

# Discussion about symmetric and exchangability

Tang Ze  
F1803017  
518030910431

July 2, 2020

## 1 Introduction

In the class we've introduced Hewitt-Savage 0-1 law, which is about symmetric random variables, without proof. Here I combine with some reference and write two proofs.

## 2 Two proofs of Hewitt-Savage 0-1 law

**Theorem 1. Hewitt-Savage 0-1 law** *If  $X_1, X_2, \dots$  are independent and identically distributed and  $A \in \mathcal{E}$ , then  $P(A) \in \{0, 1\}$ .*

*Proof.* Let  $\mathcal{T}_n = \sigma(X_1, X_2, \dots, X_n)$ . Let  $B_n = E(A|\mathcal{T}_n)$ . Then by the Lévy Upwards Theorem  $B_n$  converges in  $\mathcal{L}^1$  to  $A$ . Now define  $\mathcal{T}_n = \sigma(X_n, X_{n+1}, \dots)$ . Let  $C_n = E(A|\mathcal{T}_n)$ , then by the Lévy Upwards Theorem  $D = E(A|\mathcal{T}) = \lim_{n \rightarrow \infty} C_n$  in  $\mathcal{L}^1$ . By another application of the Lévy Downwards Theorem  $E(B_n|\mathcal{T}_m)$  converges to  $E(B_n|\mathcal{T})$  as  $m$  increases.

Now fix  $\epsilon > 0$ . Then for all  $n > N, m > M$  and  $n > m$  we have

$$E(|B_n - A|) < \epsilon/3$$

$$E(|C_m - D|) < \epsilon/3$$

$$E(|B_n - E(B_n|\mathcal{T}_m)|) < \epsilon/3$$

Let  $\rho$  be a permutation that exchanges  $1, \dots, n$  with  $m, \dots, n + m - 1$ . Then we get

$$B_n = E(A|\mathcal{T}_n) = E((A \circ \rho)|\mathcal{T}_n)$$

Now we are conditioning on variables that are in the slots  $m, \dots, m + n - 1$  and we take  $E(A|\rho(X_m, \dots, X_{m+n-1}))$  instead by invoking Fubini's theorem. But this  $\sigma$ -algebra is a subset of  $\mathcal{T}_m$ . Thus

$$E(B_n|\mathcal{T}_m)E(E((A \circ \rho)|\mathcal{T}_n)|\mathcal{T}_m) = E(A|\mathcal{T}_m)$$

Meanwhile we have  $E(|B_n - C_m|) < \epsilon/3$ , and by the triangle inequality  $E(|A - D|) < \epsilon/3$ .  $\square$

The Lévy Upwards Theorem is a powerful tool for reconstructing a random variable from various observations of it. But we can also prove without it and make more use of symmetry.

*Proof.* Let  $A \in \mathcal{E}$ . As the proof of Kolmogorov's 0-1 Law, we will show that  $A$  is independent of itself, i.e.  $P(A) = P(A \cap A) = P(A)P(A)$  so that

$$P(A) \in \{0, 1\} \quad (1)$$

We mark  $(A - B) \cup (B - A)$  as  $A \Delta B$ , which is the symmetric difference.

Let  $A_n \in \sigma(X_1, \dots, X_n)$  so that

$$P(A_n \Delta A) \rightarrow 0$$

$A_n$  can be written as  $\{\omega : (\omega_1, \dots, \omega_n) \in B_n\}$  with  $B_n \in \mathcal{S}^n$ . Let

$$\pi(j) = \begin{cases} j + n & \text{if } 1 \leq j \leq n \\ j - n & \text{if } n + 1 \leq j \leq 2n \\ j & \text{if } j \geq 2n + 1 \end{cases}$$

Since  $\pi^2$  is identity, we have  $\pi^{-1} = \pi$ . Therefore we don't need to worry about write  $\pi^{-1}$  or  $\pi$ . Since the coordinates are independent and identically distributed, so are the permuted coordinates. Thus

$$P(\omega : \omega \in A_n \Delta A) = P(\omega : \pi\omega \in A_n \Delta A) \quad (2)$$

Now we have  $\{\omega : \pi\omega \in A\} = \{\omega : \omega \in A\}$ , since  $A$  is permutable, and

$$\{\omega : \pi\omega \in A_n\} = \{\omega : (\omega_{n+1}, \dots, \omega_{2n}) \in B_n\}$$

We use  $A'_n$  to denote the last event then we get

$$\{\omega : \pi\omega \in A_n \Delta A\} = \{\omega : \omega \in A'_n \Delta A\} \quad (3)$$

Combine (2) and (3), we have

$$P(A_n \Delta A) = P(A'_n \Delta A) \quad (4)$$

It's obvious that

$$|P(B) - P(C)| \leq |P(B \Delta C)|$$

thus (4) implies  $P(A_n), P(A'_n) \rightarrow P(A)$ . Now  $A - C \subset (A - B) \cup (B - C)$  and with a similar inequality for  $C - A$  implies  $A \Delta C \subset (A \Delta B) \cup (B \Delta C)$ .

