

Multivariate Gaussian distribution

NTP

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0.1 1- PROOF NORMALIZED.

$$p(x|\mu, \sigma^2) = \frac{1}{(2\pi)^{\frac{D}{2}} |\Sigma|^{\frac{1}{2}}} \cdot \exp \left\{ \frac{-1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \right\}$$

Đặt

$$\begin{aligned} \Delta^2 &= \frac{-1}{2} (x - \mu)^T \cdot \Sigma^{-1} \cdot (x - \mu) \\ &= \frac{-1}{2} (x^T - \mu^T) \cdot \Sigma^{-1} \cdot (x - \mu) \\ &= \frac{-1}{2} x^T \Sigma^{-1} x + \frac{1}{2} x^T \Sigma^{-1} \mu + \frac{1}{2} \mu^T \Sigma^{-1} x - \frac{1}{2} \mu^T \Sigma^{-1} \mu \end{aligned}$$

Trong đó: *

$$\frac{1}{2} \mu^T \Sigma^{-1} \mu = \text{constant}$$

vì $(1^*D) \cdot (D^*D) \cdot (D^*1) = \text{Constant}$
*

$$\frac{1}{2} x^T \Sigma^{-1} \mu + \frac{1}{2} \mu^T \Sigma^{-1} x$$

Ta có: $(\frac{1}{2} x^T \Sigma^{-1} \mu)^T = \frac{1}{2} \mu^T (\Sigma^{-1})^T x$ (since $(ABC)^T = C^T B^T A^T$)

mà Σ đối xứng nên $\Sigma^{-1} = (\Sigma^{-1})^T = \Sigma$ và $\frac{1}{2} x^T \Sigma^{-1} \mu = \text{constant}$ (do $(1^*D) \cdot (D^*D) \cdot (D^*1)$) Vậy $\Delta = \frac{-1}{2} x^T \Sigma^{-1} x + x^T \Sigma^{-1} \mu + \text{constant}$

Ta có: Σ Due to symmetric :

$$\Sigma = \sum_{i=1}^D \lambda_i u_i u_i^T \Rightarrow \Sigma^{-1} = \sum_{i=1}^D \frac{1}{\lambda_i} u_i u_i^T$$

Proof: vì u_i là các eigenvectors nên theo tính chất $u_i^T \cdot u_i = 1 \Rightarrow \sum_{i=1}^D u_i \cdot u_i^T = I$ and: Σ đối xứng nên các eigenvalues sẽ là số thực. and: $\Sigma^T = \Sigma^{-1}$

$$\Delta^2 = \frac{-1}{2} (x - \mu)^T \cdot \Sigma^{-1} \cdot (x - \mu)$$

$$(1) = \frac{-1}{2} (x - \mu)^T \sum_{i=1}^D \frac{1}{\lambda_i} u_i u_i^T (x - \mu)$$

Set: $y_i = u_i^T (x - \mu)$ Since y_i có giá trị là một số thực (do $(1^*D)(D^*1)$ nên $y_i = y_i^T = (u_i^T (x - \mu))^T = (x - \mu)^T u_i$)

$$\begin{aligned} (1) &\Leftrightarrow \Delta^2 = \frac{-1}{2} \sum_{i=1}^D \frac{1}{\lambda_i} (x - \mu)^T u_i u_i^T (x - \mu) \\ &= \sum_{i=1}^D \frac{-y_i^2}{2\lambda_i} \end{aligned}$$

$$|\Sigma|^{\frac{1}{2}} = \prod_{j=1}^D \lambda_j^{\frac{1}{2}}$$

(trong đó λ_j là các eigenvalues và $\det(A)$ bằng tích các eigenvalues)

