# Chapter 9. Comparing Two Population Means

# 9.1 Introduction

- In Chapter 8, we considered a single population distribution
  - For example, we estimated the mean of a single population distribution
- But it is important to make a comparison between two population distributions

• Two-sample problem:

Set of data observations from population A

$$x_1, x_2, x_3, \dots, x_n$$

Additional set of observations from population B

$$y_1, y_2, y_3, ..., y_m$$

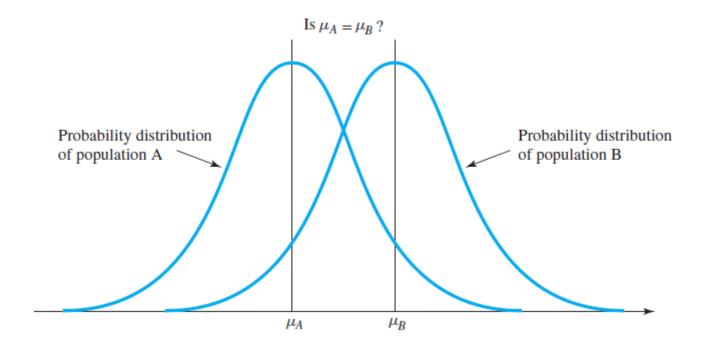
- Sample of data observations  $x_i$  are independent observations from the unknown probability distribution of A
- Sample of data observations  $y_i$  are independent observations from the unknown probability distribution of B
- Sample sizes *n* and *m* need not be equal (but experiments are often designed to have equal sample sizes)

• Example:

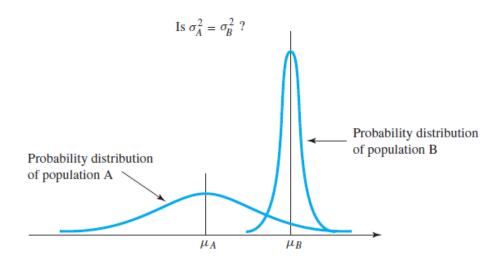
Population 1: Height of students in Korea

Population 2: Height of students in Europe

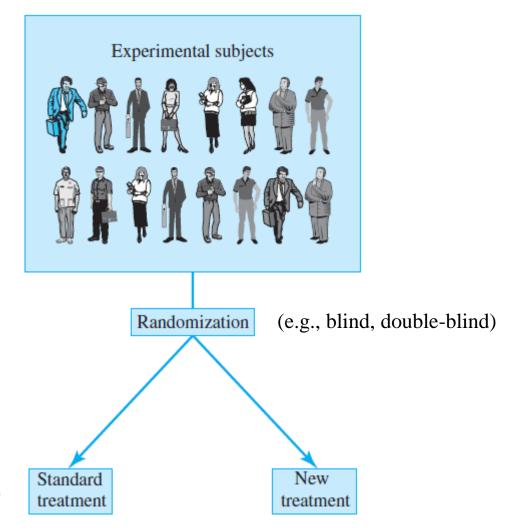
- In general, interested in assessing evidence that there is a difference between the two probability distributions
- One approach is to compare the means of the two probability distributions



- If we find that  $\mu_A \neq \mu_B$ , then this indicates that the two probability distributions are different
- If we find evidence that  $\mu_A = \mu_B$ ,
  - we may conclude that the two probability distributions may be identical
  - or we may further compare the variances of the two data sets



• Example:

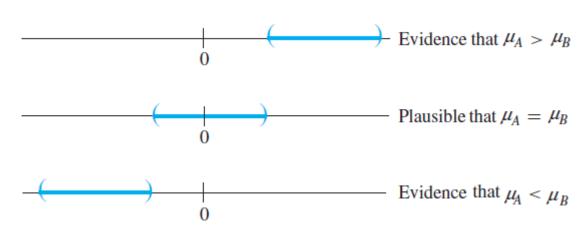


Control group

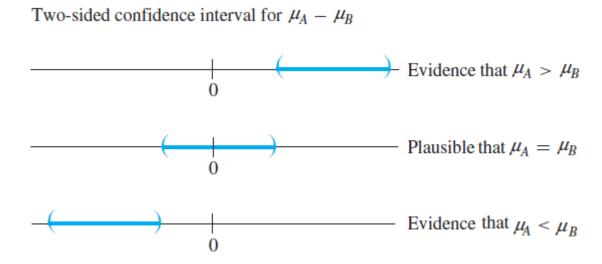
- How do we compare  $\mu_A$  and  $\mu_B$ ?
  - → Since we want to see if the two are the same, we construct a confidence interval for

$$\mu_A - \mu_B$$

Two-sided confidence interval for  $\mu_A - \mu_B$ 



- How do we compare  $\mu_A$  and  $\mu_B$ ?
  - $\rightarrow$  Since we want to see if the two are the same, we construct a confidence interval for  $\mu_A \mu_B$
- We are interested whether this confidence interval contains zero



