# *Method*

## *Materials*

This study is part of the results obtained in the LAIC survey. This survey consists of 98 items and 12 classification items. There are 50 core items and 48 mantle items in LAIC. The core items concern to the measurement of attitudes about science while the mantle items analyse individual characteristics mediated by the social environment. In this piece of work, we have selected 26 items from both core and mantle to measure the image of manipulated science. The selection of variables is both the result of a process of decanting, which selects variables by iterating different models until the achievement of a suitable model, and of an analysis of the available theory based on previous studies, that allows to filter the best models by confronting them with state of the art research. Table 1 gives a complete overview of the items used in the research, as well as indications of their abbreviations, their full wording and their membership to latent constructs.

The CTS Unit of CIEMAT designed the core items used in the model, therefore the four items that make up the manipulated science measure, *Nu15*, *Nu16*, *Nu36* and *Nu44*, are an original creation. However, the mantle items came from a wide variety of sources. Following the order of the items found in Table 1: the items that make up the dogmatism measure come from two different studies. Items *P1*, *P2*, *P3* and *D3* are part of a study on actively open-minded thinking (Stanovich and Toplak, 2019). On the other hand, *D1* and *D2* were drawn from an article assessing dogmatic behaviour among students (Altemeyer, 2002). The items reflecting conspiracy thinking come from two studies: *Cons1* and *Cons2* from a survey on COVID-19 and attitudes towards vaccination (FECYT, 2021). Whereas, *Cons3* and *Cons4* were selected from a generic conspiracy beliefs scale (Drinkwater et al, 2020). Ideology is expressed as a construct with two components: a traditional ideological self-placement scale, *Selfpos*, and a worldview indicator. The worldview indicators *idconserva1*, *idconserva2*, *idconserva3* and *idconserva4* are part of a scale on ideological consistency (Pew Research Center, 2014). In addition, *np1*, *np2r* and *npe2r* come from a revision of the New Environmental Paradigm Scale (Dunlap et al, 2000). Finally, the item *Universalismor* was extracted from a value scale (Schwartz, 2006), *idprogre2r* comes from a study on ideology (Draca and Schwarz, 2018) and the CIEMAT research team developed *idprogre3r*.

All the scales used in this study use an 11-point response scale, with two exceptions: *Selfpos* that uses a 10-point scale and worldview, which is an aggregate indicator product of the sum of its items that ranges from 0 to 100.

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| Table 1. Items and constructs of the model | | |
| Construct | Item Name | Item |
| Manipulated | *Nu15* | Science is rigged |
|  | *Nu16* | Science is governed by hidden interests |
|  | *Nu36* | Official science manipulates facts to protect its interests |
|  | *Nu44* | Science is politicized |
| Dogmatism | *P1* | In this world, there are two main groups of people, the good and the bad. |
|  | *P2* | There are many paths, but only one is the right one. |
|  | *P3* | There are two types of people, those who think like me and those who think differently. |
|  | *D1* | There is only one right way of doing things. |
|  | *D2* | People who disagree with me are wrong. |
|  | *D3* | My beliefs are too important to abandon, even if I am given good arguments against them. |
| Conspiracy | *Cons1* | Government agencies are watching us |
|  | *Cons2* | There are secret organisations that determine political decisions. |
|  | *Cons3* | There are ways to access the truth that science tries to hide. |
|  | *Cons4* | There is official science on the one hand and the real science on the other. |
| Ideology | *Worldview* | - |
|  | *Selfpos* | When talking about politics, the expressions left and right are commonly used. Where do you place yourself on a scale where 1 means extreme left and 10 means extreme right? |
| Worldview | *idconserva1* | Social inequality is necessary for the country to prosper |
|  | *idconserva2* | Immigrants are a burden on our country because they take away our jobs, housing and health care. |
|  | *idconserva3* | The current problems will be solved by the market. |
|  | *idconserva4* | Government investment in public services is a waste of money. |
|  | *npe1* | The ecological crisis has been exaggerated |
|  | *npe3r* | The impacts of industry jeopardise the balance of nature |
|  | *npe2r* | The impact of our actions on nature has disastrous consequences. |
|  | *Universalismor* | Everyone should have the same opportunities |
|  | *idprogre2r* | Government must take responsibility for protecting the whole population. |
|  | *idprogre3r* | Things would be better in Spain if there were less inequality. |

## *Ethics statement*

The CTS Unit adheres to the CIEMAT code of ethics, which can be found [here](https://www.ciemat.es/portal.do?TR=A&IDR=1&identificador=945). In addition, the company responsible for the fieldwork IKERFEL adheres to the best practice standards set by ESOMAR, Insights Association, as well as complying with requirements such as ISO 20252 and ISO 26362.

