

Fiscal Rule Stretching in Europe During Financial Market Stress and Crises

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Abstract

Elected governments have incentives to stretch accounting rules to classify loss-making and/or indebted endeavors, such as public industries and pension schemes, as off of the public balance sheet. Doing so improves the appearance of the government to cost-conscious voters as fiscally restrained in the near-term. We expect rule stretching to be especially prevalent during financial crises given their typical expense and the type of policy options available to assist troubled financial institutions, such as bad banks nationalization are often hard to classify as in or outside of the government sector. To test this proposition, we examine revisions that to government debt and deficit figures made by the European statistical agency—Eurostat. These revisions frequently occur because this politically independent agency re-classifies organizations as within the government sector, when a national government had originally classified it as off. We find that debt figures are more likely to be revised upward for years close to national elections. This effect is strengthened further by financial market stress.

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1 Introduction

2 Political budget cycles and fiscal gimmicks

De Castro, Pérez and Rodríguez-Vives (2013) Alt, Lassen and Wehner (2014)

3 Cost-shifting during financial crises

Gandrud and Hallerberg (2016)

4 Hypotheses

H_1 : Debt revisions will be greater for years closer to national government elections.

H_2 : Debt revisions will be greater for years when there are required (non-endogenous) elections.

H_3 : The effects predicted by H_1 and H_2 will be stronger when a country also has high credit provision stress.

5 Empirical tests

5.1 Eurostat revisions

To test these hypotheses we gathered all of the revisions that Eurostat made to EU member debt and deficit figures from 2003 through 2013.¹ Eurostat publishes revisions bi-annually—typically once at the end of April and again again in late October. These revisions cover government finance statistics released within the previous four years. As such, the unit of analysis in the following regressions is biannual revision point for each fiscal year. For every year that government statistics are revised, there are up to seven revision points.²

We created a variable of *cumulative revisions for debts and deficits* over these four year periods and used them as our dependent variables. For debt revisions, it ranged from -1.1 and 12.7 percent of GDP. For deficit revisions it ranged from -4.5 and 1.1 percent of GDP. It is important to note that all of these revisions were not due to revisions in the denominator, i.e. GDP. Eurostat reports these types of changes separately. As De Castro, Pérez and Rodríguez-Vives (2013) found for similar Eurostat data in shorter pre-financial crisis

¹PDF files with the figures were downloaded from <http://ec.europa.eu/eurostat/news/news-releases>.

²The first revisions occur October of the initial reporting year.

sample, there is a clear tendency for debt statistics to be revised upward and deficit statistics downward, indicating that policies are more expensive than initially reported.

5.2 Right-hand variables

To test whether governments are more likely to stretch the rules closer to elections, we use Gandrud’s (2015) *years to election* variable. The election year is recorded as zero. Not only do we expect that governments are more likely to use rule stretching as elections approach, but that rule stretching should be more prevalent when governments are not able to choose the when the election occurs so as to present themselves in the best fiscal light to voters. As such, we also include a dummy variable that is one for *required elections* (e.g. non-endogenous elections) and zero for all other years. The variable is from Brender and Drazen (2008). It was updated and corrected by Hallerberg and Wehner. We expect that that revisions will be greater for years when there is a non-endogenous election.

To examine how responding to financial market stress may exacerbate governments’ fiscal accounting rule stretching behavior, we included Gandrud and Hallerberg’s (2015) continuous “FinStress” measure of real-time perceptions of financial market stress. This measure is created by analyzing monthly content from Economist Intelligence Unit texts on banking and financial markets using a statistical method called kernel principal component analysis. Unlike previous post-hoc dichotomous measures of financial crisis (e.g. measures compiled by Laeven and Valencia, 2012; Reinhart and Rogoff, 2010), the authors argue this measure captures what we are most interested in when trying to understand policy choices: what stress policy-makers perceived and so responded to at the time. It also does not rely on ad hoc methods of determining when a crisis ended, but instead charts its perceived intensity over time. To make this measure comparable with our other variables, we found yearly averages. FinStress is able to vary between zero and one, with higher values indicating more perceived financial market stress. In the sample it varies between 0.12 and 0.76.

Because hypothesize that the effect of election timing on rule-stretching will increase at higher levels of financial market stress, we will focus on interacting FinStress with the election timing and non-endogenous elections variables.

As Eurostat has more time to examine member state government policies, e.g. how separate they are in practice from the government sector, we would expect on average that the cumulative revisions will grow over the course of the four year period during which they are revised. So we also include a variable counting the years since the original year that the revised figure is from on the right-hand side.

