Divide-and-Conquer Fusion: Methods for unifying distributed analyses $$_{\mbox{\scriptsize Ryan Chan}}$$

Murray Pollock (Newcastle), Hongsheng Dai (Essex), Adam Johansen (Warwick), Gareth Roberts (Warwick) Poster session: Tuesday 28th June (19:00-22:00)

28 June 2022



The Alan Turing Institute

Outline

Introduction to Fusion methodologies

What is the Fusion problem?

Monte Carlo Fusion

Limitations of Monte Carlo Fusion

Divide-and-Conquer Generalised Monte Carlo Fusion

Divide-and-Conquer Generalised Bayesian Fusion

Examples

Logistic regression

Negative Binomial regression

Concluding remarks and future directions

• Target fusion density:

$$f(\mathbf{x}) \propto \prod_{c=1}^{C} f_c(\mathbf{x})$$

- No general analytical approach
- Monte Carlo: assume we can sample $\mathbf{x}^{(c)} \sim f_c(\mathbf{x})$
- Applications:
 - Expert elicitation: combining views of multiple experts
 - Big Data (by construction)
 - Partitioning large datasets to make them more manageable
 - Inference in privacy settings

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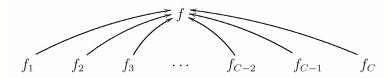
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Fork-and-join

The fork-and-join approach:



- Several fork-and-join methods have been developed (typically for Bayesian inference for large dataset applications):
 - Kernel density averaging (KDEMC) [Neiswanger et al., 2014]
 - Weierstrass sampler (WRS) [Wang and Dunson, 2013]
 - Consensus Monte Carlo (CMC) [Scott et al., 2016]
- A primary weakness of these methods is that the recombination is inexact in general and involve approximations
 - CMC is exact if all sub-posteriors are Gaussian
 - All theory is asymptotic in the number of observations
- However, Monte Carlo Fusion [Dai et al., 2019] (and subsequently Bayesian Fusion [Dai et al., 2021]) is exact in the sense it targets the correct fusion density

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Proposition

Suppose that $p_c(\mathbf{y}|\mathbf{x}^{(c)})$ is the transition density of a stochastic process with stationary distribution $f_c^2(\mathbf{x})$. The (C+1)d-dimensional (fusion) density proportional to the integrable function

$$g\left(\mathbf{x}^{(1)},\ldots,\mathbf{x}^{(C)},\mathbf{y}\right) \propto \prod_{c=1}^{C} \left[f_c^2\left(\mathbf{x}^{(c)}\right) \cdot \rho_c\left(\mathbf{y} \middle| \mathbf{x}^{(c)}\right) \cdot \frac{1}{f_c\left(\mathbf{y}\right)}\right]$$

admits the marginal density f for y.

Main idea: If we can sample from g, then we can can obtain a draw from the fusion density $(y \sim f)$

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Main idea: If we can sample from g, then we can can obtain a draw from the fusion density $(\mathbf{y} \sim f)$

- There are many possible choices for $p_c(\mathbf{y}|\mathbf{x}^{(c)})$
- Let $p_c(\mathbf{y}|\mathbf{x}^{(c)}) := p_{T,c}(\mathbf{y}|\mathbf{x}^{(c)})$, the transition density of the d-dimensional (double) Langevin (DL) diffusion processes $\mathbf{X}_t^{(c)}$ from $\mathbf{x}^{(c)}$ to \mathbf{y} for $c=1,\ldots,C$, for a pre-defined time T>0 given by

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(where $W_t^{(c)}$ is d-dimensional Brownian motion and abla is the gradient operator over x)

- Has stationary distribution $f_c^2(x)$
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Extended target density:

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$$h\left(\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(C)}, \mathbf{y}\right) \propto \prod_{c=1}^{C} \left[f_c\left(\mathbf{x}^{(c)}\right) \right] \cdot \exp\left(-\frac{C \cdot \|\mathbf{y} - \bar{\mathbf{x}}\|^2}{2T}\right)$$

- $\bar{x} = \frac{1}{C} \sum_{c=1}^{C} x^{(c)}$
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• Simulation from *h* is easy:

