



# Lecture 1

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**OPIM**  
**5603**  
Fall  
2019

## Introduction to Probability Distributions

Slide 1.1



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Slide 1.2

# Selected Topics

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1. Random Variables and their Prob. Distributions
2. Expected Value ("Central Tendency")
3. Variance ("Spread")
4. The Normal Distribution
5. Normal Approximation to Binomial Distribution
6. Continuous Distributions

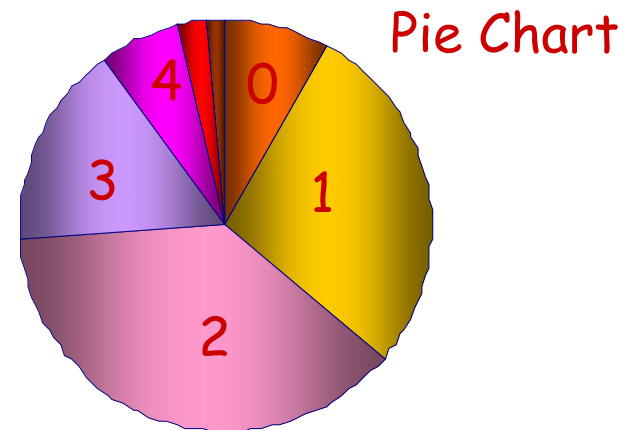
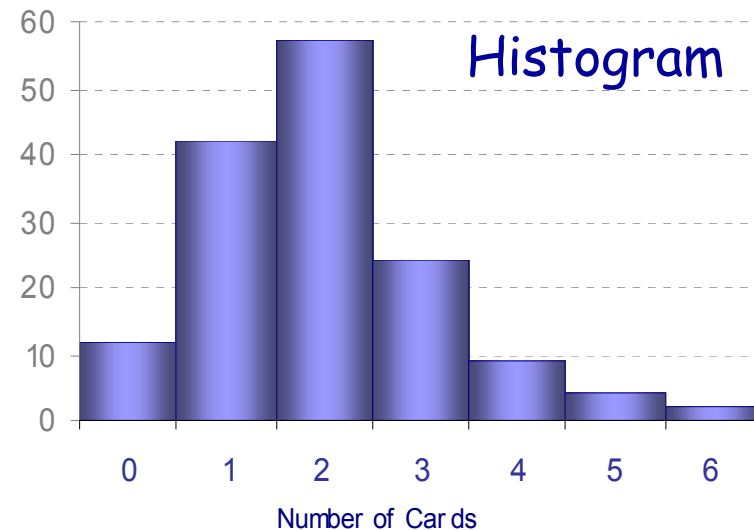


# A glance into 'Representation of data'

**Example:** In a survey the college students were asked how many credit cards they own. The results are reported in the table:

# Cards	# Students
0	12
1	42
2	57
3	24
4	9
5	4
6	2

Total: 150





# Random Variable

A **random variable** is a rule that assigns a **numeric value** and a **probability** to an outcome of a chance experiment.

- **Finite discrete** – assumes only finitely many values.

**Example:** Rolling a die.

- **Infinite discrete** – assumes infinitely many values that may be arranged in a sequence.

**Example:** Counting die rolls until the outcome is 6.

- **Continuous** – assumes values that make up an interval of real numbers.

**Examples:** Time between arrivals of two customers.

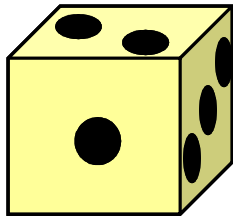
Tomorrow's temperature at noon.



# Prob. Distribution of a Random Variable

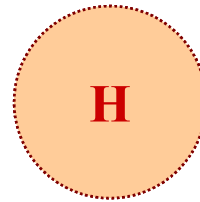
Examples: Probability distributions of:

(a) a die roll,



$x$	$P(D = x)$
1	$1/6$
2	$1/6$
3	$1/6$
4	$1/6$
5	$1/6$
6	$1/6$

(b) a coin toss.



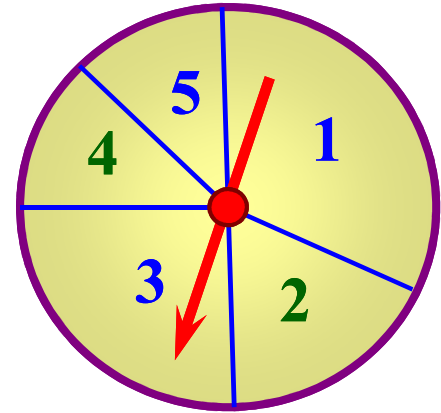
Let  $C$  denote a random variable.

$C$  must have numerical values, so we agree on:

**Tail = 0, Head = 1**

$x$	$P(C = x)$
0	$1/2$
1	$1/2$

(c) a hand spin.



$x$	$P(H = x)$
1	$1/3$
2	$1/6$
3	$1/4$
4	$1/8$
5	$1/8$



# Prob. Distribution of a Random Variable

**Example:** Random variable  $X$  assumes (only) the values

$-8, -3, -1, 0, 1, 4, 6$

(hence a finite discrete random variable).

Its probability distribution is given by:

$x$	$-8$	$-3$	$-1$	$0$	$1$	$4$	$6$
$P(X=x)$	$0.13$	$0.15$	$0.17$	$0.20$	$0.15$	$0.11$	$0.09$

Find

(a)  $P(X \leq 0) = P(\{-8, -3, -1, 0\}) = 0.13 + 0.15 + 0.17 + 0.2 = 0.65$

(b)  $P(-3 \leq X \leq 1) = P(\{-3, -1, 0, 1\}) = 0.67$



## Credit Cards example revisited

Students were asked how many credit cards they own.  $X$  is the random variable representing the number of cards and the results are below.

$x$	#Students	$P(X=x)$
0	12	0.08
1	42	0.28
2	57	0.38
3	24	0.16
4	9	0.06
5	4	0.02666
6	2	0.01333

$$\frac{12}{150}$$

Probability  
Distribution

Total: 150



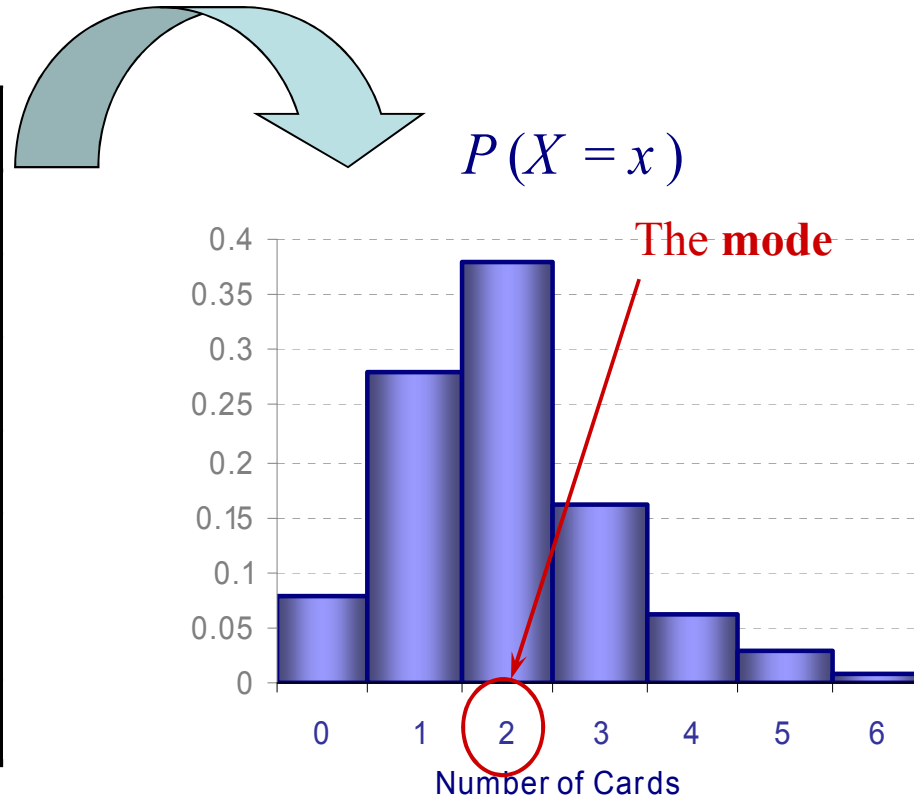
# Histogram revisited

A way to represent a probability distribution of a random variable graphically.

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Credit card results:

$x$	$P(X=x)$
0	0.08
1	0.28
2	0.38
3	0.16
4	0.06
5	0.02666
6	0.01333







# Mean, Median, Mode

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The **average (mean)** of the  $n$  numbers  $x_1, x_2, \dots, x_n$  is defined as

$$\bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n}$$

The **median** is the middle value in a set of data that is arranged in increasing or decreasing order. For an even number of data points the median is the average of the middle two.

The **mode** is the most frequent number in a set of data.



## Example

The quiz scores for a particular student are given below:

22, 25, 20, 18, 12, 20, 24, 20, 20, 25, 24, 25, 18

Find the mean, median and mode.

**Mean:**  $\frac{\text{sum of entries}}{\text{number of data points}} = \frac{273}{13} = 21$

**Median:** Sort the numbers:

Middle number = 20

12, 18, 18, 20, 20, 20, 20, 22, 24, 24, 25, 25, 25

**Mode** (most frequent): 20 (occurs 4 times)



# Expected Value of a Discrete Random Variable

Let  $X$  be a random variable that assumes the values  $x_1, x_2, \dots, x_n$  with associated probabilities  $p_1, p_2, \dots, p_n$ , respectively.

Then the **expected value (mean)** of  $X$ , denoted by  $E(X)$ , is

$$E(X) = x_1 p_1 + x_2 p_2 + \dots + x_n p_n$$

**Example:** Let  $D$  be the random variable recording the outcome of the single roll of a fair die. Find the expected value of  $D$ .

Solution: The probability distribution is



$x$	$P(D = x)$
1	$1/6$
2	$1/6$
3	$1/6$
4	$1/6$
5	$1/6$
6	$1/6$

$$\begin{aligned}\text{Mean: } E(D) &= x_1 p_1 + x_2 p_2 + \dots + x_6 p_6 \\ &= 1 \cdot 1/6 + 2 \cdot 1/6 + 3 \cdot 1/6 + \\ &\quad + 4 \cdot 1/6 + 5 \cdot 1/6 + 6 \cdot 1/6 \\ &= 21/6 = 3.5\end{aligned}$$



## Example

Use the table to find out the expected number of credit cards that a student will own.

Solution: Let  $X$  be the random variable recording the number of credit cards students have. The probability distribution of  $X$  is:

The expected value:

$$\begin{aligned} E(X) &= x_1 p_1 + x_2 p_2 + \dots + x_n p_n \\ &= 0 \cdot 0.08 + 1 \cdot 0.28 + 2 \cdot 0.38 + 3 \cdot 0.16 + \\ &\quad + 4 \cdot 0.06 + 5 \cdot 0.02666 + 6 \cdot 0.01333 \\ &= 1.97333 \end{aligned}$$

$x$	# Students	$P(X=x)$
0	12	0.08
1	42	0.28
2	57	0.38
3	24	0.16
4	9	0.06
5	4	0.02666
6	2	0.01333



## Example

The quiz scores for a particular student are given below:

22, 25, 20, 18, 12, 20, 24, 20, 20, 25, 24, 25, 18

Find the expected value of the random variable  $S$  that measures this student quiz performance.