## *Participants and procedure*

LAIC started with a sample of 4671 Spanish residents over 16 years old from a panel sample provided by [IKERFEL](https://www.ikerfel.es/), a field company specialising in social research. The source of participants for LAIC was an online panel with more than 2,500,000 Spanish residents. This panel uses the sociodemographic variables of gender, age, educational level, household size and geographical area to achieve a more reliable picture of Spanish society.

After passing the IKERFEL quality filters and dropping voluntary withdrawals from the questionnaire without completing it, a sample of 2698 participants was obtained.

This sample was also reduced to 1981 observations leaving only those participants who had filled in all the mantle and core items, as a result the variables assessed in our models had no missing cases. Therefore, the final sample consists of 980 men (49.47%), 996 women (50.28%) and 5 individuals (0.25%) declaring themselves as Other. In terms of the other socio-demographic variables, the sample is more skewed towards older participants with a level of education below short-cycle tertiary education. Table 2 shows the age and educational level in the sample. The anonymized raw database, the data cleaning R code, the cleaned database, the R code used for the statistical analysis and a document with a recollection of all the tables and plots used in the analysis can be found here: <https://github.com/IkerSoriaCIEMAT/LAIC_CTS_CIEMAT>

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| Table 2. Distribution of age and education level in the sample | | | |  |  |
| Age groups | Frequencies | Percentages | Education level | Frequencies | Percentages |
| 16-24 | 247 | 12.47% | Lower secondary education | 392 | 19.79% |
| 25-34 | 277 | 13.98% | Upper secondary education | 515 | 26.00% |
| 35-44 | 373 | 18.83% | Post-secondary non-tertiary education | 209 | 10.55% |
| 45-54 | 417 | 21.05% | Short-cycle tertiary education or bachelor | 641 | 32.36% |
| 55-64 | 338 | 17.06% | Master’s or equivalent level and above | 220 | 11.11% |
| 65 years of age and older | 329 | 16.61% | NAs | 4 | 0.20% |

# *Results*

## *Initial data analysis*

It is important to mention that the first analyses we performed, mainly Shapiro-Wilk test, recommended to assess univariate normality (Razali and Wah, 2011: 32), and Rosner's test for the presence of extreme cases (Rosner, 1983: 171), indicated univariate non-normality and the presence of outliers, the test results can be found in the Table A1 of Appendix A. These tests inclined the research team towards the use of robust location and scale parameters that were better fitted for this situation (Pérez, 2016: 76). The trimmed mean as an alternative to the mean and the NMAD as an alternative to the standard deviation were therefore adopted.

A complete account of the trimmed mean, NMAD, minimum value, maximum value and reliability scores is provided in Table 3. Regarding the reliability scores most of the scales show adequate values in both Alpha and Omega, with one exception the Ideology scale.