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- 1. Simulate $\mathbf{x}^{(c)} \sim f_c(\mathbf{x})$ independently
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Rejection sampling - acceptance probability

Acceptance probability:

$$\frac{g(\boldsymbol{x}^{(1)},\ldots,\boldsymbol{x}^{(C)},\boldsymbol{y})}{h(\boldsymbol{x}^{(1)},\ldots,\boldsymbol{x}^{(C)},\boldsymbol{y})} \propto \rho_0 \cdot \rho_1$$

where

$$\begin{cases} \rho_0 \coloneqq e^{-\frac{C\sigma^2}{2T}}, & \sigma^2 = \frac{1}{C} \sum_{c=1}^C \left\| \boldsymbol{x}^{(c)} - \bar{\boldsymbol{x}} \right\|^2 \\ \rho_1 \coloneqq \mathbb{E}_{\bar{\mathbb{W}}} \left(\prod_{c=1}^C \left[\exp\left\{ -\int_0^T \left(\phi_c \left(\boldsymbol{X}_t^{(c)} \right) - \boldsymbol{\Phi}_c \right) \mathrm{d}t \right\} \right] \right) \end{cases}$$

where $\overline{\mathbb{W}}$ denotes the law of C independent Brownian bridges $X_t^{(1)}, \dots, X_t^{(C)}$ with $X_0 = x^{(c)}$ and $X_T^{(c)} = y$

• Trade-off with choice of T: as T increases, ρ_0 increases, but this results in ρ_1 to be small (might typically decrease exponentially with T)

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Monte Carlo Fusion

Monte Carlo Fusion - Summary

- Aim: Sample from g (admits marginal density f for y)
- Proposal:

$$h\left(\mathbf{x}^{(1)},\ldots,\mathbf{x}^{(C)},\mathbf{y}\right) \propto \prod_{c=1}^{C} \left[f_c\left(\mathbf{x}^{(c)}\right)\right] \cdot \exp\left(-\frac{C \cdot \|\mathbf{y} - \bar{\mathbf{x}}\|^2}{2T}\right)$$

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Limitations of Monte Carlo Fusion

- Robustness: there is a lack of robustness when:
 - sub-posterior correlation increases
 - C increases
 - d increases
 - combining conflicting sub-posteriors
- Aim: To construct a fusion algorithm / framework to alleviate some of these limitations (see Dai et al. [2021]; Chan et al. [2021] for full details)

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The Generalised Monte Carlo Fusion (GMCF) approach

Problem: Scalability with sub-posterior correlation

- Recall we have the flexibility to choose different p_c (transition density of stochastic process with f_c^2 invariant density)
- Now, we choose p_c to be the transition density of the d-dimensional (double) Langevin (DL) diffusion processes $\boldsymbol{X}_t^{(c)}$ with covariance matrix, $\boldsymbol{\Lambda}_c$ from $\boldsymbol{x}^{(c)}$ to \boldsymbol{y} for $c=1,\ldots,C$, over [0,T] given by

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Constructing an importance sampler

- Switch to importance sampler for the extended target density $g(\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(C)}, \mathbf{y})$:
 - Rejection sampling can be wasteful
 - We will subsequently embed this approach within a SMC algorithm
- Consider an alternative proposal density *h* for the extended target *g*:

$$h\left(\boldsymbol{x}^{(1)},\dots,\boldsymbol{x}^{(C)},\boldsymbol{y}\right) \propto \prod_{c=1}^{C} \left[f_c\left(\boldsymbol{x}^{(c)}\right)\right] \cdot \exp\left\{-\frac{(\boldsymbol{y}-\tilde{\boldsymbol{x}})^{\intercal}\boldsymbol{\Lambda}^{-1}(\boldsymbol{y}-\tilde{\boldsymbol{x}})}{2T}\right\},$$

