

Advancing Wasserstein Convergence Analysis of Score-Based Models: Insights from Discretization and Second-Order Acceleration

Yifeng Yu^{*}, Lu Yu[†]

^{*} Tsinghua University, [†] City University of Hong Kong

November 6th, 2025

① Background

② Main Results

③ References

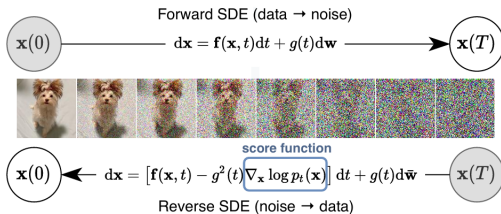
1 Background

2 Main Results

3 References

Diffusion Model

- *Diffusion models* have become a pivotal framework in modern generative modeling, achieving notable success across fields such as image generation, natural language processing, and computational biology. These models **add noise to data** via a forward process and learn to reverse it, **reconstructing data from noise**.
- A widely adopted formulation of diffusion models is the **score-based generative model (SGM)**, implemented using stochastic differential equations (SDEs) [SSDK⁺20].



Framework

Forward process:

$$dX_t = f(X_t, t) dt + g(t) dB_t,$$

Backward process:

$$dY_t = [-f(Y_t, T - t) + g(T - t)^2 \nabla \log p_{T-t}(Y_t)] dt + g(T - t) dW_t.$$

For clarity, we adopt the simplest possible choice in this work by setting $f(X_t, t) = -X_t/2$ and $g(t) = 1$. This results in the Ornstein-Uhlenbeck process, which is described by the following SDE:

$$dX_t = -\frac{1}{2}X_t dt + dB_t. \quad (1)$$

Then, diffusion models generate new data by reversing the SDE (1), which leads to the following backward SDE

$$dX_t^{\leftarrow} = \frac{1}{2}(X_t^{\leftarrow} + 2\nabla \log p_{T-t}(X_t^{\leftarrow})) dt + dW_t, \quad (2)$$

Score Matching

$$\begin{aligned} & \underset{\theta \in \Theta}{\text{minimize}} \quad \mathbb{E}[\|s_{\theta}(t, X_t) - \nabla \log p_t(X_t)\|^2], \\ & \iff \underset{\theta \in \Theta}{\text{minimize}} \quad \mathbb{E}[\text{tr}(\nabla s_{\theta}(x)) + \frac{1}{2}\|s_{\theta}(x)\|_2^2]. \end{aligned}$$

Denoising score matching. First perturbs the data point x with a pre-specified noise distribution $q_{\sigma}(\tilde{x}|x)$ and then employs score matching to estimate the score of the perturbed data distribution $q_{\sigma}(\tilde{x}) := \int q_{\sigma}(\tilde{x}|x)p_{\text{data}}(x) \, dx$.

$$\mathbb{E}_{q_{\sigma}(\tilde{x}|x)p_{\text{data}}(x)}[\|s_{\theta}(\tilde{x}) - \nabla_{\tilde{x}} \log q_{\sigma}(\tilde{x}|x)\|_2^2].$$

Sliced score matching. p_v is a simple distribution of random vectors, e.g., the multivariate standard normal.

$$\mathbb{E}_{p_v} \mathbb{E}_{p_{\text{data}}} \left[v^{\top} \nabla s_{\theta}(x) v + \frac{1}{2} \|s_{\theta}(x)\|_2^2 \right].$$

Discretization Schemes

$$X_{t+h}^{\leftarrow} = X_t^{\leftarrow} + \int_0^h \gamma(T - (t + v), X_{t+v}^{\leftarrow}) dv + \Delta_h W_t.$$

- **Euler-Maruyama scheme:**

$$\vartheta_{n+1}^{\text{EM}} = (1 + h/2)\vartheta_n^{\text{EM}} + hs_*(T - nh, \vartheta_n^{\text{EM}}) + \sqrt{h}\xi_n.$$

- **Exponential Integrator:**

$$\vartheta_{n+1}^{\text{EI}} = e^{\frac{h}{2}}\vartheta_n^{\text{EI}} + 2(e^{\frac{h}{2}} - 1)s_*(T - nh, \vartheta_n^{\text{EI}}) + \sqrt{e^h - 1}\xi_n.$$

Randomized Midpoint Method:

$$\begin{aligned} X_{t+h}^{\leftarrow} &= X_t^{\leftarrow} + \int_0^h \gamma(T - t - v, X_{t+v}^{\leftarrow}) dv + \Delta_h W_t \\ &\approx X_t^{\leftarrow} + h\gamma(T - t - hU, X_{t+hU}^{\leftarrow}) + \Delta_h W_t. \end{aligned}$$

- **Vanilla Midpoint Randomization:**

Step 1 $\xi'_n, \xi''_n \sim \mathcal{N}(\mathbf{0}, I_d)$, $U_n \sim U[0, 1]$. $\xi_n = \sqrt{U_n} \xi'_n + \sqrt{1 - U_n} \xi''_n$.