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| Table 3. Distribution of variables and reliability scores | | | | |  |  |  |
| Construct | Item Name | Trimmed mean\* | NMAD | Min | Max | α | ω |
| Manipulated | *Nu15* | 4.44 | 2.97 | 0 | 10 | 0.83 | 0.85 |
|  | *Nu16* | 3.95 | 2.97 | 0 | 10 |  |  |
|  | *Nu36* | 4.92 | 2.97 | 0 | 10 |  |  |
|  | *Nu44* | 5.97 | 1.48 | 0 | 10 |  |  |
| Dogmatism | *P1* | 5.52 | 4.45 | 0 | 10 | 0.75 | 0.84 |
|  | *P2* | 4.03 | 2.97 | 0 | 10 |  |  |
|  | *P3* | 3.59 | 4.45 | 0 | 10 |  |  |
|  | *D1* | 3.15 | 2.97 | 0 | 10 |  |  |
|  | *D2* | 1.28 | 1.48 | 0 | 10 |  |  |
|  | *D3* | 4.56 | 2.97 | 0 | 10 |  |  |
| Conspiracy | *Cons1* | 7.11 | 2.97 | 0 | 10 | 0.72 | 0.75 |
|  | *Cons2* | 6.14 | 2.97 | 0 | 10 |  |  |
|  | *Cons3* | 4.73 | 2.97 | 0 | 10 |  |  |
|  | *Cons4* | 5.02 | 2.97 | 0 | 10 |  |  |
| Ideology | *Worldview* | 22.86 | 16.47 | 0 | 100 | 0.55 | 0.55 |
|  | *Selfpos* | 4.78 | 1.48 | 1 | 10 |  |  |
| Worldview | *idconserva1* | 1.36 | 1.48 | 0 | 10 | 0.80 | 0.84 |
|  | *idconserva2* | 2.34 | 2.97 | 0 | 10 |  |  |
|  | *idconserva3* | 3.42 | 2.97 | 0 | 10 |  |  |
|  | *idconserva4* | 0.79 | 0.00 | 0 | 10 |  |  |
|  | *npe1* | 2.52 | 2.97 | 0 | 10 |  |  |
|  | *npe3r* | 1.70 | 2.97 | 0 | 10 |  |  |
|  | *npe2r* | 1.16 | 1.48 | 0 | 10 |  |  |
|  | *Universalismor* | 0.66 | 0.00 | 0 | 10 |  |  |
|  | *idprogre2r* | 1.01 | 1.48 | 0 | 10 |  |  |
|  | *idprogre3r* | 1.59 | 2.97 | 0 | 10 |  |  |
| \*Trimmed at 0.2 | | | | | | | |

Table 4 provides the correlations between the scales used in this study. We decided to use the Kendall Rank Correlation (Kendall, 1938: 81) because it does not require parametric data and because it can be used with both continuous and ordinal variables.

Regarding the correlations themselves, all the scales were significantly correlated, with positive correlations ranging between *r*=.14 and *r*=.5.

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| Table 4. Kendall Rank Correlation Coefficients of the scales in the model. | | | | |
|  | Manipulated | Dogmatism | Conspiracy | Selfpos |
| Manipulated | 1 |  |  |  |
| Dogmatism | 0.20\*\*\* | 1 |  |  |
| Conspiracy | 0.50\*\*\* | 0.26\*\*\* | 1 |  |
| Selfpos | 0.20\*\*\* | 0.24\*\*\* | 0.14\*\*\* | 1 |
| \*=p<0.05 \*\*=p<0.01 \*\*\*=p<0.001 | | |  |  |

## *Structural equation model (SEM) rationale and assumptions.*

After analysing various methods of statistical analysis, including multiple linear regression, we deemed SEM to be the appropriate method to explain the inherent relationships between the latent and manifest variables (Urbano, 2013: 147) under our theoretical framework.

Prior to the development of the SEM, we applied two tests to check the assumptions of multivariate normality required for the use of the technic in this dataset. The Mardia test (Korkmaz et al, 2014: 152) showed a statistically significant result in both Kurtosis (p<0.001) and Skewness (p<0.001). The Energy test (Székely and Rizzo, 2004: 1254), with 10,000 bootstrap replicates, was also statistically significant (p<0.001). With these two tests, we acknowledged the non-normality of the multivariate distribution and therefore to address the lack of multivariate normality we decided to apply 10,000 bootstrap replicates to our model (Kline, 2015: 42).

The other assumptions of SEM, sample size, multicolinearity and outliers, were also assessed prior to the development of the model. Regarding the sample size, and following the literature (Wolf et al, 2013: 921), we found that with 1981 observations, a p-value threshold of 0.05 and a theoretical model with four latent variables, the requirement for the number of observations was fulfilled in this case.

We paid close attention to any possible multicolinearity between the variables in the model. However, the correlations between manifest variables, that can be found at Table B1 in Appendix B, showed that no variable was correlated at a level greater than or equal to *r*=.85 (Weston and Gore, 2006: 735). In addition to correlations, we obtained the variance inflation factors (VIFs) of different models using as predictors the variables in the model and a dummy binary variable as dependent variable to test strong relations between variables that may have been overlooked in the correlations (Franke, 2010: 1). None of the VIFs for any of the variables in each of the models tested showed a value higher than 2.2 well below the suggested 2.5 VIF value of significant colinearity (Johnston et al, 2018: 1958), the VIFs results are displayed on Appendix B Table B2. Because of this analysis, we concluded that the multicolinearity assumption was not a concern in this particular study.