$$\begin{aligned} p(y) &= \frac{1}{(2\pi)^{\frac{D}{2}} \prod_{j=1}^D \lambda_j^{\frac{1}{2}}} \cdot e^{\sum_{j=1}^D \frac{-y_j^2}{2\lambda_j}} \\ &= \prod_{j=1}^D \left(\frac{1}{2\pi\lambda_j} \right)^{\frac{1}{2}} e^{\sum_{j=1}^D \frac{-y_j^2}{2\lambda_j}} \\ &= \prod_{j=1}^D \left(\frac{1}{2\pi\lambda_j} \right)^{\frac{1}{2}} \cdot \prod_{j=1}^D e^{\frac{-y_j^2}{2\lambda_j}} \\ &= \prod_{j=1}^D \left(\frac{1}{2\pi\lambda_j} \right)^{\frac{1}{2}} e^{\frac{-y_j^2}{2\lambda_j}} \\ \int_{-\infty}^{\infty} p(y) dy &= \prod_{j=1}^D \int_{-\infty}^{\infty} \frac{1}{(2\pi\lambda_j)^{\frac{1}{2}}} \cdot e^{\frac{-y_j^2}{2\lambda_j}} dy_j \end{aligned}$$

+) pick $j = 1$ and set $\lambda = \sigma^2 \rightarrow \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2}} \cdot e^{\frac{-y^2}{2\sigma^2}} = 1$. Thus:

$$\int_{-\infty}^{\infty} p(y) dy = 1 \Rightarrow \text{Normality}$$

0.2 PROOF Conditional Gauss distribution

Set $x = \begin{pmatrix} x_a \\ x_b \end{pmatrix}$ and $\mu = \begin{pmatrix} \mu_a \\ \mu_b \end{pmatrix}$ and $\Sigma = \begin{bmatrix} \Sigma_{aa} & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_{bb} \end{bmatrix}$ Đặt: matrix $A = \Sigma^{-1} = \begin{bmatrix} ccA_{aa} & A_{ab} \\ A_{ba} & A_{bb} \end{bmatrix}$

Since Σ đối xứng nên Σ_{aa} và Σ_{bb} cũng đối xứng Tìm $p(x_a|x_b)$ Ta có:

$$\begin{aligned} \Delta^2 &= \frac{-1}{2} (x - \mu)^T \cdot \Sigma^{-1} \cdot (x - \mu) \\ &= \frac{-1}{2} \cdot \begin{pmatrix} x_a - \mu_a \\ x_b - \mu_b \end{pmatrix}^T \cdot \begin{pmatrix} A_{aa} & A_{ab} \\ A_{ba} & A_{bb} \end{pmatrix} \cdot \begin{pmatrix} x_a - \mu_a \\ x_b - \mu_b \end{pmatrix} \end{aligned}$$

Set: $(x_a - \mu_a)^T = C_1$ $(x_b - \mu_b)^T = C_2$ $(x_a - \mu_a) = D_1$ $(x_b - \mu_b) = D_2$ thay vào Δ^2 có:

$$\begin{aligned} &= \frac{-1}{2} \cdot \begin{pmatrix} C_1 \\ C_2 \end{pmatrix}^T \cdot \begin{pmatrix} A_{aa} & A_{ab} \\ A_{ba} & A_{bb} \end{pmatrix} \cdot \begin{pmatrix} D_1 \\ D_2 \end{pmatrix} \\ &= \frac{-1}{2} ((C_1 A_{aa} + C_2 A_{ba} \quad C_1 A_{ab} + C_2 A_{bb}) \cdot \begin{pmatrix} D_1 \\ D_2 \end{pmatrix}) \\ &= \frac{-1}{2} [(C_1 A_{aa} + C_2 A_{ba}) \cdot D_1 + (C_1 A_{ab} + C_2 A_{bb}) \cdot D_2] \\ &= \frac{-1}{2} (C_1 A_{aa} D_1 + C_2 A_{ba} D_1 + C_1 A_{ab} D_2 + C_2 A_{bb} D_2) \end{aligned}$$

$$\Delta^2 = \frac{-1}{2} (x_a - \mu_a)^T A_{aa} (x_a - \mu_a) - \frac{1}{2} (x_b - \mu_b)^T A_{ba} (x_a - \mu_a) - \frac{1}{2} (x_a - \mu_a)^T A_{ab} (x_b - \mu_b) - \frac{1}{2} (x_b - \mu_b)^T A_{bb} (x_b - \mu_b)$$