Table 1: Linear Regression Estimation of Debt Revisions

| | <i>Dependent variable:</i> | | | | |
|-----------------------------|----------------------------|---------------------|----------------------|----------------------|----------------------|
| | Cumulative Debt Revisions | | | | |
| | (1) | (2) | (3) | (4) | (5) |
| Yrs. Since Original | 0.218*** (0.040) | 0.175*** (0.038) | 0.219*** (0.040) | 0.217*** (0.040) | 0.174*** (0.038) |
| Yrs. to Election | -0.119*** (0.032) | | -0.119*** (0.032) | 0.296** (0.141) | |
| Required Election | | 0.733*** (0.181) | | | -3.638*** (0.773) |
| FinStress | | | 0.081 (0.379) | 1.411** (0.581) | -1.225*** (0.378) |
| Yrs. to Elect. * FinStress | | | | -0.875*** (0.290) | |
| Required Elect. * FinStress | | | | | 9.264*** (1.603) |
| Constant | 1.081*** (0.231) | 0.418* (0.214) | 1.038*** (0.305) | 0.485 (0.355) | 1.125*** (0.291) |
| Country FE? | Yes | Yes | Yes | | |
| Observations | 1,430 | 1,263 | 1,430 | 1,430 | 1,263 |
| R ² | 0.249 | 0.287 | 0.249 | 0.254 | 0.308 |
| Adjusted R ² | 0.233 | 0.271 | 0.233 | 0.237 | 0.291 |
| Residual Std. Error | 1.508 (df = 1400) | 1.367 (df = 1234) | 1.508 (df = 1399) | 1.504 (df = 1398) | 1.348 (df = 1232) |

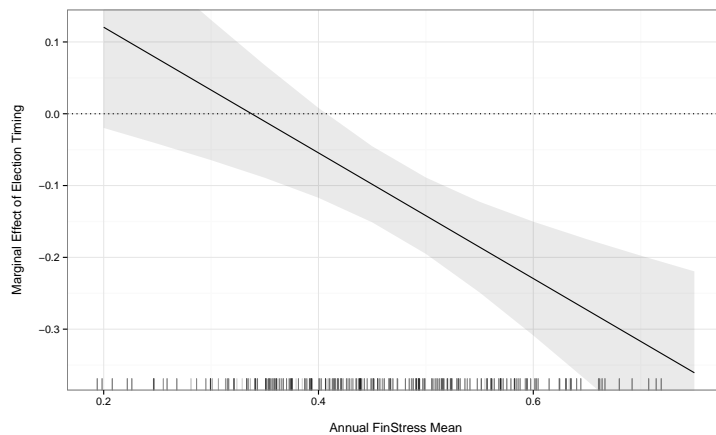
Note: *p<0.1; **p<0.05; ***p<0.01

Table 2: Linear Regression Estimation of Deficit Revisions

| | <i>Dependent variable:</i> | | | | |
|-----------------------------|------------------------------|----------------------|----------------------|----------------------|----------------------|
| | Cumulative Deficit Revisions | | | | |
| | (1) | (2) | (3) | (4) | (5) |
| Yrs. Since Original | -0.058*** (0.017) | -0.042*** (0.015) | -0.057*** (0.017) | -0.057*** (0.017) | -0.040*** (0.015) |
| Yrs. to Election | 0.035** (0.013) | | 0.035** (0.014) | 0.042 (0.061) | |
| Required Election | | -0.042 (0.071) | | | 0.155 (0.306) |
| FinStress | | | 0.020 (0.162) | 0.043 (0.249) | 0.298** (0.150) |
| Yrs. to Elect. * FinStress | | | | -0.015 (0.125) | |
| Required Elect. * FinStress | | | | | -0.402 (0.635) |
| Constant | -0.246** (0.099) | -0.235*** (0.084) | -0.256** (0.130) | -0.266* (0.152) | -0.394*** (0.115) |
| Country FE? | Yes | Yes | Yes | | |
| Observations | 1,430 | 1,263 | 1,430 | 1,430 | 1,263 |
| R ² | 0.217 | 0.289 | 0.217 | 0.217 | 0.292 |
| Adjusted R ² | 0.201 | 0.273 | 0.200 | 0.200 | 0.274 |
| Residual Std. Error | 0.645 (df = 1400) | 0.535 (df = 1234) | 0.645 (df = 1399) | 0.645 (df = 1398) | 0.534 (df = 1232) |

Note: *p<0.1; **p<0.05; ***p<0.01

Figure 1: Marginal Effect of Election Timing (years to election) at Various Levels of Financial Market Stress on Debt Revisions



Shaded area represents 95% confidence interval.

