

# Tutorial 4 (STAT8019)

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**Differentiate different methods:**

(1) Factor Analysis and Linear Regression

**Orthogonal factor model:**

$$\mathbf{X} = \boldsymbol{\mu} + \mathbf{L}\mathbf{F} + \boldsymbol{\varepsilon}, \text{Cov}(\boldsymbol{\varepsilon}) = \boldsymbol{\Psi}$$

Goal: Try to figure out the structure of the distribution of  $\mathbf{X}$ ,  $P(\mathbf{X}) = P(\mathbf{X}|\boldsymbol{\mu}, \mathbf{L}, \boldsymbol{\Psi})$ .

**Regression model:**

$$Y = \beta_0 + \boldsymbol{\beta}^T \mathbf{x} + \varepsilon, \text{Var}(\varepsilon) = \sigma^2$$

Goal: Try to figure out the linear relationship between  $Y$  and  $\mathbf{X}$  based on the conditional distribution of  $Y$  given  $\mathbf{X} = \mathbf{x}$ ,  $P(Y|\mathbf{x}) = P(Y|\mathbf{x}, \beta_0, \boldsymbol{\beta}, \sigma^2)$ .

(2) Principal Component Analysis and Probabilistic PCA

**PCA** is a model-free method based on the following compression-reconstruction linear transformations:

$$\mathbf{Z} = \mathbf{D}\mathbf{L}, \mathbf{D} \approx \mathbf{D}^* = \mathbf{Z}\mathbf{L}^T$$

where  $\mathbf{D} \in R^{n \times p}$ ,  $\mathbf{Z} \in R^{n \times m}$ ,  $\mathbf{L} \in R^{p \times m}$ ,  $\mathbf{L}^T \mathbf{L} = \mathbf{I}$ .

Goal: Try to find  $\mathbf{L}$ , or more generally, a compression-reconstruction method, so that we loss the least information in reconstructed data  $\mathbf{D}^*$ .

**Probabilistic PCA model:**

$$\mathbf{X} = \boldsymbol{\mu} + \mathbf{L}\mathbf{F} + \boldsymbol{\varepsilon}, \text{Cov}(\boldsymbol{\varepsilon}) = \sigma^2 \mathbf{I}$$

It's different from PCA.

(3) Principal Component Method and Maximum Likelihood Method

They are two estimation methods for factor model.

**Principal component method** is based on eigen-decomposition :

$$\boldsymbol{\Sigma} = \mathbf{V}\mathbf{D}\mathbf{V}^T = \mathbf{V}\mathbf{D}^{1/2}\mathbf{D}^{1/2}\mathbf{V}^T \approx \mathbf{L}\mathbf{L}^T.$$

**Maximum likelihood method** added a new assumption to the factor model:  $\mathbf{F}$  and  $\boldsymbol{\varepsilon}$  are jointly normal, and so is  $\mathbf{X}$ . And it estimate  $\mathbf{L}$  and  $\boldsymbol{\Psi}$  based on MLE.

**Q1:**

a).

Assume

$$\mathbf{Z}^* = \begin{bmatrix} Z_1^* \\ Z_2^* \\ Z_3^* \end{bmatrix} = \begin{bmatrix} 0.9 \\ 0.7 \\ 0.5 \end{bmatrix} F_1 + \begin{bmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \varepsilon_3 \end{bmatrix} = \mathbf{L}F_1 + \boldsymbol{\varepsilon},$$

here we try to prove  $Cov(\mathbf{Z}^*) = Cov(\mathbf{Z}) = \boldsymbol{\Sigma}$ .

From the independence assumption between  $F_1$  and  $\boldsymbol{\varepsilon}$ , we can see that  $Cov(\mathbf{Z}^*) = \mathbf{L}Var(F_1)\mathbf{L}^T + Cov(\boldsymbol{\varepsilon}) = \mathbf{L}\mathbf{L}^T + \boldsymbol{\Psi}$ .

So we only need to check  $\mathbf{L}\mathbf{L}^T + \boldsymbol{\Psi} = \boldsymbol{\Sigma}$ .

$$\begin{bmatrix} 0.9 \\ 0.7 \\ 0.5 \end{bmatrix} \begin{bmatrix} 0.9 & 0.7 & 0.5 \end{bmatrix} + \begin{bmatrix} 0.19 & 0 & 0 \\ 0 & 0.51 & 0 \\ 0 & 0 & 0.75 \end{bmatrix} = \begin{bmatrix} 0.81 & 0.63 & 0.45 \\ 0.63 & 0.49 & 0.35 \\ 0.45 & 0.35 & 0.25 \end{bmatrix} + \begin{bmatrix} 0.19 & 0 & 0 \\ 0 & 0.51 & 0 \\ 0 & 0 & 0.75 \end{bmatrix} = \begin{bmatrix} 1 & 0.63 & 0.45 \\ 0.63 & 1 & 0.35 \\ 0.45 & 0.35 & 1 \end{bmatrix} = \boldsymbol{\Sigma}$$

b).

Remember the definition of communality: variance of the  $i$ th variable contributed by all the common factors.

So here we have  $h_i^2 = l_{i1}^2$ ,  $i = 1, 2, 3$ .

	$Z_1$	$Z_2$	$Z_3$
$h_i^2$	0.81	0.49	0.25
$h_i^2/Var(Z_i)$	81%	49%	25%

c).

Here we try to calculate  $Cor(Z_i, F_1) = Cov(Z_i, F_1) / \sqrt{Var(Z_i)Var(F_1)}$ ,  $i = 1, 2, 3$ .