Solution: The frequency table and prob. distribution of  $S$  are given by

$$\begin{aligned} E(X) &= x_1 p_1 + x_2 p_2 + \dots + x_n p_n \\ &= 12 \cdot \frac{1}{13} + 18 \cdot \frac{2}{13} + 20 \cdot \frac{4}{13} + \\ &\quad + 22 \cdot \frac{1}{13} + 24 \cdot \frac{2}{13} + 25 \cdot \frac{3}{13} \\ &= \frac{12 + 36 + 80 + 22 + 48 + 75}{13} \\ &= \frac{273}{13} = 21 \end{aligned}$$

$x$	# quizzes	$P(S=x)$
12	1	$\frac{1}{13}$
18	2	$\frac{2}{13}$
20	4	$\frac{4}{13}$
22	1	$\frac{1}{13}$
24	2	$\frac{2}{13}$
25	3	$\frac{3}{13}$



## Example

Your friend tosses a fair coin. If the outcome is a Head, you win \$5. Otherwise you lose \$5. What is your expected win?

Solution: Let  $W$  be the random variable recording your winnings in a single toss of a fair coin. The probability distribution of  $W$  is given by:

$x$	$P(W = x)$
-5	$1/2$
5	$1/2$

Expected value is:  $E(W) = x_1 p_1 + x_2 p_2 = -5 \cdot 1/2 + 5 \cdot 1/2 = 0$ .

A game in which the expected win is 0 is called a **fair game**.



## Example

What is the expected win for a \$1 bet on red in a single roll of American roulette?

**Note:** The American roulette wheel has 38 numbered fields, two of which are green (0 and 00), 18 red and 18 black.



Solution: Let  $R$  be the random variable recording your winnings from a \$1 bet on red in a single roll of American roulette.

The probability distribution of  $R$  is given by:

$x$	$P(W = x)$
-1	$20/38$
1	$18/38$

Expected value is:

$$E(R) = x_1 p_1 + x_2 p_2 = -1 \cdot \frac{20}{38} + 1 \cdot \frac{18}{38} = -\frac{2}{38} = -\frac{1}{19}.$$

Expected **loss** is \$0.052632, i.e., about 5.3 cents per \$1 bet.



# Variance and Standard Deviation

**Variance** is a measure of the spread of the data. The larger the variance, the larger the spread.

Suppose a random variable has a probability distribution

$x$	$x_1$	$x_2$	$x_3$	$\dots$	$x_n$
$P(X=x)$	$p_1$	$p_2$	$p_3$	$\dots$	$p_n$

and expected value  $E(X) = \mu$ .

The **variance** of a random variable  $X$  is defined by:

$$\text{Var}(X) = p_1 (x_1 - \mu)^2 + p_2 (x_2 - \mu)^2 + \dots + p_n (x_n - \mu)^2 = E((X - \mu)^2)$$

The **standard deviation** of a random variable  $X$  is defined as a square root of the variance:  $\sigma = \sqrt{\text{Var}(X)}$ .

It measures the spread of the data using *the same unit* as the data.





# Example

The daily sales of *Impalas* at two *Chevrolet* dealerships are given:

Shiny Chevy Ltd.

Chevy Rules Co.

# cars sold	7	8	9
Frequency	62	106	62

# cars sold	2	3	4	5	6	7	9	10	11	23
Frequency	5	10	14	17	22	28	59	49	24	2

Find the variance and standard deviation of their daily sales.

Note: Both dealerships sold the same number of cars during 230 days: 1840.

Solution: Let  $S$  be the random variable recording the daily sales at *Shiny Chevy*.

The probability distribution of  $S$  is:

$x$	7	8	9
Frequency	62	106	62
$P(S = x) \approx$	0.27	0.46	0.27

Expected value  $\mu = p_1 \cdot x_1 + p_2 \cdot x_2 + p_3 \cdot x_3 \approx 0.27 \cdot 7 + 0.46 \cdot 8 + 0.27 \cdot 9 = 8$

$$\begin{aligned} \text{Var}(S) &= p_1(x_1 - \mu)^2 + p_2(x_2 - \mu)^2 + p_3(x_3 - \mu)^2 \\ &\approx 0.27 \cdot (-1)^2 + 0.46 \cdot 0^2 + 0.27 \cdot 1^2 \\ &\approx 0.53913 \end{aligned}$$

$$\sigma = \sqrt{\text{Var}(S)} \approx 0.73426$$



## Example (cont)

Let  $C$  be the random variable recording the daily sales at *Chevy Rules*.  
The probability distribution of  $C$  is:

$x$	2	3	4	5	6	7	9	10	11	23
Frequency	5	10	14	17	22	28	59	49	24	2
$P(C = x) \approx$	0.02	0.04	0.06	0.07	0.1	0.12	0.26	0.21	0.1	0.01

Expected value  $\mu = p_1 \cdot x_1 + \dots + p_{10} \cdot x_{10} = \dots = 8$

$$\text{Var}(C) = p_1(x_1 - \mu)^2 + \dots + p_{10}(x_{10} - \mu)^2 = \dots \approx 8.01739$$

$$\sigma = \sqrt{\text{Var}(C)} \approx 2.8315$$



## Example (conclusion)

Shiny Chevy Ltd.

$x$	7	8	9
Frequency	62	106	62

$$\mu_S = E(S) = 8$$

$$\text{Var}(S) \approx 0.53913$$

$$\sigma_S \approx 0.73426$$

Chevy Rules Co.

$x$	2	3	4	5	6	7	9	10	11	23
Frequency	5	10	14	17	22	28	59	49	24	2

$$\mu_C = E(C) = 8$$

$$\text{Var}(C) \approx 8.01739$$

$$\sigma_C \approx 2.8315$$

Conclusion: These two probability distributions have the same mean yet significantly different variances. Variance (i.e., standard deviation) measures the spread of data around its mean.



# Bernoulli Random Variable

A random variable with outcomes 0 and 1 is called *Bernoulli variable* (17<sup>th</sup> century Swiss mathematician Jacob Bernoulli).

The probability of outcome 1 is denoted by  $p$ .

The probability of 0 is  $q = 1 - p$  (i.e.,  $p + q = 1$ ).

Expected value of a Bernoulli variable is:

$$\mu = E(X) = x_1 p_1 + x_2 p_2 = 1 \cdot p + 0 \cdot q = p.$$

$x$	$P(X=x)$
1	$p$
0	$1 - p$

The variance of a Bernoulli variable is:

$$\begin{aligned} \text{Var}(X) &= p_1(x_1 - \mu)^2 + p_2(x_2 - \mu)^2 = p \cdot (1 - p)^2 + q \cdot (0 - p)^2 \\ &= pq^2 + qp^2 = pq(q + p) = pq. \end{aligned}$$

A *biased coin toss* experiment is modeled by Bernoulli variable.

*Independent* repetitions of this experiment are called *Binomial Trials*.



# Binomial (Bernoulli) Trials

A **Binomial Trial** has the properties:

1. Number of trials in the experiment is fixed,
2. The only outcomes are **success** and **failure**,
3. In each trial the **success** probability is the same, and
4. The trials are independent of each other.

In a binomial trial in which the probability of **success** in any trial is  $p$ , the probability of exactly  $k$  **successes** in  $n$  independent trials is given by

$$C(n, k) p^k (1 - p)^{n-k}$$

where  $C(n, k) = \binom{n}{k} = \frac{n!}{k!(n-k)!}$

Probability Distribution  
(Density)

$k$	$P(X = k)$
0	$q^n$
1	$C(n, 1) p q^{n-1}$
2	$C(n, 2) p^2 q^{n-2}$
3	$C(n, 3) p^3 q^{n-3}$
...	...
$n - 1$	$C(n, n-1) p^{n-1} q$
$n$	$p^n$



# Why combinatorial coefficient $C(n, k)$ ?

Take for example  $n = 5$  and  $k = 2$ : in five repeated independent experiments we want to list all outcomes that have exactly two successes.

Here's one:  $S S F F F$  Others are obtained by shuffling  $S$ 's and  $F$ 's:

$S F S F F$

$S F F S F$

$S F F F S$

$F S S F F$

$F S F S F$

$F S F F S$

$F F S S F$

$F F S F S$

$F F F S S$

How many are there?

$$C(5, 2) = \frac{5!}{2! \cdot (5 - 2)!} = \frac{5 \cdot 4 \cdot \cancel{3} \cdot \cancel{2} \cdot 1}{2 \cdot 1 \cdot \cancel{3} \cdot \cancel{2} \cdot 1} = 10$$



# Mean, Variance, and Standard Deviation

## of a Binomial Random Variable $X$

If  $X$  is a binomial random variable associated with a binomial experiment consisting of  $n$  trials with probability of **success**  $p$  and probability of **failure**  $q$ , then the mean, variance, and standard deviation of  $X$  are

$$\mu = E(X) = np \quad \text{Var}(X) = npq \quad \sigma_X = \sqrt{npq}$$

**Example:** Five cards are drawn, with replacement, from a standard 52-card deck. If drawing a club is considered a success, find the mean, variance, and standard deviation of the number of successes  $X$ .

$$p = \frac{1}{4}, \quad q = 1 - \frac{1}{4} = \frac{3}{4}$$

$$\mu = np = 5 \left( \frac{1}{4} \right) = 1.25$$

$$\text{Var}(X) = npq = 5 \left( \frac{1}{4} \right) \left( \frac{3}{4} \right) = 0.9375 \quad \sigma_X = \sqrt{npq} = \sqrt{0.9375} \approx 0.968$$



## Example

If the probability of a student successfully passing the class (D or better) is 0.82, find the probability that given 8 students

(a) all 8 pass:  $C(8,8) \cdot 0.82^8 \cdot 0.18^0 \approx 0.2044$

(b) none pass:  $C(8,0) \cdot 0.82^0 \cdot 0.18^8 \approx 0.0000011$

(c) at least 6 pass. Means: 6, **or** 7, **or** 8 'successes':

$$\begin{aligned} & C(8,6) \cdot 0.82^6 \cdot 0.18^2 + \\ & + C(8,7) \cdot 0.82^7 \cdot 0.18^1 + \\ & + C(8,8) \cdot 0.82^8 \cdot 0.18^0 \\ & \approx 0.2758 + 0.3590 + 0.2044 = 0.8392 \end{aligned}$$





## Example revisited

If the probability of a student successfully passing the class (D or better) is **0.82**, find the probability that given **800** students at least 650 pass.

