## 1 Differential Privacy on graphs

### 1.1 Edge density

**Proposition 1.1.** The sufficient and necessary level of the noise to  $\varepsilon$ -differentiably private estimate the edge density of a graph (where the set of graphs have a "known" degree distribution, scales like  $\frac{\max_i d_i}{n^2 \varepsilon}$ . In paricular for  $\rho$ -sparse graphon like  $\frac{\rho_n}{n \varepsilon}$ .

## 1.2 SBMs

**Definition 1.2.** We say that two (undirected) graphs  $G_1$  and  $G_2$  have distance 1 if  $G_1$  is generated from  $G_2$  by deleting a node from  $G_1$  and all its adjacent edges. Let d be the natural metric induced by this operation.

We start with k-Stochastic Block Models where the clustering sizes can vary arbitrarily.

**Proposition 1.3.** Let  $M, \Delta \in \mathbb{N}$  and  $\varepsilon, \eta > 0$  such that  $M > 2^{\Delta \varepsilon + 1}$ .

Suppose that  $Q_1, \ldots, Q_M \in [0,1]^{k \times k}$  induce k-SBM's with the property that  $\min_{i,j} ||Q_i - Q_j||_2 \ge \eta$  and any two  $Q_i, Q_j$  induce almost surely graphs that are at most  $\Delta$ -close. Then any  $\varepsilon$ -differential private mechanism M should have rate at least  $\eta/2$ .

*Proof.* Same idea as in the survey. Also k-close Q's imply k-close graphs (check this with a coupling possible argument).

**Proposition 1.4.** Let k > 2 (possibly not necessary). The rate for  $\varepsilon$ -diff approximating the matrix Q from the SBM is at least of the order  $\frac{k}{n\varepsilon} = \frac{1}{\varepsilon} \sqrt{\frac{k^2}{n^2}}$ . That is for any  $\varepsilon$ -differential private estimator  $\hat{Q}$  it holds  $\|\hat{Q} - Q\|_2^2 = \Omega(\frac{1}{\varepsilon}\sqrt{\frac{k^2}{n^2}})$ .

*Proof.* Choose a family F of  $2^{k/2}$  sets  $S \subset \{0,1\}^k$  with  $|S\triangle T| \ge k/4$  for any two  $S \ne T$  both elements of F. (check constants).

Then choose Q=0 with the block structure that the first k/2 vertices correspond to a block of size  $\lambda>0$  and after these we have blocks of size  $2(n-\lambda k)/k$ . The parameter  $\lambda>0$  will be tuned later. We enumerate the blocks so that the first k/2 blocks correspond to the sizes  $\lambda>0$ .

Now for every  $S \in F$ , we define  $Q_S$  we define the k-SBM probability matrix such that for the block corresponding to  $i \in S$ , we adjust the inwards and outwards probabilities to be equal to 1, that is  $Q_S = 1_{b(i) \in S, j \in [n]}$ , where by b(i) we refer to the block that vertex i belongs.

Now clearly every pair  $Q_S, Q_T$  is at most  $\lambda k$ -away and  $\|Q_S - Q_T\|_2^2 \ge c' |S\triangle T| n = c\lambda k n$ . Hence by setting  $\lambda = \frac{1}{7\varepsilon}$  we have that  $M := |F| > 2^{\Delta \varepsilon + 1}$  since k > 2 we are done.

Remark: The above argument using the result from Gao et al [16] can be boosted for the equipartition case to  $\lambda \rho$  rate. In particular, no reasonable estimation seems possible in this case with differentiable private estimation. **Differential private algorithms can not give the clustering partition!** 

**Proposition 1.5.** There exists a constant c > 0 small enough such that if  $\varepsilon < c \frac{k^2}{n}$  then the trivial estimator is the optimal  $\varepsilon$ -differentiable private estimator for estimating the probability matrix of a k-SBM.

*Proof.* Using the Vashanorv-Gilber bound (probabilistic method) we can find  $M = 2^{\Omega(k^2)} > 2^{\varepsilon n+1}$  matrices  $B_i \in \{0,1\}^{k \times k}$  such that  $\inf_{i \neq j} \|B_i - B_j\|_2 = \Omega(1)$ . Since each two k-SBMs induces graphs that are n-close we get the result using the Proposition 1.3.

**Proposition 1.6.** There exists c > 0 such that if  $\varepsilon < c \log k$ , then the trivial estimator is the optimal  $\varepsilon$ -differentiable private estimator for estimating the probability matrix of the induced graph of a k-SBM.

*Proof.* Indeed in pages 26-27 of Gao et al, we see the existence of  $\exp(cn \log k)$  probability matrices of graphs induced by balances k-SBMs with pairwise  $\ell_2$  distance constant. The result follows from the obvious variant of Proposition 1.3.

**Proposition 1.7.** Suppose for some graphon W we see the  $A = G_n(W)$  and then we want to  $\varepsilon$ -differential private estimate the graphon W given A, by producing say the estimator  $\hat{W}$ . Then if  $1 > \varepsilon > \frac{1}{n}$ , it necessarily holds for some W that  $\mathbb{E}\left[\delta_2(\hat{W}, W)\right] \geq \Omega(\frac{1}{n\varepsilon})$ .

