

# On Martingale Posteriors, Gaussian Processes & Pseudo Data

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# References

Discussions based on many papers. To start I recommend reading:

- Fong, Edwin, Chris Holmes, and Stephen G. Walker. "**Martingale posterior distributions.**" *Journal of the Royal Statistical Society Series B: Statistical Methodology* 85.5 (2023): 1357-1391.
- Leibfried, Felix, et al. "**A tutorial on sparse Gaussian processes and variational inference.**" *arXiv preprint arXiv:2012.13962* (2020).
- Ritter, Hippolyt, et al. "**Sparse uncertainty representation in deep learning with inducing weights.**" *Advances in Neural Information Processing Systems* 34 (2021): 6515-6528.

# Two ways to construct posterior

The task:

given data  $y_{1:N}$ , construct a predictor  $p(y_{N+1:\infty} | y_{1:N})$

Solution 1: Explicit Bayesian modelling (de Finetti)

- Build a model:

Prior:  $p(\theta)$

Likelihood:  $p(y|\theta)$

- Compute posterior:

$$p(\theta | y_{1:N}) \propto \prod_{n=1}^N p(y_n | \theta) p(\theta)$$

- Bayesian predictive inference:

$$p(y_{N+1:\infty} | y_{1:N}) = \int \prod_{n=N+1}^{\infty} p(y_n | \theta) p(\theta | y_{1:N}) d\theta$$



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Solution 2: Direct construction of predictive (Doob)

- Define a 1-step predictor (with some conditions):

$$\{p(\cdot | y_{1:N})\}_{n>N} \Rightarrow p(y_{N+1:\infty} | y_{1:N}) = \underbrace{\prod_{n=N+1}^{\infty} p(y_n | y_{1:n-1})}$$

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- Consider the following definition:

$$Y_{N+1:M} \sim \{p(\cdot | y_{1:N})\}_{n>N}$$
$$p_M(y | Y_{N+1:M}) := \hat{p}(y | y_{1:N}, Y_{N+1:M}) \quad (\text{empirical dist.})$$

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- Consider the following definition:

$$Y_{N+1:M} \sim \{p(\cdot | y_{1:N})\}_{n > N} \quad \left\{ \begin{array}{l} y_{1:N} \\ Y_{N+1:M} \text{ is an empirical dist.} \end{array} \right\}$$

$$p_M(y | Y_{N+1:M}) := \hat{p}(y | y_{1:N}, Y_{N+1:M})$$

- Construct the (finite) **Martingale posterior**  $p_M(\theta_M | y_{1:N})$  and take  $M \rightarrow \infty$ :

$$\theta_M \sim p_M(\theta_M | y_{1:N}) \Leftrightarrow \underbrace{Y_{N+1:M}}_{\sim} \sim \{p(\cdot | y_{1:N})\}_{n > N},$$

$$\theta_M = \operatorname{argmin}_{\theta} KL[p_M(y | Y_{N+1:M}) \| p(y | \theta)]$$

# How do they connect to each other?

- Remember the steps for sampling from the Martingale posterior:
  - Specify a predictive model  $\{p(\cdot | y_{1:N})\}_{n>N}$
  - Sample  $Y_{N+1:M} \sim \{p(\cdot | y_{1:N})\}_{n>N}$
  - Fit a model  $p(y|\theta)$  to dataset  $\{y_{1:N}, Y_{N+1:M}\}$  with MLE
  - Consider the distribution of the MLE estimator when  $M \rightarrow \infty$

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  - Consider the distribution of the MLE estimator when  $M \rightarrow \infty$
- Conditions to make this procedure work:
  - The limit exists (a.s., wrt. dist of  $Y_{N+1:\infty}$ ):

$$p_\infty(y|Y_{N+1:\infty}) = \lim_{M \rightarrow \infty} p_M(y|Y_{N+1:M}),$$



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  - Consider the distribution of the MLE estimator when  $M \rightarrow \infty$
- Conditions to make this procedure work:
  - The limit exists (a.s., wrt. dist of  $Y_{N+1:\infty}$ ):
  - Bias-free:

$$p_\infty(y|Y_{N+1:\infty}) = \lim_{M \rightarrow \infty} p_M(y|Y_{N+1:M}),$$

$$E[p_\infty(y|Y_{N+1:\infty})] = \hat{p}(y|y_{1:N})$$

# How do they connect to each other?

- Now let's use the Bayesian predictive dist. from Solution 1:

$$p(y_{N+1:M}|y_{1:N}) = \int \prod_{n=N+1}^M p(y_n|\theta)p(\theta|y_{1:N}) d\theta$$

The following two sampling methods are equivalent:

Joint sampling:  $\theta \sim p(\theta|y_{1:N})$ ,  $Y_{N+1:M} \sim p(\cdot|\theta)$  i.i.d.

