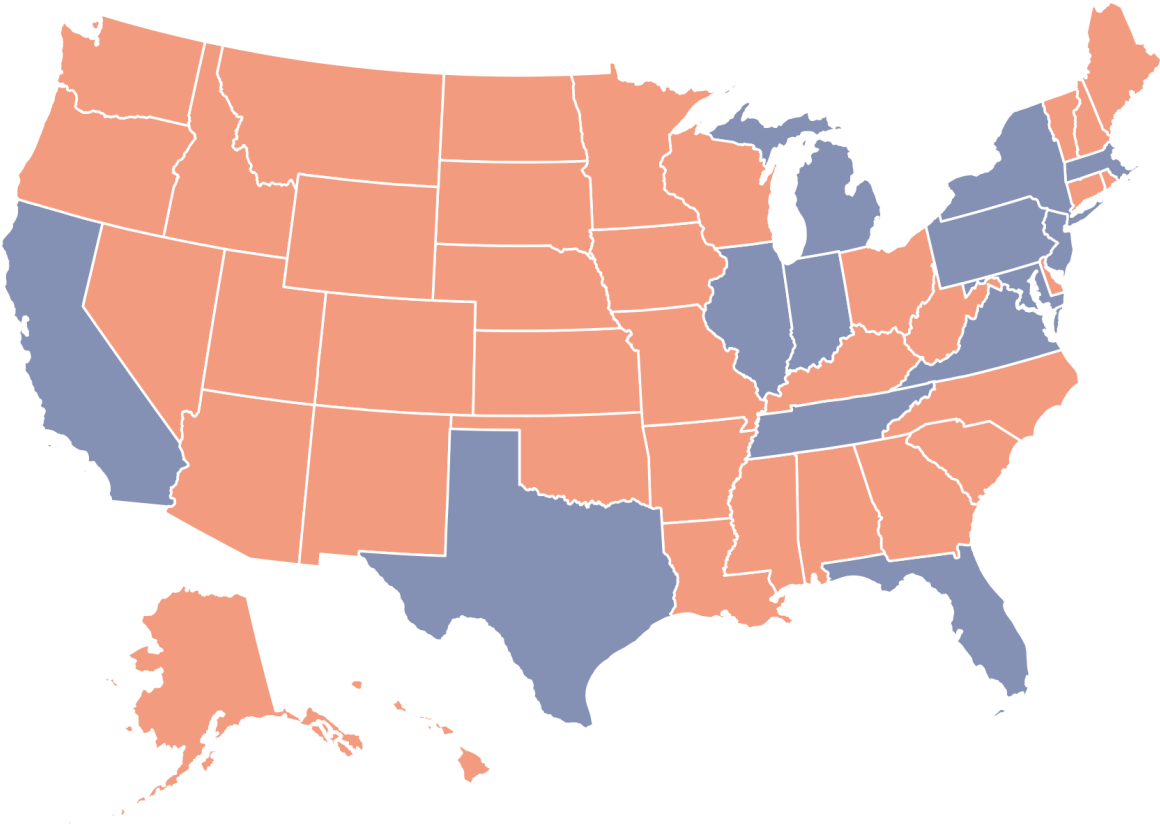
**A**



**B**



# Results: Estimated contribution to COVID-19 deaths over time

