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$$p(x) = \begin{cases} \frac{e^{-x/\theta}}{\theta}, & x \geq 0 \\ 0, & x < 0. \end{cases} \quad \theta > 0 \quad F(x) = \int_{-\infty}^x p(x) dx = \begin{cases} 1 - e^{-x/\theta}, & x \geq 0 \\ 0, & x < 0. \end{cases}$$

$$n=3 \quad \theta: \tilde{\theta}_1 = \bar{x}, \tilde{\theta}_2 = X_{(2)}.$$

$$M[\xi] = \int_0^{+\infty} x e^{-x/\theta} \cdot \frac{1}{\theta} dx = \frac{1}{\theta} (-\theta x e^{-x/\theta}) \Big|_0^{+\infty} + \theta \int_0^{+\infty} e^{-x/\theta} dx = \\ = \int_0^{+\infty} e^{-x/\theta} dx = \theta e^{-x/\theta} \Big|_0^{+\infty} = \theta$$

$$M[\xi^2] = \int_0^{+\infty} x^2 e^{-x/\theta} \cdot \frac{1}{\theta} dx = \frac{1}{\theta} \int_0^{+\infty} x^2 e^{-x/\theta} dx = \frac{1}{\theta} (-\theta x^2 e^{-x/\theta}) \Big|_0^{+\infty} + \\ + 2\theta \int_0^{+\infty} x e^{-x/\theta} dx = 2\theta^2$$

$$D[\xi] = M[\xi^2] - M^2[\xi] = \theta^2$$

$\tilde{\theta}_1$

$$M[\tilde{\theta}_1 | (X_n)] = M\left[\sum_{i=1}^3 X_i\right] \cdot \frac{1}{3} = \frac{1}{3} \cdot 3 M[\xi] = \theta \text{ - верно.}$$

$\tilde{\theta}_2$

$$p(y) = n p(y) \binom{n-1}{n-1} (1 - F(y))^{n-1} (F(y))^{k-1} = \binom{n-1}{k-1} (1 - e^{-y/\theta})^{n-k} (e^{-y/\theta})^{k-1} \\ M[\tilde{\theta}_2] = M[X_{(2)}] = \int_0^{+\infty} y p(y) dy = \int_0^{+\infty} \frac{6}{\theta} (4e^{-y/\theta} - ye^{-y/\theta}) dy = \\ = \frac{6}{\theta} \left( -\frac{4}{2} ye^{-y/\theta} \Big|_0^{+\infty} + \frac{4}{2} ye^{-y/\theta} \Big|_0^{+\infty} + \frac{\theta}{2} \int_0^{+\infty} e^{-y/\theta} dy - \frac{\theta}{2} \int_0^{+\infty} e^{-y/\theta} dy \right) = \\ = 6 \left( \frac{1}{2} \frac{6}{2} (-e^{-y/\theta}) \Big|_0^{+\infty} - \frac{1}{2} \frac{\theta}{2} (-e^{-y/\theta}) \Big|_0^{+\infty} \right) = \frac{5\theta}{6} \Rightarrow \tilde{\theta}_2 \text{ - неверно.}$$

$$\hat{\theta}_1' = \frac{5}{6} \tilde{\theta}_2 \text{ - верно.}$$



2)  $\bar{\theta}_1$ :

$$D[\bar{\theta}_1] = D\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \frac{1}{n^2} \cdot 32 D[x] = \frac{\theta^2}{n}$$

$$\begin{aligned} \bar{\theta}_1 &= M[\bar{\theta}_1^2] = M\left[\left(\frac{1}{n} \sum_{i=1}^n x_i\right)^2\right] = \frac{1}{n^2} \int_0^{+\infty} (x^2 e^{-\frac{x}{\theta}} - x^2 e^{-\frac{2x}{\theta}}) dx = \\ &= \frac{1}{n^2} \left( -\theta x^2 e^{-\frac{x}{\theta}} \Big|_0^{+\infty} + \frac{2}{n} \int_0^{+\infty} x e^{-\frac{x}{\theta}} dx - 2 \cdot \frac{1}{n} \int_0^{+\infty} x e^{-\frac{2x}{\theta}} dx \right) = \\ &= \frac{1}{n^2} \left( \frac{2}{n} \int_0^{+\infty} x e^{-\frac{x}{\theta}} dx - \frac{2}{n} \int_0^{+\infty} x e^{-\frac{2x}{\theta}} dx \right) = \frac{1}{n^2} \left( \frac{2}{n} \cdot \theta^2 - \frac{2}{n} \cdot \frac{\theta^2}{2} \right) = \frac{1}{n^2} \cdot \theta^2 \end{aligned}$$

$$D[\bar{\theta}_1] = M[\bar{\theta}_1^2] - M^2[\bar{\theta}_1] = \frac{1}{n^2} \theta^2$$

$$D[\bar{\theta}_1] = \frac{36}{25} \cdot \frac{1}{16} \theta^2 = \frac{1}{25} \theta^2$$

