Bayesian Sensitivity Analysis for a Missing Data Model

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Outline

1 Sensitivity Analysis

2 Missing data model

3 Results



Observation	Outcome
1	15
2	3
3	*
4	*
5	5
:	:



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1	15
2	3
3	*
4	*
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:	:

Solutions

- Assume data is missing completely at random
- Data imputation



What if we assume MCAR, but in reality it does not hold?



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Sensitvity Analysis

How robust are study conclusion to violations of assumptions?



How?

Statistics

Assign a parameter to the assumption



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Problems

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- Difficult to give intuition for practicioners



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Robustness

- At what values of the sensitivity parameter do study conclusions not hold anymore?
- Size is an indicator of robustness



Bayesian sensitivity analysis

Why Bayesian?

- Put all your intuition into the prior for the sensitivity parameter
- The posterior distribution acts as a single summary of your beliefs



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Problem

- Not a lot of theory available
- Wat if we are "close" to the truth with our parameter, will our conclusion converge to the truth?



Observations

$$Y_1,\ldots,Y_n\stackrel{\mathsf{iid}}{\sim} P$$



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Interest

 $\mathbb{E}[g(Y)]$



Sensitivity parameter

Conditional distributions

$$Y \mid R = 1 \sim P_1,$$

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Consequence

$$P = pP_1 + (1-p)P_0$$



Sensitivity parameter ¹

$$P_0(A) = \frac{\int_A e^{\mathbf{q}} dP_1}{\int e^{\mathbf{q}} dP_1}$$



¹D. O. Scharfstein, M. J. Daniels, and J. Robins (2003). "Incorporating prior beliefs about selection bias into the analysis of randomized trials with missing outcomes". In: Biostatistics 4.4, pp. 495-512

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Example

$$q(y) = \alpha \log(y), \quad \alpha \in \mathbb{R}$$



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Intuition?

$$logit Pr(R = 0 | Y) = \eta + q(Y)$$



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When $q(y) = \alpha \log(y)$

- α = 0: MCAR
 - α > 0: Higher outcomes are more likely to not be observed
 - α < 0: Lower outcomes are more likely to not be observed



Prior choices

Parametrization p, P_1

$$p \sim Beta(\cdot, \cdot)$$

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We want to show

Bernstein-von Mises theorem for both parametrizations

$$\sqrt{n}\left(\chi_q(p, P_1) - \chi_q(\hat{p}, \hat{P_1})\right) \mid R^{(n)}, X^{(n)}, q \rightsquigarrow ?$$
 a.s.



Relation

Relation between P and P_1

$$P_{1}(A) = \frac{\int_{A} \frac{1}{1 + e^{\eta + q(y)}} dP(y)}{\int \frac{1}{1 + e^{\eta + q(y)}} dP(y)}$$



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Prior

 $P \sim DP(a)$ is equivalent to $P_1 \sim NEGP(a, b)$



BvM for normalized extended gamma processes

Theorem

Let $X_1,...,X_n \mid P \sim P$ with $P \sim \mathsf{NEGP}(a,b)$ and let \mathcal{F} be a Donsker-class. Let P_0 be the true distribution of the data. Under assumption later to be specified

$$\sqrt{n}(P-\mathbb{P}_n)\mid X^{(n)} \rightsquigarrow B_{P_0},$$

in $\ell^{\infty}(\mathcal{F})$ almost surely $[P_0^{\infty}]$, where B_{P_0} is a P_0 -Brownian bridge.



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Assumptions

- *b* is a positive, bounded measurable function
- a is an atomless, finite measure
- \mathcal{F} has envelope function F with $\int Fda < \infty$
- There exist r,q>0 with 1/q+1/r<1/2 such that $P_0F^r<\infty$ and $P_0b^{-q}<\infty$



Proof strategy

- The posterior of P is a mixture (over λ) of NCRMs
- ullet Show mixing density concentrates on big λ 's
- Show the continuous part of the NCRM vanishes
- Use multiplier central limit theorem



To do

- Fix assumption on prior for η
- Show what happens when we put a prior on $q(\alpha)$
- Extend model with covariates!
- ...