Step 2 With the initialization $\vartheta_0^{\text{REM}} \sim \hat{p}_T$, define

$$\vartheta_{n+U}^{\text{REM}} = \vartheta_n^{\text{REM}} + hU_n\gamma(T - nh, \vartheta_n^{\text{REM}}) + \sqrt{hU_n}\xi'_n,$$

$$\vartheta_{n+1}^{\text{REM}} = \vartheta_n^{\text{REM}} + h\gamma(T - (n + U_n)h, \vartheta_{n+U}^{\text{REM}}) + \sqrt{h}\xi_n.$$

- **Exponential Integrator with Midpoint Randomization:**

Step 1 $\xi'_n, \xi''_n \sim \mathcal{N}(\mathbf{0}, I_d)$, $U_n \sim U[0, 1]$. $\xi_n = \rho_n \xi'_n + \sqrt{1 - \rho_n^2} \xi''_n$ with

$$\rho_n = e^{\frac{h(1+U_n)}{2}} (1 - e^{-hU_n}) \left[(e^{hU_n} - 1)(e^h - 1) \right]^{-1/2}.$$

Step 2 With the initialization $\vartheta_0^{\text{REI}} \sim \hat{p}_T$, define

$$\vartheta_{n+U}^{\text{REI}} = e^{hU_n/2} \vartheta_n^{\text{REI}} + 2(e^{hU_n/2} - 1)s_*(T - nh, \vartheta_n^{\text{REI}}) + \sqrt{e^{hU_n} - 1} \xi'_n,$$

$$\vartheta_{n+1}^{\text{REI}} = e^{h/2} \vartheta_n^{\text{REI}} + he^{(1-U_n)h/2} s_*(T - (n + U_n)h, \vartheta_{n+U}^{\text{REI}}) + \sqrt{e^h - 1} \xi_n.$$

1 Background

2 Main Results

3 References

Wasserstein Convergence Analysis

Assumption 1

The target density p_0 is m_0 -strongly log-concave, and the score function $\nabla \log p_0$ is L_0 -Lipschitz.

Assumption 2

There exists a constant $M_1 > 0$ such that for $n = 0, 1, \dots, N-1$,

$$\sup_{nh \leq t, s \leq (n+1)h} \|\nabla \log p_{T-t}(x) - \nabla \log p_{T-s}(x)\| \leq M_1 h(1 + \|x\|), \quad \forall x.$$

Assumption 3

Given a small $\varepsilon_{sc} > 0$, the score estimator satisfies

$$\sup_{0 \leq n \leq N} \|\nabla \log p_{T-nh}(\vartheta_n) - s_*(T - nh, \vartheta_n)\|_{\mathbb{L}_2} \leq \varepsilon_{sc}.$$

Theorem 1 (EM, EI)

Suppose that Assumptions 1, 2 and 3 hold, it holds that

$$W_2(\mathcal{L}(\vartheta_N^\alpha), p_0) \lesssim e^{-m_{\min} T} \|X_0\|_{\mathbb{L}_2} + \mathcal{C}_1^\alpha \sqrt{dh} + \mathcal{C}_2^\alpha \varepsilon_{sc},$$

where $\mathcal{C}_1^{\text{EM}} = \frac{L_{\max}+1/2}{m_{\min}-1/2}$, $\mathcal{C}_2^{\text{EM}} = \frac{1}{m_{\min}-1/2}$, and $\mathcal{C}_1^{\text{EI}} = \frac{L_{\max}}{m_{\min}-1/2}$, $\mathcal{C}_2^{\text{EI}} = \frac{1}{m_{\min}-1/2}$, with $m_{\min} = \min(1, m_0)$ and $L_{\max} = 1 + L_0$.

Corollary 2

Given a small $\varepsilon > 0$ and $\varepsilon_{sc} = \mathcal{O}(\varepsilon)$, the Wasserstein distance satisfies $W_2(\mathcal{L}(\vartheta_N^\alpha), p_0) < \varepsilon$, $\alpha \in \{\text{EM}, \text{EI}\}$ after

$N = \mathcal{O}\left(\frac{d}{\varepsilon^2} \log\left(\frac{\sqrt{d}}{\varepsilon}\right)\right)$ iterations, provided that $T = \mathcal{O}\left(\log\left(\frac{\sqrt{d}}{\varepsilon}\right)\right)$

and $h = \mathcal{O}\left(\frac{\varepsilon^2}{d}\right)$.

