ML-HW2: Mathematical Derivation

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GitLink: https://github.com/IaTsai/ML-HW2

HW2-1: Naive Bayes Classifier

1. Mathematical Model

Naive Bayes models the class probability:

$$P(Y|X) \propto P(Y) \prod_{i} P(X_i|Y) \tag{1}$$

- P(Y) is the prior (estimated from training samples).
- $P(X_i|Y)$ is the likelihood (discrete/continuous).
- All probabilities are computed in log-space to avoid underflow.

2.1 Discrete Mode Implementation

Implemented Code:

```
image = np.zeros((10, image_size, 32), dtype=np.int32)
...
image[label][j][grayscale // 8] += 1
...
if likelihood == 0:
    likelihood = np.min(image[j][k][np.nonzero(image[j][k])])
...
prob[j] += np.log(prior[j] / train_size)
prob[j] += np.log(likelihood / image_sum[j][k])
```

Explanation:

- Pixel grayscale is mapped to bin index, each pixel value (0-255) into 32 bins (each bin represents 8 levels).
- Count pixel frequencies for each class from training data.
- Frequencies stored in a 3D array [class][pixel][bin].
- **Debug:** If the bin count is zero, use smoothing to avoid crashing the log by substituting a minimal non-zero value.
- Classifier sums the log prior and log-likelihoods over all pixels.

2.2 Continuous Mode Implementation

Implemented Code:

Explanation:

- MLE estimates μ and σ^2 from training data.
- Gaussian PDF (in log form) computes $P(X_i|Y)$.
- Sum all pixel-wise log-likelihoods and add the class prior; the class with the highest score is the predicted result.

Log-Likelihood under Gaussian Naive Bayes

Example: Given a single test sample (flattened pixel array), and estimated Gaussian parameters (mean & variance) for each class:

• For each class j, sum the log prior and the pixel-wise log-likelihoods:

$$\log P(Y=j) + \sum_{i} \log P(X_i \mid Y=j) \tag{2}$$

Code illustration:

This additive formulation across pixels accounts for the overall class score. The class with the highest (least negative) final prob[j] becomes the prediction.

4. Visualization and Output

```
print('Posterior (in log scale):')
print(f'Prediction: {pred}, Ans: {answer}\n')
DrawImagination(...)
```

5. Execution Interface

- python 2-1_naive_bayes_classifier.py --train: Trains Discrete or Continuous mode.
- python 2-1_naive_bayes_classifier.py: Loads saved results and displays imaginations.

HW2-2: Beta-Binomial Bayesian Online Learning

Note: In accordance with the assignment's requirements, no library-based sampling (e.g., numpy.random.beta) is used. All distributions are computed analytically, and likelihoods are implemented manually.

1.1 Mathematical Proof

Given:

- Prior: Beta(a, b)
- Data: m successes (1s), n failures (0s)

Posterior Update:

$$P(p|D) \propto p^{a+m-1} (1-p)^{b+n-1}$$
 (3)

$$a' = a + m, \quad b' = b + n \tag{4}$$

This posterior update can be derived by applying Bayes' theorem:

$$P(p|D) \propto P(D|p) \cdot P(p)$$
 (5)

Assuming:

- Binomial likelihood: $P(D|p) = p^m(1-p)^n$
- Beta prior: $P(p) = \frac{1}{B(a,b)}p^{a-1}(1-p)^{b-1}$

Then:

$$P(p|D) \propto p^{m} (1-p)^{n} \cdot p^{a-1} (1-p)^{b-1} = p^{a+m-1} (1-p)^{b+n-1}$$
(6)

Therefore, the posterior is also a Beta distribution: Beta(a + m, b + n)

Case Example

Example 1: Uniform Prior (a = 0, b = 0)

Input:

```
input a: 0
input b: 0
case 1: 010101010101010101
```

Step-by-step:

- m = 11 (number of '1's)
- n = 11 (number of '0's)

Posterior update:

$$a' = a + m = 0 + 11 = 11, \quad b' = b + n = 0 + 11 = 11$$
 (7)

Output:

```
Likelihood: 0.16818809509277344

Beta prior: a = 0 b = 0

Beta posterior: a = 11 b = 11
```

Example 2: Informative Prior (a = 10, b = 1)

Input:

```
input a: 10
input b: 1
case 1: 010101010101010101
```

Step-by-step:

- m = 11
- n = 11

Posterior update:

$$a' = a + m = 10 + 11 = 21, \quad b' = b + n = 1 + 11 = 12$$
 (8)

Output:

```
Likelihood: 0.16818809509277344

Beta prior: a = 10 b = 1

Beta posterior: a = 21 b = 12
```

Code Illustration

```
import sys, math
a = int(input("input a: "))
b = int(input("input b: "))
with open(sys.argv[1]) as f:
    for line in f:
        line = line.strip()
        count1 = line.count('1')
        count0 = line.count('0')
        likelihood = (math.gamma(a + b) / (math.gamma(a) * math.gamma(b
           ))) * \
                     (math.gamma(a + count1) * math.gamma(b + count0) /
                      math.gamma(a + b + count1 + count0))
        print("Likelihood:", likelihood)
        print(f"Beta prior: a = {a} b = {b}")
        a += count1
        b += count0
        print(f"Beta posterior: a = {a} b = {b}\n")
```

Note: This Gamma-function-based implementation of the likelihood follows the theoretical Beta-Binomial formulation. However, the actual implementation in 2-2_online_learning.py computes the likelihood using the standard Binomial model:

$$Likelihood = C(N, m) \cdot P^m \cdot (1 - P)^{N - m}$$
(9)

where P = m/N, and C(N, m) is the binomial coefficient.

