# Chapter 5: Bayesian Inference for Poisson Distribution

## 강의 목표

- ▶ 푸아송분포를 중심으로 베이지안 추론의 이해.
- ▶ Parameter Estimation (모수 추정).
  - ▶ Point Estimation (점추정).
  - ▶ Confidence Interval (구간추정).
- ▶ Prediction (예측).

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$$f(X_1, X_2, ..., X_n \mid \theta) = \prod_{i=1}^n f(x \mid \theta) \propto \theta^{\sum x_i} e^{-n\theta}.$$

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▶ Gamma distribution with  $\sum x_i + 1$  and n.

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▶ It is Gamma Dist. with  $\alpha = \sum x_i + a + 1$  and  $\beta = n$ .

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Posterior Dist:

$$P(\theta\mid x_1,...,x_n) \propto \theta^{\sum x_i} \mathrm{e}^{-(n+b)\theta}, \quad \theta > 0.$$

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Posterior Dist:

$$P(\theta \mid x_1, ..., x_n) \propto \theta^{\sum x_i} e^{-(n+b)\theta}, \quad \theta > 0.$$

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- ▶ It is Gamma Dist. with  $\alpha = \sum x_i + a$  and  $\beta = n + b$ .
- ▶ If the prior distribution is gamma, the posterior is gamma.

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▶ The data get weighted more heavily as  $n \to \infty$ .

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- ▶ We can use the Bayesian approach to update our information about the parameter(s) of interest **sequentially** as new data become available.
- Suppose we formulate a prior for our parameter  $\theta$  and observe a random sample  $x_1$ .
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$$\pi(\theta \mid x_1) \propto p(\theta) L(\theta \mid x_1).$$

 $\triangleright$  Suppose that we observe a new (independent) sample  $x_2$ .



$$\pi(\theta \mid x_1, x_2) \propto p(\theta) L(\theta \mid x_1, x_2)$$

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We can use our previous posterior as the new prior and derive a new posterior:

$$\pi(\theta \mid x_1, x_2) \propto \rho(\theta) L(\theta \mid x_1, x_2)$$

$$= \rho(\theta) L(\theta \mid x_1) L(\theta \mid x_2)$$

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Note this is the same posterior we would have obtained had  $x_1$  and  $x_2$  arrived at the same time.

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- Note this is the same posterior we would have obtained had x₁ and x₂ arrived at the same time.
- ► This "sequential updating" process can continue indefinitely in the Bayesian setup.



두 도시 에서 차량 통행량 등 주위의 교통화경이 비슷한 교차로를 하나씩 선택하여 매주 발생한 교통사고 건수를 1년 동안 조사하였다. 첫 번째 도시 에서는 직진 후 좌회전 신호를 사용하고 두 번째 도시 에서는 좌회전 후 직진 신호를 사용한다. 교통사고 건수는 독립적으로 포아송 분포를따 른다고 가정한다. 교통통제 등의 이유로 조사를 할 수 없었던 기간을 제외하고 다음 표와 같은 조사 결과를 얻었다.

교통사고 건수	0	1	2	3	4	5	6	7
City 1의 사고 건수	7	14	13	8	4	2	2	0
City 2의 사고 건수	4	13	15	6	2	2	3	1

$$n_1 = 50$$
,  $\sum x_{1i} = 102$ ,  $\bar{x}_1 = 2.04$   
 $n_2 = 46$ ,  $\sum x_{2i} = 104$ ,  $\bar{x}_2 = 2.26$ 

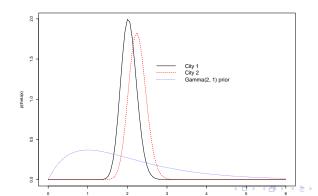
- ▶ 두 도시의 실제 평균 교통사고 건수  $\theta_1$ 과  $\theta_2$ 에 대하여 Gamma(2,1)의 사전분포를 가정하자.
- ▶ 그렇다면 다음과 같은 posterior distribution을 찾을 수 있다.