- How do we compare  $\mu_A$  and  $\mu_B$ ?
  - $\rightarrow$  Since we want to see if the two are the same, we construct a confidence interval for  $\mu_A \mu_B$
- We are interested whether this confidence interval contains zero
- The confidence interval is centered at  $\bar{x} \bar{y}$  (from our samples)

- How do we compare  $\mu_A$  and  $\mu_B$ ?
  - → Another approach is to perform a hypothesis test

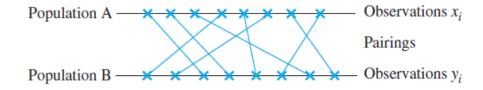
$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 

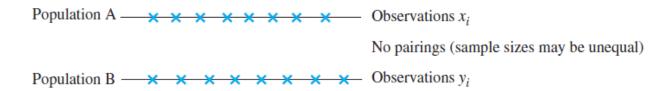
• Then, a small p-value will indicate that the null hypothesis is not plausible

# Paired vs. Independent Samples

- Two data sets may be paired samples or independent samples
- Paired samples may alleviate variability from outside factors

#### Paired Samples





• Paired samples can be expressed as

$$(x_1, y_1), (x_2, y_2), \ldots, (x_n, y_n)$$

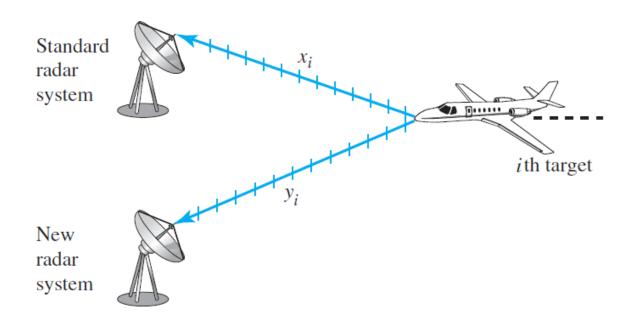
- Samples sizes from the two populations must be equal
- Comparison between the two is then based upon the pairwise differences:

$$z_i = x_i - y_i \qquad 1 \le i \le n$$

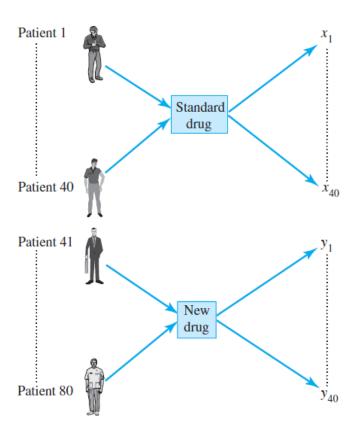
#### • Example:

|  | Day 1   | Day 2   | Difference  |
|--|---|---|---|
| patient 1<br>patient 2<br>patient 3<br>patient 4 | standard drug $x_1$ new drug $y_2$ standard drug $x_3$ new drug $y_4$ | new drug $y_1$ standard drug $x_2$ new drug $y_3$ standard drug $x_4$ | $z_1 = x_1 - y_1$ $z_2 = x_2 - y_2$ $z_3 = x_3 - y_3$ $z_4 = x_4 - y_4$ |
| patient 39 patient 40                            | :<br>standard drug x <sub>39</sub><br>new drug y <sub>40</sub>        | :<br>new drug y <sub>39</sub><br>standard drug x <sub>40</sub>        | $ \vdots  z_{39} = x_{39} - y_{39}  z_{40} = x_{40} - y_{40} $          |

#### • Example:



• Example:



What if the second set of patients happen to be a group more receptive to drugs?

# 9.2 Analysis of Paired Samples

- Analysis of paired samples  $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$  is performed by reducing the problem to a one-sample problem
  - → Calculate the differences

$$z_i = x_i - y_i$$
  $1 \le i \le n$ 

- Data observations  $z_i$  can be taken to be independent identically distributed observations from some probability distribution with mean  $\mu$
- Then, one-sample techniques from Chapter 8 can be applied to

$$z_1, \ldots, z_n$$

• Then, the parameter  $\mu$  is the average difference between A and B

$$\mu = \mu_A - \mu_B$$

- Positive  $\mu$  indicates that the mean population of A is larger than the mean population of B
- Negative  $\mu$  indicates that the mean population of A is smaller than the mean population of B
- Often valuable to test:

$$H_0: \mu = 0$$
 versus  $H_A: \mu \neq 0$ 

→ If the null hypothesis is plausible, then there is no sufficient evidence that the means of A and B are different

• Suppose each observation obtained from population A is thought of as below

$$x_i = \mu_A + \gamma_i + \epsilon_i^A$$

- $\mu_A$ : effect of population A
- $\gamma_i$ : effect of subject i
- $\epsilon_i^A$ : random error (with expectation = 0)

