

# Identifiability of Generalized Hypergeometric Distribution (GHD) DAG Models

Gunwoong Park, Hyewon Park

University of Seoul, Statistics

2019-01-03

# Outline

- Introduction
  - ▶ Directed graphical model
- Generalized Hypergeometric Distribution (GHD) DAG Models
  - ▶ Identifiability
- Algorithm
  - ▶ Statistical Guarantees
- Numerical Experiments

# Introduction

# Introduction: Sports data

Why do we learn the graphical models?

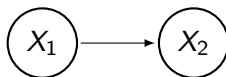


# What is the directed graph?



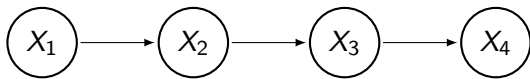
- $G = (V, E)$
- $V$  : A set of nodes, e.g.  $V = \{1, 2\}$ .
- $E$  : A set of directed edges, e.g.  $E = \{(1, 2)\}$ .

# What is the directed graphical model?



- $X := (X_j)_{j \in V}$ , a set of random variables, e.g.  $X = \{X_1, X_2\}$ .
- $X_1$  and  $X_2$  are correlated with each other.

# Notations for directed graphical model

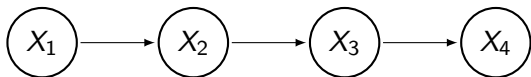


- $X_{Pa(j)}$  : A parents set of  $X_j$ , e.g.  $X_{Pa(2)} = \{X_1\}$ .
- $X_{De(j)}$  : A set of all descendants of  $X_j$ , e.g.  $X_{De(2)} = \{X_3, X_4\}$
- $X_{Nd(j)}$  : A set of non-descendants of  $X_j$ ,  
e.g.  $X_{Nd(2)} = \{X_1\}$ ,  $X_{Nd(3)} = \{X_1, X_2\}$ .

# Generalized Hypergeometric Distribution (GHD) DAG Models



# Directed Acyclic Graph (DAG)



- A directed graph with no directed cycles.
- DAG model has the factorization,

$$P(G) = P(X_1, X_2, \dots, X_p) = \prod_{j=1}^p P(X_j \mid X_{\text{Pa}(j)}).$$

- In this graph,

$$P(G) = P(X_1)P(X_2 \mid X_1)P(X_3 \mid X_2)P(X_4 \mid X_3).$$

# Probability generating function

- A power series representation of the probability mass function of the random variable.

$$G(s) = \mathbb{E}(s^x) = \sum_{x=0}^{\infty} p(x)s^x.$$

- Example: A Poisson random variable with rate parameter  $\lambda$ .

$$G(s) = \mathbb{E}(s^x) = \sum_{x=0}^{\infty} \frac{\lambda^x e^{-\lambda}}{x!} s^x = e^{\lambda(s-1)}.$$

# Generalized hypergeometric function

$${}_pF_q[a_1, \dots, a_p; b_1, \dots, b_q; \theta] := \sum_{j \geq 0} \frac{\langle a_1 \rangle^j \cdots \langle a_p \rangle^j \theta^j}{\langle b_1 \rangle^j \cdots \langle b_q \rangle^j j!}.$$

- $\langle a \rangle^j = a(a+1) \cdots (a+j-1)$ , the rising factorial,  
e.g.  $\langle 2 \rangle^4 = 2(2+1)(2+2)(2+3) = 120$ .
- $(a)_j = a(a-1) \cdots (a-j+1)$ , the falling factorial,  
e.g.  $(5)_3 = 5(5-1)(5-2) = 60$ .
- $\langle a \rangle^0 = (a)_0 = 1$ .