Means: 650, 651, 652, ..., 799, or 800 ‘successes’:

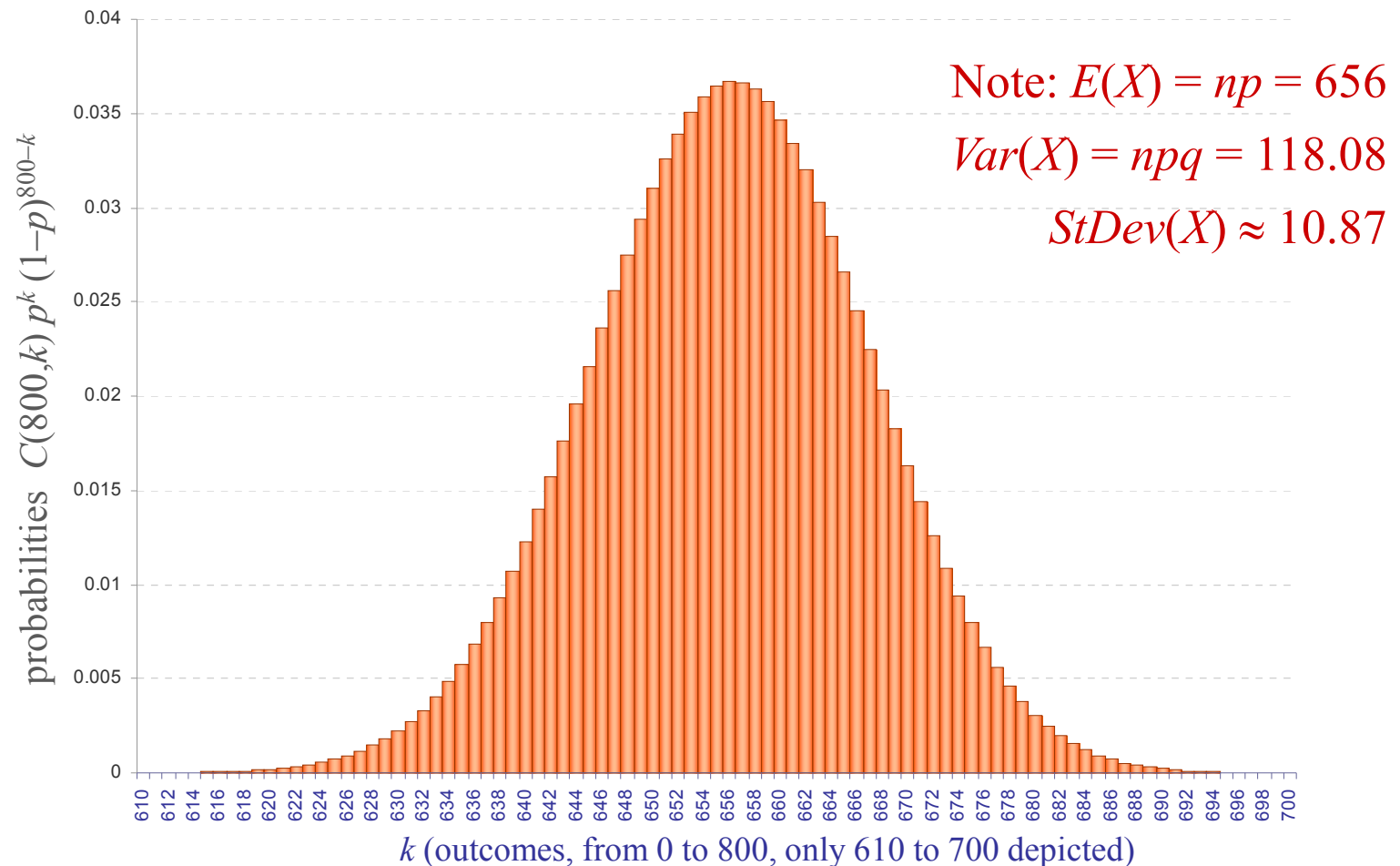
$$\begin{aligned} & C(800, 650) \cdot 0.82^{650} \cdot 0.18^{150} + \\ & + C(800, 651) \cdot 0.82^{651} \cdot 0.18^{149} + \\ & + C(800, 652) \cdot 0.82^{652} \cdot 0.18^{148} + \\ & \quad \dots \\ & + C(800, 799) \cdot 0.82^{799} \cdot 0.18^1 + \\ & + C(800, 800) \cdot 0.82^{800} \cdot 0.18^0 \end{aligned}$$

Pretty cumbersome computation! But easy with  $R$  ( $\approx 0.72722$ ).



# Histogram of Binomial distribution

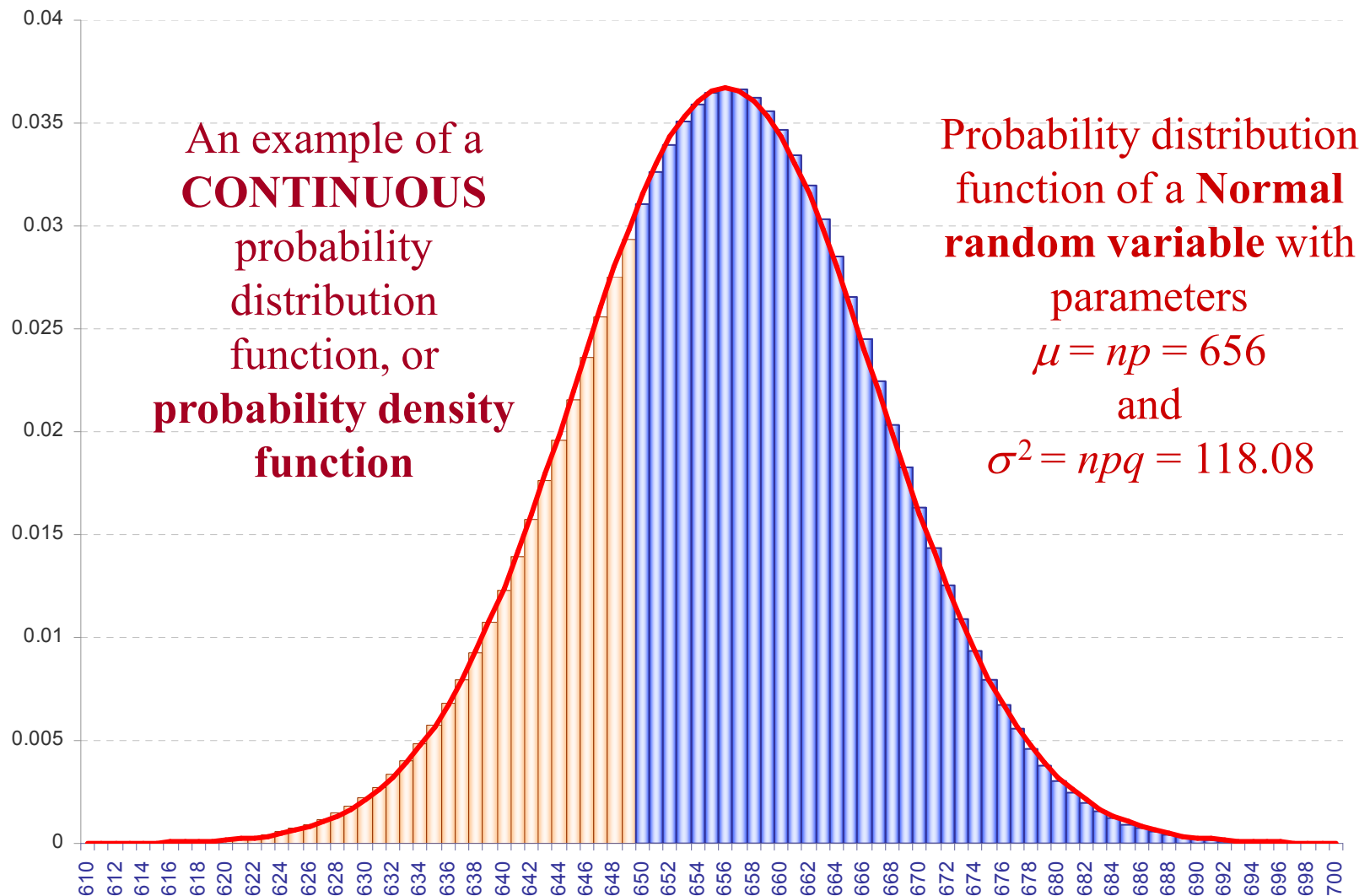
If we compute the probability distribution (table) for the binomial random variable  $X \sim B(n, p)$ , with  $n = 800$  and  $p = 0.82$  (from previous example) and visualize the resulting values with a histogram, we get:





# Example revisited

Use histogram to depict  $P(\text{at least } 650 \text{ pass})$     Answer: sum of the blue 'bars'



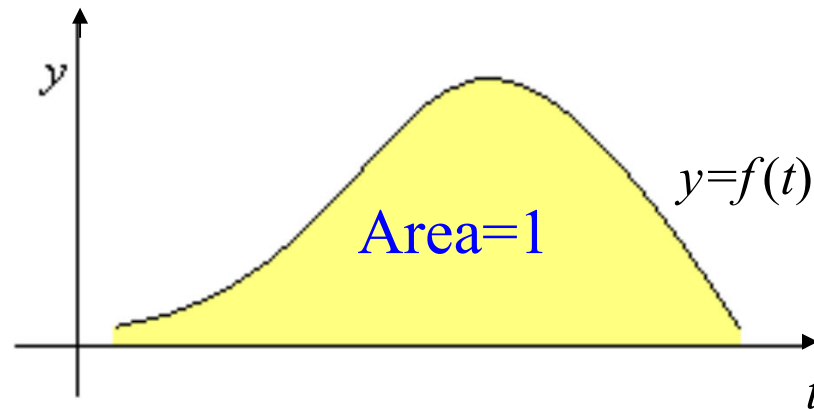


# Probability Density Function

A **probability density function**  $f$  defines a continuous probability distribution and coincides with the interval of values taken on by the random variable associated with an experiment.

A **pdf** must satisfy:

- $f(t) \geq 0$  for all  $t$  in  $(-\infty, +\infty)$ , and
- the area of the region between the graph of  $f$  and the  $t$ -axis is equal to 1.