$$\pi(\theta \mid n_1 = 50, \sum x_{1i} = 102) \sim Gamma(2 + 102, 1 + 50),$$
  
 $\pi(\theta \mid n_2 = 46, \sum x_{2i} = 104) \sim Gamma(2 + 104, 1 + 46).$ 

```
x1 = rep(c(0, 1, 2, 3, 4, 5, 6), c(7, 14, 13, 8, 4, 2, 2))
x2 = rep(c(0, 1, 2, 3, 4, 5, 6, 7), c(4, 13, 15, 6, 2, 2, 3, 1))
a = 2: b = 1
n1 = length(x1); s1 = sum(x1)
n2 = length(x2); s2 = sum(x2)
postmean.theta1 = (a+s1)/(b+n1)
postmean.theta2 = (a+s2)/(b+n2)
### plot the posterior
par(mfrow=c(1, 1))
theta \leftarrow seq(0, 6, length=100)
plot(theta, dgamma(theta, a+s1, b+n1), type="l", xlab="theta", ylab="p(theta|x)
lines(theta, dgamma(theta, a+s2, b+n2), lty=2, col = "red")
lines(theta, dgamma(theta, a, b), lty=3, col = "blue")
legend ( 2.5, 1.5, legend=c (paste ("City 1"), paste("City 2"),
paste("Gamma(2, 1) prior")), cex = 1.3, lty=c(1, 2, 3), col=c(1, 2, 4),
  btv="n")
                                                  4□▶ 4個▶ 4 분 ▶ 4 분 ▶ 1 분 9 9 0 0
```

## **Traffic Example**

- ▶ City 1의 사고 발생 건수가 City 2에 비해 작다.
- ▶ 사후 분포의 분산이 사전 분포의 분산보다 작다.
- ▶ Likelihood의 영향으로 사후 분포들이 구간 (1.5, 3)이외에는 매우 비슷한다.



$$f(x_{n+1} \mid x_1,...,x_n) = \int f(x_{n+1} \mid \theta, x_1, x_2,...,x_n) \pi(\theta \mid x_1, x_2,...,x_n)$$

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=  $\int f(x_{n+1} \mid \theta) \pi(\theta \mid x_1, x_2, ..., x_n) d\theta$ 

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$$= \int \frac{\theta^{x_{n+1}} e^{-\theta}}{x_{n+1}!} \times \frac{(b+n)^{a+\sum x_i}}{\Gamma(a+\sum x_i)} \theta^{a+\sum x_i-1} e^{-(b+n)\theta} d\theta$$

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$$= \frac{(b+n)^{a+\sum x_i}}{x_{n+1}! \Gamma(a+\sum x_i)} \int \theta^{a+\sum x_i+x_{n+1}-1} e^{-(b+n+1)\theta} d\theta$$

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$$= \frac{(b+n)^{a+\sum x_i}}{x_{n+1}! \Gamma(a+\sum x_i)} \times \frac{\Gamma(a+\sum x_i+x_{n+1})}{(b+n+1)^{a+\sum x_i+x_{n+1}}}$$

$$f(x_{n+1} \mid x_1, ..., x_n) = \int f(x_{n+1} \mid \theta, x_1, x_2, ..., x_n) \pi(\theta \mid x_1, x_2, ..., x_n)$$

$$= \int f(x_{n+1} \mid \theta) \pi(\theta \mid x_1, x_2, ..., x_n) d\theta$$

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$$= \frac{(b+n)^{a+\sum x_i}}{x_{n+1}! \Gamma(a+\sum x_i)} \int \theta^{a+\sum x_i + x_{n+1}-1} e^{-(b+n+1)\theta} d\theta$$

$$= \frac{(b+n)^{a+\sum x_i}}{x_{n+1}! \Gamma(a+\sum x_i)} \times \frac{\Gamma(a+\sum x_i + x_{n+1})}{(b+n+1)^{a+\sum x_i + x_{n+1}}}$$

$$\propto \frac{1}{x_{n+1}!} \times \frac{\Gamma(a+\sum x_i + x_{n+1})}{(b+n+1)^{a+\sum x_i + x_{n+1}}}$$

$$f(x_{n+1}\mid x_1,...,x_n)$$

$$f(x_{n+1} \mid x_1, ..., x_n) = \begin{pmatrix} a + \sum x_i + x_{n+1} - 1 \\ a + \sum x_i - 1 \end{pmatrix} \left( \frac{b+n}{b+n+1} \right)^{a+\sum x_i} \left( \frac{1}{b+n+1} \right)^{x_{n+1}} = \begin{pmatrix} a + \sum x_i + x_{n+1} - 1 \\ x_{n+1} \end{pmatrix} \left( \frac{1}{b+n+1} \right)^{x_{n+1}} \left( \frac{b+n}{b+n+1} \right)^{a+\sum x_i}.$$

$$Pr(X = x) = {x+r-1 \choose x} p^x (1-p)^r \text{ for } x = 0, 1, 2, ...$$

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Hence, the prediction distribution is