# Generalized Hypergeometric Distributions (GHD)

- A family of GHDs has a special form of probability generating functions expressed in terms of the generalized hypergeometric series.

$$\begin{aligned} G(s) &= {}_pF_q[a_1, \dots, a_p; b_1, \dots, b_q; \theta(s-1)] \\ &= \sum_{j \geq 0} \frac{\langle a_1 \rangle^j \cdots \langle a_p \rangle^j (\theta(s-1))^j}{\langle b_1 \rangle^j \cdots \langle b_q \rangle^j j!} \end{aligned}$$

- Example: Poisson distribution

$$G(s) = e^{\lambda(s-1)} = \sum_{j \geq 0} \frac{(\lambda(s-1))^j}{j!} = {}_0F_0[; ; \lambda(s-1)]$$

# Examples of the hypergeometric distribution

Distributions	p.g.f. $G(s)$	Parameters
Poisson	${}_0F_0[; ; \lambda(s-1)]$	$\lambda > 0$
Hyper-Poisson	${}_1F_1[1; b; \lambda(s-1)]$	$\lambda > 0$
Binomial	${}_1F_0[-N; ; -p(s-1)]$	$N, p > 0$
Negative Binomial	${}_1F_0[k; ; p(s-1)]$	$k, p > 0$
Poisson Beta	${}_1F_1[a; a+b; \lambda(s-1)]$	$a, b, \lambda > 0$
Negative Binomial Beta	${}_2F_1[k, a; a+b; \lambda(s-1)]$	$k, a, b, \lambda > 0$
STERRED Geometric	${}_2F_1[1, 1; 2; q(s-1)/(1-q)]$	$1 > q > 0$
Shifted UNSTERRED Poisson	${}_1F_1[2; 1; \lambda(s-1)]$	$1 \geq \lambda > 0$

# GHD DAG models

## Definition: GHD DAG Models

For each  $j \in V$ ,  $X_j \mid X_{\text{Pa}(j)}$  has the following probability generating function

$$G(s; a(j), b(j)) = {}_{p_j}F_{q_j}[a(j); b(j); \theta(X_{\text{Pa}(j)})(s - 1)]$$

where  $a(j) = (a_{j1}, \dots, a_{jp_j})$ ,  $b(j) = (b_{j1}, \dots, b_{jq_j})$ , and  $\theta : \mathcal{X}_{\text{Pa}(j)} \rightarrow \mathbb{R}$ .

- The conditional distribution of each node given its parents belongs to a family of GHDs.
- The parameter depends only on its parents.

# Proposition of GHD DAG models

## Proposition: Constant Moments Ratio (CMR) Property

Consider a GHD DAG model. Then for any  $j \in V$  and any integer  $r = 2, 3, \dots$ , there exist a  $r$ -th factorial constant moments ratio (CMR) function

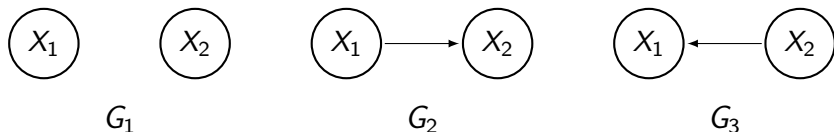
$$f_j^{(r)}(x; a(j), b(j)) = x^r \prod_{i=1}^{p_j} \left( \frac{(a_{ji}+r-1)_r}{a_{ji}^r} \right) \prod_{k=1}^{q_j} \left( \frac{b_{jk}^r}{(b_{jk}+r-1)_r} \right) \text{ such that}$$

$$\begin{aligned} \mathbb{E}((X_j)_r \mid X_{\text{Pa}(j)}) &= \mathbb{E}(X_j(X_j - 1) \cdots (X_j - r + 1) \mid X_{\text{Pa}(j)}) \\ &= f_j^{(r)}(\mathbb{E}(X_j \mid X_{\text{Pa}(j)}); a(j), b(j)), \end{aligned}$$

as long as  $\max X_j \geq r$ .

# GHD DAG models: Identifiability

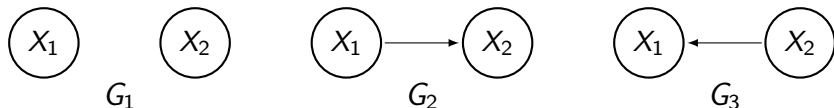
Is it possible to recover a graph from distribution?