The strategy used to deal with outliers was somewhat more complex. Following a comprehensive guide on this topic in SEM (Aguinis et al, 2013: 292), we opted to test our model without outliers and with outliers to see if there were any differences in the results. We found no significant divergences between the two models in any of the metrics used to assess model fit, nonetheless the indices and estimates of the outlier free model can be found at the Appendix B Tables B3, B4 and B5 for further enquiries. In addition, and following the aforementioned guidance (Aguinis et al, 2013: 293), our approach using robust resampling techniques also had the indirect result of reducing any effects of outliers in the estimators.

## *SEM fit indices.*

Under all the above mentioned assumptions, the fit indices of our model were as follow: 𝜒2 = 401.447, df = 89, p < 0.001; RMSEA = 0.042, CI [.038, .046]; SRMR = 0.038; CFI = 0.968; TLI = 0.957; GFI = 0.975; AGFI = 0.961.

The p-value of the 𝜒2 test was found to be statistically significant. However, concerning this significant value two constrains need to be highlighted: non-normality and sample size. Previous studies suggest that non-normality often has a negative effect on the value of the 𝜒2 test that can lead to the rejection of models that otherwise could offer relevant information (Hayduk et al, 2007: 847). Earlier on this paper, both Mardia and Energy tests shown that the data follows a non-normal distribution and these results support the conclusion that the 𝜒2 test may suffer from a biased negative effect as a by-product of the lack of normality.

Sample size, in like manner to non-normality, can also create an artificial statistically significant value of 𝜒2. Research on this topic has shown that small-scale differences between model and data when using large samples can increase the risk of a significant result in 𝜒2 (Kline, 2015: 201). Consequently, we find important to acknowledge that the rather large dataset used for this model, with more than 1900 observations, may be another factor contributing to the significant result of 𝜒2.

With these clarifications, we do not want to disregard as whole the result of the 𝜒2 test, but rather to highlight that the results of this test are mediated by a series of factors that can lead to the rejection of a model when rejection is not necessary. In short, we can not rely solely on the result of this test on account of the abovementioned constrains and hence the need for the goodness-of-fit indices.

Overall, the results of the model's goodness-of-fit indices are adequate albeit far from perfect. The RMSEA is below the threshold of 0.06 that the literature mentions (Hooper et al, 2008) as well as the SRMR that is below its 0.08 cut-off (Hooper et al, 2008). Both statistics show that the model fits well the data that there are no major differences between the observed data and the model.

The CFI and TLI of the model surpassed the cut-off value of 0.95 (Hu and Bentler, 1999). With respect to GFI and AGFI the values of these two goodness-of-fit indices where above their traditional thresholds, GFI above 0.9 (Hooper et al., 2008) and AGFI above the 0.9 cut-off (Schermelleh-engel et al., 2003). Both of these two indices were of great importance for this model with non-normal data due to their robustness when dealing with this type of distributions (Ainur et al., 2017) as they added a layer of security in the acceptance of the model.