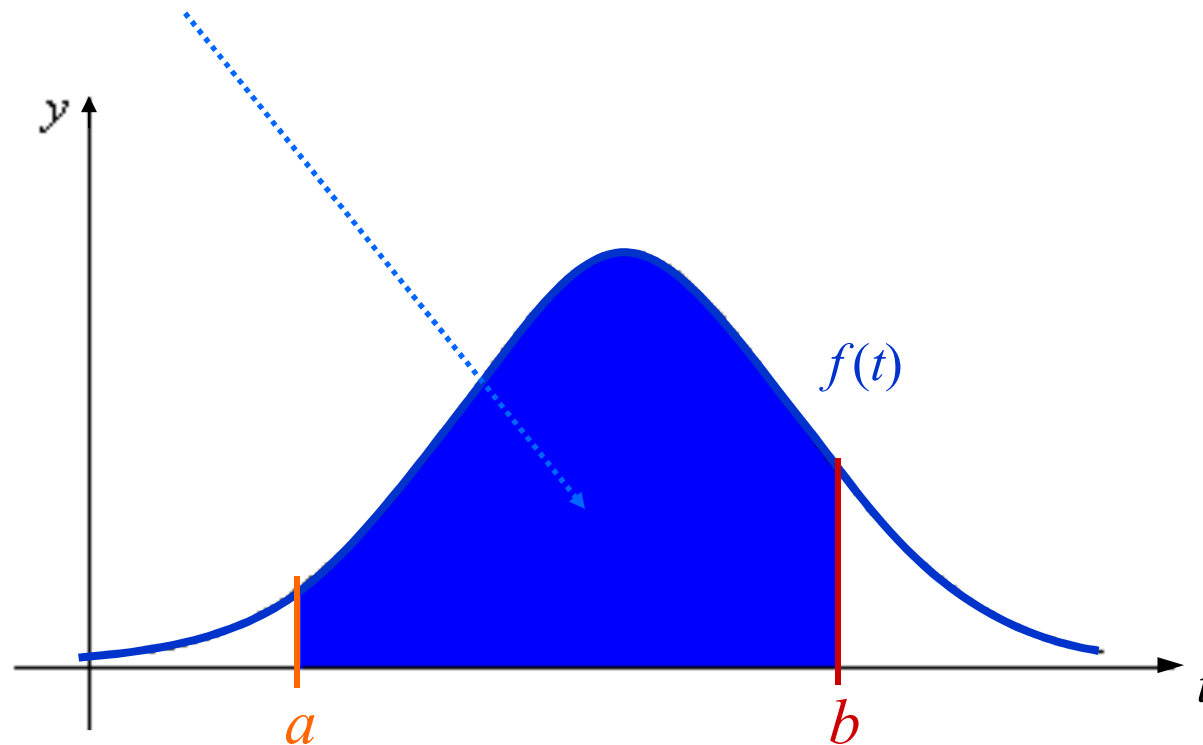


# Probability Density Function

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$P(a < X \leq b)$  is given by the area of the shaded region.



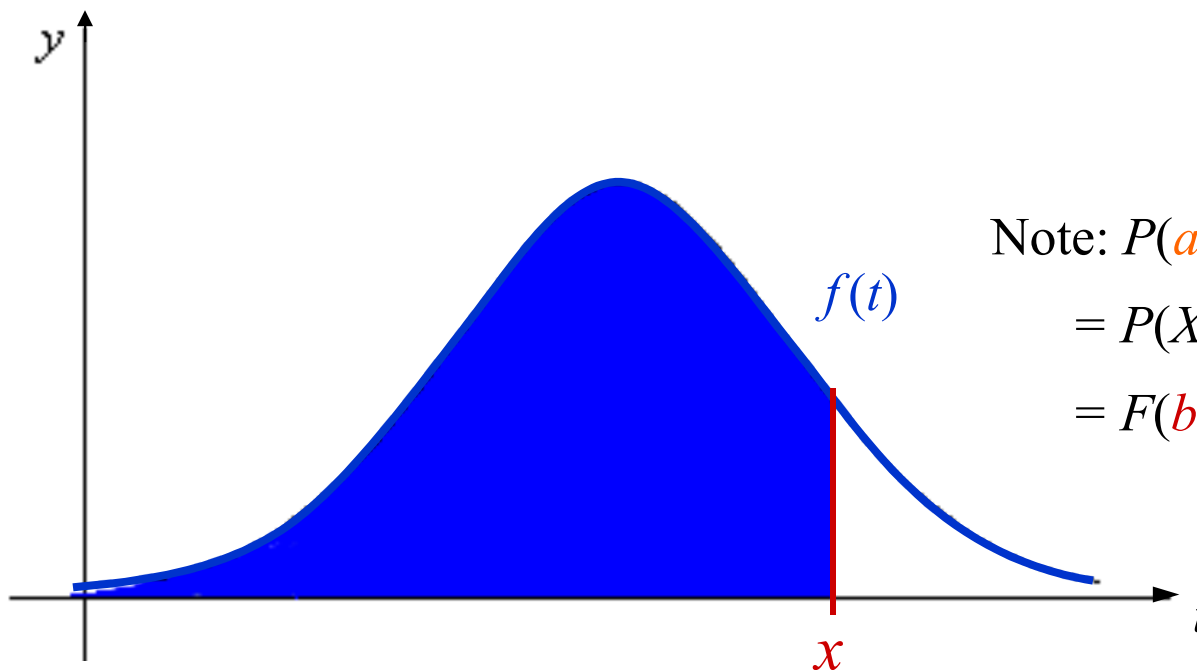
$$\text{Basic calculus: } P(a < X \leq b) = \int_a^b f(t) dt$$



# Cumulative Distribution Function

A **cumulative distribution function (CDF)**  $F$  associated with a probability density function  $f$  is 'defined' by

$$F(x) = \text{area under } f \text{ over the interval } (-\infty, x].$$



$$\begin{aligned} \text{Note: } P(a < X \leq b) &= \\ &= P(X \leq b) - P(X \leq a) = \\ &= F(b) - F(a) \end{aligned}$$

Given a random variable  $X$  with a pdf  $f$ , we have

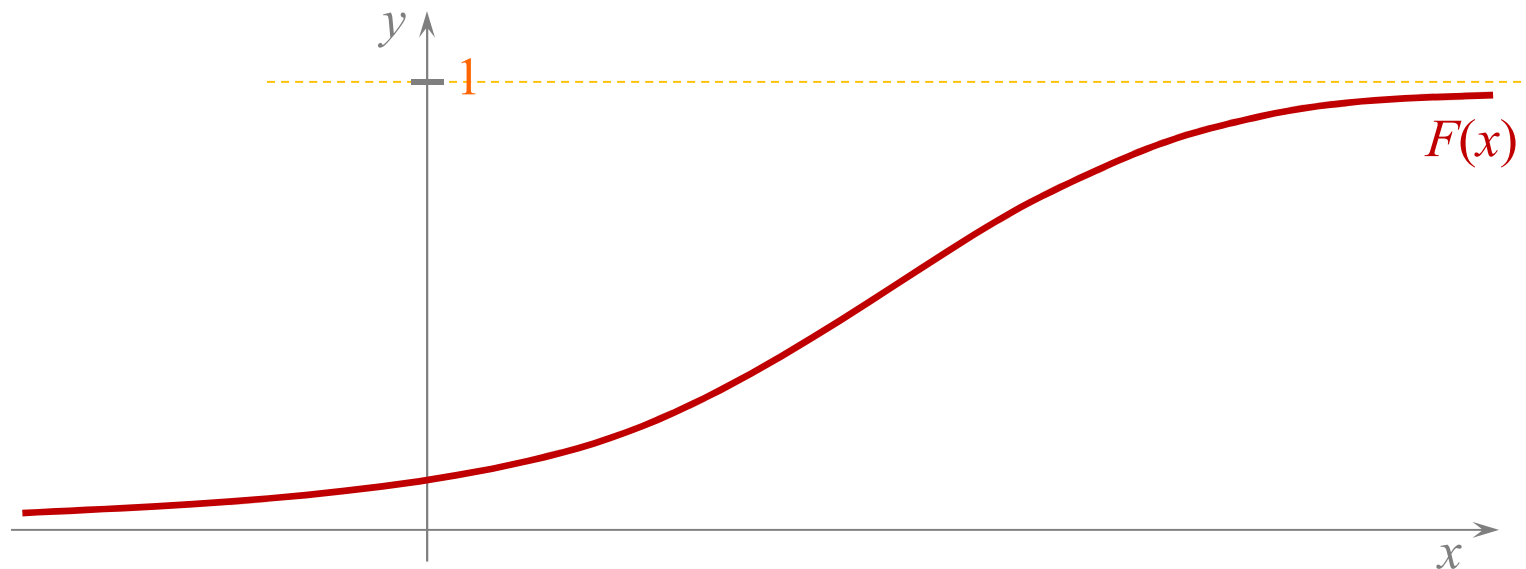
$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt$$



# Properties of CDF

A **CDF** must satisfy:

- $F(x) \geq 0$  for all  $x$  in  $(-\infty, +\infty)$ ,
- $F$  is increasing\* on  $(-\infty, +\infty)$ ,
- $\lim F(x) = 0$ , as  $x \rightarrow -\infty$ , and
- $\lim F(x) = 1$ , as  $x \rightarrow +\infty$ .



(\*)  $F$  needs not be strictly increasing, i.e., it can be constant on some intervals.



# Normal Distribution

**Normal (Gaussian) distributions** are a special class of continuous probability density functions. Many phenomena have probability density functions that are normal.

The graph of this distribution is called a **normal** or **bell curve**.

The probability density function associated with the normal curve:

$$f(t) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{t-\mu}{\sigma}\right)^2}$$

For  $\mu = 0$  and  $\sigma = 1$  we have the **standard normal** pdf:

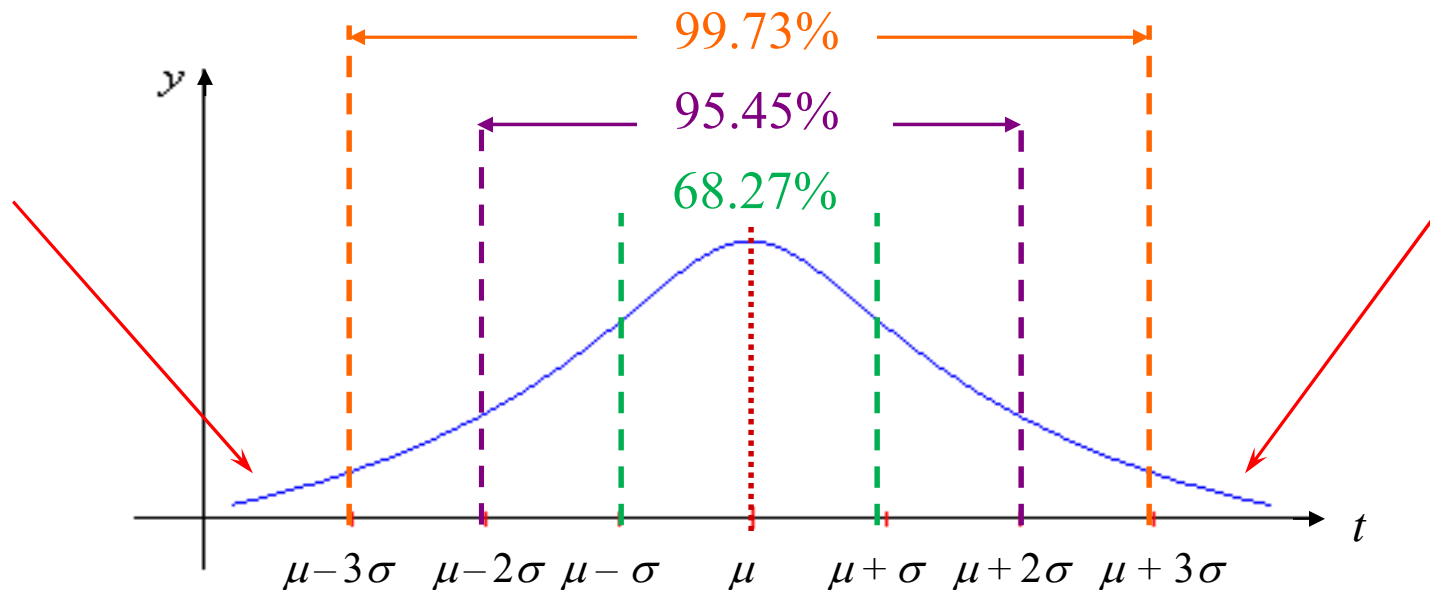
$$f(t) = \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}}$$





