

1

Given:

$$f(y; \theta) = \frac{y^{\phi-1} \theta^\phi e^{-y\theta}}{\Gamma(\phi)} \quad (\phi \text{ known}) \quad [1]$$

$$f(y; \theta) = \binom{y-r+1}{r-1} \theta^r (1-\theta)^y \quad (r \text{ known}) \quad [2]$$

1.a

[1]:

Likelihood function:

$$\begin{aligned} L(\theta; y_1, \dots, y_n) &= \prod_{i=1}^n \frac{y_i^{\phi-1} \theta^\phi e^{-y_i \theta}}{\Gamma(\phi)} \\ &= \frac{(\prod_{i=1}^n y_i)^{\phi-1} \theta^{n\phi} \exp(-\sum_{i=1}^n y_i \theta)}{\Gamma(\phi)^n} \end{aligned}$$

Log-likelihood:

$$\begin{aligned} l(\theta; y_1, \dots, y_n) &= \ln(L(\theta; y_1, \dots, y_n)) \\ &= (\phi-1) \sum_{i=1}^n \ln(y_i) + n\phi \ln(\theta) - \sum_{i=1}^n y_i \theta - n \ln \Gamma(\phi) \end{aligned}$$

Taking the derivative of the log-likelihood:

$$\begin{aligned} U(\theta) &= \frac{dl}{d\theta} \\ &= \frac{n\phi}{\theta} - \sum_{i=1}^n y_i \end{aligned}$$

For MLE of θ , $\hat{\theta}$:

$$\begin{aligned} \frac{dl}{d\theta} &= 0 \\ \frac{n\phi}{\theta} - \sum_{i=1}^n y_i &= 0 \\ \frac{\theta}{n\phi} &= \frac{1}{\sum_{i=1}^n y_i} \\ \theta &= \frac{n\phi}{\sum_{i=1}^n y_i} \\ \therefore \hat{\theta} &= \frac{n\phi}{\sum_{i=1}^n y_i} \end{aligned}$$

[2]:

Likelihood function:

$$\begin{aligned} L(\theta; y_1, \dots, y_n) &= \prod_{i=1}^n \left[\binom{y_i - r + 1}{r - 1} \theta^r (1 - \theta)^{y_i} \right] \\ &= \left[\prod_{i=1}^n \binom{y_i - r + 1}{r - 1} \right] \theta^{nr} (1 - \theta)^{\sum_{i=1}^n y_i} \end{aligned}$$

Log-likelihood:

$$\begin{aligned} l(\theta; y_1, \dots, y_n) &= \ln(L(\theta; y_1, \dots, y_n)) \\ &= \text{const.} + nr \ln \theta + \sum_{i=1}^n y_i \ln(1 - \theta) \end{aligned}$$

Taking the derivative of the log-likelihood:

$$\begin{aligned} U(\theta) &= \frac{dl}{d\theta} \\ &= \frac{nr}{\theta} - \frac{\sum_{i=1}^n y_i}{1 - \theta} \end{aligned} \tag{1}$$

For MLE of θ , $\hat{\theta}$:

$$\begin{aligned} \frac{dl}{d\theta} &= 0 \\ \frac{nr}{\theta} - \frac{\sum_{i=1}^n y_i}{1 - \theta} &= 0 \\ (1 - \theta)nr - \theta \sum_{i=1}^n y_i &= 0 \\ \theta \left(nr + \sum_{i=1}^n y_i \right) &= nr \\ \theta &= \frac{nr}{nr + \sum_{i=1}^n y_i} \\ \therefore \hat{\theta} &= \frac{nr}{nr + \sum_{i=1}^n y_i} \end{aligned}$$

1.b

[1]:

Given:

$$\begin{aligned}U(\theta) &= \frac{n\phi}{\theta} - \sum_{i=1}^n y_i \\U'(\theta) &= \frac{-n\phi}{\theta^2}\end{aligned}\tag{2}$$