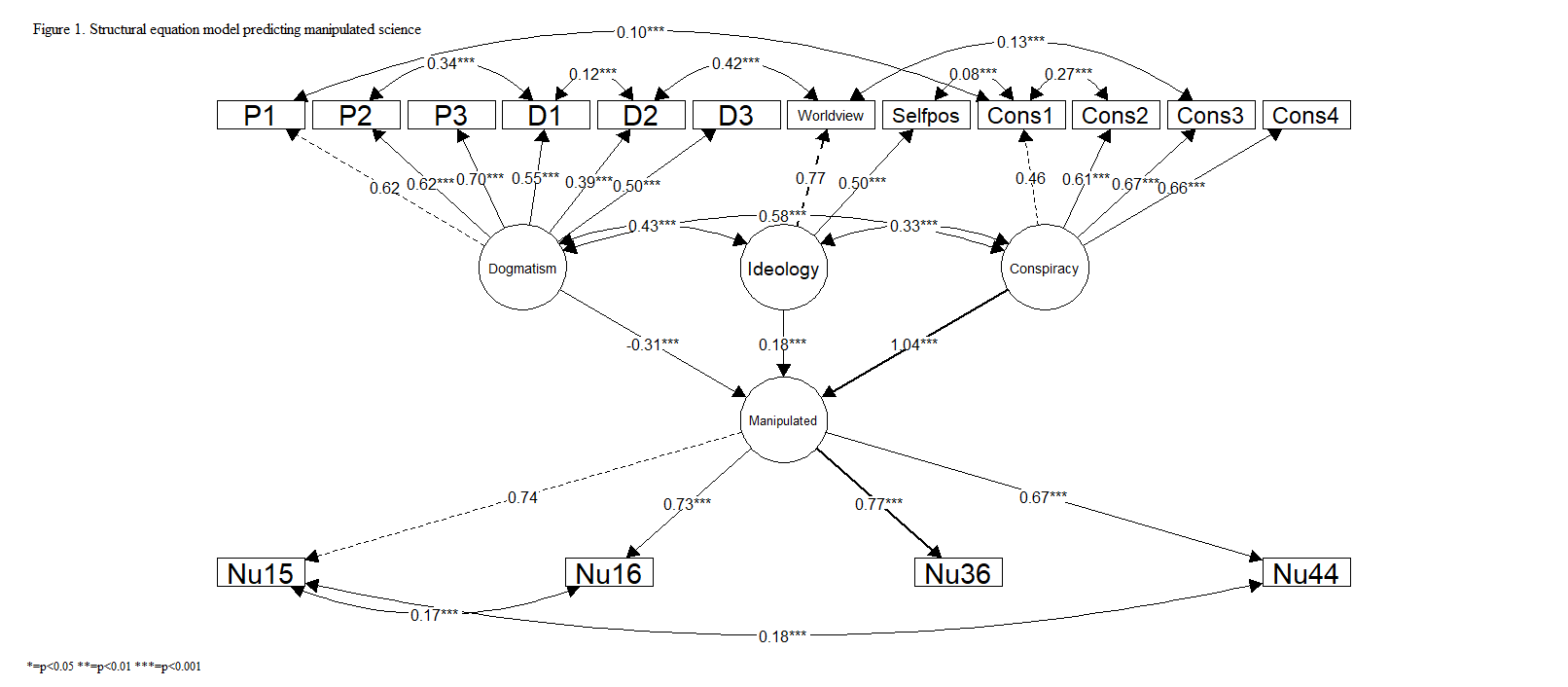
## *SEM parameter estimates and correlations.*

The structure of the model with the standardized estimates and their significance level is displayed in Figure 1. All the relations between manifest variables, correlations and paths between latent variables are statistically significant (p<0.01). This model shows a very strong positive relationship between Manipulated and Conspiracy (β=1.04) and a lesser positive relation between Manipulated and Ideology (β=.18). On the other hand, there seems to be a negative relation between Manipulated and Dogmatism (β=-.31). Altogether, these three latent variables explain 91.7% of the variance of Manipulated.

Regarding the correlations of the latent variables we found positive strong correlations between Dogmatism and Ideology, *r*=.43, and Dogmatism and Conspiracy, *r*=.58. Ideology and Conspiracy were also correlated, *r*=.33, although to a lesser extent. These correlations suggest that individuals that are more conservative will tend also to have strong conspiracy beliefs and a dogmatic mentality in clear opposition to more progressive individuals.

Leaving aside that a significant part of the Dogmatism items were correlated with each other, as were the Conspiracy items, the correlations between items of different latent variables show relationships of interest that are worth highlighting. Between these manifest variables, higher values in Worldview (i.e. higher conservative worldview) were positively correlated with D2 (*People who disagree with me are wrong*), with a correlation of *r*=.42. D2 is one of the key measures of Dogmatism and its correlation with Worldview indicates, again, a relationship between conservative views and a tendency to view other positions not as the other side of the coin but as fundamentally incorrect. We also found out that Worldview was also slightly correlated with Cons3 (*There are ways to access the truth that science tries to hide*), *r*=.13. By linking, as with latent variables, conservatism and conspiratorial ideas.

Less well correlated were Cons1 (*Government agencies are watching us*) and both P1 (*In this world there are two main groups of people, the good and the bad*), *r*=.1, and Selfpos (*When talking about politics, the expressions left and right are commonly used. Where do you place yourself on a scale where 1 means extreme left and 10 means extreme right?*), *r*=.08*.* These correlations hint a very tenous relation between a Manichean interpretation of social reality and state surveillance and to an even more tenuous relationship between right-wing thinking and the aforementioned state surveillance.



