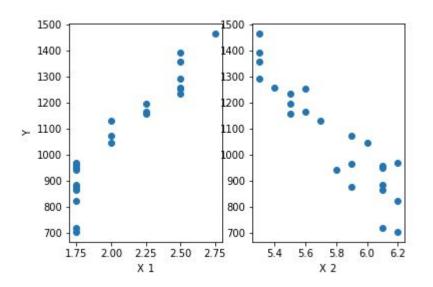
Multivariate Regression

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With more predictors



2.75	5.3	1464
2.50	5.3	1394
2.50	5.3	1357
2.50	5.3	1293
2.50	5.4	1256
2.50	5.6	1254
2.50	5.5	1234
2.25	5.5	1195
2.25	5.5	1159
2.25	5.6	1167
2.00	5.7	1130
2.00	5.9	1075
2.00	6.0	1047
1.75	5.9	965
1.75	5.8	943
1.75	6.1	958
1.75	6.2	971
1.75	6.1	949
1.75	6.1	884
1.75	6.1	866
1.75	5.9	876
1.75	6.2	822
1.75	6.2	704
1.75	6.1	719

Generally speaking...

$$Y = eta_0 + x^{(1)}eta_1 + x^{(2)}eta_2 + \dots + x^{(p-1)}eta_{p-1} + \epsilon$$

$$x = \begin{bmatrix} 1 & x_1^{(1)} & \cdots & x_1^{(p-1)} \\ 1 & x_2^{(1)} & \cdots & x_2^{(p-1)} \\ \vdots & \vdots & \vdots & \vdots \\ 1 & x_n^{(1)} & \cdots & x_n^{(p-1)} \end{bmatrix} = \begin{bmatrix} 1 & x_1^\top \\ 1 & x_2^\top \\ \vdots & \vdots \\ 1 & x_n^\top \end{bmatrix} \in \mathbb{R}^{n \times p}$$

Recall our road map...

x - predictor (random) variables Y - response random variable

- Build your model:
 - 1) relationship: $Y = x\beta + \epsilon$, $\mathbb{E}[\epsilon] = 0$, $Var[\epsilon] = \sigma^2 I_n$.
 - 2) preference: choose $\widehat{oldsymbol{eta}}$ to minimize mean squared error
- ullet Estimate your model parameters: $(x_1,y_1),(x_2,y_2),\ldots,(x_n,y_n)$
 - 1) using observed data to express your preference: $Q = \sum_{i=1}^{n} (y_i x_i^{\top} \beta)^2$
 - 2) get parameters estimation for your model:
- Understand your model:
 - 1) properties of estimations:
 - 2) predictions: $\widehat{Y}_0 = x_0^{\top} \widehat{\beta}, \quad \widehat{y}_0 = x_0^{\top} b.$

Taking partial derivatives wrt a vector

Linear forms. If $f(\mathbf{X}) = \mathbf{X}^T \mathbf{a}$, with \mathbf{a} not a function of \mathbf{X} , then

$$\nabla(\mathbf{X}^T\mathbf{a}) = \mathbf{a}$$

Quadratic forms. Let C be a $p \times p$ matrix which is not a function of X, and consider the quadratic form $\mathbf{X}^T \mathbf{C} \mathbf{X}$. (You can check that this is scalar.) The gradient is

$$\nabla(\mathbf{X}^T\mathbf{C}\mathbf{X}) = (\mathbf{C} + \mathbf{C}^T)\mathbf{X}.$$

Back to optimize $Q(\beta) = \|y - x\beta\|^2$

$$\widehat{eta} = (x^ op x)^{-1} x^ op Y$$

$$\|\widehat{\sigma}_{LS}^2 = rac{1}{n-p}\|Y-\widehat{Y}\|^2 = rac{1}{n-p}\|Y-x\widehat{eta}\|^2$$

Collinear

In order to invert $x^{ op}x$ we need on of the following:

- Its determinant is non-zero.
- It is of "full column rank", meaning all of its columns are linearly independent.
- It is of "full row rank", meaning all of its rows are linearly independent.

Collinear in terms of data if:

- If n < p.
- If one of the predictor variables is constant.
- If two of the predictor variables are proportional to each other.
- If two of the predictor variables are otherwise linearly related.