The Fisher information:

$$\begin{aligned}I_\theta &= -\mathbb{E}(U') \\&= -\mathbb{E}\left(\frac{-n\phi}{\theta^2}\right) \\&= \frac{n\phi}{\theta^2} \\\therefore \text{Var}(\hat{\theta}) &= I_\theta^{-1} \\&= \frac{\theta^2}{n\phi}\end{aligned}$$

[2]:

Given:

$$\begin{aligned}U(\theta) &= \frac{nr}{\theta} - \frac{\sum_{i=1}^n y_i}{1-\theta} \\U'(\theta) &= \frac{-nr}{\theta^2} - \frac{\sum_{i=1}^n y_i}{(1-\theta)^2}\end{aligned}$$

For Fisher information:

$$\begin{aligned}I_\theta &= -\mathbb{E}(U') \\&= -\mathbb{E}\left(\frac{-nr}{\theta^2} - \frac{\sum_{i=1}^n y_i}{(1-\theta)^2}\right) \\&= \frac{nr}{\theta^2} + \frac{\sum_{i=1}^n \mathbb{E}(y_i)}{(1-\theta)^2} \\&= \frac{nr}{\theta^2} + \frac{\sum_{i=1}^n \theta r}{(1-\theta)^3} \\&= \frac{nr}{\theta^2} + \frac{nr\theta}{(1-\theta)^3} \\&= \frac{nr(1-\theta)^3 + nr\theta^3}{\theta^2(1-\theta)^3} \\\therefore \text{Var}(\hat{\theta}) &= I_\theta^{-1} \\&= \frac{\theta^2(1-\theta)^3}{nr(1-\theta)^3 + nr\theta^3}\end{aligned}$$

2

Given:

$$f(y; \theta) = \theta e^{-y\theta}$$

Likelihood:

$$\begin{aligned} L(\theta; y_1, \dots, y_n) &= \prod_{i=1}^n [\theta e^{-y_i \theta}] \\ &= \theta^n \exp \left(-\theta \sum_{i=1}^n y_i \right) \end{aligned}$$

Log-likelihood:

$$\begin{aligned} l(\theta; y_1, \dots, y_n) &= \ln(L(\theta; y_1, \dots, y_n)) \\ &= n \ln \theta - \theta \sum_{i=1}^n y_i \end{aligned}$$

Score:

$$\begin{aligned} U(\theta) &= \frac{dl}{d\theta} \\ &= \frac{n}{\theta} - \sum_{i=1}^n y_i \end{aligned}$$

Claim: $\mathbb{E}(U) = 0$

Proof:

$$\begin{aligned} \mathbb{E}(U) &= \mathbb{E} \left(\frac{n}{\theta} - \sum_{i=1}^n y_i \right) \\ &= \frac{n}{\theta} - \sum_{i=1}^n \mathbb{E}(y_i) \\ &= \frac{n}{\theta} - \sum_{i=1}^n \frac{1}{\theta} \\ &= \frac{n}{\theta} - \frac{n}{\theta} \\ &= 0 \\ \therefore \mathbb{E}(U) &= 0 \end{aligned}$$

Claim: $Var(U) = \mathbb{E}(U^2) = -\mathbb{E}(U')$

Proof:

$$\begin{aligned} Var(U) &= Var\left(\frac{n}{\theta} - \sum_{i=1}^n y_i\right) \\ &= Var\left(\sum_{i=1}^n y_i\right) \\ &= \sum_{i=1}^n Var(y_i) \\ &= \sum_{i=1}^n \frac{1}{\theta^2} \\ &= \frac{n}{\theta^2} \end{aligned}$$

$$\begin{aligned} \mathbb{E}(U^2) &= Var(U) + \mathbb{E}(U)^2 \\ &= \frac{n}{\theta^2} - 0^2 \\ &= \frac{n}{\theta^2} \end{aligned}$$

$$\begin{aligned} -\mathbb{E}(U') &= -\mathbb{E}\left(\frac{-n}{\theta^2}\right) \\ &= \frac{n}{\theta^2} \end{aligned}$$

$$\therefore Var(U) = \mathbb{E}(U^2) = -\mathbb{E}(U')$$