 Suppose each observation obtained from population A is thought of as below

$$x_i = \mu_A + \gamma_i + \epsilon_i^A$$

• Similarly, suppose each observation obtained from population B is thought of as below

$$y_i = \mu_B + \gamma_i + \epsilon_i^B$$

- $\mu_B$ : effect of population B
- $\gamma_i$ : effect of subject i
- $\epsilon_i^B$ : random error (with expectation = 0)

$$x_i = \mu_A + \gamma_i + \epsilon_i^A$$

$$y_i = \mu_B + \gamma_i + \epsilon_i^B$$

• Then, the difference becomes

$$z_i = \mu_A - \mu_B + \epsilon_i^{AB}$$

where the error term is

$$\epsilon_i^{AB} = \epsilon_i^A - \epsilon_i^B$$

• Since the error term is an observation from a distribution with a zero expectation, the differences  $z_i$  are consequently observations from a distribution with expectation

$$\mu = \mu_A - \mu_B$$

which does not depend on the subject effect  $\gamma_i$ 

• *Example*: (Heart rate reduction from new drug)

$$H_0: \mu = 0$$
 versus  $H_A: \mu \neq 0$ 

| Patient | Standard drug | New drug | $z_i = x_i - y_i$ | Patient | Standard drug | New drug | $z_i = x_i - y_i$ |
|---------|---------------|----------|-------------------|---------|---------------|----------|-------------------|
| 1       | 28.5          | 34.8     | -6.3              | 21      | 27.0          | 25.3     | 1.7               |
| 2       | 26.6          | 37.3     | -10.7             | 22      | 33.1          | 34.5     | -1.4              |
| 3       | 28.6          | 31.3     | -2.7              | 23      | 28.7          | 30.9     | -2.2              |
| 4       | 22.1          | 24.4     | -2.3              | 24      | 33.7          | 31.9     | 1.8               |
| 5       | 32.4          | 39.5     | -7.1              | 25      | 33.7          | 36.9     | -3.2              |
| 6       | 33.2          | 34.0     | -0.8              | 26      | 34.3          | 27.8     | 6.5               |
| 7       | 32.9          | 33.4     | -0.5              | 27      | 32.6          | 35.7     | -3.1              |
| 8       | 27.9          | 27.4     | 0.5               | 28      | 34.5          | 38.4     | -3.9              |
| 9       | 26.8          | 35.4     | -8.6              | 29      | 32.9          | 36.7     | -3.8              |
| 10      | 30.7          | 35.7     | -5.0              | 30      | 29.3          | 36.3     | -7.0              |
| 11      | 39.6          | 40.4     | -0.8              | 31      | 35.2          | 38.1     | -2.9              |
| 12      | 34.9          | 41.6     | -6.7              | 32      | 29.8          | 32.1     | -2.3              |
| 13      | 31.1          | 30.8     | 0.3               | 33      | 26.1          | 29.1     | -3.0              |
| 14      | 21.6          | 30.5     | -8.9              | 34      | 25.6          | 33.5     | -7.9              |
| 15      | 40.2          | 40.7     | -0.5              | 35      | 27.6          | 28.7     | -1.1              |
| 16      | 38.9          | 39.9     | -1.0              | 36      | 25.1          | 31.4     | -6.3              |
| 17      | 31.6          | 30.2     | 1.4               | 37      | 23.7          | 22.4     | 1.3               |
| 18      | 36.0          | 34.5     | 1.5               | 38      | 36.3          | 43.7     | -7.4              |
| 19      | 25.4          | 31.2     | -5.8              | 39      | 33.4          | 30.8     | 2.6               |
| 20      | 35.6          | 35.5     | 0.1               | 40      | 40.1          | 40.8     | -0.7              |

Heart rate reductions data set (% reduction in heart rate)

• Example: (Heart rate reduction from new drug)

From the collected data:  $\bar{z} = -2.655$ ,  $s_z = 3.730$ , n = 40

Then, with  $H_0: \mu = 0$ 

$$t = \frac{\sqrt{n}(\bar{z} - \mu)}{s} = \frac{\sqrt{40} \times (-2.655)}{3.730} = -4.50$$

and

$$p$$
-value =  $2 \times P(X > 4.50) \simeq 0.0001$ 

where the random variable X has a t-distribution with 39 degrees of freedom

This analysis reveals that it is *not* plausible that  $\mu = 0$ , and so the experimenter can conclude that there is evidence that the new drug has a different effect from the standard drug.

• Example: (Heart rate reduction from new drug)

From the critical point  $t_{0.005,39} = 2.7079$ , a 99% two-sided confidence interval for the difference between the average effects of the drugs is

$$\mu = \mu_A - \mu_B \in \left(\bar{z} - \frac{t_{0.005,39}s}{\sqrt{40}}, \bar{z} + \frac{t_{0.005,39}s}{\sqrt{40}}\right)$$

$$= \left(-2.655 - \frac{2.7079 \times 3.730}{\sqrt{40}}, -2.655 + \frac{2.7079 \times 3.730}{\sqrt{40}}\right)$$

$$= (-4.252, -1.058)$$

Consequently, based upon this data set the experimenter can conclude that the new drug provides a reduction in a patient's heart rate of somewhere between 1% and 4.25% more on average than the standard drug.