- We can distinguish  $G_2$  and  $G_3$  from  $G_1$ .
- How can we know the direction of the edge?
- We exploit the CMR property for model identifiability.

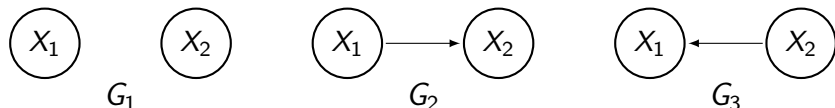


## Example : The CMR property for Poisson



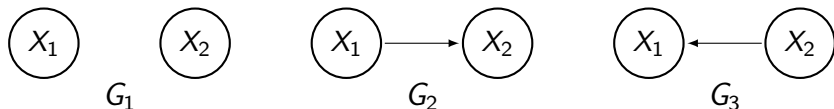
- $G_1 : X_1 \sim \text{Poisson}(\lambda_1), X_2 \sim \text{Poisson}(\lambda_2), X_1 \perp\!\!\!\perp X_2$ .
- $G_2 : X_1 \sim \text{Poisson}(\lambda_1), X_2 \mid X_1 \sim \text{Poisson}(\theta_2(X_1))$ .
- $G_3 : X_2 \sim \text{Poisson}(\lambda_2), X_1 \mid X_2 \sim \text{Poisson}(\theta_1(X_2))$ .
- $\theta_1, \theta_2 : \mathbb{N} \cup \{0\} \rightarrow \mathbb{R}^+$ , arbitrary positive functions.

## Example : The CMR property for Poisson



- We exploit the CMR property for Poisson,  
 $\mathbb{E}((X_j)_r) = \mathbb{E}(X_j)^r$  for any positive integer  $r \in \{2, 3, \dots\}$ .
- For  $G_1$ ,  $\mathbb{E}((X_1)_r) = \mathbb{E}(X_1)^r$  and  $\mathbb{E}((X_2)_r) = \mathbb{E}(X_2)^r$ .

## Example : The CMR property for Poisson



- For  $G_2$ ,  $\mathbb{E}((X_1)_r) = \mathbb{E}(X_1)^r$ , while

$$\begin{aligned}\mathbb{E}((X_2)_r) &= \mathbb{E}(\mathbb{E}((X_2)_r \mid X_1)) = \mathbb{E}(\mathbb{E}(X_2 \mid X_1)^r) \\ &> \mathbb{E}(\mathbb{E}(X_2 \mid X_1))^r = \mathbb{E}(X_2)^r, \quad (\text{by Jensen's inequality})\end{aligned}$$

as long as  $\mathbb{E}(X_2 \mid X_1)$  is not a constant.

- For  $G_3$ ,  $\mathbb{E}((X_2)_r) = \mathbb{E}(X_2)^r$  and  $\mathbb{E}((X_1)_r) > \mathbb{E}(X_1)^r$ .
- **Now we can distinguish  $G_1$ ,  $G_2$  and  $G_3$  by testing whether a moments ratio  $\mathbb{E}((X_j)_r)/\mathbb{E}(X_j)^r$  is greater than or equal to 1.**

# GHD DAG models: Identifiability

## Identifiability

### Assumptions: Identifiability Conditions

- For a given GHD DAG model, the conditional distribution of each node given its parents is known.
- For any node  $j \in V$ ,  $\mathbb{E}(X_j \mid X_{\text{Pa}(j)})$  is non-degenerated.

### Theorem: Identifiability

- Under the identifiability conditions, the class of GHD DAG models is identifiable.

# GHD DAG models: Identifiability

- For any node  $j \in V$ ,  $\mathbb{E}((X_j)_r) = \mathbb{E}(f_j^{(r)}(\mathbb{E}(X_j \mid X_{\text{Pa}(j)}); a(j), b(j)))$
- For any non-empty  $\text{Pa}_0(j) \subset \text{Pa}(j)$  and  $S_j \subset \text{Nd}(j) \setminus \text{Pa}_0(j)$ ,

$$\begin{aligned}\mathbb{E}((X_j)_r) &= \mathbb{E}(\mathbb{E}(f_j^{(r)}(\mathbb{E}(X_j \mid X_{\text{Pa}(j)}); a(j), b(j)) \mid X_{S_j})) \\ &> \mathbb{E}(f_j^{(r)}(\mathbb{E}(X_j \mid X_{S_j}); a(j), b(j))),\end{aligned}$$

because the CMR function is strictly convex.