3. Output Format

```
Likelihood: ...

Beta prior: a = X b = Y

Beta posterior: a = X' b = Y'
```

HW2-3: Mathematical Derivation

This section provides complete mathematical proofs for two conjugate prior relationships required in the assignment:

- 1. Beta-Binomial Conjugation
- 2. Gamma-Poisson Conjugation

Proof 1: Beta-Binomial Conjugation

Objective: Show that the Beta distribution acts as a conjugate prior to the Binomial likelihood, including deriving the posterior distribution.

Problem Setup

• Likelihood: Binomial distribution

$$P(D|p) = \binom{N}{m} p^m (1-p)^n \tag{10}$$

where N = m + n (total trials), m = successes, n = failures

• Prior: Beta distribution

$$P(p) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} p^{a-1} (1-p)^{b-1} = \frac{1}{B(a,b)} p^{a-1} (1-p)^{b-1}$$
(11)

• Goal: Prove that the posterior is also a Beta distribution

Derivation

Given:

- Prior: Beta(a, b)
- Data: m successes (1s), n failures (0s)
- Total trials: N = m + n

Step 1: Write out Posterior (Bayes' Theorem)

$$P(p|D) \propto P(D|p) \cdot P(p)$$
 (12)

Step 2: Binomial Likelihood

$$P(D|p) = \binom{N}{m} p^m (1-p)^n \tag{13}$$

$$\propto p^m (1-p)^n \tag{14}$$

(The binomial coefficient $\binom{N}{m}$ is a constant and can be ignored)

Step 3: Beta Prior

$$P(p) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} p^{a-1} (1-p)^{b-1}$$
(15)

$$\propto p^{a-1} (1-p)^{b-1} \tag{16}$$

(The normalization constant can be ignored)

Step 4: Calculate Posterior

$$P(p|D) \propto P(D|p) \cdot P(p)$$
 (17)

$$\propto [p^{m}(1-p)^{n}] \cdot \left[p^{a-1}(1-p)^{b-1}\right]$$

$$= p^{m+a-1}(1-p)^{n+b-1}$$
(18)

$$= p^{m+a-1}(1-p)^{n+b-1} (19)$$

Step 5: Identify Posterior Distribution

$$P(p|D) \propto p^{a'-1} (1-p)^{b'-1} \tag{20}$$

where:

$$a' = a + m \tag{21}$$

$$b' = b + n \tag{22}$$

Conclusion

The Beta distribution is a **conjugate prior** for the Binomial likelihood. The parameter update rules are:

- a' = a + m (shape parameter increases by number of successes)
- b' = b + n (shape parameter increases by number of failures)

Therefore:

Posterior = Beta
$$(a+m,b+n)$$
 (23)

This is why in online learning (HW2-2), we can directly update Beta parameters without recomputing the entire distribution.

Proof 2: Gamma-Poisson Conjugation

Objective: Show that the Gamma distribution acts as a conjugate prior to the Poisson likelihood, including deriving the posterior distribution.

Problem Setup

• Likelihood: Poisson distribution

$$P(X|\lambda) = \frac{\lambda^x}{x!} e^{-\lambda} \tag{24}$$

• Prior: Gamma distribution

$$P(\lambda) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} \lambda^{\alpha - 1} e^{-\beta \lambda}$$
 (25)

• Goal: Prove that the posterior is also a Gamma distribution

Derivation

Given:

- Data: Observed n events with total count $X = \sum_{i=1}^{n} x_i$
- Prior: Gamma(α, β)

Step 1: Write out Posterior (Bayes' Theorem)

$$P(\lambda|X) \propto P(X|\lambda) \cdot P(\lambda)$$
 (26)

Step 2: Poisson Likelihood (n independent observations)

$$P(X|\lambda) = \prod_{i=1}^{n} \left[\frac{\lambda^{x_i}}{x_i!} e^{-\lambda} \right]$$
 (27)

$$= \left[\prod_{i=1}^{n} \frac{1}{x_i!}\right] \lambda^{\sum x_i} e^{-n\lambda}$$
 (28)

$$\propto \lambda^X e^{-n\lambda}$$
 (29)

(The factorial terms are constants and can be ignored)

Step 3: Gamma Prior

$$P(\lambda) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} \lambda^{\alpha - 1} e^{-\beta \lambda} \tag{30}$$

$$\propto \lambda^{\alpha - 1} e^{-\beta \lambda}$$
 (31)

(The normalization constant can be ignored)

Step 4: Calculate Posterior

$$P(\lambda|X) \propto P(X|\lambda) \cdot P(\lambda)$$
 (32)

$$\propto \left[\lambda^X e^{-n\lambda}\right] \cdot \left[\lambda^{\alpha - 1} e^{-\beta\lambda}\right] \tag{33}$$

$$= \lambda^{X+\alpha-1} e^{-(n+\beta)\lambda} \tag{34}$$

Step 5: Identify Posterior Distribution

$$P(\lambda|X) \propto \lambda^{\alpha'-1} e^{-\beta'\lambda} \tag{35}$$

where:

$$\alpha' = \alpha + X = \alpha + \sum_{i=1}^{n} x_i \tag{36}$$

$$\beta' = \beta + n \tag{37}$$

Conclusion

The Gamma distribution is a **conjugate prior** for the Poisson likelihood. The parameter update rules are:

- $\alpha' = \alpha + \sum x_i$ (shape parameter increases by total observed events)
- $\beta' = \beta + n$ (rate parameter increases by number of observations)

Therefore:

Posterior = Gamma(
$$\alpha + \sum x_i, \beta + n$$
) (38)

Applications

Gamma-Poisson conjugation is commonly used in:

- Event count analysis (e.g., website visits per hour)
- Failure rate estimation (e.g., equipment failures per day)
- Text analysis (e.g., word frequency per document)