Sequential sampling:  $Y_m \sim p(\cdot|Y_{N+1,m-1}, y_{1:N})$ ,  $m = N + 1:M$

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Two solutions are the same:

- If  $Y_{N+1:M}$  is generated by  $p(y|\theta_0)$  using  $\theta_0 \sim p(\theta|y_{1:N})$ , then under identifiability conditions, the MLE solution converges:

$$\theta_M = \operatorname{argmin}_{\theta} KL[p_M(y|Y_{N+1:M})||p(y|\theta)],$$

$$\theta_M \rightarrow \theta_0 \text{ as } M \rightarrow \infty$$

$$p_M(y|Y_{N+1:M}) = \frac{1}{n} \sum_{m=N+1}^M \mathbb{1}(y = Y_m)$$

$$p(\theta|y_{1:N}) \rightarrow \theta_0 \rightarrow \{Y_{N+1:\infty}, y_{1:N}\} \rightarrow \theta_0$$

$$\sim p(\theta|y_{1:N})$$

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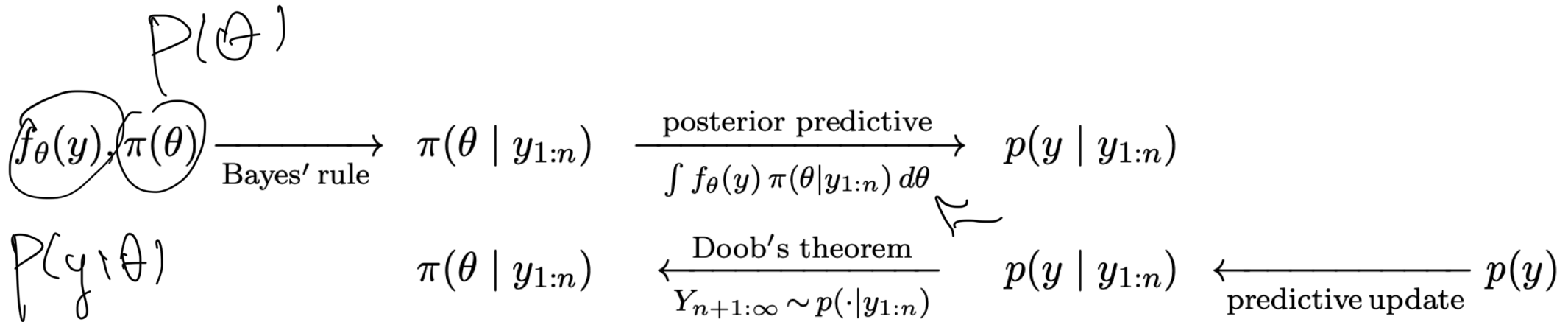
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then under identifiability conditions, the MLE solution converges:  
 $\theta_M = \operatorname{argmin}_{\theta} KL[p_M(y|Y_{N+1:M})||p(y|\theta)]$ ,  $\theta_M \rightarrow \theta_0$  as  $M \rightarrow \infty$

- Unbiasedness is satisfied since  $Y_{N+1:M}$  are conditionally identically distributed  
(Berti et al. 2004)

