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$\{x \sim R[0, \theta]\}$ распределение на отрезке $[0, \theta]$.

$$\hat{\theta}_1 = 2\bar{x} \quad \hat{\theta}_2 = x_{\min} \quad \hat{\theta}_3 = x_{\max} \quad \hat{\theta}_4 = x \cdot \frac{\sum_{k=1}^n x_k}{n-1}$$

$$D[\hat{\theta}_1] = \frac{1}{9} [COV] M[\hat{\theta}_1] = \frac{\theta^2}{9} \quad D[\hat{\theta}_2] = \frac{\theta^2}{n}$$

a) $\hat{\theta}_1 = 2\bar{x} \quad M[\hat{\theta}_1] = \theta \Rightarrow$ несмещ.

$$D[\hat{\theta}_1] = D\left[\frac{2}{n} \sum_{i=1}^n x_i\right] = \frac{4}{n^2} \sum_{i=1}^n D[x_i] = \frac{4}{n^2} \cdot n \cdot \frac{\theta^2}{3} = \frac{4\theta^2}{3n} \rightarrow 0 \text{ как } n \rightarrow \infty$$

$$\hat{\theta}_2 = x_{\min} \quad \varphi(y) = 1 - (1 - F(y))^n$$

$$\varphi(y) = 1 - (1 - F(y))^n \quad f(y) = n(1 - F(y))^{n-1} \cdot \frac{1}{\theta} \cdot \{COV\}$$

$$M[\hat{\theta}_2] = \int_0^\theta n(1 - \frac{y}{\theta})^{n-1} \cdot \frac{1}{\theta} dy = \int_0^1 n(1-t)^{n-1} dt = \left\{ \begin{matrix} t = 1 - \frac{y}{\theta} \\ dt = -\frac{1}{\theta} dy \end{matrix} \right\} =$$

$$= - \int_1^0 n(1-t)^{n-1} dt = \int_0^1 n(1-t)^{n-1} dt = \int_0^1 n t^{n-1} dt = \int_0^1 n t^{n-1} dt =$$

$$= \theta \left(1 - \frac{1}{n+1}\right) = \frac{\theta}{n+1} \Rightarrow \text{смещен}$$

$$\hat{\theta}_3 = (n+1)x_{\min}$$

$$D[\hat{\theta}_3] = (n+1)^2 D[\hat{\theta}_2]$$

$$M[\hat{\theta}_3] = \int_0^\theta n(1 - \frac{y}{\theta})^{n-1} \cdot \frac{1}{\theta} dy = \left\{ \begin{matrix} t = 1 - \frac{y}{\theta} \\ dt = -\frac{1}{\theta} dy \end{matrix} \right\} =$$

$$= - \int_1^0 n(1-t)^{n-1} dt = \int_0^1 n(1-t)^{n-1} dt = \int_0^1 n t^{n-1} dt = 2 \int_0^1 n t^{n-1} dt +$$

$$+ \int_0^1 n t^{n-1} dt = \theta \left[\frac{n}{n+1} - 2 \frac{n}{n+1} + 1 \right] =$$

$$= \theta \left[\frac{n(n+1) - 2n(n+1) + (n+1)(n+2)}{(n+1)(n+2)} \right] = \theta \left[\frac{2}{(n+1)(n+2)} \right] = \frac{2\theta^2}{(n+1)(n+2)}$$

$$D[\hat{\theta}_2] = \frac{2\theta^2}{(n+1)(n+2)} - \frac{\theta^2}{(n+1)^2} = \frac{\theta^2(2n+2-n-2)}{(n+1)^2(n+2)} = \frac{n\theta^2}{(n+1)^2(n+2)} \rightarrow 0 \text{ как } n \rightarrow \infty$$

$$D[\hat{\theta}_3] = \frac{1\theta^2}{n+2} \rightarrow 0 \text{ как } n \rightarrow \infty$$

$$\hat{\theta}_4 \text{ не смещ.}$$

$$\forall \theta > 0 \quad \forall \varepsilon > 0 \quad P(|\hat{\theta}_4 - \theta| > \varepsilon) \rightarrow 0 \text{ как } n \rightarrow \infty$$

$$P(|\hat{\theta}_4 - \theta| > \varepsilon) = P(\hat{\theta}_4 > \theta + \varepsilon) = P((n+1)x_{\min} > \theta + \varepsilon) =$$

$$= P(x_{\min} > \frac{\theta + \varepsilon}{n+1}) = 1 - P(x_{\min} \leq \frac{\theta + \varepsilon}{n+1}) =$$

$$= 1 - (1 - (1 - F(\frac{\theta + \varepsilon}{n+1}))^n) = (1 - \frac{\theta + \varepsilon}{\theta(n+1)})^n \rightarrow e^{-\frac{\theta + \varepsilon}{\theta}} > 0 \Rightarrow \text{не смещ.}$$

$$\hat{\theta}_4 \text{ не смещ.}$$

$$P(|\hat{\theta}_4 - \theta| > \varepsilon) = P(\hat{\theta}_4 \leq \theta - \varepsilon) + P(\hat{\theta}_4 \geq \theta + \varepsilon) =$$

$$= P(x_{\min} < \theta - \varepsilon) = \Phi(\theta - \varepsilon) = 1 - (1 - \frac{\theta - \varepsilon}{\theta})^n =$$

$$= 1 - (\frac{\varepsilon}{\theta})^n \rightarrow 1 \text{ как } n \rightarrow \infty$$

$$\hat{\theta}_5 = x_{\max} \quad M[\hat{\theta}_5] = \int_0^\theta n(\frac{z}{\theta})^{n-1} dz = \frac{n}{n+1} \theta \Rightarrow \text{смещен}$$

$$\hat{\theta}_5' = \frac{n+1}{n} x_{\max} \quad D[\hat{\theta}_5'] = (\frac{n+1}{n})^2 D[\hat{\theta}_5]$$

$$D[\hat{\theta}_5] = \frac{n}{n+2} \theta^2 - (\frac{n}{n+1})^2 \theta^2 = \theta^2 \left[\frac{n(n+1) - n^2(n+2)}{(n+1)^2(n+2)} \right] = \theta^2 \left[\frac{n}{(n+1)^2(n+2)} \right] \rightarrow 0 \text{ как } n \rightarrow \infty$$

$$D[\hat{\theta}_5'] = \frac{\theta^2}{n(n+2)} \rightarrow 0 \text{ как } n \rightarrow \infty$$