• *Example*: (Radar detection systems)

$$H_0: \mu \geq 0$$
 versus  $H_A: \mu < 0$ 

| Target | Standard radar system | New radar system |                   |
|--------|-----------------------|------------------|-------------------|
|        | $x_i$                 | Уi               | $z_i = x_i - y_i$ |
| 1      | 48.40                 | 51.14            | -2.74             |
| 2      | 47.73                 | 46.48            | 1.25              |
| 3      | 51.30                 | 50.90            | 0.40              |
| 4      | 50.49                 | 49.82            | 0.67              |
| 5      | 47.06                 | 47.99            | -0.93             |
| 6      | 53.02                 | 53.20            | -0.18             |
| 7      | 48.96                 | 46.76            | 2.20              |
| 8      | 52.03                 | 54.44            | -2.41             |
| 9      | 51.09                 | 49.85            | 1.24              |
| 10     | 47.35                 | 47.45            | -0.10             |
| 11     | 50.15                 | 50.66            | -0.51             |
| 12     | 46.59                 | 47.92            | -1.33             |
| 13     | 52.03                 | 52.37            | -0.34             |
| 14     | 51.96                 | 52.90            | -0.94             |
| 15     | 49.15                 | 50.67            | -1.52             |
| 16     | 48.12                 | 49.50            | -1.38             |
| 17     | 51.97                 | 51.29            | 0.68              |
| 18     | 53.24                 | 51.60            | 1.64              |
| 19     | 55.87                 | 54.48            | 1.39              |
| 20     | 45.60                 | 45.62            | -0.02             |
| 21     | 51.80                 | 52.24            | -0.44             |
| 22     | 47.64                 | 47.33            | 0.31              |
| 23     | 49.90                 | 51.13            | -1.23             |
| 24     | 55.89                 | 57.86            | -1.97             |

Radar detection systems data set (distance of target in miles when detected)

• Example: (Radar detection systems)

From the collected data:  $\bar{z} = -0.261$ ,  $s_z = 1.305$ , n = 24

Then, with  $H_0: \mu \geq 0$ 

$$t = \frac{\sqrt{n}(\bar{z} - \mu)}{s} = \frac{\sqrt{24} \times (-0.261)}{1.305} = -0.980$$

and

$$P(X \le -0.980) = 0.170$$

where the random variable X has a t-distribution with 23 degrees of freedom

Large p-value  $\rightarrow$  This data set does not provide sufficient evidence to establish that the new radar system is better than the standard system

# 9.3 Analysis of Independent Samples

- Independent (unpaired) samples
  - n observations  $x_i$  from population A
  - m observations  $y_i$  from population B
- Goal: Inference on the difference between population means,  $\mu_A \mu_B$
- Point estimate:  $\bar{x} \bar{y}$

|                              | Sample size | Sample mean      | Sample standard deviation |
|------------------------------|-------------|------------------|---------------------------|
| Population A<br>Population B | n<br>m      | $ar{x} \\ ar{y}$ | $s_x$ $s_y$               |

- Point estimate:  $\bar{x} \bar{y}$
- For confidence intervals, we need the point estimate and also the standard error
- What is the standard error of  $\bar{x} \bar{y}$ ?

- Point estimate:  $\bar{x} \bar{y}$
- What about the standard error of this estimate?
  - Since  $Var(\bar{x}) = \sigma_A^2/n$  and  $Var(\bar{y}) = \sigma_B^2/m$ , where  $\sigma_A^2$  and  $\sigma_B^2$  are the two population variances, this point estimate has a standard error

s.e.
$$(\bar{x} - \bar{y}) = \sqrt{\frac{\sigma_A^2}{n} + \frac{\sigma_B^2}{m}}$$

• But this includes the population variances of A and B

s.e.
$$(\bar{x} - \bar{y}) = \sqrt{\frac{\sigma_A^2}{n} + \frac{\sigma_B^2}{m}}$$

- What can we do?
  - Use the sample variances of A and B:

s.e.
$$(\bar{x} - \bar{y}) = \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}$$

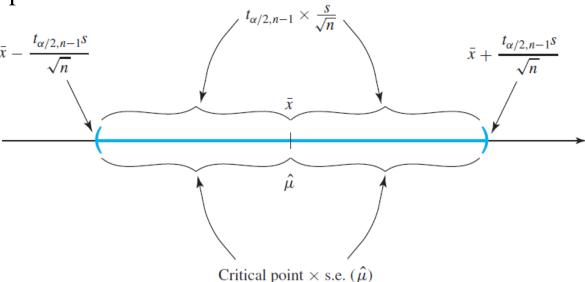
This is the general approach for independent samples