Assumption 4

There exists a constant $\varepsilon_{sc} > 0$ such that for any $u \in [0, 1]$ and $n = 0, \dots, N$,

$$\|\nabla \log p_{T-(n+u)h}(\vartheta_{n+u}^\alpha) - s_*(T - (n+u)h, \vartheta_{n+u}^\alpha)\|_{\mathbb{L}_2} \leq \varepsilon_{sc}.$$

Theorem 3 (REM, REI)

Suppose that Assumptions 1, 2 and 4 hold, then for $\alpha \in \{\text{REM}, \text{REI}\}$,

$$W_2(\mathcal{L}(\vartheta_N^\alpha), p_0) \lesssim e^{-m_{\min} T} \|X_0\|_{\mathbb{L}_2} + \mathcal{C}_1^\alpha(d) \sqrt{h} + \mathcal{C}_2^\alpha \varepsilon_{sc},$$

where $\mathcal{C}_1^{\text{REM}}(d) = \frac{\sqrt{d/3} L_{\max} + 1/2 \sqrt{3}}{m_{\min} - 1/2}$, $\mathcal{C}_2^{\text{REM}} = \frac{3}{m_{\min} - 1/2}$ and

$\mathcal{C}_1^{\text{REI}}(d) = \frac{\sqrt{d/3} L_{\max}}{(m_{\min} - 1/2)}$, $\mathcal{C}_2^{\text{REI}} = \frac{3}{m_{\min} - 1/2}$, with L_{\max} and m_{\min} as defined in Theorem 1.

Second-order Acceleration

$$dx_t = \gamma(T - t, x_t) dt + \sigma dW_t,$$

We assume that $\gamma(t, x) \in C^{1,3}(\mathbb{R}_+ \times \mathbb{R}^d)$ and approximate it by a linear function in both state and time within each discretization step. Applying Itô's formula to $\gamma(T - t, x)$, we derive the following approximation for $\gamma(T - t, x_t) - \gamma(T - s, x_s)$

$$\left[\frac{\sigma^2}{2} \frac{\partial^2 \gamma}{\partial x^2}(T - s, x_s) - \frac{\partial \gamma}{\partial t}(T - s, x_s) \right] (t - s) + \frac{\partial \gamma}{\partial x}(T - s, x_s)(x_t - x_s).$$

This allows us to express $\gamma(T - t, x_t)$ in the following form

$$\gamma(T - t, x_t) \approx \gamma(T - s, x_s) + L_s(x_t - x_s) + M_s(t - s),$$

with

$$L_s = \frac{\partial \gamma}{\partial x}(T - s, x_s), \quad M_s = \frac{\sigma^2}{2} \frac{\partial^2 \gamma}{\partial x^2}(T - s, x_s) - \frac{\partial \gamma}{\partial t}(T - s, x_s).$$

$$x_t = \vartheta_n^{\text{SO}} + \int_{nh}^t \left(\frac{1}{2} \vartheta_n^{\text{SO}} + \nabla \log p_{T-nh}(\vartheta_n^{\text{SO}}) + L_n(x_u - \vartheta_n^{\text{SO}}) + M_n(u - nh) \right) du + \int_{nh}^t dW_u$$

where

$$L_n = \frac{1}{2} I_d + \nabla^2 \log p_{T-nh}(\vartheta_n^{\text{SO}}),$$

$$M_n = \frac{1}{2} \sum_{j=1}^d \frac{\partial^2}{\partial x_j^2} \nabla \log p_{T-nh}(\vartheta_n^{\text{SO}}) - \frac{\partial}{\partial t} \nabla \log p_{T-nh}(\vartheta_n^{\text{SO}}).$$

$$\begin{aligned} \vartheta_{n+1}^{\text{SO}} &= \vartheta_n^{\text{SO}} + s_*^{(L)}(T-nh, \vartheta_n^{\text{SO}})^{-1} \left(e^{s_*^{(L)}(T-nh, \vartheta_n^{\text{SO}})h} - I_d \right) \left(\frac{1}{2} \vartheta_n^{\text{SO}} + s_*(T-nh, \vartheta_n^{\text{SO}}) \right) \\ &\quad + s_*^{(L)}(T-nh, \vartheta_n^{\text{SO}})^{-2} \left(e^{s_*^{(L)}(T-nh, \vartheta_n^{\text{SO}})h} - s_*^{(L)}(T-nh, \vartheta_n^{\text{SO}})h - I_d \right) s_*^{(M)}(T-nh, \vartheta_n^{\text{SO}}) \\ &\quad + \int_{nh}^{(n+1)h} e^{s_*^{(L)}(T-nh, \vartheta_n^{\text{SO}})[(n+1)h-t]} dW_t. \end{aligned}$$