# *Discussion*

**Appendix A. Univariate normality and outlier detection**

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| Table A1. Shapiro and Rosner tests | | | | | | |
| Item Name | W | p-value | Observation number | R | Lambda | Outlier |
| *Nu15* | 0.96 | <.001 | 33 | 2.06 | 4.20 | 0 |
|  |  |  | 61 | 2.06 | 4.20 | 0 |
|  |  |  | 64 | 2.07 | 4.20 | 0 |
| *Nu16* | 0.94 | <.001 | 24 | 2.14 | 4.20 | 0 |
|  |  |  | 40 | 2.14 | 4.20 | 0 |
|  |  |  | 61 | 2.14 | 4.20 | 0 |
| *Nu36* | 0.95 | <.001 | 33 | 1.80 | 4.20 | 0 |
|  |  |  | 61 | 1.80 | 4.20 | 0 |
|  |  |  | 64 | 1.81 | 4.20 | 0 |
| *Nu44* | 0.95 | <.001 | 47 | 2.27 | 4.20 | 0 |
|  |  |  | 50 | 2.27 | 4.20 | 0 |
|  |  |  | 59 | 2.27 | 4.20 | 0 |
| *P1* | 0.92 | <.001 | 32 | 1.63 | 4.20 | 0 |
|  |  |  | 47 | 1.63 | 4.20 | 0 |
|  |  |  | 52 | 1.64 | 4.20 | 0 |
| *P2* | 0.93 | <.001 | 40 | 1.90 | 4.20 | 0 |
|  |  |  | 61 | 1.90 | 4.20 | 0 |
|  |  |  | 64 | 1.91 | 4.20 | 0 |
| *P3* | 0.89 | <.001 | 24 | 1.77 | 4.20 | 0 |
|  |  |  | 34 | 1.77 | 4.20 | 0 |
|  |  |  | 46 | 1.77 | 4.20 | 0 |
| *D1* | 0.91 | <.001 | 34 | 2.20 | 4.20 | 0 |
|  |  |  | 56 | 2.20 | 4.20 | 0 |
|  |  |  | 114 | 2.21 | 4.20 | 0 |
| *D2* | 0.80 | <.001 | 98 | 3.31 | 4.20 | 0 |
|  |  |  | 114 | 3.32 | 4.20 | 0 |
|  |  |  | 167 | 3.33 | 4.20 | 0 |
| *D3* | 0.94 | <.001 | 56 | 1.88 | 4.20 | 0 |
|  |  |  | 64 | 1.89 | 4.20 | 0 |
|  |  |  | 68 | 1.89 | 4.20 | 0 |
| *Cons1* | 0.92 | <.001 | 8 | 2.68 | 4.20 | 0 |
|  |  |  | 59 | 2.68 | 4.20 | 0 |
|  |  |  | 60 | 2.69 | 4.20 | 0 |
| *Cons2* | 0.94 | <.001 | 32 | 2.02 | 4.20 | 0 |
|  |  |  | 59 | 2.02 | 4.20 | 0 |
|  |  |  | 60 | 2.02 | 4.20 | 0 |
| *Cons3* | 0.95 | <.001 | 24 | 1.86 | 4.20 | 0 |
|  |  |  | 26 | 1.86 | 4.20 | 0 |
|  |  |  | 47 | 1.86 | 4.20 | 0 |
| *Cons4* | 0.94 | <.001 | 50 | 1.73 | 4.20 | 0 |
|  |  |  | 85 | 1.73 | 4.20 | 0 |
|  |  |  | 102 | 1.73 | 4.20 | 0 |
| *Worldview* | 0.96 | <.001 | 498 | 4.76 | 4.20 | 1 |
|  |  |  | 1280 | 3.88 | 4.20 | 0 |
|  |  |  | 233 | 3.82 | 4.20 | 0 |
| *Selfpos* | 0.96 | <.001 | 60 | 2.38 | 4.20 | 0 |
|  |  |  | 114 | 2.39 | 4.20 | 0 |
|  |  |  | 184 | 2.39 | 4.20 | 0 |
| *idconserva1* | 0.79 | <.001 | 46 | 3.00 | 4.20 | 0 |
|  |  |  | 98 | 3.00 | 4.20 | 0 |
|  |  |  | 134 | 3.01 | 4.20 | 0 |
| *idconserva2* | 0.85 | <.001 | 27 | 2.25 | 4.20 | 0 |
|  |  |  | 42 | 2.26 | 4.20 | 0 |
|  |  |  | 50 | 2.26 | 4.20 | 0 |
| *idconserva3* | 0.93 | <.001 | 37 | 2.54 | 4.20 | 0 |
|  |  |  | 259 | 2.55 | 4.20 | 0 |
|  |  |  | 275 | 2.55 | 4.20 | 0 |
| *idconserva4* | 0.72 | <.001 | 98 | 3.32 | 4.20 | 0 |
|  |  |  | 134 | 3.33 | 4.20 | 0 |
|  |  |  | 193 | 3.34 | 4.20 | 0 |
| *npe1* | 0.87 | <.001 | 46 | 2.20 | 4.20 | 0 |
|  |  |  | 64 | 2.20 | 4.20 | 0 |
|  |  |  | 81 | 2.21 | 4.20 | 0 |
| *npe3r* | 0.86 | <.001 | 28 | 3.75 | 4.20 | 0 |
|  |  |  | 184 | 3.76 | 4.20 | 0 |
|  |  |  | 221 | 3.78 | 4.20 | 0 |
| *npe2r* | 0.81 | <.001 | 221 | 4.26 | 4.20 | 1 |
|  |  |  | 233 | 4.28 | 4.20 | 1 |
|  |  |  | 498 | 4.31 | 4.20 | 1 |
| *Universalismor* | 0.74 | <.001 | 498 | 5.02 | 4.20 | 1 |
|  |  |  | 572 | 5.05 | 4.20 | 1 |
|  |  |  | 904 | 5.08 | 4.20 | 1 |
| *idprogre2r* | 0.78 | <.001 | 68 | 4.49 | 4.20 | 1 |
|  |  |  | 75 | 4.51 | 4.20 | 1 |
|  |  |  | 328 | 4.54 | 4.20 | 1 |
| *idprogre3r* | 0.83 | <.001 | 49 | 3.43 | 4.20 | 0 |
|  |  |  | 175 | 3.44 | 4.20 | 0 |
|  |  |  | 192 | 3.46 | 4.20 | 0 |