- To search a smallest conditioning set  $S_j$  for each node  $j$  such that the moments ratio  $\frac{\mathbb{E}((X_j)_r)}{\mathbb{E}(f_j^{(r)}(\mathbb{E}(X_j \mid X_{S_j})))} = 1$ .

# Algorithm

# Algorithm

The Moments Ratio Scoring (MRS) algorithm consist of two steps.

- 1 Estimate the skeleton of graph.
  - ▶ Any sceleton learning algorithms.  
Examples) GES, MMHC.
- 2 Find an ordering.
  - ▶ Using moments ratio scores.
  - ▶ Find an element which has the smallest momet ratio score.

# Moments ratio score

- Moments ratio:

$$\frac{\mathbb{E}((X_j)_r)}{\mathbb{E}(f_j^{(r)}(\mathbb{E}(X_j | X_{S_j})))}.$$

- If all samples are less than  $r$ , the moments ratio is 0.

- Moments ratio score :

$$\hat{\mathcal{S}}_r(j) := \frac{\hat{\mathbb{E}}(X_j^r)}{f_j^{(r)}(\hat{\mathbb{E}}(X_j)) - \sum_{k=0}^{r-1} s(r, k) \hat{\mathbb{E}}(X_j^k)},$$

where  $s(r, k)$  is Stirling numbers of the first kind.



# Algorithm: Statistical guarantees

## Assumptions

For all  $j \in V$ , any non-empty  $\text{Pa}_0(j) \subset \text{Pa}(j)$ , and  $S_j \subset \text{Nd}(j) \setminus \text{Pa}_0(j)$ ,

(A1) There exists a positive constant  $M_{\min} > 0$  such that

$$\frac{\widehat{\mathbb{E}}(X_j^r \mid X_{S_j})}{f_j^{(r)}(\widehat{\mathbb{E}}(X_j \mid X_{S_j})) - \sum_{k=0}^{r-1} s(r, k) \widehat{\mathbb{E}}(X_j^k \mid X_{S_j})} > 1 + M_{\min}.$$

(A2) There exist a positive constant  $V_1$  such that

$$\mathbb{E}(\exp(X_j) \mid X_{\text{Pa}(j)}) < V_1.$$

# Algorithm: Statistical guarantees

## Theorem: Recovery of the ordering

In regularity conditions, there exist constant  $C_\epsilon > 0$  for any  $\epsilon > 0$  such that if sample size is sufficiently large

$n > C_\epsilon \log^{2r+d}(\max(n, p))(\log(p) + \log(r))$ , the MRS algorithm with the  $r$ -th moments ratio scores recovers the ordering with high probability:

$$P(\hat{\pi} \in \mathcal{E}(\pi)) \geq 1 - \epsilon.$$

- $d$  : The maximum indegree of the graph.
- $\mathcal{E}(\pi)$  : The set of all the orderings that are consistent with the DAG  $G$ .
- The MRS algorithm accurately estimates a true ordering with high probability, if  $n = \Omega(\log^{2r+d}(\max(n, p)) \log(p))$ .

# Numerical Experiments

# Simulation settings

- Two sets of simulation study using 150 realizations of  $p$ -node random GHD DAG models.
  - ▶ Poisson DAG models: The conditional distribution of each node given its parents is Poisson.
  - ▶ Hybrid DAG models: The conditional distributions are sequentially Poisson, Binomial with  $N = 3$ , hyper-Poisson with  $b = 2$ , and Binomial with  $N = 3$ .
- $d = 2$ .