# Independent Samples: t-Intervals

• *Recall*: The basic structure of *t*-intervals

$$\mu \in (\hat{\mu} - \text{critical point} \times \text{s.e.}(\hat{\mu}), \hat{\mu} + \text{critical point} \times \text{s.e.}(\hat{\mu}))$$

• For example:



# Independent Samples: t-Intervals

```
\mu \in (\hat{\mu} - \text{critical point} \times \text{s.e.}(\hat{\mu}), \hat{\mu} + \text{critical point} \times \text{s.e.}(\hat{\mu}))
where \mu = \mu_A - \mu_B
```

- So for independent samples, we need to find the following values:
  - $\blacksquare$   $\hat{\mu}$
  - s.e.  $(\hat{\mu})$
  - critical point

$$\mu \in (\hat{\mu} - \text{critical point} \times \text{s.e.}(\hat{\mu}), \hat{\mu} + \text{critical point} \times \text{s.e.}(\hat{\mu}))$$

where 
$$\mu = \mu_A - \mu_B$$

- Since our goal is the difference  $\mu_A \mu_B$ 
  - $\hat{\mu} \rightarrow \bar{x} \bar{y}$

• s. e. 
$$(\hat{\mu}) \to \text{s. e.} (\bar{x} - \bar{y}) = \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}$$

critical point: ?

- It turns out that we can still use the t-distribution
- But the degree of freedom is calculated as:

$$\nu = \frac{\left(\frac{s_x^2}{n} + \frac{s_y^2}{m}\right)^2}{\frac{s_x^4}{n^2(n-1)} + \frac{s_y^4}{m^2(m-1)}}$$

Round down to the nearest integer

$$\mu \in (\hat{\mu} - \text{critical point} \times \text{s.e.}(\hat{\mu}), \hat{\mu} + \text{critical point} \times \text{s.e.}(\hat{\mu}))$$

where 
$$\mu = \mu_A - \mu_B$$

- Since our goal is the difference  $\mu_A \mu_B$ 
  - $\hat{\mu} \rightarrow \bar{x} \bar{y}$

• s. e. 
$$(\hat{\mu}) \to \text{s. e.} (\bar{x} - \bar{y}) = \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}$$

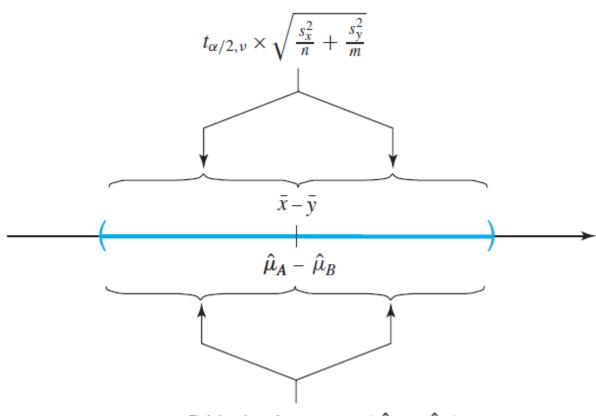
• critical point:  $t_{\alpha/2,v}$ 

 $\mu \in (\hat{\mu} - \text{critical point} \times \text{s.e.}(\hat{\mu}), \hat{\mu} + \text{critical point} \times \text{s.e.}(\hat{\mu}))$ 



A two-sided  $1 - \alpha$  level confidence interval for  $\mu_A - \mu_B$  is therefore

$$\mu_A - \mu_B \in \left( \bar{x} - \bar{y} - t_{\alpha/2,\nu} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}, \bar{x} - \bar{y} + t_{\alpha/2,\nu} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}} \right)$$



Critical point  $\times$  s.e.  $(\hat{\mu}_A - \hat{\mu}_B)$ 

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

What is the 99% two-sided confidence interval?

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

What is the 99% two-sided confidence interval?

1) First, calculate the degrees of freedom:

$$\nu = \frac{\left(\frac{3.438^2}{24} + \frac{3.305^2}{34}\right)^2}{\frac{3.438^4}{24^2 \times 23} + \frac{3.305^4}{34^2 \times 33}} = 48.43 \implies 48$$

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

What is the 99% two-sided confidence interval?

1) First, calculate the degrees of freedom:

$$\nu = \frac{\left(\frac{3.438^2}{24} + \frac{3.305^2}{34}\right)^2}{\frac{3.438^4}{24^2 \times 23} + \frac{3.305^4}{34^2 \times 33}} = 48.43 \implies 48$$

2) Second, find the critical point:  $t_{0.005,48} = 2.6822$ 

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

What is the 99% two-sided confidence interval?