**Appendix B. SEM assumptions.**



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| Table B2. Regression results of the multicollinearity tests | | |
| A. Coefficients and VIFs of the model with Manipulated Science variables | | |
| Variables | Coefficients | VIFs |
| Intercept | 0.54 | ꟷ |
| Nu15 | 0.00 | 2.10 |
| Nu16 | 0.00 | 1.80 |
| Nu36 | 0.00 | 1.77 |
| Nu44 | -0.01 | 1.66 |
| B. Coefficients and VIFs of the model with Dogmatism variables | | |
| Variables | Coefficients | VIFs |
| Intercept | 0.53 | ꟷ |
| P1 | 0.00 | 1.35 |
| P2 | -0.01 | 1.67 |
| P3 | 0.00 | 1.48 |
| D1 | -0.01 | 1.61 |
| D2 | 0.00 | 1.17 |
| D3 | 0.00 | 1.20 |
| C. Coefficients and VIFs of the model with Ideology variables | | |
| Variables | Coefficients | VIFs |
| Intercept | 0.49 | ꟷ |
| Worldview | 0.00 | 1.17 |
| Selfpos | 0.00 | 1.17 |
| D. Coefficients and VIFs of the model with Conspiracy variables | | |
| Variables | Coefficients | VIFs |
| Intercept | 0.52 | ꟷ |
| Cons1 | 0.00 | 1.33 |
| Cons2 | 0.00 | 1.52 |
| Cons3 | 0.00 | 1.33 |
| Cons4 | -0.01 | 1.36 |

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| Table B3. Goodness of Fit Indices of the SEM model with no outliers. | | | | | | | | | |
| Model | χ2 | df | RMSEA | 90% CI | SRMR | CFI | TLI | GFI | AGFI |
| SEM model no outliers | 401.447 | 89 | 0.04 | 0.038 - 0.046 | 0.04 | 0.97 | 0.96 | 0.98 | 0.96 |