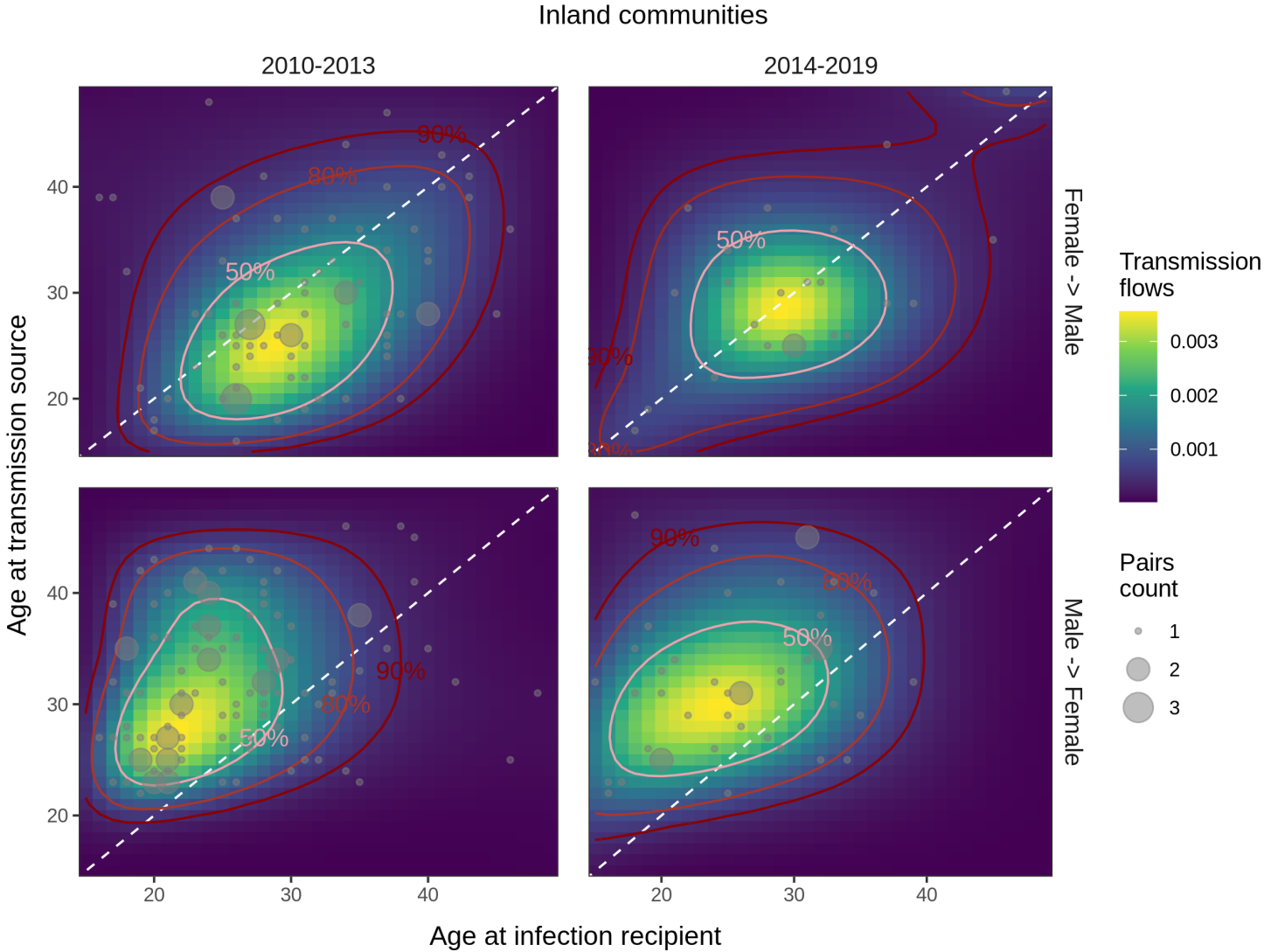


Estimated change in the contribution of 65+ to COVID-19 weekly deaths two months after vaccination started