3) Then, the interval becomes:

$$\mu_A - \mu_B \in \left(9.005 - 11.864 - 2.6822\sqrt{\frac{3.438^2}{24} + \frac{3.305^2}{34}}, \right.$$

$$9.005 - 11.864 + 2.6822\sqrt{\frac{3.438^2}{24} + \frac{3.305^2}{34}}\right)$$

$$= (-5.28, -0.44)$$

• Example: Suppose we have two sets of data

**A**: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

As we can see from the two-sided confidence interval, zero is not included in the interval.

• One-sided confidence intervals are also similar:

$$\mu_A - \mu_B \in \left(-\infty, \bar{x} - \bar{y} + t_{\alpha,\nu} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}\right)$$

and

$$\mu_A - \mu_B \in \left(\bar{x} - \bar{y} - t_{\alpha,\nu} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}, \infty\right)$$

• *Recall*: For one sample data

$$\mu \in \left(-\infty, \bar{x} + \frac{t_{\alpha, n-1}s}{\sqrt{n}}\right)$$
  $\mu \in \left(\bar{x} - \frac{t_{\alpha, n-1}s}{\sqrt{n}}, \infty\right)$ 

• Consider the **two-sided hypothesis testing** problem:

$$H_0: \mu_A - \mu_B = \delta$$
 versus  $H_A: \mu_A - \mu_B \neq \delta$ 

• Now the *t*-statistic can be written as

$$t = \frac{\bar{x} - \bar{y} - \delta}{\sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}}$$

• *Recall*: For one-sample data

$$t = \frac{\sqrt{n}(\bar{x} - \mu_0)}{s} = \frac{(\bar{x} - \mu_0)}{s/\sqrt{n}} = \frac{(\bar{x} - \mu_0)}{s. e. (\hat{\mu})}$$

• Consider the **two-sided hypothesis testing** problem:

$$H_0: \mu_A - \mu_B = \delta$$
 versus  $H_A: \mu_A - \mu_B \neq \delta$ 

• When testing if A and B have the same population mean, we set  $\delta = 0$ 

$$t = \frac{\bar{x} - \bar{y}}{\sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}}$$

For the two-sided hypothesis testing problem

$$H_0: \mu_A - \mu_B = \delta$$
 versus  $H_A: \mu_A - \mu_B \neq \delta$ 

for some fixed value  $\delta$  of interest (usually  $\delta = 0$ ), the appropriate t-statistic is

$$t = \frac{\bar{x} - \bar{y} - \delta}{\sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}}$$

The two-sided p-value is calculated as

$$p$$
-value =  $2 \times P(X > |t|)$ 

where the random variable X has a t-distribution with  $\nu$  degrees of freedom, and a size  $\alpha$  hypothesis test accepts the null hypothesis if

$$|t| \leq t_{\alpha/2,\nu}$$

and rejects the null hypothesis when

$$|t| > t_{\alpha/2,\nu}$$

• Example (continued): Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

How do we do the following hypothesis test?

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 

(In other words, the above is the same as saying that  $\delta = 0$  in the below test)

$$H_0: \mu_A - \mu_B = \delta$$
 versus  $H_A: \mu_A - \mu_B \neq \delta$ 

• Example (continued): Suppose we have two sets of data

**A**: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

Hypothesis test:

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 

The *t*-statistic is

$$t = \frac{\bar{x} - \bar{y} - \delta}{\sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}} = \frac{\bar{x} - \bar{y}}{\sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}} = \frac{9.005 - 11.864}{\sqrt{\frac{3.438^2}{24} + \frac{3.305^2}{34}}} = -3.169$$

• *Example (continued)*: For the hypothesis test:

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 

The two-sided p-value is therefore

$$p$$
-value =  $2 \times P(X > 3.169)$ 

where the random variable X has a t-distribution with degrees of freedom

$$v = \frac{\left(\frac{3.438^2}{24} + \frac{3.305^2}{34}\right)^2}{\frac{3.438^4}{24^2 \times 23} + \frac{3.305^4}{34^2 \times 33}} = 48.43$$

Using the integer value  $\nu = 48$  gives

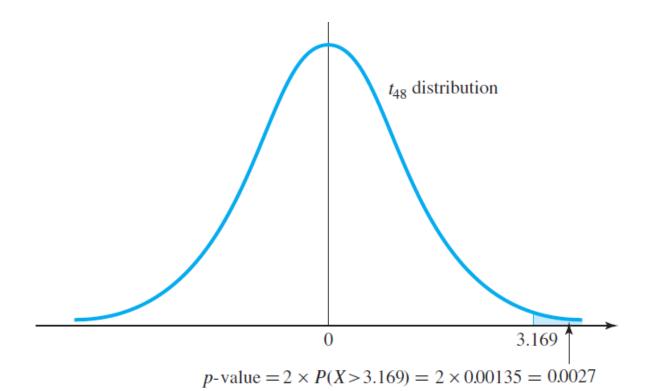
$$p$$
-value  $\simeq 2 \times 0.00135 = 0.0027$ 



Null hypothesis is rejected and conclude that  $\mu_A \neq \mu_B$ 

• *Example (continued)*: For the hypothesis test:

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 



• Example (continued): Suppose we have two sets of data

**A**: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

*Recall*: The two-sided confidence interval did not include zero.