$\tilde{\theta}_n$ no exp. $\forall \theta > 0 \quad \forall \varepsilon > 0$

$$P(|\tilde{\theta}_n - \theta| > \varepsilon) = P(X_{\max} \leq \theta - \varepsilon) + P(X_{\max} \geq \theta + \varepsilon) =$$

$$= F^n(\theta - \varepsilon) = \left(\frac{\theta - \varepsilon}{\theta}\right)^n \xrightarrow{n \rightarrow \infty} 0; \quad \varepsilon \geq \theta \rightarrow (0)^n \xrightarrow{n \rightarrow \infty} 0 \quad \text{--- conf.}$$

$\tilde{\theta}_n'$ no exp.

$$\forall \theta > 0 \quad \forall \varepsilon > 0 \quad P(|\tilde{\theta}_n' - \theta| > \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

$$P(|\tilde{\theta}_n' - \theta| > \varepsilon) = P(\tilde{\theta}_n' \leq \theta - \varepsilon) + P(\tilde{\theta}_n' \geq \theta + \varepsilon)$$

P_1 P_2

$$P_1: P(X_{\max} \leq \frac{\theta - \varepsilon}{1 + \frac{1}{n}}) = \left(F\left(\frac{\theta - \varepsilon}{1 + \frac{1}{n}}\right)\right)^n$$

$$\varepsilon > \theta: 0^n \rightarrow 0, n \rightarrow \infty$$

$$\text{occeco: } \left(\frac{\theta - \varepsilon}{\theta(1 + \frac{1}{n})}\right)^n = \left(\frac{1 - \frac{\varepsilon}{\theta}}{1 + \frac{1}{n}}\right)^n \xrightarrow{n \rightarrow \infty} 0 \quad \left. \begin{matrix} P_1 \rightarrow 0 \\ n \rightarrow \infty \end{matrix} \right\}$$

$$P_2: P(X_{\max} \geq \frac{\theta + \varepsilon}{1 + \frac{1}{n}}) = 1 - \left(F\left(\frac{\theta + \varepsilon}{1 + \frac{1}{n}}\right)\right)^n = 1 - \left(\frac{1 + \frac{\varepsilon}{\theta}}{1 + \frac{1}{n}}\right)^n$$

$$\frac{\theta + \varepsilon}{1 + \frac{1}{n}} > \theta: (1 - 1)^n \rightarrow 0 \quad n \rightarrow \infty \quad \Rightarrow P_2 \rightarrow 0$$

$n \rightarrow \infty$

$$\exists n_0: \forall n > n_0 \quad \frac{n(\theta + \varepsilon)}{n+1} > \theta$$

$$\Rightarrow P_1 + P_2 \xrightarrow{n \rightarrow \infty} 0 \quad \Rightarrow \text{conf.}$$

$$\tilde{\theta}_n = X_1 + \frac{1}{n-1} \sum_{k=2}^n X_k$$

$$M[\tilde{\theta}_n] = M\left[X_1 + \frac{1}{n-1} \sum_{k=2}^n X_k\right] = M[X_1] + \frac{1}{n-1} \sum_{k=2}^n M[X_k] = \frac{\theta}{2} + \frac{\theta}{2} = \theta \Rightarrow \text{Классиф.}$$

$$D[\tilde{\theta}_n] = D\left[X_1 + \frac{1}{n-1} \sum_{k=2}^n X_k\right] = D\left[\left\{\frac{1}{n-1} \cdot (n-1) D[X_k]\right\}\right] = \frac{\theta^2}{12} + \frac{\theta^2}{12(n-1)} \approx$$

$$\approx \frac{n\theta^2}{12(n+1)} \xrightarrow{n \rightarrow \infty} 0 \quad \Rightarrow \text{good var. не рад.}$$

$\{X_1, \dots, X_n\}$ - независимы и одинаково распределены \Rightarrow используем 3.5.4 Хитчинса

$$\frac{1}{n} \sum_{k=1}^n X_k \xrightarrow{P} M[X] \Rightarrow \frac{1}{n-1} \sum_{k=2}^n X_k \xrightarrow{P} M[X] = \theta \Rightarrow \tilde{\theta}_n = X_1 + \frac{1}{n-1} \sum_{k=2}^n X_k \xrightarrow{P} X_1 + \frac{\theta}{2} \Rightarrow \text{не сход.}$$

б)

$$\Rightarrow \text{conf: } \tilde{\theta}_n = \bar{X} \quad \tilde{\theta}_n' = \frac{n-1}{n} X_{\max}$$

$$D[\tilde{\theta}_n] = \frac{\theta^2}{12n} \quad D[\tilde{\theta}_n'] = \frac{\theta^2}{12(n-1)}$$

$$\frac{\theta^2}{12(n-1)} < \frac{\theta^2}{12n} \quad \forall \theta > 0 \quad \forall n > 1 \Rightarrow$$

$$\Rightarrow \tilde{\theta}_n' \text{ --- лучше чем } \tilde{\theta}_n$$