# Normal Curve Properties

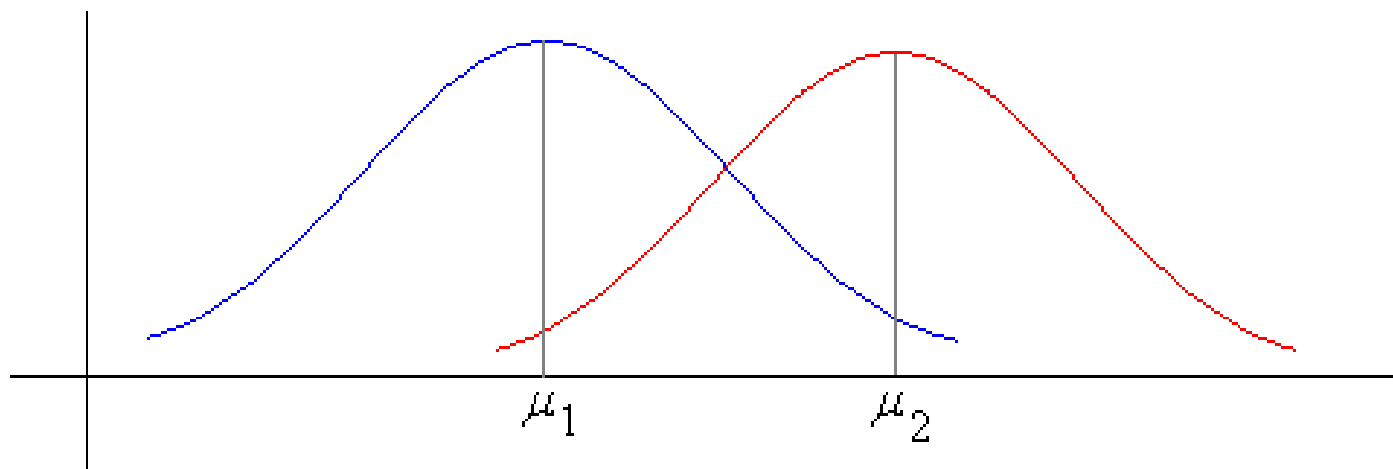
1. The area under the curve is 1.
2. The peak is at  $t = \mu$ , and the curve is symmetric with respect to the vertical line  $t = \mu$ .
3. The curve lies above and approaches the  $t$ -axis.
4. 68.27% of the area lies within  $(\mu - \sigma, \mu + \sigma)$ ,  
95.45% within  $(\mu - 2\sigma, \mu + 2\sigma)$ ,  
99.73% within  $(\mu - 3\sigma, \mu + 3\sigma)$ .



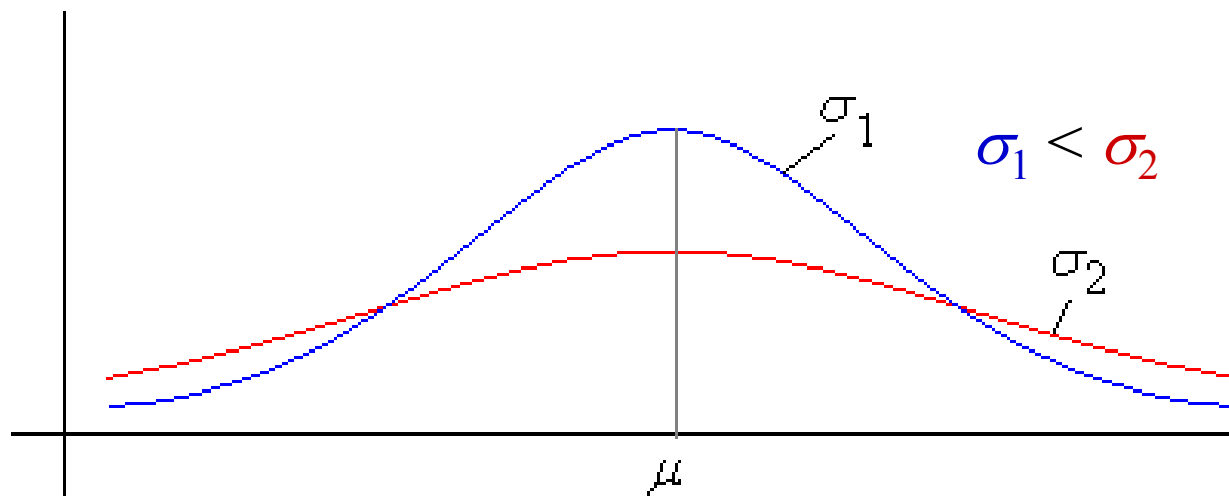


# Normal Curves

Normal curves with same  $\sigma$  and different  $\mu$ 's.



Normal curves with same  $\mu$  but different  $\sigma$ 's'.



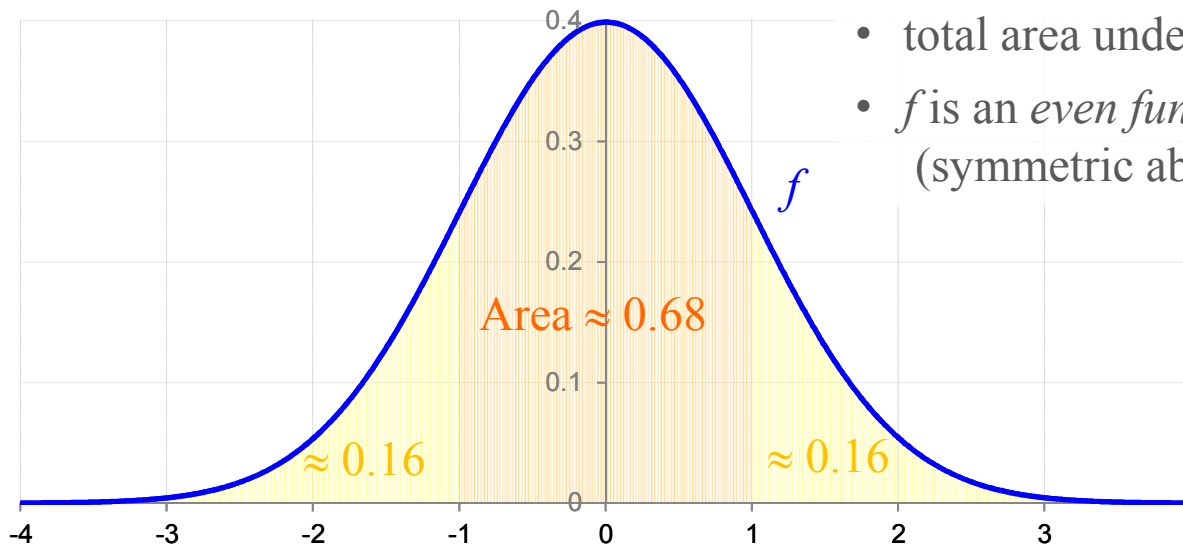


# Standard Normal Distribution

Typically denoted by  $Z$ :  $\mu = 0$  and  $\sigma = 1$ .

pdf:  $f(t) = \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} = \text{dnorm}(t)$  *R function*

- $f(t) > 0$  for all  $t$  in  $(-\infty, +\infty)$ ,
- total area under graph of  $f$  is 1,
- $f$  is an *even function* (symmetric about the y-axis)



CDF:  $F(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt = \text{pnorm}(x)$  *R function*

**Note:**  $F(-1) \approx 0.16$

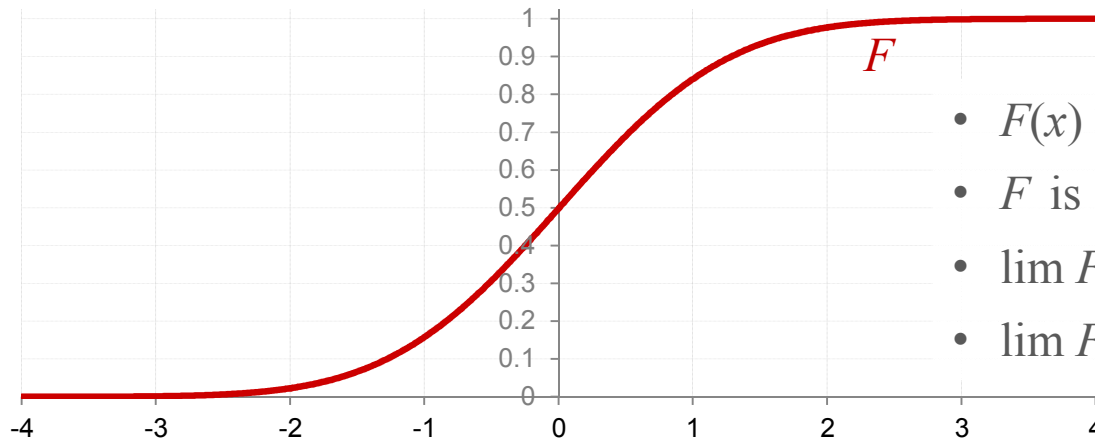
$F(1) \approx 0.16 + 0.68$

$F(0) = 0.5$



# Standard Normal Distribution cont.

CDF:  $F(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt = \text{pnorm}(x)$



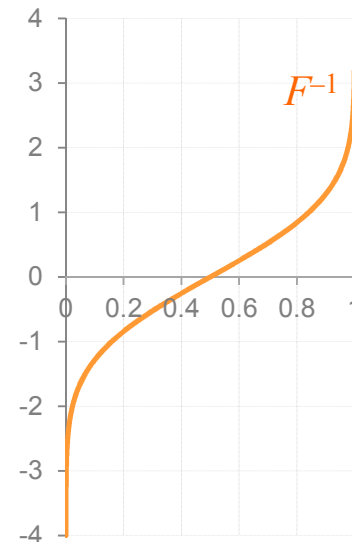
- $F(x) > 0$  for all  $x$  in  $(-\infty, +\infty)$ ,
- $F$  is strictly increasing,
- $\lim F(x) = 0$  as  $x \rightarrow -\infty$ ,
- $\lim F(x) = 1$  as  $x \rightarrow +\infty$ .

CDF inverse:  $F^{-1}(y) = \text{qnorm}(y)$  R function

$$x = F^{-1}(y) \text{ for } y \text{ in } (0,1)$$

*if and only if*

$$F(x) = y$$





# Normal distribution in *R*

*R* provides *density* (pdf), *distribution function* (CDF), *quantile function* (CDF<sup>-1</sup>) and *random number generator* for the normal distribution with parameters  $\mu$  (*mean*) and  $\sigma$  (*sd*).