# Two philosophies on Bayesian inference

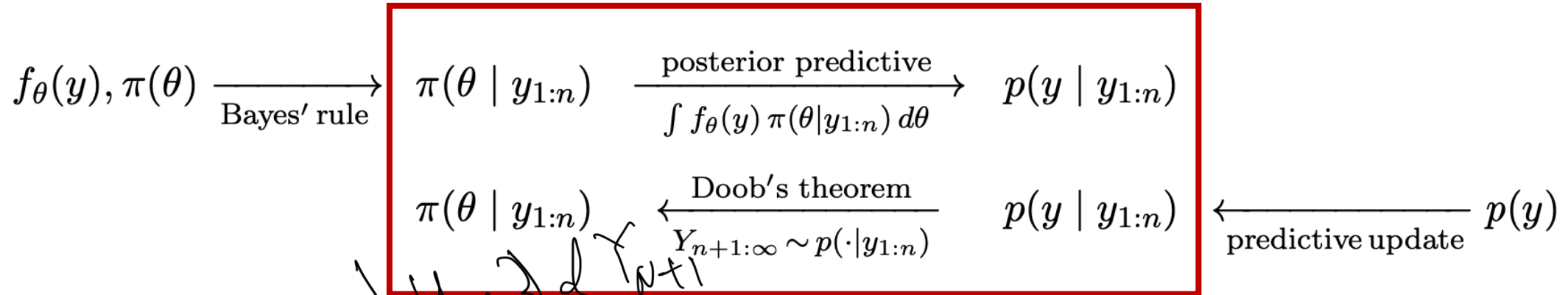
- Solution 1 (de Finetti): explicit parameter dist. modelling
  - Specify prior  $p(\theta)$
  - Specify likelihood  $p(y|\theta)$
  - Posterior  $p(\theta|y_{1:N})$  from Bayes' rule
- Solution 2 (Doobs): predictive distribution modelling
  - Specify a sequence of predictive dist.  $\{p(\cdot | y_{1:N})\}_{n>N}$
  - Specify likelihood  $p(y|\theta)$
  - Posterior  $p(\theta|y_{1:N})$  implicitly defined by the following procedure
    - Sample more data  $Y_{N+1:M}$  from the predictive dists.
    - Fit  $p(y|\theta)$  to the augmented dataset by MLE
    - Repeat the above two steps multiple times and have a set of samples for  $\theta$

# Two philosophies on Bayesian inference



# Two philosophies on Bayesian inference

1-1 correspondence if  $\{p(\cdot | y_{1:N})\}$  and  $f_\theta(y)$  (or in our notation  $p(y|\theta)$ ) are consistent!



$$\begin{aligned}
 & \int p(y_{n+2}=y, y_{n+1}|y_{1:n}) p(y_{n+1}|y_{1:n}) p(y_{n+2}|y_{1:n}) p(y_{n+1}|y_{1:n}) p(y_{n+2}|y_{1:n}) \\
 &= \int \pi(y_{n+1}|\theta) p(y_{n+2}|\theta) p(y_{n+1}|y_{1:n}) p(y_{n+2}|y_{1:n}) d\theta \\
 &= \int \pi(y_{n+1}|\theta) p(y_{n+2}|\theta) p(y_{n+1}|y_{1:n}) p(y_{n+2}|y_{1:n}) d\theta
 \end{aligned}$$

# Not so surprising to GP researchers...

- Gaussian process modelling:

$$f \sim GP(m(\cdot), k(\cdot, \cdot)), \quad y \sim p(y|f(x))$$



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- Gaussian process modelling:

$$f \sim GP(m(\cdot), k(\cdot, \cdot)),$$

$$y \sim p(y|f(x))$$

$p(y|f(x))$   
 $f(x) \sim \mathcal{N}(f(x), \sigma^2 \mathbf{I})$

- Important:  $f$  is infinite dimensional!

- A prior on  $f$  will induce a prior on  $Y_{1:\infty}$

prior:  $p(\underline{f} | \underline{X})$

$$\underline{X} = \{x_1, \dots, x_N\}$$

$$\underline{f} = \{f(x_1), \dots, f(x_N)\}$$

$$N \rightarrow \infty$$

$$= \mathcal{N}\left(\begin{bmatrix} m(x_1) \\ \vdots \\ m(x_N) \end{bmatrix}, K_{\underline{X}\underline{X}}\right)$$

$$K_{\underline{X}\underline{X}} = \begin{bmatrix} k(x_1, x_1) & k(x_1, x_2) & \dots \\ \vdots & \ddots & \ddots \\ k(x_N, x_1) & k(x_N, x_2) & \dots \end{bmatrix}$$

$$p(Y_{1:\infty} | \underline{X}) = \int \prod_n p(Y_N | f(x_n)) p(\underline{f} | \underline{X}) d\underline{f}$$

$$= \mathcal{N}\left(\begin{bmatrix} m(x_1) \\ \vdots \end{bmatrix}, K_{\underline{X}\underline{X}} + \sigma^2 \mathbf{I}\right)$$

