The use/abuse of copulas in actuarial science and finance

Abstract

This is an assignment for the actuarial models course. The assignment is to summarize [EMS02], [FV98], [DE] and discuss the following:

- The purpose is to understand the impact of the assumption regarding the dependence structure between risk factors.
- This is done by means of the concept of copulas.
- In particular, we study the impact of misused copulas and correlation in the valuation of collateralized debt obligations (CDO's).

1 Overview

[EMS02] introduces linear dependency, copulas, comonotonicity and rank correlation which already are introduced in the course. Also spherical/elliptical distributions and tail dependence are introduced. Spherical/elliptical distributions generalize the normal distribution and has good properties for common dependency measures. Tail dependence quantifies dependency in the tails. They also discuss some common dependency fallacies and simulation of copulas.

[FV98] also introduces copulas, present examples, simulation and fitting of copulas. Main examples are joint mortality modeling and modeling insurance company indemnity claims. And later also introducing stochastic orders and distortion functions which we also covered in class.

[DE] also introduces copulas but mainly talks about 2007-2008 financial crisis. It mainly discusses modeling CDO's/trenching, the impact of not modeling dependence correctly specifically default clustering and illustrating with modeling default probability of bonds under identical pairwise correlation assumption.

2 Introduction to new concepts

2.1 Spherical/elliptical distributions

Spherical/elliptical distributions have a spherical/elliptical symmetry. I.e. we can characterize a spherical distributions X as follows:

$$X =_d RU. (1)$$

with R a positive random variable and U independent of R a random vector uniformly distributed on the unit sphere. Elliptical distributions can be obtained as affine transformations of spherical distributions. Elliptical distributions are fully characterized by their mean, covariance matrix and the characteristic function of normalized R (normalize such that the mean = 1) also called generator when the covariance exists $(E[R^2] < \infty)$. So it is a semi-parametric family of distributions with limited 1 dimensional non-parametric random variable.

Here are some obvious properties of elliptical distributions:

- An affine transformation of an elliptical distribution is also elliptical.
- The sum in the set of elliptical distributions with the same generator is closed.
- Marginal and conditional distributions of the components of elliptical distributions are elliptical. The intuition for this is the same as for normal distributions, intersection between ellipsoids and planes are also ellipsoids.

In [EMS02] they show that for linear portfolios where individual risk together are jointly elliptical distributed, risk measures lose structure, simplifying risk management tasks. Specifically they show that VaR is equivalent to variance risk analysis. This emphasizes that the assumption of elliptical distributions is a strong assumption and normal distributions are even stronger.

2.2 Tail dependence

Making assumptions is dangerous. Data trades-off with assumptions. Determining high dimensional structure requires a lot of data. Dependence is a high dimensional structure and at tails we have little data both by definition. So naturally making assumptions about the dependence structure at the tails is a risky business. Tail dependence is a way to quantify the dependence in the tails. The upper tail dependence (λ) between X and Y is defined as follows:

$$\lim_{\alpha \to 1^{-}} P[Y > F_2^{-1}(\alpha) \mid X > F_1^{-1}(\alpha)] = \lambda.$$
 (2)

When X and Y are continuous distribution we can express this in terms of their copula :

$$\lim_{\alpha \to 1^{-}} P[Y > F_2^{-1}(\alpha) \mid X > F_1^{-1}(\alpha)] \tag{3}$$

$$= \lim_{\alpha \to 1^{-}} P[U_2 > \alpha \mid U_1 > \alpha] \tag{4}$$

$$= \lim_{\alpha \to 1-} \frac{P[U_1 > \alpha, U_2 > \alpha]}{P[U_1 > \alpha]} \tag{5}$$

$$= \lim_{\alpha \to 1-} \frac{1 - P[(U_1 > \alpha, U_2 > \alpha)^c]}{1 - \alpha}$$
 (6)

$$= \lim_{\alpha \to 1^{-}} \frac{1 - P[(U_1 \le \alpha) \cup (U_2 \le \alpha)]}{1 - \alpha}$$
 (7)

$$= \lim_{\alpha \to 1-} \frac{1 - P[U_1 \le \alpha] - P[U_2 \le \alpha] + P[U_1 \le \alpha, U_2 \le \alpha]}{1 - \alpha}$$
 (8)

$$= \lim_{\alpha \to 1-} \frac{1 - 2\alpha + C(\alpha, \alpha)}{1 - \alpha}.$$
 (9)

[EMS02] show different ways to calculate tail dependence. Because tail dependence can be expressed in terms of the copula, it is preserved under bijective transformations of individual distributions. In the case of elliptical distributions, upper tail dependence is 0 when $\forall t > 0 \in \mathbb{R} : P[R > t] > 0$ (R is unbounded) and there is no collinearity. This can be proven by linearly transforming the marginals to be the same and using the fact that the conditional distribution is unbounded converging to 0.

$$\lim_{\alpha \to 1^{-}} P[Y > F_2^{-1}(\alpha) \mid X > F_1^{-1}(\alpha)]$$
 (10)

$$= \lim_{\alpha \to 1^{-}} P[Y > F^{-1}(\alpha) \mid X > F^{-1}(\alpha)]$$
 (11)

$$= \lim_{z \to \infty} P[Y > z \mid X > z] \tag{12}$$

$$=0 (13)$$

I.e. elliptical distributions or normal distribution have no tail dependence and are unable to approximate dependence of extreme events.

References

- [DE] Catherine Donnelly and Paul Embrechts. "THE DEVIL IS IN THE TAILS: ACTUARIAL MATHEMATICS AND THE SUBPRIME MORTGAGE CRISIS". en. In: ().
- [EMS02] Paul Embrechts, Alexander J. McNeil, and Daniel Straumann. "Correlation and Dependence in Risk Management: Properties and Pitfalls". en. In: Risk Management. Ed. by M. A. H. Dempster. 1st ed. Cambridge University Press, Jan. 2002, pp. 176–223. ISBN: 978-0-521-78180-0 978-0-511-61533-7 978-0-521-16963-9. DOI: 10.1017/CB09780511615337.008. URL: https://www.cambridge.org/core/product/identifier/CB09780511615337A013/type/book_part (visited on 04/21/2024).
- [FV98] Edward W. Frees and Emiliano A. Valdez. "Understanding Relationships Using Copulas". en. In: North American Actuarial Journal 2.1 (Jan. 1998), pp. 1–25. ISSN: 1092-0277, 2325-0453. DOI: 10.1080/10920277. 1998.10595667. URL: http://www.tandfonline.com/doi/abs/10.1080/10920277.1998.10595667 (visited on 04/05/2024).