## Usage:

<code>dnorm(x, mean=0, sd=1, log=FALSE)</code>	pdf
<code>pnorm(q, mean=0, sd=1, lower.tail=TRUE, log.p=FALSE)</code>	CDF
<code>qnorm(p, mean=0, sd=1, lower.tail=TRUE, log.p=FALSE)</code>	CDF <sup>-1</sup>
<code>rnorm(n, mean=0, sd=1)</code>	random number generator

Using the same naming convention, *R* provides these functions for many other common parametric distributions:

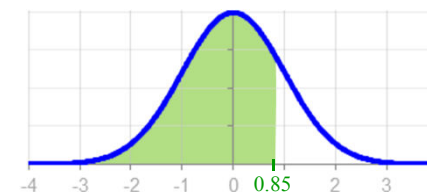
d_____pdf	q_____quantile
p_____CDF	r_____random number generator



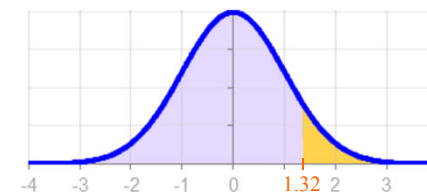
## Simple examples

**Example:** Let  $Z$  be the standard normal variable. Find:

(a)  $P(Z < 0.85) = (\text{area to the left of } 0.85)$   
 $= F(0.85) = \text{pnorm}(0.85) \approx 0.8023$

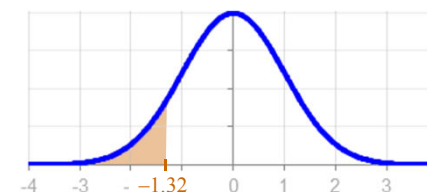


(b)  $P(Z > 1.32) = (\text{area to the right of } 1.32)$   
 $= 1 - (\text{area to the left of } 1.32)$   
 $= 1 - F(1.32) = 1 - \text{pnorm}(1.32) \approx 0.0934$



Alternatively, using the fact that pdf  $f$  is an even function, the area to the right of 1.32 is the same as the area to the left of  $-1.32$ :

$$P(Z > 1.32) = P(Z < -1.32)$$
$$= \text{pnorm}(-1.32) \approx 0.0934$$

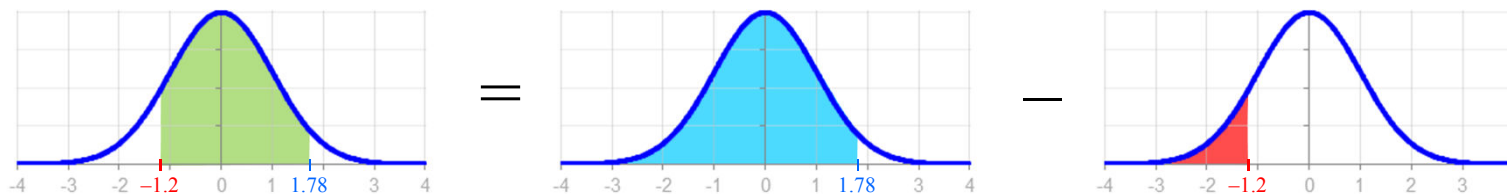


Furthermore:  $\text{pnorm}(1.32, \text{lower.tail}=\text{FALSE}) \approx 0.0934$



## Examples cont.

(c)  $P(-1.2 < Z < 1.78) =$  (area to the right of  $-1.2$  and to the left of  $1.78$ )



$=$  (area left of  $1.78$ ) minus (area left of  $-1.2$ )

$= F(1.78) - F(-1.2)$

$= \text{pnorm}(1.78) - \text{pnorm}(-1.2)$

$\approx 0.9625 - 0.1151 = 0.8474$



# Expected Value and Variance

OPIM  
5603  
Fall  
2019

**Definitions:** A random variable is *continuous* if it has the pdf.  
(if so, the pdf is *the derivative* of its CDF)

Given a continuous random variable  $X$  with a pdf  $f$ , we have

$$\begin{aligned}\text{Expected Value } EX &= \int_{-\infty}^{\infty} t f(t) dt, \\ \text{Variance } VarX &= E(X - EX)^2.\end{aligned}$$

Both values are real numbers (variance is non-negative).

**Properties:** Given a random variable  $X$  and a real number  $c$ :

$$(1) E(X + c) = EX + c \quad (3) Var(X + c) = VarX$$

$$(2) E(cX) = cEX \quad (4) Var(cX) = c^2 VarX$$

Note: (3) and (4) follow  
from (1) and (2),  
respectively.

Furthermore: for any random variables  $X$  and  $Y$ , (5)  $E(X + Y) = EX + EY$

if  $X$  and  $Y$  are *independent\**, (6)  $Var(X + Y) = VarX + VarY$

\* will be introduced in the next chapter.





## Additional properties of $N(\mu, \sigma^2)$

Given  $X \sim N(\mu, \sigma^2)$  ( $X$  has a  $N(\mu, \sigma^2)$  distribution),

$$(1) \quad EX = \int_{-\infty}^{\infty} t f(t) dt = \frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} t e^{-\frac{1}{2}\left(\frac{t-\mu}{\sigma}\right)^2} dt = (\text{substitutions, ...}) = \mu.$$

Note: This justifies naming the parameter  $\mu$  the *mean*.

$$(2) \quad \text{Var}X = E(X - EX)^2 = E(X - \mu)^2 = \dots = \sigma^2.$$

Hence the parameter  $\sigma$  is named the *standard deviation*.

(3) For any real numbers  $a \neq 0$  and  $b$ ,  $Y = aX + b$  is a normal random variable.

Furthermore,  $Y \sim N(a\mu + b, (a\sigma)^2)$ .

$$(4) \quad Z = \frac{X - \mu}{\sigma} \sim N(0, 1). \text{ (Standard Normal)}$$



## Examples

**Example:** Suppose  $X \sim N(3, 2)$ . What is the distribution of  $2X - 5$ ?

Solution:  $2X - 5$  is a normal random variable, by (3) on slide 1.41.

Notice that  $\mu = 3$ ,  $\sigma^2 = 2$ , and with notation used in (3),  $a = 2$  and  $b = -5$ .

Hence  $2X - 5 \sim N(a\mu + b, a^2\sigma^2) \sim N(2 \cdot 3 + (-5), 2^2 \cdot 2) \sim N(1, 8)$ .

Note:  $E(2X - 5) = E(2X) - 5 = 2E(X) - 5 = 2 \cdot 3 - 5 = 1$  (by (1), (2) on 1.40)

$Var(2X - 5) = Var(2X) = 2^2 Var(X) = 4 \cdot 2 = 8$  (by (3), (4) on 1.40)

**Example:** A particular rash has shown up at an elementary school. It has been determined that the length of time that the rash will last is normally distributed with  $\mu = 6$  days and  $\sigma = 1.5$  days. Find the probability that for a student selected at random, the rash will last for less than 3 days.

$$P(X < 3) = P(X - 6 < 3 - 6) = P\left(\frac{X - 6}{1.5} < \frac{3 - 6}{1.5}\right) = \text{pnorm}(-2) \approx 0.02275.$$

Standard Normal                      -2

Alternatively:  $P(X < 3) = P(N(6, 1.5^2) < 3) = \text{pnorm}(3, 6, 1.5) \approx 0.02275$ .



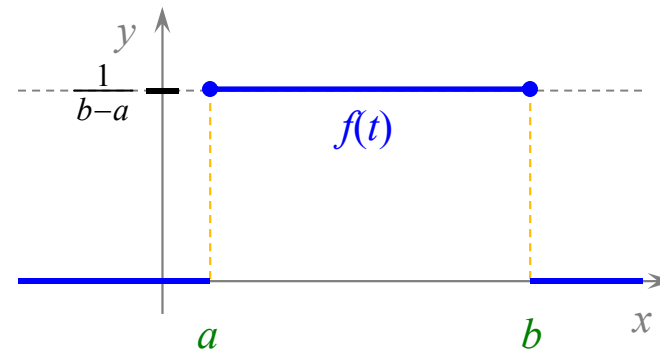
# Uniform Distribution

**Uniform distribution** is the simplest of continuous distributions.

It has two parameters: lower bound  $a$  and upper bound  $b$ , and it is usually denoted as  $U(a,b)$  or  $Unif(a,b)$ .

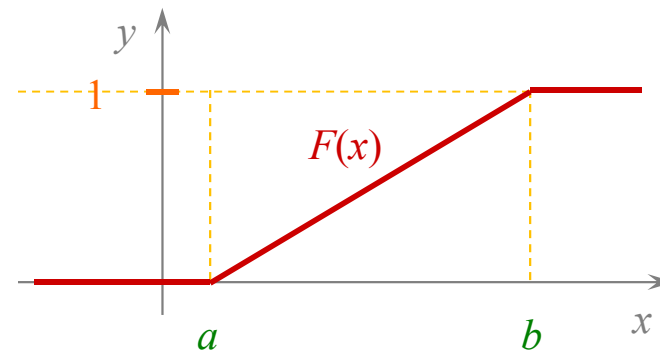
Probability density function:

$$f(t) = \begin{cases} \frac{1}{b-a}, & a \leq t \leq b \\ 0, & \text{otherwise} \end{cases}$$



Cumulative distribution function:

$$F(x) = \begin{cases} 0, & x < a \\ \frac{x-a}{b-a}, & a \leq x \leq b \\ 1, & x > b \end{cases}$$



$$\text{Expected value} = \text{Median} = \frac{a+b}{2}$$

$$\text{Variance} = \frac{(b-a)^2}{12}$$



## Examples: Properties (1) - (6) on 1.40

**Example:** Suppose that  $X \sim Unif(1,7)$ . Compute:

(a)  $E(X + 3) = \text{prop.}(1) = E(X) + 3 = 3 + 4 = 7.$

(b)  $E(4X) = \text{prop.}(2) = 4E(X) = 4 \cdot 4 = 16.$

(c)  $Var(X + 2) = \text{prop.}(3) = Var(X) = \frac{(7-1)^2}{12} + 2 = 3 + 2 = 5.$

(d)  $Var(5X) = \text{prop.}(4) = 5^2 \cdot Var(X) = 25 \cdot 3 = 75.$

(e)  $Var(9 - 2X) = \text{prop.}(3) = Var(-2X) = \text{prop.}(4) = (-2)^2 \cdot Var(X) = 4 \cdot 3 = 12.$

**Example:** Suppose that  $X \sim N(3,2)$  and  $Y \sim Unif(5,8)$ . Compute  $E(X + Y)$ .

Solution:  $E(X + Y) = \text{prop.}(5) = E(X) + E(Y) = 3 + \frac{5+8}{2} = 3 + 6.5 = 9.5$

**Example:** Assume that  $X$  and  $Y$  above are independent. Compute:

(a)  $Var(X + Y) = \text{prop.}(6) = Var(X) + Var(Y) = 2 + \frac{(8-5)^2}{12} = 2 + 0.75 = 2.75$