# Not so surprising to GP researchers...

- In fact: this idea works for any explicit Bayesian models
- The structure of prior on  $Y_{1:\infty}$

Solution 1:

$$p(\theta), \quad \underline{\underline{p(y|\theta)}}$$

$$p(Y_{1:N}) = \int p(\theta) \prod_n p(Y_n | \theta) d\theta$$
$$p(Y_{N+1:m} | Y_{1:N}) = \frac{p(Y_{N+1:m}, Y_{1:N})}{\underline{\underline{p(Y_{1:N})}}}$$

# GP people is smart about augmenting data

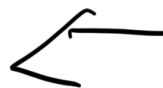
- GP posterior is expensive to compute -  $O(N^3)$  time complexity

$$p(f|y_{1:N}) = GP(m_{post}(\cdot), k_{post}(\cdot, \cdot))$$

- Sparse GP approximation: use augmented “pseudo data”

$$\rightarrow p(\underline{f}, \underline{u} | \underline{X}, \underline{Z}) \quad \text{pseudo data} \quad u = f(z)$$

$$\rightarrow p(f|y_{1:N}) \approx q(f) \triangleq \int p(\underline{f} | \underline{u}) q(\underline{u}) d\underline{u}$$



# The augmented prior idea goes beyond GPs


- Consider a probabilistic model

Prior:  $p(\theta)$

Likelihood:  $p(y|\theta)$

- Augmentations for predictive inference:

- Augment the prior: define  $\pi(\theta, U)$  such that

$$\int \pi(\theta, U) dU = p(\theta)$$


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- Now consider the posterior predictive:

$$p(y|y_{1:N}) = \int \underbrace{p(y|\theta)p(\theta|y_{1:N})}_{\text{posterior}} d\theta = \frac{\underbrace{p(y, y_{1:N})}_{\text{joint}}}{\underbrace{p(y_{1:N})}_{\text{marginal}}}$$

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- Now consider the posterior predictive:

$$p(y|y_{1:N}) = \int p(y|\theta) p(\theta|y_{1:N}) d\theta = \frac{p(y, y_{1:N})}{p(y_{1:N})}$$

$$p(y_{1:N}) = \int \prod_{n=1}^N p(y_n|\theta) \underbrace{p(\theta)}_{\pi(\theta, U)} d\theta = \int \prod_{n=1}^N p(y_n|\theta) \underbrace{\pi(\theta, U)}_{\pi(\theta|U) \hat{\pi}(\theta)} d\theta dU \quad \swarrow$$

$$\pi(\theta, U) = \pi(\theta|U) \hat{\pi}(\theta) \quad \int \prod_{n=1}^N p(y_n|\theta) \pi(\theta|U) d\theta$$

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- Now consider the posterior predictive:

Key: predictive  $p(y|y_{1:N})$  remains the same regardless of using  $p(\theta|y_{1:N})$  or  $p(U|y_{1:N})$

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- Now consider the posterior predictive:

You can then think about how to do variational inference then ;)

$$p(y|y_{1:n}) \approx \int p(y|\theta) q(\theta) d\theta \quad q(\theta) \approx p(\theta|y_{1:n})$$
$$\approx \int p(y|u) q(u) du, \quad q(u) \approx p(u|y_{1:n})$$





# Take aways

- ML people cares about predictions (only?)
- This prompts us to think about more direct approaches
  - Function-space uncertainty (de Finetti)
  - Martingale posterior (Doob)
- Auxiliary variables (e.g., pseudo data) as flexible approach to assist!
  - Gaussian processes
  - General (approximate) Bayesian inference

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  - Gaussian processes
  - General (approximate) Bayesian inference
- Bonus: other Bayesian methods that involves auxiliary variables
  - Hamiltonian Monte Carlo
  - Polya-Gamma Augmentation
  - ...