- significant increase
- no significant change
- significant decrease

**Other applications**

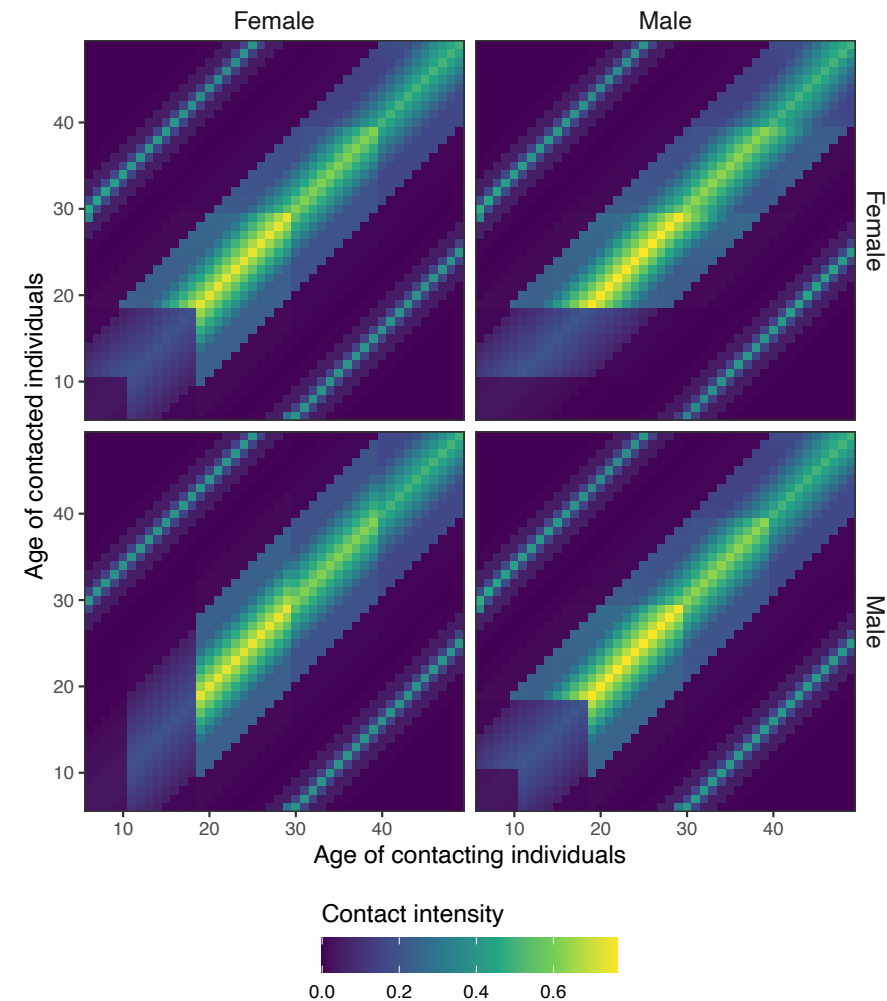
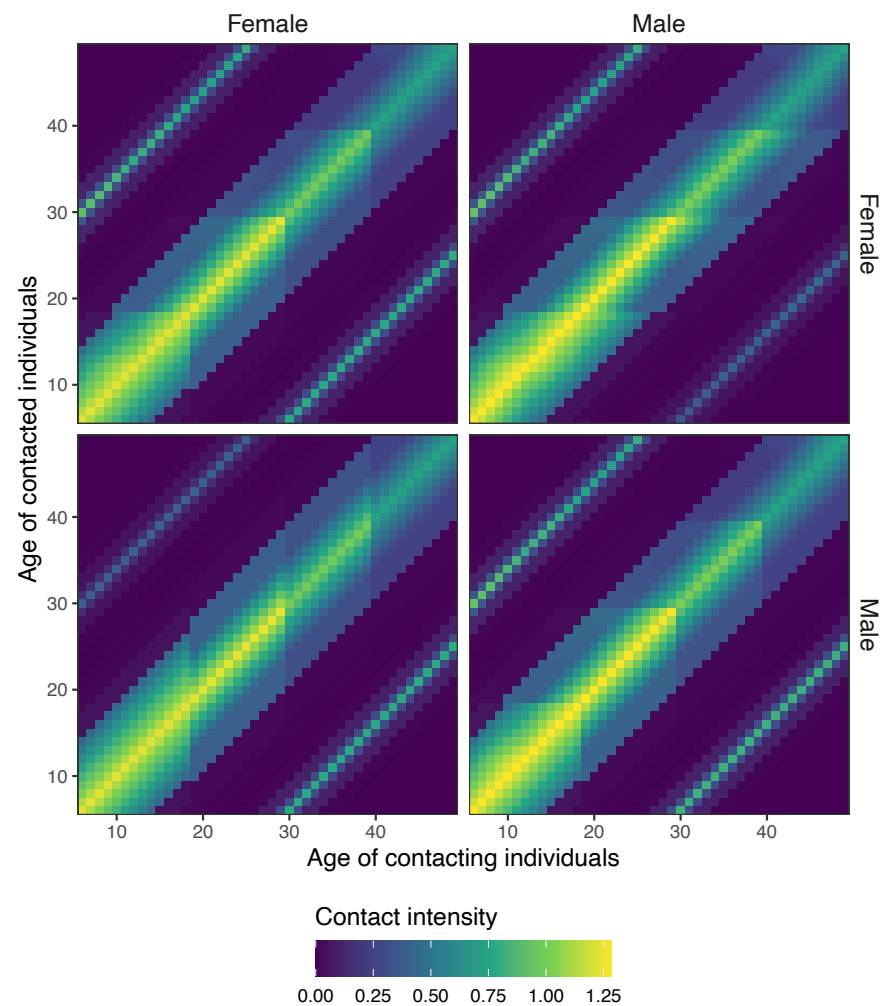
# Estimating shifting HIV transmission dynamics in Rakai, Uganda





# Inferring human contact patterns and contact dynamics in Germany

COVIMOD social contact surveys data collected in Germany during the COVID-19 outbreak.



## **Discussion**

- We develop a novel a low-rank Gaussian Process (GP) projected by regularised B-splines. This projection defines a **new GP with attractive smoothness and computational efficiency properties**.
- Simulation analyses and benchmark results show that the B-splines projected GP may perform better than un-regularised B-splines and Bayesian P-splines, and equivalently well as a standard GP at considerably lower runtimes.
- This approach is **versatile**, applicable to many problems and easy to implement. We provide **templated stan files**.
- The B-splines projected GP priors are likely an appealing addition to the arsenal of Bayesian regularising priors.

# Thank you

## Authors

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## Funding

Imperial College London COVID-19 Response Fund  
EPSRC CDT in Modern Statistics and Statistical Machine Learning  
Imperial College London presidential scholarship  
Bill & Melinda Gates Foundation  
UK Medical Research Council  
NIHR Health Protection Research Unit in Modelling Methodology  
Community Jameel

- Pre-print  
<https://arxiv.org/abs/2106.12360>
- Data and code including templated stan files CC-BY-4.0  
<https://github.com/ImperialCollegeLondon/BSplinesProjectedGPs>