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Hence, the prediction distribution is

$$NB\left(a+\sum_{i=1}^{n}x_{i},\frac{1}{b+n+1}\right).$$

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▶ 예측분산은 사후분산보다 (b + n + 1) 곱한 만큼 크다.



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$$= \mathbb{E}(\theta \mid x_1, ..., x_n).$$

► Hence, the expected prediction is the same as the posterior expectation.

▶ Variance: using the fact  $Var(X_{n+1} | \theta) = \theta$ .

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▶ Variance: using the fact  $Var(X_{n+1} \mid \theta) = \theta$ .

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$$= Var(\theta \mid x_1, ..., x_n) + \mathbb{E}(\theta \mid x_1, ..., x_n)$$

▶ Variance: using the fact  $Var(X_{n+1} \mid \theta) = \theta$ .

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$$+ \mathbb{E}(Var(X_{n+1} \mid \theta, x_1, ..., x_n) \mid x_1, ..., x_n)$$

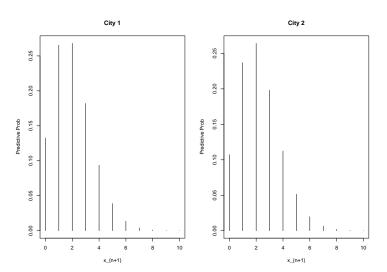
$$= Var(\theta \mid x_1, ..., x_n) + \mathbb{E}(\theta \mid x_1, ..., x_n)$$

► Hence, the variance of prediction distribution is larger than the the variance of the posterior expectation.

# **Predictive Probability**

```
> ## Ch 6
> #predictive distribution of X_{n+1}
> x1=c(rep(0,7),rep(1,14),rep(2,13),rep(3,8),rep(4,4),rep(5,2),
       rep(6,2))
+
> x2=c(rep(0,4),rep(1,13),rep(2,15),rep(3,6),rep(4,2),rep(5,2),
       rep(6,3), rep(7,1))
> a = 2:b = 1
> n1 = length(x1); s1 = sum(x1)
> n2 = length(x2); s2 = sum(x2)
> x = seq(0,10)
> par(mfrow=c(1, 2))
> plot(x,dnbinom(x,size=a+s1,prob=(b+n1)/(b+n1+1)), xlab="x_{n+1}",
       ylab="Predictive Prob", type="h", main="City 1")
> plot(x,dnbinom(x,size=a+s2,prob=(b+n2)/(b+n2+1)), xlab="x_{n+1}",
       ylab="Predictive Prob" ,type="h" , main="City 2" )
+
```

# **Predictive Probability**



▶ 두 도시의 실제 평균 교통사고 건수  $\theta_1$ 과  $\theta_2$ 에 대하여 다음과 같은 posterior distribution을 찾았다.

$$\pi(\theta \mid n_1 = 50, \sum x_{1i} = 102) \sim \textit{Gamma}(2 + 102, 1 + 50),$$
  
 $\pi(\theta \mid n_2 = 46, \sum x_{2i} = 104) \sim \textit{Gamma}(2 + 104, 1 + 46).$ 

▶ 두 도시의 실제 평균 교통사고 건수  $\theta_1$ 과  $\theta_2$ 에 대하여 다음과 같은 posterior distribution을 찾았다.

$$\pi(\theta \mid n_1 = 50, \sum x_{1i} = 102) \sim \textit{Gamma}(2 + 102, 1 + 50),$$
  
 $\pi(\theta \mid n_2 = 46, \sum x_{2i} = 104) \sim \textit{Gamma}(2 + 104, 1 + 46).$ 

▶ Gamma 분포는 이미 알려져 있지만, 관련 statistics는 여전히 찾기 어렵다.