This is consistent with the hypothesis test rejecting the null hypothesis that the two means are the same.

A one-sided hypothesis testing problem

$$H_0: \mu_A - \mu_B \leq \delta$$
 versus  $H_A: \mu_A - \mu_B > \delta$ 

has a p-value

$$p$$
-value =  $P(X > t)$ 

and a size  $\alpha$  hypothesis test accepts the null hypothesis if

$$t \leq t_{\alpha,\nu}$$

and rejects the null hypothesis if

$$t > t_{\alpha,\nu}$$

A size  $\alpha$  test for the one-sided hypotheses

$$H_0: \mu \le \mu_0 \quad \text{versus} \quad H_A: \mu > \mu_0$$

rejects the null hypothesis when

$$t > t_{\alpha,n-1}$$

and accepts the null hypothesis when

$$t \leq t_{\alpha,n-1}$$

Similarly, the one-sided hypothesis testing problem

$$H_0: \mu_A - \mu_B \ge \delta$$
 versus  $H_A: \mu_A - \mu_B < \delta$ 

has a *p*-value

$$p$$
-value =  $P(X < t)$ 

and a size  $\alpha$  hypothesis test accepts the null hypothesis if

$$t \geq -t_{\alpha,\nu}$$

and rejects the null hypothesis if

$$t < -t_{\alpha,\nu}$$

A size  $\alpha$  test for the one-sided hypotheses

$$H_0: \mu \ge \mu_0$$
 versus  $H_A: \mu < \mu_0$  rejects the null hypothesis when

$$t < -t_{\alpha,n-1}$$

and accepts the null hypothesis when

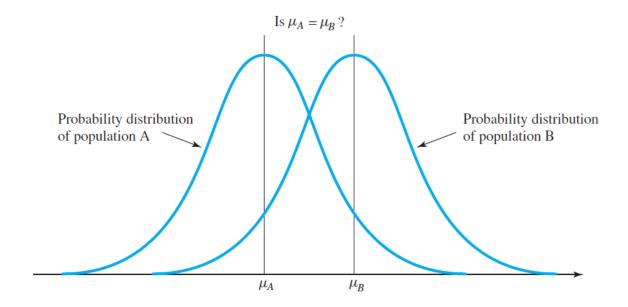
$$t \geq -t_{\alpha,n-1}$$

|  | Hypothesis Testing   |  |   | Confidence Intervals  |   |   |
|--|--|--|---|---|---|---|
|  | $H_0: \mu_A - \mu_B = \delta$ versus $H_A: \mu_A - \mu_B \neq \delta$<br>Significance Levels |  |   | $\left(\bar{x} - \bar{y} - t_{\alpha/2, \nu} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}, \bar{x} - \bar{y} + t_{\alpha/2, \nu} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}\right)$ |   |   |
|  |  |  |   | Confidence Levels   |   |   |
| p-Value                                      | $\alpha = 0.10$  | $\alpha = 0.05$  | $\alpha = 0.01$                                     | $1-\alpha=0.90$   | $1-\alpha=0.95$   | $1-\alpha=0.99$   |
| $\geq 0.10$ $0.05-0.10$ $0.01-0.05$ $< 0.01$ | accept $H_0$ reject $H_0$ reject $H_0$ reject $H_0$  | accept $H_0$<br>accept $H_0$<br>reject $H_0$<br>reject $H_0$ | accept $H_0$ accept $H_0$ accept $H_0$ reject $H_0$ | contains $\delta$ does not contain $\delta$ does not contain $\delta$ does not contain $\delta$   | contains $\delta$ contains $\delta$ does not contain $\delta$ does not contain $\delta$ | contains $\delta$ contains $\delta$ contains $\delta$ does not contain $\delta$ |

The procedures with independent samples discussed so far are known as
 two-sample t-tests without a pooled variance estimate

• What if we want to assume that the two variance are the same?

$$\sigma_A^2 = \sigma_B^2$$



• If we assume that the variances are equal to a common value  $\sigma^2$ , then

$$\sigma_A^2 = \sigma_B^2 = \sigma^2$$

and we can estimate  $\sigma^2$  with

$$\hat{\sigma}^2 = s_p^2 = \frac{(n-1)s_x^2 + (m-1)s_y^2}{n+m-2}$$

pooled variance estimate

• If we assume that the variances are equal to a common value  $\sigma^2$ , then

$$\sigma_A^2 = \sigma_B^2 = \sigma^2$$

and we can estimate  $\sigma^2$  with

$$\hat{\sigma}^2 = s_p^2 = \frac{(n-1)s_x^2 + (m-1)s_y^2}{n+m-2}$$

Also, the standard error

s.e.
$$(\bar{x} - \bar{y}) = \sqrt{\frac{\sigma_A^2}{n} + \frac{\sigma_B^2}{m}} = \sigma\sqrt{\frac{1}{n} + \frac{1}{m}}$$

can be estimated as

s.e.
$$(\bar{x} - \bar{y}) = s_p \sqrt{\frac{1}{n} + \frac{1}{m}}$$

• When a pooled variance estimate is employed, p-values and critical points are calculated from a t-distribution with n + m - 2 degrees of freedom