*Proof.* Assume not. Then for all W using Cauchy-Shwartz we get that

$$\mathbb{E}\left[\mathbb{P}\left(|\|\hat{W}\|_{1} - \|W\|_{1}| < \frac{1}{100n\varepsilon}|A\right)\right] \ge 1 - \delta$$

But then we can choose  $W_1 = 0$  and  $W_2 = 1(\min\{x,y\} < \frac{1}{n\varepsilon})$  and let  $A_1, A_2$  the graphs they induce. Hence it holds

$$\mathbb{E}\left[\mathbb{P}\left(|\|\hat{W}\|_{1}| < \frac{1}{100n\varepsilon}|A_{1}\right)\right] \ge 1 - \delta$$

$$\mathbb{E}\left[\mathbb{P}\left(|\|\hat{W}\|_{1} - \frac{1}{n\varepsilon}| < \frac{1}{100n\varepsilon}|A_{2}\right)\right] \ge 1 - \delta$$

So for  $X(A) = \mathbb{P}\left(|\|\hat{W}\|_1| < \frac{1}{100n\varepsilon}|A\right)$  it holds  $\mathbb{E}[X(A_1)] \geq 1 - \delta$  and  $\mathbb{E}[X(A_2)] \leq \delta$ . But from  $\varepsilon$ -differentiable privacy we have a.s. with the obvious coupling that  $\frac{X(A_1)}{X(A_2)} \leq 2^{\varepsilon U}$  a.s. where U follows Binomial $(n, \frac{1}{n\varepsilon})$ . Hence,

$$1 - \delta \le \mathbb{E}[X(A_1)]^2 \le \mathbb{E}\left[\left[\frac{X(A_1)}{X(A_2)}\right]^2\right] \mathbb{E}\left[X(A_2)^2\right] \le \mathbb{E}\left[2^{2\varepsilon U}\right] \delta \le O(1)\delta,$$

a contradiction.

Remark: Assuming our sparsity  $\rho > \frac{1}{\varepsilon n}$  this min-max bound transfers in this case as well. Important Observation for the Borgs et al question:

We need to to control the  $k \times k$  matrices with distances  $\inf_{\pi}$  introduced in Borgs et al [14].

**Question 1.8.** Given  $\varepsilon$  find the maximum  $\eta > 0$  such that for some  $\Delta > 0$  and  $M > 2^{\varepsilon \frac{n}{k} \Delta}$  there exist  $B_1, \ldots, B_M \in [0, 1]^{k \times k}$  such that for all  $i \neq j$ ,  $\inf_{\pi} \|B_{i,\pi} - B_j\|_2 \geq \eta$  and  $B_i, B_j$  are  $\Delta$ -close.

# 2 Differential Privacy and the Lipschitz Property

Consider  $(\mathcal{G}_n, \delta_v)$  the set of all graphs with n vertices treated as a metric space with the vertex-distance and  $(\mathcal{M}, D_{\infty})$  the set of probability measures in [0, 1] with the Borel  $\sigma$ -field treated as metric space with the  $D_{\infty}$  distance. We remind that for measures  $\mu, \mu', D_{\infty}(\mu, \mu') = \|\log \mu - \log \mu'\|_{\infty}$ .

**Proposition 2.1.** A mapping  $\mu: (\mathcal{G}_n, \delta_v) \to (\mathcal{M}, D_{\infty})$  corresponds to an  $\varepsilon$ -differential private mechanism if and only if  $\mu$  is  $\varepsilon$ -Lipschitz with respect to  $\delta_v$  and  $D_{\infty}$ .

**Proposition 2.2.** Suppose that for some  $H_n \subset \mathcal{G}_n$  a function  $\hat{\mu}: (H_n, \delta_v) \to (\mathcal{M}, D_\infty)$  is  $\varepsilon$ -Lipschitz for some  $\varepsilon > 0$ . Then we can extend the function to a  $\mu: (\mathcal{G}_n, \delta_v) \to (\mathcal{M}, D_\infty)$  such that it is  $2\varepsilon$ -Lipschitz and for every  $G \in H_n$ ,  $\mu(G) = \hat{\mu}(G)$ .

*Proof.* We define for every  $G \in \mathcal{G}_n$ ,

$$d\mu(G) \propto \inf_{G' \in H_n} 2^{\varepsilon \delta_v(G, G')} d\hat{\mu}.$$

Both the properties follow.

#### 2.1 The minimax rate

Given the Proposition the rate we want to find is

$$R = \min_{\mu: (\mathcal{G}_n, \delta_v) \to (\mathcal{M}, D_\infty) \in -Lipschitz} \max_{p \in [0, 1]} \mathbb{E}_{G \sim G_{n, p}, \hat{p} \sim \mu_G}[|\hat{p} - p|]$$

## 2.2 An upper bound

First we define for every  $p \in [0,1]$  we define the distribution over [0,1] given by

$$\hat{\mu}_p(q) \propto 2^{\epsilon \min\{\frac{n^2}{\log n}|p-q|,n\}}$$

for  $q \in [0,1]$ . Now we define  $H_n \subseteq \mathcal{G}_n$  to be a subset of all the graphs such that

$$\forall p \in [0, 1], \frac{1}{n^{10}} \le \mathbb{P}_{G \sim G_{n,p}} (G \notin H_n) \le \frac{1}{n}.$$

Then for any  $G \in \mathcal{G}_n$  we set  $\mu : (\mathcal{G}_n, \delta_v) \to (\mathcal{M}, D_\infty)$  the mapping given by

$$\mu_G(q) \propto \inf_{G' \in H} \left[ 2^{\epsilon \delta_v(G,G')} \hat{\mu}_{e(G')}(q) \right],$$

for  $q \in [0, 1]$ .

Claim 2.3. The map  $\mu_G$  is  $2\varepsilon$ -D.P.

*Proof.* Follows from the definition.

For typical  $G_{n,p}, G_{n,q}$  it holds  $\delta_V(G, G') \ge \min\{\frac{n^2}{\log n} | p - q|, n\}$ .