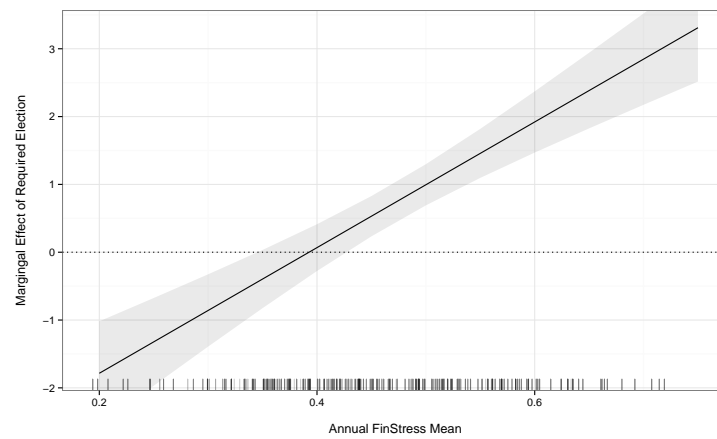
5.3 Results

Tables 1 and 2 showing results from linear regressions with cumulative debt and deficit revisions, respectively, as the dependent variable. All models include country fixed effects.

We can see in Table 2 that election timing is estimated to have a negative effect on debt revisions in models where it is not interacted with FinStress. As the election timing variable is larger, the further away the next election is, this means that government debt figures for years closer to elections tend to be revised more. This finding is similar in direction and magnitude to what De Castro, Pérez and Rodríguez-Vives (2013) found with similar data over a shorter time-span.

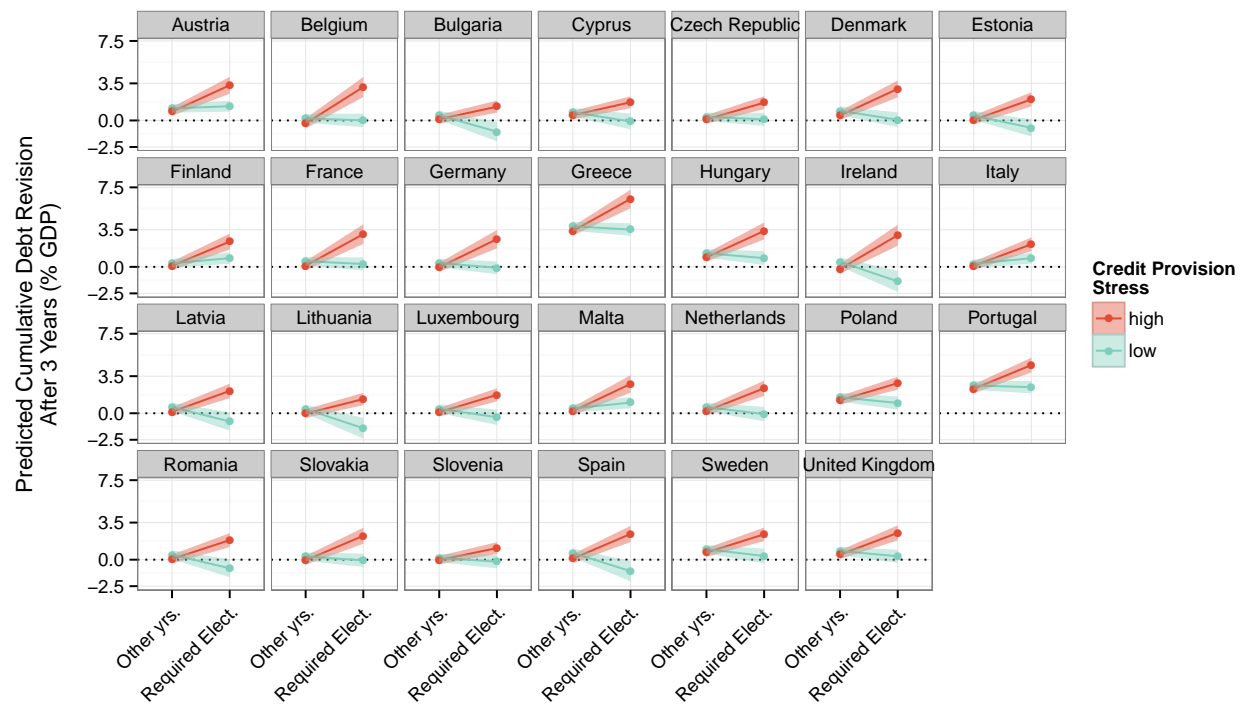
6 Conclusion

Figure 2: Marginal Effect of Non-Endogenous Elections at Various Levels of Financial Market Stress on Debt Revisions



Shaded area represents 95% confidence interval.

Figure 3: Predicted Debt Revisions in Three Years After Publication for Years with Non-Endogenous Elections vs. All Other Years



High and Low stress values refer to country minimum and maximum FinStress scores in the sample.

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