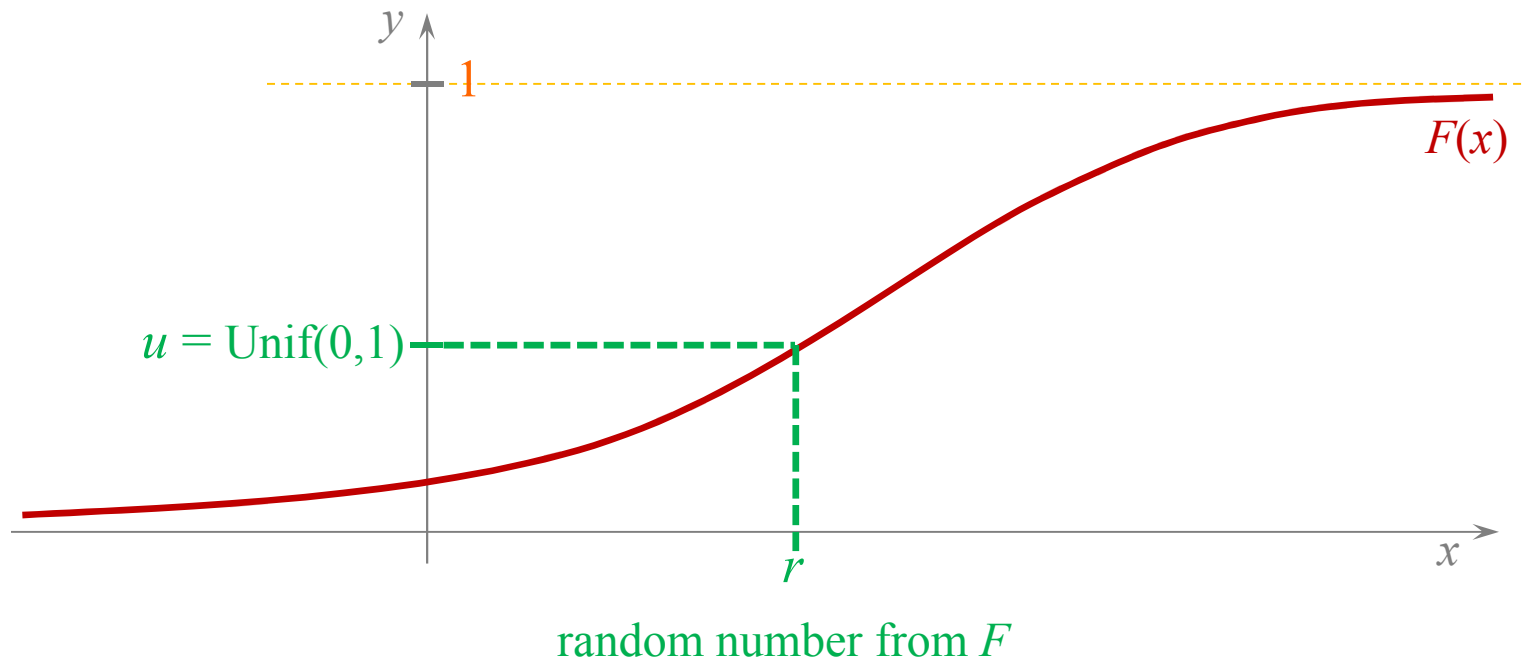
(b)  $Var(2X - 4Y) = \text{prop.}(6) = Var(2X) + Var(-4Y)$  ( $2X$  and  $-4Y$  are also independent)  
 $= \text{prop.}(4) = 2^2 \cdot Var(X) + (-4)^2 \cdot Var(Y) = 4 \cdot 2 + 16 \cdot 0.75 = 8 + 12 = 20$



# Generating random numbers from a CDF

Given a CDF  $F$  one can easily generate random numbers from this distribution, assuming:

- $F$  is an *invertible* function, and
- Uniform random number generator is available.



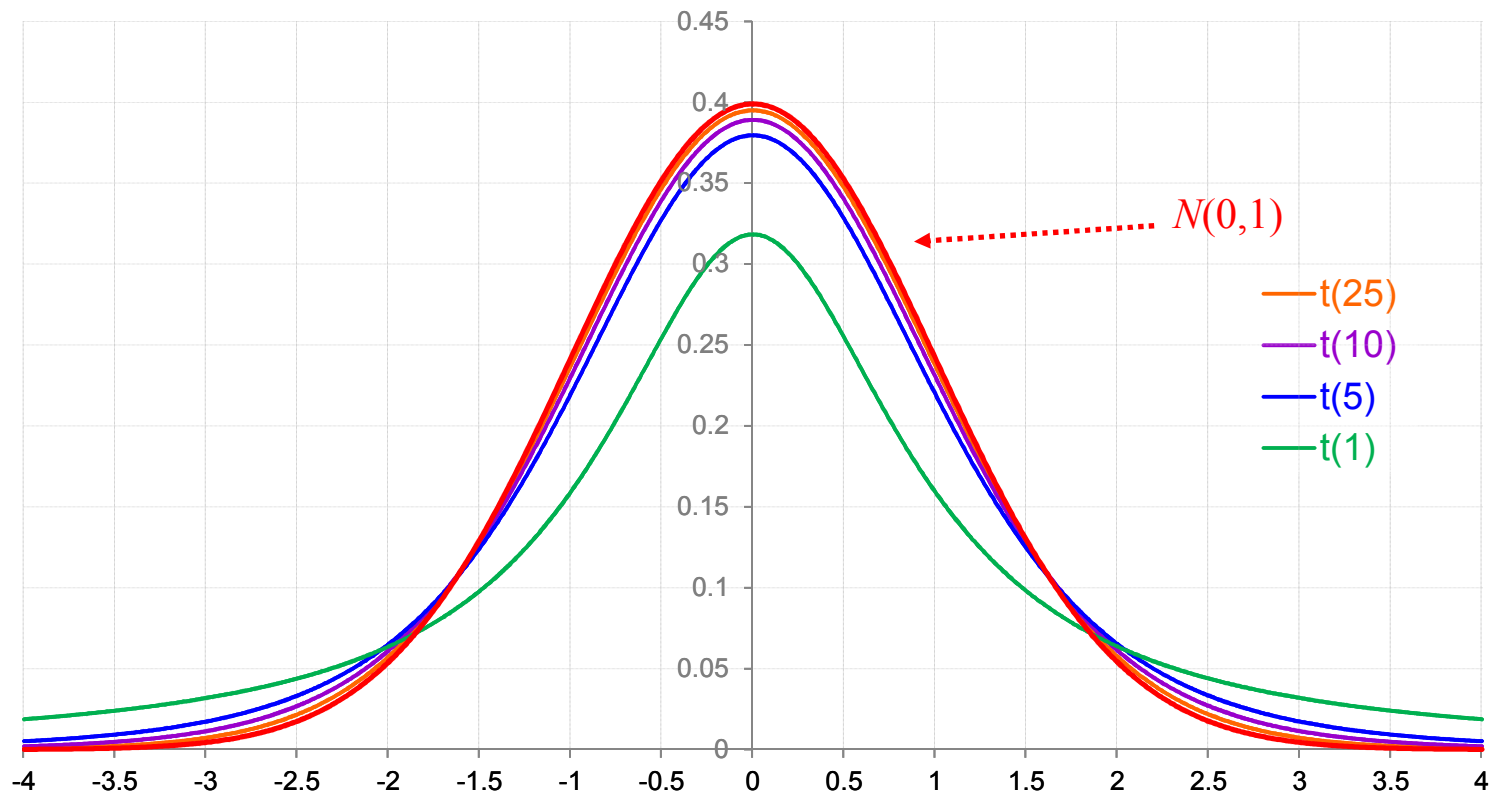
Hence  $r = F^{-1}(u)$ , where  $u$  is a uniform random number in  $[0,1]$ .



# Student $t$ distribution(s)

Student  $t$  distribution with  $n$  degrees of freedom is a continuous distribution with a probability density function

$$f(x) = \frac{1}{\sqrt{n\pi}} \frac{\Gamma\left(\frac{n+1}{2}\right)}{\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}}, \quad x \in (-\infty, \infty).$$

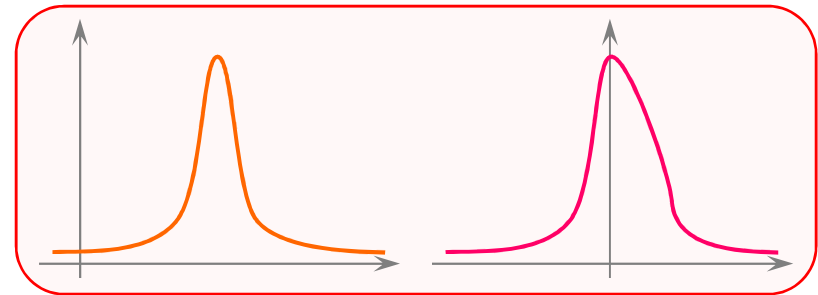
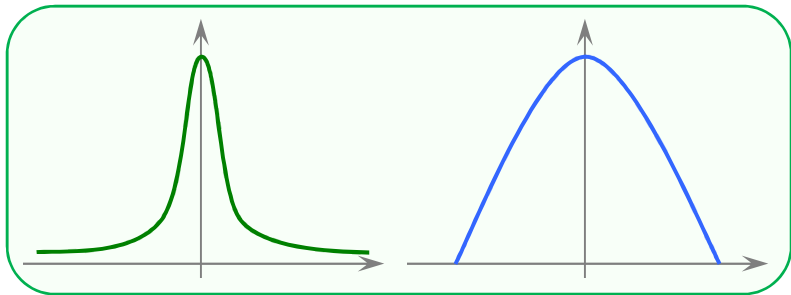




# Symmetric distributions

$N(0,1)$  and  $t$ -distributions are *symmetric*.

Random variable  $X$  is *symmetric* if  $X$  and  $-X$  have the same distribution. Equivalently, its pdf must be symmetric around the  $y$ -axis, i.e., must be an even function.



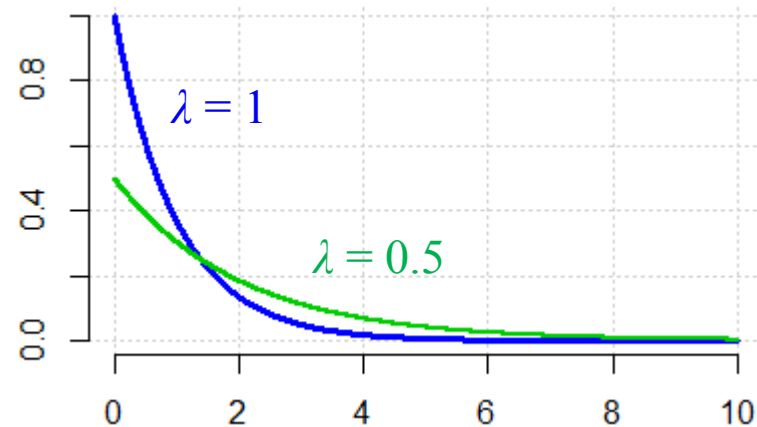


# Exponential Distribution

Exponential distribution has only one parameter: *rate*  $\lambda > 0$ .