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Importance weights

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where $\phi_c(\mathbf{x}) := \frac{1}{2} \left(\nabla \log f_c(\mathbf{x})^{\mathsf{T}} \mathbf{\Lambda}_c \nabla \log f_c(\mathbf{x}) + \mathrm{Tr}(\mathbf{\Lambda}_c \nabla^2 \log f_c(\mathbf{x})) \right)$, with $\mathbf{W}_{\mathbf{\Lambda}_c}$ denoting the law of a Brownian bridge $\{\mathbf{X}_t^{(c)}, t \in [0, T]\}$ with $\mathbf{X}_0^{(c)} := \mathbf{x}^{(c)}, \ \mathbf{X}_T^{(c)} := \mathbf{y}$ and covariance matrix $\mathbf{\Lambda}_c$

Scalability with sub-posterior correlation

In our Generalised Monte Carlo Fusion setting:

- Able to incorporate covariance / correlation information within our proposals and through p_c and h (in MCF $\Lambda_c = \mathbb{I}_d$ for c = 1, ..., C)
- Unfortunately no longer have i.i.d. draws from f but now have weighted samples to approximate f (later embed within divide-and-conquer SMC [Lindsten et al., 2017] framework)

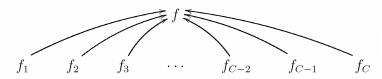
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Problem: Scalability with C

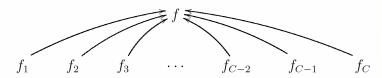
The (Generalised) Monte Carlo Fusion algorithm implies a fork-and-join approach:



- Not necessarily the most efficient way to combine sub-posteriors
- ullet For MCF, acceptance probabilities typically decrease geometrically with C

Problem: Scalability with C

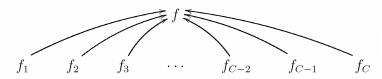
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- Not necessarily the most efficient way to combine sub-posteriors
- For MCF, acceptance probabilities typically decrease geometrically with C

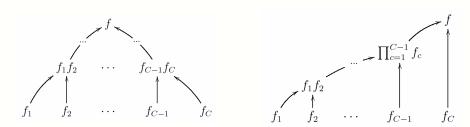
Problem: Scalability with C

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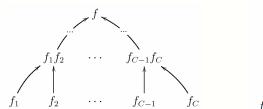
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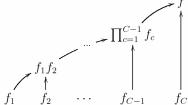
- Solution: Divide-and-Conquer Monte Carlo Fusion
 - We could perform fusion in a proper divide-and-conquer framework
 - i.e. a fork-and-join method is recursively applied
 - Two possible choices are balanced-binary (left) and progressive (right) trees



Note: Other trees are possible

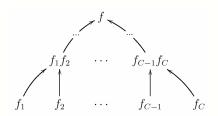
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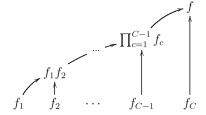




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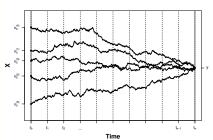


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Generalised Bayesian Fusion

Problem: Robustness to conflicting sub-posteriors

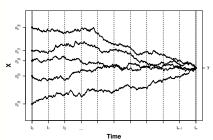
- Generalising the Bayesian Fusion approach of Dai et al. [2021]
- Recall choosing a value T > 0 for MCF can be hard:
 - Want to make T large so that ρ_0 is large but this makes ρ_1 smaller (since we have to simulate a diffusion over a longer time horizon T)
- Solution: Introduce temporal partition of *T*
 - Have the flexibility to choose T large enough for initialisation, while being able to have small intervals in the partition



Generalised Bayesian Fusion

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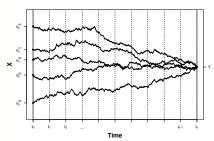
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Generalised Bayesian Fusion

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Examples

- We compare our methodology with the approximate methodologies KDEMC [Neiswanger et al., 2014], WRS [Wang and Dunson, 2013] and CMC [Scott et al., 2016]
- To compare methods we calculate the integrated absolute distance metric

$$JAD = rac{1}{2d} \sum_{i=1}^d \int \left| \hat{f}(\pmb{x}_j) - f(\pmb{x}_j)
ight| \, \mathrm{d}\pmb{x}_j \in [0,1]$$

where $\hat{f}(x_j)$ is the marginal density for x_j based on the method applied (computed using a kernel density estimate) and $f(x_j)$ is target marginal density