Multiple Regression Isn't Just a Bunch of Simple Regressions

Suppose the real model is $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \epsilon$. (Nothing turns on p = 2, it just keeps things short.) What would happen if we did a simple regression of Y on just X_1 ? We know that the optimal (population) slope on X_1 is

$$\frac{\operatorname{Cov}\left[X_{1},Y\right]}{\operatorname{Var}\left[X_{1}\right]}$$

Let's substitute in the model equation for Y:

$$\frac{\text{Cov}[X_{1}, Y]}{\text{Var}[X_{1}]} = \frac{\text{Cov}[X_{1}, \beta_{0} + \beta_{1}X_{1} + \beta_{2}X_{2} + \epsilon]}{\text{Var}[X_{1}]}$$

$$= \frac{\beta_{1}\text{Var}[X_{1}] + \beta_{2}\text{Cov}[X_{1}, X_{2}] + \text{Cov}[X_{1}, \epsilon]}{\text{Var}[X_{1}]}$$

$$= \beta_{1} + \frac{\beta_{2}\text{Cov}[X_{1}, X_{2}] + 0}{\text{Var}[X_{1}]}$$

$$= \beta_{1} + \beta_{2}\frac{\text{Cov}[X_{1}, X_{2}]}{\text{Var}[X_{1}]}$$

Multiple Regression Isn't Just a Bunch of Simple Regressions

Dep. Varia Model: Method: Date: Time: No. Obserr Df Residua Df Model:	F vations: als:	Least Squar ri, 02 Oct 20 15:43:	res F-star 120 Prob 15 Log-L 24 AIC: 22 BIC:	ared: R-squared: tistic: (F-statist: ikelihood:		0.851 0.844 125.4 1.49e-10 -139.14 282.3 284.6	Dep. Var Model: Method:		OL	s F-sta	R-squared: tistic:		0.898 0.888 92.07
Covariance	e Type:	nonrobu					Date:	F	ri, 02 Oct 202		(F-statisti		4.04e-11
	coef	std err	t	P> t	[0.025	0.9751	Time:	rvations:	20:31:0	-	ikelihood:		-134.61 275.2
							No. Obse	rvations:	2				278.8
const	4471.3393	304.254	14.696	0.000	3840.354	5102.324	Df Model	uais:	2	DIC:			2/0.0
x1	-588.9621	52.602	-11.196	0.000	-698.053	-479.871	Covarian	ce Type:	nonrobus	<u> </u>			
Dep. Varia	able:		y R-squa	red:					HOHLODUS				
Model:		0		-squared:		0.870		coef	std err	t	P> t	[0.025	0.975]
Method:		Least Squar				155.0				,		[0.025	
Date:	Fi	ri, 02 Oct 20		F-statisti		1.95e-11	const	1798.4040	899.248	2.000	0.059	-71.685	3668.493
Time:		20:28:		kelihood:		-136.94	v 1	-250.1466	117.950	-2.121	0.046	-495.437	-4.856
No. Observ			24 AIC:			277.9			111.367	3.103	0.046	113.940	577.140
Df Residua Df Model:	ils:		22 BIC:			280.2	x_2	345.5401	111.30/	3.103	0.005	113.940	5//.140
Covariance	Type:	nonrobu	1 s+										
========	: 19pe.	.========	========			.=======							
	coef	std err	t	P> t	[0.025	0.975]							
const	-99.4643	95.210	-1.045	0.308	-296.918	97.990							

0.000

Dep. Variable:		y R-sq	R-squared:				
Model:	OI	S Adj.	Adj. R-squared: F-statistic:				
Method:	Least Square	s F-st					
Date:	Fri, 02 Oct 202	0 Prob	(F-statisti	c):	4.04e-11		
Time:	20:31:0	4 Log-	Log-Likelihood:				
No. Observations:	2	4 AIC:			275.2		
Df Residuals:	2	BIC:	BIC:				
Df Model:		2					
Covariance Type:	nonrobus	it					
COE	f std err	t	P> t	[0.025	0.975		
const 1798.404	0 899.248	2.000	0.059	-71.685	3668.493		
x 1 -250.146	6 117.950	-2.121	0.046	-495.437	-4.856		
x 2 345.540	1 111.367	3.103	0.005	113.940	577.140		

Something more about the coding

• Multivariate linear regression in R:

$$lm(y \sim x_1 + x_2 + x_3, data=dataset)$$

- Multivariate linear regression in Python:
 - import statsmodels.api as sm import statsmodels
 - Intercept is not set as default!X = sm.add_constant(X)
 - model = sm.OLS(Y, X).fit()