• 이 예제에서 주요 목적은  $\theta_1$ 과  $\theta_2$  차이가 얼마인지 찿는데에 있다. (i.e.,  $\theta_1 - \theta_2$ ).

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  - ▶ Posterior expectation for  $\theta_1 \theta_2$  given data.
  - ▶ Posterior variance for  $\theta_1 \theta_2$  given data.

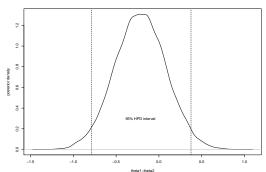
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  - ▶ Posterior expectation for  $\theta_1 \theta_2$  given data.
  - Posterior variance for  $\theta_1 \theta_2$  given data.
- ▶ Monte Carlo Method를 통해 두 parameter들의 차이에 관련된 통계량을 찾을 수 있다

▶ For notational convenience, let  $\eta = \theta_1 - \theta_2$ .

```
a =2; b = 1
n1 = 50; s1 = 102; n2 = 46; s2 = 104;
nsim = 30000
theta1.sim = rgamma(nsim,a+s1,b+n1)
theta2.sim = rgamma(nsim,a+s2,b+n2)
eta=theta1.sim- theta2.sim
mean(eta)
[1] -0.2155787
var( eta)
[1] 0.08880491
```

- ightharpoonup 이론  $\mathbb{E}(\eta \mid \mathsf{data}) = -0.2161$ .
- ▶ 이론 (η | data) = 0.0875.

```
HPD=HPDsample(eta)
par(mfrow=c(1,1))
plot(density(eta), type="l", xlab= "theta1-theta2",
ylab="posterior density", main="")
abline( v= HPD, lty=2)
text(mean(eta),0.3, "95% HPD interval" )
```



## Monte Carlo Method for Prediction

▶ 예측 분포 X<sub>n+1</sub> | X<sub>1</sub>,...,X<sub>n</sub>은 Posterior 분포보다 더욱 복잡한 형태를 가진 경우가 많다.

### Monte Carlo Method for Prediction

- ▶ 예측 분포  $X_{n+1} \mid X_1, ..., X_n$ 은 Posterior 분포보다 더욱 복잡한 형태를 가진 경우가 많다.
- ▶ Monte Carlo Method는  $\mathbb{E}(X_{n+1} \mid X_1,...,X_n)$ 과  $Var(X_{n+1} \mid X_1,...,X_n)$  를 추정하는데 큰 도움을 준다.

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- Recall that

$$f(X_{n+1} \mid X_1, ..., X_n) = \mathbb{E}(f(X_{n+1} \mid \theta) \mid X_1, ..., X_n).$$

### First Method

1. N개의  $\theta_i$ 를 랜덤하게 생성.

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Estimation

i. 
$$\widehat{\mathbb{E}}(X_{n+1} \mid X_1, ..., X_n) = \frac{1}{N} \sum_{i=1}^{N} x_{n+1,i}$$

ii. 
$$\widehat{\text{Var}}(X_{n+1} \mid X_1, ..., X_n) = \frac{1}{N} \sum_{i=1}^{N} x_{n+1,i}^2 - \left(\frac{1}{N} \sum_{i=1}^{N} x_{n+1,i}\right)^2$$

### Second Method: Rao-Blackwellization

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Estimation

i. 
$$\widehat{\mathbb{E}}(X_{n+1} \mid X_1, ..., X_n) = \sum_{\text{all } x_{n+1}} x_{n+1} \hat{f}(x_{n+1} \mid x_1, ..., x_n)$$

ii. 
$$\operatorname{Var}(X_{n+1} \mid X_1, ..., X_n) = \sum_{\text{all } x_{n+1}} x_{n+1}^2 \hat{f}(x_{n+1} \mid x_1, ..., x_n) - \left(\sum_{\text{all } x_{n+1}} x_{n+1} \hat{f}(x_{n+1} \mid x_1, ..., x_n)\right)^2.$$

# **Comparisons**

- ▶ Rao-Blackwellization는 probability mass function 을 추측. 또 그 확률을 통해 mean과 variance를 예측.
- ▶ 첫번째 방법이 더 복잡하나, 더 정확한 경향이 있다.