$D[\bar{\theta}_1] < D[\bar{\theta}_2] \Rightarrow \bar{\theta}_1$  - более точ. оц.

$$3) \hat{\theta}(\theta) = M\left[\left(\frac{\partial}{\partial \theta} \ln p(x, \theta)\right)^2\right] = M\left[\left(\frac{\partial}{\partial \theta} \ln p(x, \theta)\right)^2\right]$$

$$= M\left[\frac{1}{\theta^2} \left(-\frac{x}{\theta} - \ln \theta\right)^2\right] = M\left[\frac{1}{\theta^2} \left(\frac{x^2}{\theta^2} + \frac{2x}{\theta} + \ln^2 \theta\right)\right] =$$

$$= M\left[\frac{1}{\theta^2} \left(\frac{x^2}{\theta^2} + \frac{2x}{\theta} + \ln^2 \theta\right)\right] = M\left[\frac{x^2}{\theta^4} + \frac{2x}{\theta^3} + \frac{\ln^2 \theta}{\theta^2}\right] = \frac{1}{\theta^4} M[x^2] + \frac{2}{\theta^3} M[x] + \frac{\ln^2 \theta}{\theta^2}$$

$$+ \frac{1}{\theta^2} = \frac{1}{\theta^4} \cdot \frac{1}{\theta} + \frac{2}{\theta^3} \cdot \frac{1}{\theta} + \frac{1}{\theta^2} = \frac{1}{\theta^2}$$

$\bar{\theta}_1$ :

$$D[\bar{\theta}_1] = \frac{\theta^2}{n}, \text{ где } n \in \mathbb{N} \text{ и } n \geq 1 \text{ и } \theta \in (0, +\infty)$$

$$(a, b) \subset (0, +\infty) \Rightarrow \forall \theta \in (a, b) \Rightarrow \frac{\theta^2}{n} \leq \frac{b^2}{n} \Rightarrow$$

$\Rightarrow \bar{\theta}_1$  - не точ. оц.

$\bar{\theta}_2$ :

$$D[\bar{\theta}_2] = \frac{1}{n^2} \theta^2 - \text{где } n \in \mathbb{N} \text{ и } n \geq 1 \text{ и } \theta \in (0, +\infty)$$

$$(a, b) \subset (0, +\infty) \Rightarrow \forall \theta \in (a, b) \Rightarrow \frac{1}{n^2} \theta^2 \leq \frac{1}{n^2} b^2 \Rightarrow$$

$\Rightarrow \bar{\theta}_2$  - не точ. оц.

$$p(x, \theta) = P(x, \theta) - \text{где } x \in \mathbb{R} \text{ и } \theta \in (0, +\infty)$$

$$\frac{\partial}{\partial \theta} p(x, \theta) = \frac{x e^{-\frac{x}{\theta}}}{\theta^2} - \frac{e^{-\frac{x}{\theta}}}{\theta^2}$$

$$\int_0^{+\infty} \frac{\partial}{\partial \theta} p(x, \theta) dx = \int_0^{+\infty} \left( \frac{x e^{-\frac{x}{\theta}}}{\theta^2} - \frac{e^{-\frac{x}{\theta}}}{\theta^2} \right) dx = \left( -\frac{x}{\theta^2} e^{-\frac{x}{\theta}} \right) \Big|_0^{+\infty} +$$

$$+ \int_0^{+\infty} \frac{1}{\theta^2} e^{-\frac{x}{\theta}} dx = \int_0^{+\infty} \frac{1}{\theta^2} e^{-\frac{x}{\theta}} dx = 0$$

$$\frac{\partial}{\partial \theta} \int_0^{+\infty} p(x, \theta) dx = \int_0^{+\infty} \frac{\partial}{\partial \theta} p(x, \theta) dx$$

$$\hat{\theta}(\theta) = \frac{1}{\theta^2} \in C[0, +\infty)$$

$$\hat{\theta}(\theta) > 0 \quad \forall \theta \in (0, +\infty)$$

$\Rightarrow$  не точ. оц.  $\Rightarrow$  можно использовать н.б. критерия - Rao

$$\forall \theta \in (0, +\infty) \Rightarrow D[\bar{\theta}_1] \geq \frac{1}{n^2} \theta^2 \geq \frac{1}{n^2} \theta^2$$

$$\bar{\theta}_1 \geq \bar{\theta}_2, D[\bar{\theta}_1] = \frac{1}{n^2} \theta^2 \Rightarrow \text{не точ. оц. } \bar{\theta}_1 - \text{точ.}$$

$\bar{\theta}_1 \neq \bar{\theta}_2 \Rightarrow \bar{\theta}_2$  - не точ. оц. т.к. может быть только одна оц.