trong đó; $-\frac{1}{2} (x_b - \mu_b)^T A_{bb} (x_b - \mu_b) = \text{constant}$ do $(1^*b) \cdot (b^*b) \cdot (b^*1)$ thuộc \mathbb{R} $\frac{1}{2} (x_b - \mu_b)^T A_{ba} (x_a - \mu_a) = \frac{1}{2} (x_a - \mu_a)^T A_{ab} (x_b - \mu_b)$ do đây là 2 ma trận chuyển vị. Vậy :

$$\Delta^2 = \frac{-1}{2} (x_a - \mu_a)^T A_{aa} (x_a - \mu_a) - (x_a - \mu_a)^T A_{ab} (x_b - \mu_b) + \text{constant}$$

$$= \frac{-1}{2}x_a^T A_{aa}x_a + \frac{1}{2}x_a^T A_{aa}\mu_a + \frac{1}{2}\mu_a^T A_{aa}x_a - \frac{1}{2}\mu_a^T A_{aa}\mu_a - (x_a - \mu_a)^T A_{ab}(x_b - \mu_b) + constant$$

trong đó: $\frac{1}{2}\mu_a^T A_{aa}x_a = \frac{1}{2}x_a^T A_{aa}\mu_a$ and $\frac{1}{2}\mu_a^T A_{aa}\mu_a = constant$

$$\begin{aligned}\Delta^2 &= \frac{-1}{2}x_a^T A_{aa}x_a + x_a^T A_{aa}\mu_a - (x_a - \mu_a)^T A_{ab}(x_b - \mu_b) + constant \\ &= \frac{-1}{2}x_a^T A_{aa}x_a + x_a^T A_{aa}\mu_a - x_a^T A_{ab}(x_b - \mu_b) + \mu_a^T A_{ab}(x_b - \mu_b) + constant \\ &= \frac{-1}{2}x_a^T A_{aa}x_a + x_a^T \cdot (A_{aa}\mu_a - A_{ab}(x_b - \mu_b)) + constant\end{aligned}$$

Compare with Gaussian distribution: $\Delta = \frac{-1}{2}x^T \Sigma^{-1}x + x^T \Sigma^{-1}\mu + constant$

$$\begin{aligned}A_{aa} &= \Sigma^{-1} \Rightarrow \Sigma_{a|b} = A_{aa}^{-1} \\ \Sigma^{-1}\mu &= A_{aa}\mu_a - A_{ab}(x_b - \mu_b)\end{aligned}$$

Nhân 2 vế với $\Sigma = A_{aa}^{-1}$ ta được:

$$\Sigma \Sigma^{-1} \mu_{a|b} = A_{aa}^{-1} A_{aa} \mu_a - A_{aa}^{-1} A_{ab}(x_b - \mu_b)$$

$$\mu_{a|b} = \mu_a - A_{aa}^{-1} A_{ab}(x_b - \mu_b)$$

By using Schur complement:

$$\begin{aligned}A_{aa} &= (\Sigma_{aa} - \Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba})^{-1} \\ A_{ab} &= -(\Sigma_{aa} - \Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba})^{-1}\Sigma_{ab}\Sigma_{bb}^{-1}\end{aligned}$$

Thus: $\Sigma_{a|b} = A_{aa}^{-1} = \Sigma_{aa} - \Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba}$

$$\mu_{a|b} = \mu_a - (\Sigma_{aa} - \Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba})(-(\Sigma_{aa} - \Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba})^{-1}\Sigma_{ab}\Sigma_{bb}^{-1}(x_b - \mu_b))$$

$$\mu_{a|b} = \mu_a + \Sigma_{ab}\Sigma_{bb}^{-1}(x_b - \mu_b)$$

$$p(x_a|x_b) = N(x_{a|b}|\mu_{a|b}, \Sigma_{a|b})$$