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| --- | --- | --- | --- | --- | --- | --- | --- |
| Table B4. Loadings of the latent variables in the SEM model with no outliers. | | | | | | | |
| A. Estimates of loadings | |  |  |  |  |  |  |
| Latent variables | Measurement variables | Loadings | 95% CI of the loadings | | Standardized estimates | p-value | SMC |
| Lower | Upper |
| Manipulated | Nu15 | 1 | 1 | 1 | 0.74 | ꟷ | 0.55 |
|  | Nu16 | 1.02 | 0.96 | 1.09 | 0.73 | <.001 | 0.534 |
|  | Nu36 | 1.10 | 1.03 | 1.17 | 0.77 | <.001 | 0.597 |
|  | Nu44 | 0.86 | 0.80 | 0.91 | 0.67 | <.001 | 0.442 |
| Polarisation | P1 | 1 | 1 | 1 | 0.62 | ꟷ | 0.379 |
|  | P2 | 0.94 | 0.85 | 1.03 | 0.62 | <.001 | 0.383 |
|  | P3 | 1.17 | 1.08 | 1.27 | 0.70 | <.001 | 0.484 |
|  | D1 | 0.80 | 0.7 | 0.901 | 0.55 | <.001 | 0.302 |
|  | D2 | 0.47 | 0.4 | 0.55 | 0.39 | <.001 | 0.151 |
|  | D3 | 0.72 | 0.6 | 0.81 | 0.50 | <.001 | 0.247 |
| Ideology | Worldview | 1 | 1 | 1 | 0.77 | ꟷ | 0.59 |
|  | Selfpos | 0.09 | 0.07 | 0.11 | 0.50 | <.001 | 0.253 |
| Conspiracy | Cons1 | 1 | 1 | 1 | 0.46 | ꟷ | 0.215 |
|  | Cons2 | 1.49 | 1.34 | 1.67 | 0.61 | <.001 | 0.367 |
|  | Cons3 | 1.63 | 1.45 | 1.84 | 0.67 | <.001 | 0.45 |
|  | Cons4 | 1.64 | 1.47 | 1.85 | 0.66 | <.001 | 0.436 |
| B. Covariances |  |  |  |  |  |  |  |
| Measurement variables | | Covariances | 95% CI of the covariances | | Correlations | p-value |  |
| Lower | Upper |  |
| Nu15 | Nu44 | 0.61 | 0.37 | 0.87 | 0.18 | <.001 |  |
|  | Nu16 | 0.60 | 0.32 | 0.88 | 0.17 | <.001 |  |
| P2 | D1 | 2.06 | 1.63 | 2.48 | 0.35 | <.001 |  |
| D2 | Worldview | 9.48 | 7.89 | 11.06 | 0.42 | <.001 |  |
| P1 | Cons1 | 0.61 | 0.32 | 0.90 | 0.10 | <.001 |  |
| Cons1 | Cons2 | 1.41 | 1.09 | 1.74 | 0.27 | <.001 |  |
| D1 | D2 | 0.66 | 0.40 | 0.92 | 0.12 | <.001 |  |
| Worldview | Cons3 | 2.77 | 1.22 | 4.26 | 0.13 | <.001 |  |
| Selfpos | Cons1 | 0.35 | 0.16 | 0.54 | 0.08 | <.001 |  |

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| Table B5. Estimates of the SEM model with no outliers. | | | |  |  |  |  |
| A. Regression weights | |  |  |  |  |  |  |
| Endogenous variables | Exogenous variables | Estimate | 95% CI of the estimates | | Standarized estimates | p-value | SMC |
| Lower | Upper |
| Manipulated | Ideology | 0.03 | 0.02 | 0.04 | 0.18 | <.001 | 0.92 |
|  | Dogmatism | -0.31 | -0.41 | -0.22 | -0.31 | <.001 |  |
|  | Conspiracy | 1.77 | 1.55 | 2.04 | 1.05 | <.001 |  |
| B. Covariances of exogenous variables | | |  |  |  |  |  |
| Exogenous variables | | Covariances | 95% CI of the covariances | | Correlations | p-value |  |
| Lower | Upper |  |
| Polarisation | Ideology | 10.64 | 8.93 | 12.37 | 0.43 | <.001 |  |
|  | Conspiracy | 1.40 | 1.18 | 1.63 | 0.58 | <.001 |  |
| Ideology | Conspiracy | 4.71 | 3.67 | 5.77 | 0.33 | <.001 |  |

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