The last inequality, (4), and (1) imply

$$P(A_n \Delta A'_n) \leq P(A_n \Delta A) + P(A \Delta A'_n) \rightarrow 0$$

The last result implies

$$\begin{aligned} 0 &\leq P(A_n) - P(A_n \cap A'_n) \\ &\leq P(A_n \cup A'_n) - P(A_n \cap A'_n) \\ &= P(A_n \Delta A'_n) \rightarrow 0 \end{aligned}$$

so  $P(A_n \cap A'_n) \rightarrow P(A)$ . But  $A_n$  and  $A'_n$  are independent, so

$$P(A_n \cap A'_n) = P(A_n)P(A'_n) \rightarrow P(A)^2$$

This shows  $P(A) = P(A)^2$ , which finishes the proof of Hewitt-Savage 0-1 law.  $\square$

### 3 de Finettis Theorem

Another theorem about exchangeability is de Finettis Theorem. Suppose we have a sequence of exchangeable Bernoulli random variables, that is variables that are either 0 or 1. de Finettis theorem tells us that they are conditionally independent given the tail algebra.

**Theorem 2. de Finettis Theorem** *Let  $X_1, X_2, \dots$  be a sequence of exchangeable random variables in  $\mathcal{L}^1$ . Let  $\mathcal{T}$  be the tail algebra. Then  $E(X_1|\mathcal{T}), E(X_2|\mathcal{T}), \dots$  is a independent identically distributed sequence of random variables.*

*Proof.* Let  $\mathcal{T}_m = \sigma(X_m, X_{m+1}, \dots)$ . Then  $\mathcal{T} := \bigcap_m \mathcal{T}_m$ . Since the  $X_i$  are exchangeable,  $E(X_i|\mathcal{T})$  is an exchangeable sequence.

Independence of  $X_1, \dots, X_n$  is equivalent to  $X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n$  being independent events. Let  $f_n(x)$  be the indicator function of  $x \leq x_n$ . Then we need to show that

$$E(f_1(X_1) \dots f_n(X_n)|\mathcal{T}) = E(f_1(X_1)|\mathcal{T}) \dots E(f_n(X_n)|\mathcal{T})$$

Consider  $S_m^k = f_k(X_1) + f_k(X_2) + \dots + f_k(X_m)$ . By exchangeability and a symmetry argument we have  $E(f_k(X_i)|\mathcal{T}) = S_m^k/m$ .

Then we apply the Lévy Downward Theorem we have that

$$E(f_k(X_i)|\mathcal{T}) = \lim_{m \rightarrow \infty} S_m^k/m$$

Then we have

$$E(f_1(X_1)|\mathcal{T}) \dots E(f_n(X_n)|\mathcal{T}) = \lim_{m \rightarrow \infty} \prod_{k=1}^n S_m^k/m$$

Now we have

$$m^{-n} \prod_{k=1}^n S_m^k = m^{-n} \sum_{1 \leq m_1, \dots, m_n \leq n} \prod f_{m_i}(X_i)$$

Each term of the sum is either 0 or 1, and there are  $m^{n-1}$  terms that have two or more indices the same. As we take the limit the contributions of these terms vanishes since we are dividing by  $m^n$  and we are left with

$$\begin{aligned} & \lim_{m \rightarrow \infty} \prod_{k=1}^n S_m^k/m \\ &= \lim_{m \rightarrow \infty} \frac{1}{m(m-1) \dots (m-n+1)} \sum_{m_i \text{ nonequal}} \prod_{i=1}^n f_{m_i}(X_{m_i}) \\ &= \lim_{m \rightarrow \infty} E(f_1(X_1)f_2(X_2) \dots f_n(X_n)|\mathcal{T}) \end{aligned}$$

□

### 4 reference

*Probability, Theory and Examples* by Rick Durrett

*Some Applications of Martingales to Probability Theory* by Watson Ladd