Refer to  $\boldsymbol{\Sigma}$  and the definition of factor, we know  $Cor(Z_i, F_1) = Cov(Z_i, F_1) = Cov(l_{i1}F_1 + \varepsilon_i, F_1) = l_{i1}Var(F_1) = l_{i1}$ . (Normally,  $l_{i1}/Std(Z_i)$ )

	$Z_1$	$Z_2$	$Z_3$
$Cor(Z_i, F_1)$	0.9	0.7	0.5

Since  $Z_1$  has the greatest correlation with  $F_1$ , it carries the greatest weight in ‘naming’ the common factor. (When we want to ‘name’ or interpret a factor, we mainly count on the variables with the biggest positive or negative correlations with the factor.)

**Q2:**

a).

Remember the factor model:

$$\mathbf{X}_{p \times 1} - \boldsymbol{\mu}_{p \times 1} = \mathbf{L}_{p \times m} \mathbf{F}_{m \times 1} + \boldsymbol{\varepsilon}_{p \times 1},$$

Normally we assume  $m \leq p$  and the loading matrix  $\mathbf{L}$  has rank  $m$ . Then by manipulating the above equation, we have:

$$\begin{aligned}\boldsymbol{\varepsilon} &= \mathbf{X} - \boldsymbol{\mu} - \mathbf{L}\mathbf{F}, \\ \mathbf{F} &= (\mathbf{L}^T \mathbf{L})^{-1} \mathbf{L}^T (\mathbf{X} - \boldsymbol{\mu} - \boldsymbol{\varepsilon}),\end{aligned}$$

So the existence of linear combination of the attributes for  $\mathbf{F}$  is equivalent to the existence of that for  $\boldsymbol{\varepsilon}$ . So we only need to check whether we can get  $\boldsymbol{\varepsilon} = \mathbf{A}\mathbf{X} + \mathbf{b}$ , which is equivalent to check the properties of  $\boldsymbol{\varepsilon}$ :

$$\begin{aligned}E[\boldsymbol{\varepsilon}] &= \mathbf{A}\boldsymbol{\mu} + \mathbf{b} = \mathbf{0}, \\ Cov(\boldsymbol{\varepsilon}) &= \mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^T = \boldsymbol{\Psi}.\end{aligned}$$

Because  $\boldsymbol{\Sigma}$  and  $\boldsymbol{\Psi}$  are both positive semidefinite,  $\mathbf{A}$  and  $\mathbf{b}$  exist if and only if  $\text{Rank}(\boldsymbol{\Sigma}) \geq \text{Rank}(\boldsymbol{\Psi})$ . And this condition almost always holds because  $\boldsymbol{\Sigma}$  is normally nonsingular.

b).

Check the definition of rotation:  $\mathbf{L}^* = \mathbf{L}\mathbf{T}$ , where  $\mathbf{T}\mathbf{T}^T = \mathbf{I}$ . As the communality  $\mathbf{h}^{*2} = \text{Diag}(\mathbf{L}^* \mathbf{L}^{*T}) = \text{Diag}(\mathbf{L} \mathbf{T} \mathbf{T}^T \mathbf{L}^T) = \text{Diag}(\mathbf{L} \mathbf{L}^T) = \mathbf{h}^2$  didn't change before and after rotation, we have the following equations:

$$\begin{aligned}h_1^2 &= 0.976^2 + 0.139^2 = 0.025^2 + l_{12}^{*2}, \\ h_3^2 &= l_{31}^2 + 0.032^2 = 0.131^2 + 0.972^2, \\ h_4^2 &= 0.535^2 + 0.739^2 = l_{41}^{*2} + 0.405^2.\end{aligned}$$

So, given they are all positive,  $l_{12}^* = 0.986$ ,  $l_{31} = 0.980$ ,  $l_{41}^* = 0.818$ .

c).

Since standardized data were adopted, the variability in Flavour  $\sigma_3^2 = 1$ , and the proportion of variance of Flavour cannot be explained by the common factors should be  $\psi_3/\sigma_3^2 = (1 - h_3^2)/1 = 3.8\%$ .

d).

Proportion of total variances explained by the first factor:

$$\frac{\sum_{i=1}^5 l_{i1}^2}{5} = \frac{0.976^2 + 0.15^2 + 0.98^2 + 0.535^2 + 0.146^2}{5} = \frac{2.244}{5} = 44.8\%.$$

Proportion of total variances explained by the 2 factors:

$$\frac{\sum_{j=1}^2 \sum_{i=1}^5 l_{ij}^2}{5} = 44.8\% + \frac{0.139^2 + 0.86^2 + 0.032^2 + 0.739^2 + 0.963^2}{5} = 89.5\%.$$

e).

	Taste	Money's worth	Flavour	Good for snack	Nutrition
F1	0.025	0.873	0.131	0.818	0.974
F2	0.986	0.005	0.972	0.405	-0.016

Use 0.5 or 0.6 as threshold to determine whether a correlation is large or small.

Factor 1 represents Money's worthiness, Good for snack and Nutrition. Factor 2 represents Taste and Flavour.

f).

$H_0: \Sigma = \mathbf{L}\mathbf{L}^T + \Psi$  vs  $H_1: \Sigma =$  any other positive definite matrix.

Likelihood ratio test statistic  $R = n \ln(|\hat{\mathbf{L}}\hat{\mathbf{L}}^T + \hat{\Psi}|/|\mathbf{S}_n|) = 150 \ln(3.56/3.48) = 3.409$ .

Critical value  $\chi^2_{\alpha, [(p-m)^2 - p - m]/2} = \chi^2_{0.05, [(5-2)^2 - 5 - 2]/2} = \chi^2_{0.05, 1} = 3.8415$ .

As  $R < \chi^2_{0.05, 1}$ , we do not reject  $H_0$  at the 5% level of significance, i.e. the 2-factor model is adequate.