• Gives a measure of how accurate our samples are to our target (lower is better)

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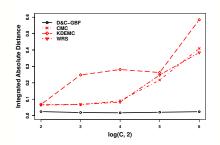
$$IAD = \frac{1}{2d} \sum_{i=1}^{d} \int \left| \hat{f}(\mathbf{x}_{j}) - f(\mathbf{x}_{j}) \right| d\mathbf{x}_{j} \in [0, 1]$$

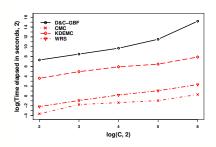
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Logistic regression

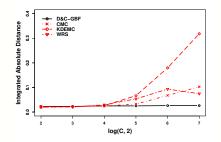
- Simulated data example with n = 1000 and d = 5 (and set N = 10000)
 - Small data size means that large data assumptions will fail
- We split the data into C=4,8,16,32,64 and apply D&C-GBF (using a balanced binary tree approach)

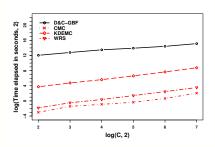




Negative Binomial regression

- Using the Bike sharing dataset (n = 17379, d = 10) (and set N = 10000)
- We split the data into C=4,8,16,32,64,128 and apply D&C-GBF (using a balanced binary tree approach)





Ongoing research questions

- Reducing the computational cost of the Fusion approach
 - Exactness comes at a cost
- Practical implementation considerations for specific applications:
 - Big data setting: evaluating ϕ_c has $\mathcal{O}(m_c)$ cost can perhaps employ sub-sampling methods to reduce this cost
 - Confidential fusion (Con-fusion): where sharing information/data between cores is not permitted
- Scalability with dimension
 - Performance with regards to dimension has improved since MCF, but not been explicitly addressed

References

- Beskos, A., Papaspiliopoulos, O., Roberts, G. O., and Fearnhead, P. (2006). Exact and computationally efficient likelihood-based estimation for discretely observed diffusion processes (with discussion). Journal of the Royal Statistical Society: Series B (Statistical Methodology), 68(3):333–382.
- Beskos, A., Roberts, G. O., et al. (2005). Exact simulation of diffusions. The Annals of Applied Probability, 15(4):2422-2444.
- Chan, R. S., Pollock, M., Johansen, A. M., and Roberts, G. O. (2021). Divide-and-Conquer Monte Carlo Fusion. Statistics e-print 2110.07265, arXiv.
- Dai, H., Pollock, M., and Roberts, G. O. (2019). Monte Carlo Fusion. Journal of Applied Probability, 56(1):174-191.
- Dai, H., Pollock, M., and Roberts, G. O. (2021). Bayesian Fusion: Scalable unification of distributed statistical analyses. Statistics e-print 2102.02123, arXiv.
- Lindsten, F., Johansen, A. M., Naesseth, C. A., Kirkpatrick, B., Schön, T. B., Aston, J. A., and Bouchard-Côté, A. (2017). Divide-and-Conquer with Sequential Monte Carlo. Journal of Computational and Graphical Statistics, 26(2):445–458.
- Neiswanger, W., Wang, C., and Xing, E. P. (2014). Asymptotically Exact, Embarrassingly Parallel MCMC. In Proceedings of the Thirtieth Conference on Uncertainty in Artificial Intelligence, UAl'14, page 623–632, Arlington, Virginia, USA. AUAI Press.
- Pollock, M., Johansen, A. M., Roberts, G. O., et al. (2016). On the exact and ε-strong simulation of (iump) diffusions. Bernoulli, 22(2):794-856.
- Scott, S. L., Blocker, A. W., Bonassi, F. V., Chipman, H. A., George, E. I., and McCulloch, R. E. (2016). Bayes and Big Data: The Consensus Monte Carlo Algorithm. International Journal of Management Science and Engineering Management, 11(2):78–88.
- Wang, X. and Dunson, D. B. (2013). Parallelizing MCMC via Weierstrass Sampler, Statistics e-print 1312,4605, arXiv.

Poster session: Tuesday 28th June (19:00-22:00)