

§ 3. 几种重要随机变量的数学期望及方差

1. 两点分布

X	0	1
p_k	$1-p$	p

$$EX=p, \quad DX = EX^2 - (EX)^2 = p - p^2 = pq \circ$$

2. 二项分布

方法1:

$$P\{X = k\} = C_n^k p^k q^{n-k}, k = 0, 1, \dots, n \circ$$

$$EX = \sum_{k=0}^n k \cdot C_n^k p^k q^{n-k} = \sum_{k=0}^n k \cdot \frac{n!}{k!(n-k)!} p^k q^{n-k}$$

$$= np \sum_{k=1}^n \frac{(n-1)!}{(k-1)!(n-1-(k-1))!} p^{k-1} q^{n-1-(k-1)}$$



$$EX = np \sum_{k=1}^n C_{n-1}^{k-1} p^{k-1} q^{n-1-(k-1)} = np \sum_{i=0}^{n-1} C_{n-1}^i p^i q^{n-1-i}$$

$$= np(p+q)^{n-1} = np$$

$$EX^2 = \sum_{k=0}^n k^2 \cdot C_n^k p^k q^{n-k} = \sum_{k=0}^n k^2 \cdot \frac{n!}{k!(n-k)!} p^k q^{n-k}$$

$$= p \sum_{k=1}^n k \frac{n!}{(k-1)!(n-k)!} p^{k-1} q^{n-k}$$

$$= p \sum_{k=1}^n (k-1) \frac{n!}{(k-1)!(n-k)!} p^{k-1} q^{n-k} + p \sum_{k=1}^n \frac{n!}{(k-1)!(n-k)!} p^{k-1} q^{n-k}$$

$$= n(n-1)p^2 \sum_{k=2}^n \frac{(n-2)!}{(k-2)!(n-2-(k-2))!} p^{k-2} q^{n-2-(k-2)} + np$$

$$= n(n-1)p^2 (p+q)^{n-2} + np = n^2 p^2 - np^2 + np$$

$$DX = EX^2 - (EX)^2 = n^2 p^2 - n p^2 + np - n^2 p^2 = np(1-p) = npq$$

方法2:

X_i 服从 (0-1) 分布, $P\{X_i = 0\} = q, P\{X_i = 1\} = p, i = 1, 2, \dots, n$
且 X_1, \dots, X_n 独立, 令 $X = X_1 + \dots + X_n$, 则 X 的可能取值为 $0, 1, \dots, n$,

$$P\{X = k\} = C_n^k p^k q^{n-k}, k = 0, \dots, n$$

$$EX = \sum_{i=1}^n EX_i = np, \quad DX = \sum_{i=1}^n DX_i = npq,$$

3. 泊松分布

设 X 服从参数为 λ 泊松分布,

$$\text{其分布律为 } P\{X = k\} = \frac{\lambda^k}{k!} e^{-\lambda}, \quad k=0, 1, \dots$$

$$EX = \sum_{k=0}^{\infty} k \frac{\lambda^k}{k!} e^{-\lambda} = \lambda e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} = \lambda e^{-\lambda} e^{\lambda} = \lambda$$



$$\begin{aligned} EX^2 &= \sum_{k=0}^{\infty} k^2 \frac{\lambda^k}{k!} e^{-\lambda} = \sum_{k=1}^{\infty} k \frac{\lambda^k}{(k-1)!} e^{-\lambda} \\ &= \sum_{k=1}^{\infty} (k-1) \frac{\lambda^k}{(k-1)!} e^{-\lambda} + \sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!} e^{-\lambda} \\ &= \lambda^2 e^{-\lambda} \sum_{k=2}^{\infty} \frac{\lambda^{k-2}}{(k-2)!} + \lambda e^{-\lambda} e^{\lambda} = \lambda^2 + \lambda \end{aligned}$$

$$DX = EX^2 - (EX)^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$$

4. 均匀分布

$$f(x) = \begin{cases} 1/(b-a), & a < x < b \\ 0, & \text{其它} \end{cases}.$$

$$EX = \int_{-\infty}^{\infty} xf(x)dx = \int_a^b x \frac{1}{b-a} dx = \frac{a+b}{2}$$



$$DX = EX^2 - (EX)^2 = \int_a^b x^2 \frac{1}{b-a} dx - \left(\frac{a+b}{2}\right)^2 = \frac{(b-a)^2}{12}$$

5. 正态分布 $X \sim N(\mu, \sigma^2)$

$$\begin{aligned} EX &= \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} (\sigma t + \mu) e^{-\frac{t^2}{2}} dt, \left(\frac{x-\mu}{\sigma} = t\right) \\ &= \frac{\sigma}{\sqrt{2\pi}} \int_{-\infty}^{\infty} t e^{-\frac{t^2}{2}} dt + \frac{\mu}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{t^2}{2}} dt = \mu \end{aligned}$$

$$\begin{aligned} DX &= E(X - \mu)^2 = \int_{-\infty}^{\infty} (x - \mu)^2 \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx, \left(\frac{x-\mu}{\sigma} = t\right) \\ &= \int_{-\infty}^{\infty} \frac{\sigma^2 t^2}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt = \frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} t^2 e^{-\frac{t^2}{2}} dt = -\frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} t de^{-\frac{t^2}{2}} \\ &= -\frac{\sigma^2}{\sqrt{2\pi}} t e^{-\frac{t^2}{2}} \Big|_{-\infty}^{\infty} + \frac{\sigma^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{t^2}{2}} dt = \sigma^2 \end{aligned}$$



$$\begin{aligned} P\{|X - \mu| \leq \sigma\} &= P\{\mu - \sigma \leq X \leq \mu + \sigma\} \\ &= \Phi\left(\frac{\mu + \sigma - \mu}{\sigma}\right) - \Phi\left(\frac{\mu - \sigma - \mu}{\sigma}\right) = \Phi(1) - \Phi(-1) = 2\Phi(1) - 1 = 0.682 \end{aligned}$$

$$\begin{aligned} P\{|X - \mu| \leq 2\sigma\} &= P\{\mu - 2\sigma \leq X \leq \mu + 2\sigma\} \\ &= 2\Phi(2) - 1 = 0.9544 \end{aligned}$$

$$\begin{aligned} P\{|X - \mu| \leq 3\sigma\} &= P\{\mu - 3\sigma \leq X \leq \mu + 3\sigma\} \\ &= 2\Phi(3) - 1 = 0.9974 \end{aligned}$$

因此，对于正态随机变量来说，它的值落在区间 $[\mu - 3\sigma, \mu + 3\sigma]$ 内几乎是肯定的。

在上一节用切比晓夫不等式估计概率有：

$$P\{|X - \mu| < 3\sigma\} \geq 0.8889$$