Assumption 5

For some constants $\varepsilon_{sc}^{(L)}, \varepsilon_{sc}^{(M)} > 0$, the estimate for high-order derivatives of the score function satisfies that

$$\sup_{0 \leq n \leq N-1} \|s_*^{(L)}(T - nh, \vartheta_n^{SO}) - L_n\|_{\mathbb{L}_2} \leq \varepsilon_{sc}^{(L)},$$
$$\sup_{0 \leq n \leq N-1} \|s_*^{(M)}(T - nh, \vartheta_n^{SO}) - M_n\|_{\mathbb{L}_2} \leq \varepsilon_{sc}^{(M)}.$$

Assumption 6

Let $\|\cdot\|_F$ denote the Frobenius norm. There exists a positive constant L_F such that

$$\|\nabla^2 \log p_t(x) - \nabla^2 \log p_t(y)\|_F \leq L_F \|x - y\|, \quad \forall x, y \in \mathbb{R}^d.$$

Assumption 7

There exists a constant $M_2 > 0$ such that, for any $n = 0, \dots, N-1$ and $t \in [nh, (n+1)h]$, it holds that

$$\|\nabla^2 \log p_{T-t}(x) - \nabla^2 \log p_{T-nh}(x)\| \leq M_2 h(1 + \|x\|), \quad \forall x \in \mathbb{R}^d.$$

Theorem 4

Suppose that Assumptions 1, 3, 5, 6 and 7 hold, then

$$\begin{aligned} W_2(\mathcal{L}(\vartheta_N^{\text{SO}}), p_0) &\lesssim e^{-m_{\min} T} \|X_0\|_{\mathbb{L}_2} + \mathcal{C}_1^{\text{SO}}(d)h \\ &\quad + \mathcal{C}_2^{\text{SO}} \left(\varepsilon_{sc} + \frac{2}{3} \sqrt{hd} \varepsilon_{sc}^{(L)} + \frac{1}{2} h \varepsilon_{sc}^{(M)} \right) \end{aligned}$$

where $\mathcal{C}_1^{\text{SO}}(d) = e^{(L_{\max}-1/2)h} \cdot \frac{\sqrt{d}(L_{\max}^{3/2} + \sqrt{2}L_F/4)}{m_{\min}-1/2}$ and

$\mathcal{C}_2^{\text{SO}} = \frac{e^{(L_{\max}-1/2)h}}{m_{\min}-1/2}$ with L_{\max} and m_{\min} as defined in Theorem 1.

Corollary 5

For a given $\varepsilon > 0$, the Wasserstein distance satisfies $W_2(\mathcal{L}(\vartheta_N^{\text{SO}}), p_0) < \varepsilon$ after $N = \mathcal{O}\left(\frac{\sqrt{d}}{\varepsilon} \log\left(\frac{\sqrt{d}}{\varepsilon}\right)\right)$ iterations, provided that $T = \mathcal{O}\left(\log\left(\frac{\sqrt{d}}{\varepsilon}\right)\right)$ and $h = \mathcal{O}\left(\frac{\varepsilon}{\sqrt{d}}\right)$.

The accelerated convergence of this method is driven by two key innovations: approximating the drift term through its **Itô expansion** rather than endpoint evaluations, and deriving **a closed-form solution** to the integral equation using the Itô formula, akin to Exponential Integrator techniques.

Numerical Studies

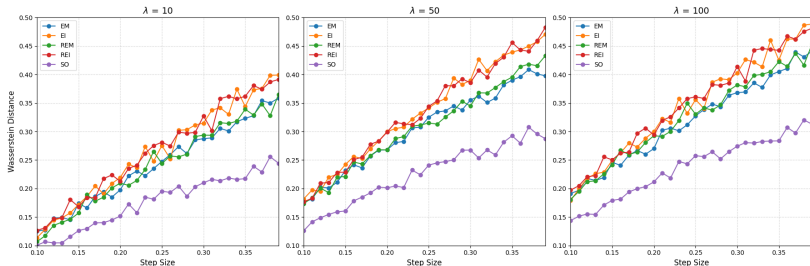
MNIST dataset: To accelerate the SO algorithm, we use Hessian-vector products (HVPs) instead of explicitly computing the Hessian.



We apply the five schemes to the posterior density of penalized logistic regression, defined by $p_0(\theta) \propto \exp(-f(\theta))$ with the potential function

$$f(\theta) = \frac{\lambda}{2} \|\theta\|^2 + \frac{1}{n_{\text{data}}} \sum_{i=1}^{n_{\text{data}}} \log(1 + \exp(-y_i x_i^\top \theta)),$$

where $\lambda > 0$ denotes the tuning parameter.



1 Background

2 Main Results

3 References

[SSDK⁺20] Yang Song, Jascha Sohl-Dickstein, Diederik P Kingma, Abhishek Kumar, Stefano Ermon, and Ben Poole. Score-based generative modeling through stochastic differential equations. *arXiv preprint arXiv:2011.13456*, 2020.

Thanks!