# Link functions and parameters

- (Hyper) Poisson parameter :

$$\theta_j(\text{Pa}(j)) = \exp(\theta_j + \sum_{k \in \text{Pa}(j)} \theta_{jk} X_k),$$

$$\theta_j \in [1, 3], \theta_{jk} \in [-1.75, -0.25] \cup [0.25, 1.75].$$

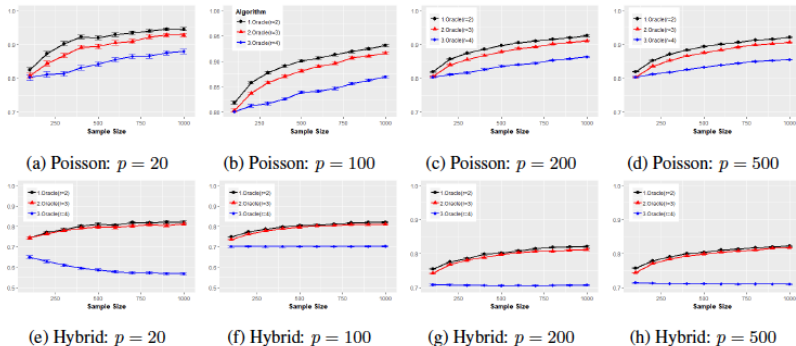
- Binomial probability :

$$p_j(\text{Pa}(j)) = \text{logit}^{-1}(\theta_j + \sum_{k \in \text{Pa}(j)} \theta_{jk} X_k),$$

$$\theta_j \in [1, 3], \theta_{jk} \in [-1.2, -0.2].$$

# Numerical experiments: Known skeleton

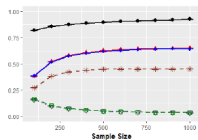
Using different values of  $r \in \{2, 3, 4\}$ .



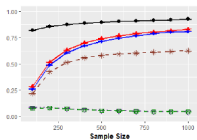
- The MRS algorithm with  $r = 2$  performs better than the MRS algorithms with  $r = \{3, 4\}$ .

# Numerical experiments: Unknown skeleton

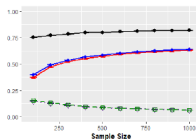
Using GES, MMHC algorithms in Step 1,  $r = 2$ ,  $p = 200$ .



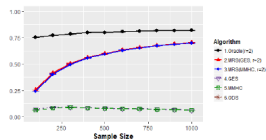
(a) Poisson: Precision



(b) Poisson: Recall



(c) Hybrid: Precision

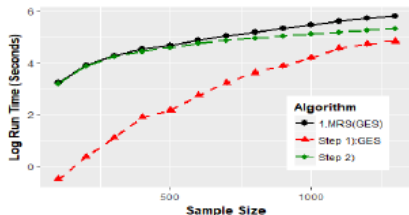


(d) Hybrid: Recall

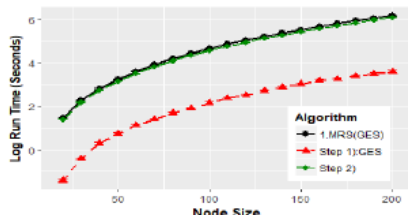
- The MRS algorithm accurately recovers the true directed edges as sample size increases.

# Computational complexity

Using GES algorithm in Step 1.



(a) Varying  $n$



(b) Varying  $p$

- (a):  $n \in \{100, 200, \dots, 1300\}$ ,  $p = 100$ .
- (b):  $p \in \{10, 20, \dots, 200\}$ ,  $n = 500$ .
- Time complexity of step 1 is  $O(n^2 p^2)$ .
- Time complexity of step 2 of MRS algorithm is  $O(np^2)$ .



# Summary

- GHD DAG models.
  - ▶ The conditional distribution of each node given its parents belongs to a family of GHDs.
  - ▶ The parameter depends only on its parents.
- We can find the ordering of the DAG using moments ratio score.
- The MRS algorithm
  - ▶ recovers the ordering more accurately as sample size increases.
  - ▶ can recover the ordering in high dimensional settings.
  - ▶ The MRS algorithm with  $r = 2$  performs better than the MRS algorithms with  $r = \{3, 4\}$ .

Thank you!