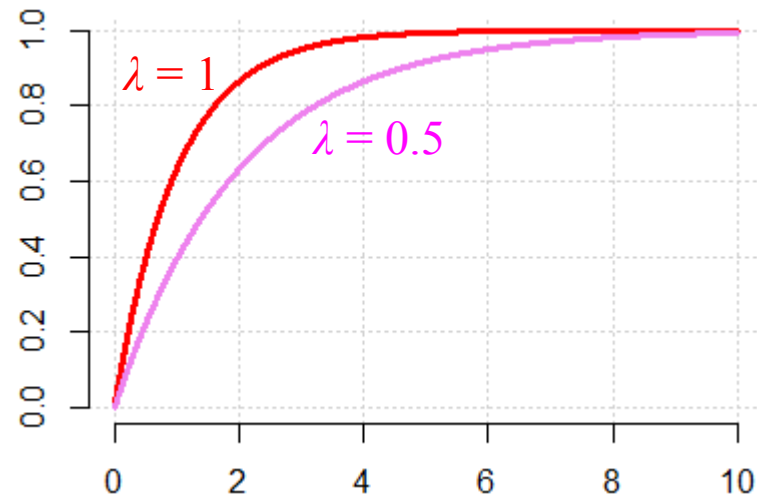
Probability density function:

$$f(t) = \begin{cases} \lambda e^{-\lambda t}, & t \geq 0 \\ 0, & \text{otherwise} \end{cases}$$



Cumulative distribution function:

$$F(x) = \begin{cases} 1 - e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$



$$\text{Expected value} = \frac{1}{\lambda} \quad \text{Median} = \frac{\ln(2)}{\lambda} \quad \text{Variance} = \frac{1}{\lambda^2}$$





# Poisson Arrival Process

*Arrival Process* is a sequence of random variables  $0 < A_1 < A_2 < \dots$  for which *interarrival times* random variables  $T_k = A_k - A_{k-1}$  (for  $k = 1, 2, \dots$  with  $A_0 = 0$ )

- 1) are positive,
- 2) have the same distribution (*identically distributed*), and
- 3) are *independent* (more on this shortly).

Notice that two or more arrivals cannot happen at exactly the same instant.

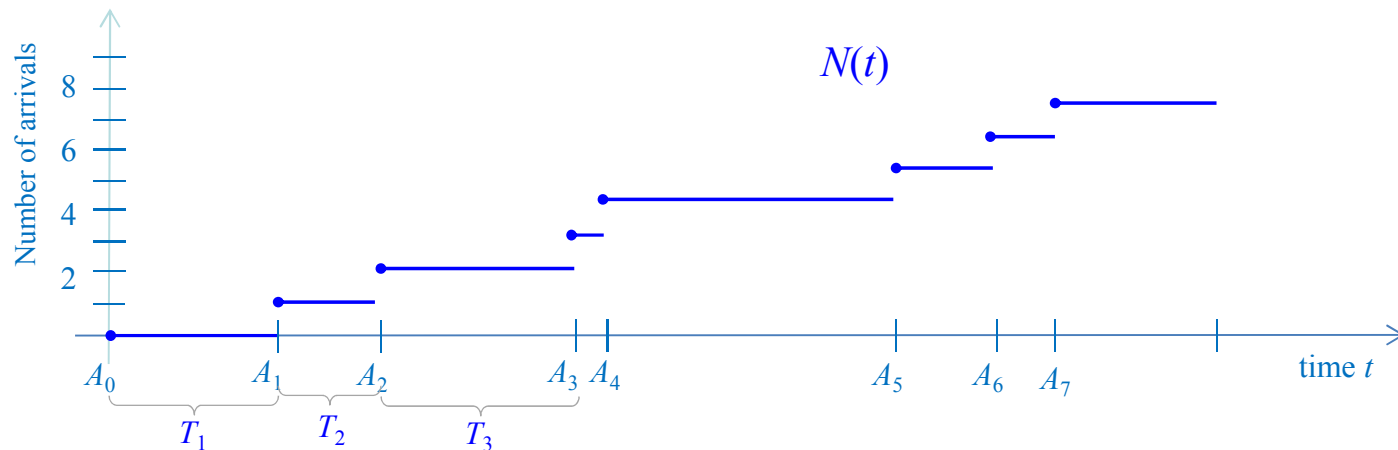


Illustration of arrival process:  $N(t)$  counts the number of arrivals by time  $t \geq 0$ .

*Poisson Arrival Process* is an arrival process whose interarrival times have  $Exp(\lambda)$  distribution.  $\lambda$  is called the Poisson Arrival Process *rate*.

The rate  $\lambda$  at which arrivals occur is constant: it cannot be higher in some intervals and lower in other intervals.

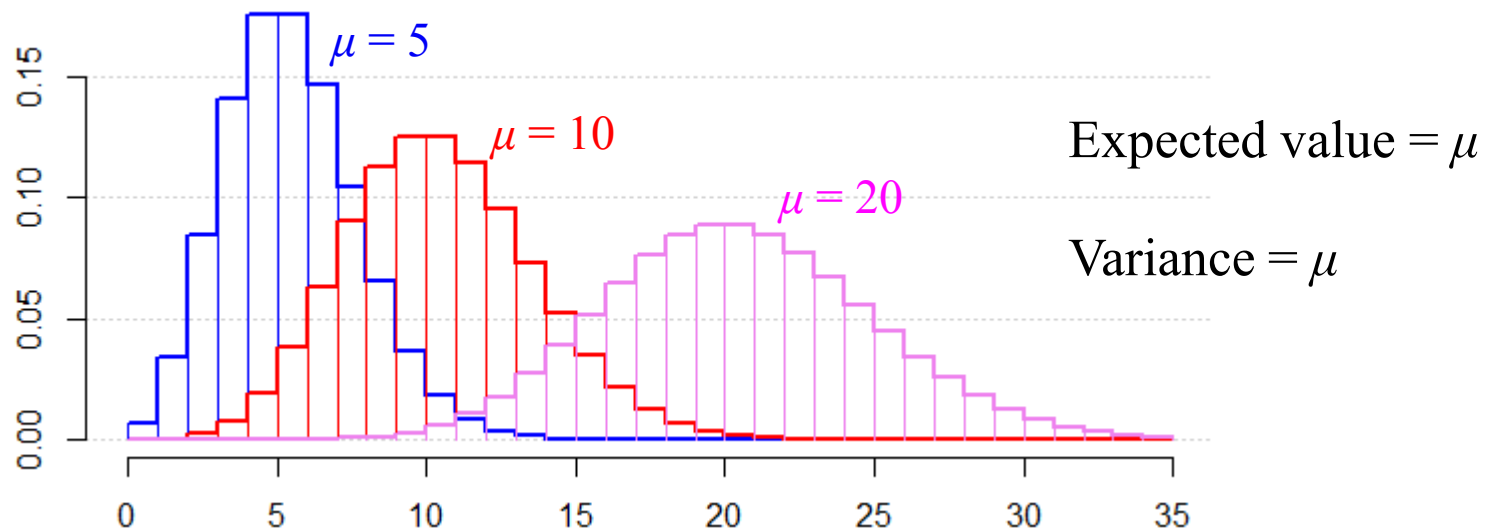


# Poisson Distribution

**Poisson distribution** is an infinite discrete distribution with a parameter  $\mu > 0$ .

Probability distribution:  $P(N = k) = e^{-\mu} \frac{\mu^k}{k!}$ ,  $k = 0, 1, 2, \dots$

Alternatively:  $N = Pois(\mu) \sim e^{-\mu} \begin{pmatrix} 0 & 1 & 2 & 3 & \dots \\ 1 & \mu & \frac{\mu^2}{2} & \frac{\mu^3}{6} & \dots \end{pmatrix}$



**Connection:** Number of arrivals of a Poisson process with rate  $\lambda$  in time interval of length  $\Delta t$  has  $Pois(\lambda \cdot \Delta t)$  distribution.

**Note:** Number of arrivals of a Poisson process depends only on the **length of the interval**, not on the interval itself (endpoints are not relevant).



## Example

**Example:** Arrival of customers at a county diner on Saturday afternoons is modeled as a *Poisson Arrival* process with a rate of 2.2 customers per minute. Let  $X$  be a random variable recording the number of arrivals between 12:45 PM and 1:05 PM.

(a) Is this variable finite discrete, infinite discrete, or continuous?

Solution: Number of customer arrivals is a non-negative integer. In principle it cannot be bounded: if you argue (for the sake of argument) that it is bounded by the population of the town, state, U.S., or the world, one could bring over the extra-terrestrials, parallel universes folks, etc., to exceed your bound however large it may be. Thus this is an infinite discrete random variable.

(b) What is the distribution of  $X$ ?

Solution: The rate of customer arrivals per minute is  $\lambda = 2.2$ . Since the customer arrival is a *Poisson Process*, the number of arrivals in time interval of length  $\Delta t$  minutes is a *Poisson random variable* with parameter  $\mu = \lambda \cdot \Delta t$ .

In this case  $\lambda = 2.2$  and  $\Delta t = 20$ , thus  $\mu = 2.2 \cdot 20 = 44$ . All told,  $X \sim \text{Pois}(44)$ .



## Example (cont.)

Example: Arrival of customers at a county diner on Saturday afternoons is modeled as a *Poisson Arrival* process with a rate of 2.2 customers per minute. Let  $X$  be a random variable recording the number of arrivals between 12:45 PM and 1:05 PM.

(c) What is the expected value of  $X$ ?

Solution:  $X \sim \text{Pois}(44)$  (part (b)), so  $E(X) = E(\text{Pois}(44)) = 44$ .

(d) What is the probability that at most 50 people (i.e., 50 or less) will arrive between 12:45 PM and 1:05 PM?

Solution: By (b),  $P(\text{at most 50 arrivals between 12:45 and 1:05 PM})$

$$= P(X \leq 50) = \text{ppois}(50, 44) \approx 0.836891$$



# Exponential Distribution is *Memoryless*

Optional

For any positive  $s, t$ :  $P(T > t + s | T > s) = P(T > t)$

Note: Conditional probability  $P(A|B) = \frac{P(A \cap B)}{P(B)}$

$$\begin{aligned} P(T > t + s | T > s) &= \frac{P((T > t + s) \cap (T > s))}{P(T > s)} \\ &= \frac{P(T > t + s)}{P(T > s)} = \frac{1 - F(t + s)}{1 - F(s)} = \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} \\ &= \frac{e^{-\lambda t} e^{-\lambda s}}{e^{-\lambda s}} = e^{-\lambda t} = 1 - F(t) = P(T > t) \end{aligned}$$

Consequence: Knowledge that arrival did not occur before time  $s$  does not yield any information about probability of its arrival between  $s$  and  $s + t$ ; this probability depends only on  $t$ .