- When a pooled variance estimate is employed, p-values and critical points are calculated from a t-distribution with n + m 2 degrees of freedom
- Therefore, a two-sided  $1 \alpha$  level confidence interval for  $\mu_A \mu_B$  is

$$\mu_A - \mu_B \in \left(\bar{x} - \bar{y} - t_{\alpha/2, n+m-2} \, s_p \sqrt{\frac{1}{n} + \frac{1}{m}}, \bar{x} - \bar{y} + t_{\alpha/2, n+m-2} \, s_p \sqrt{\frac{1}{n} + \frac{1}{m}}\right)$$

$$\mu \in (\hat{\mu} - \text{critical point} \times \text{s.e.}(\hat{\mu}), \ \hat{\mu} + \text{critical point} \times \text{s.e.}(\hat{\mu}))$$

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

These are the same data from the previous example

Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

These are the same data from the previous example

But if we assume the population variances to be the same, the common standard deviation becomes:

$$s_p = \sqrt{\frac{(n-1)s_x^2 + (m-1)s_y^2}{n+m-2}} = \sqrt{\frac{(23 \times 3.438^2) + (33 \times 3.305^2)}{24 + 34 - 2}} = 3.360$$

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_v = 3.305$ 

$$s_p = 3.360$$

How do we do the following hypothesis test?

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

 $s_p = 3.360$ 

How do we do the following hypothesis test?

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 

The *t*-statistic is

$$t = \frac{\bar{x} - \bar{y}}{s_p \sqrt{\frac{1}{n} + \frac{1}{m}}} = \frac{9.005 - 11.864}{3.360 \sqrt{\frac{1}{24} + \frac{1}{34}}} = -3.192$$

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

 $s_p = 3.360$ 

How do we do the following hypothesis test?

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 

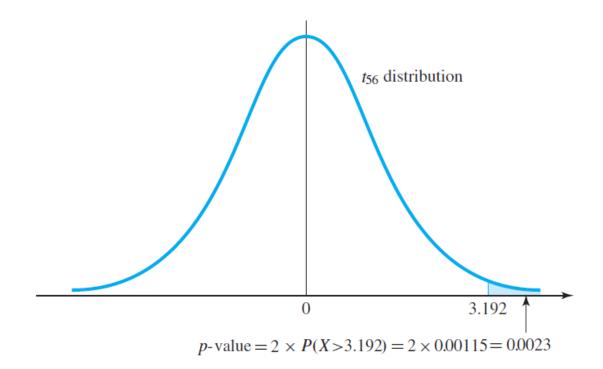
The *t*-statistic is

$$t = \frac{\bar{x} - \bar{y}}{s_p \sqrt{\frac{1}{n} + \frac{1}{m}}} = \frac{9.005 - 11.864}{3.360 \sqrt{\frac{1}{24} + \frac{1}{34}}} = -3.192$$

and the *p*-value is p-value =  $2 \times P(X > 3.192) \simeq 2 \times 0.00115 = 0.0023$ where the random variable X has a t-distribution with degrees of freedom n + m - 2 = 56,

• Example: Hypothesis test

$$H_0: \mu_A = \mu_B$$
 versus  $H_A: \mu_A \neq \mu_B$ 



• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_v = 3.305$ 

$$s_p = 3.360$$

What about a 99% two-sided confidence interval?

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

$$s_p = 3.360$$

What about a 99% two-sided confidence interval?

First, find the critical point:  $t_{0.005,56} = 2.6665$ 

## Independent Samples: Pooled Variance

• Example: Suppose we have two sets of data

A: 
$$n = 24$$
,  $\bar{x} = 9.005$ ,  $s_x = 3.438$ 

**B**: 
$$m = 34$$
,  $\bar{y} = 11.864$ ,  $s_y = 3.305$ 

$$s_p = 3.360$$

What about a 99% two-sided confidence interval?

First, find the critical point:  $t_{0.005,56} = 2.6665$ 

Then, the two-sided interval becomes:

$$\mu_A - \mu_B \in \left(9.005 - 11.864 - 2.6665 \times 3.360 \times \sqrt{\frac{1}{24} + \frac{1}{34}}, \right.$$

$$9.005 - 11.864 + 2.6665 \times 3.360 \times \sqrt{\frac{1}{24} + \frac{1}{34}}\right)$$

$$= (-5.25, -0.47)$$

# Independent Samples: Pooled Variance

A two-sided p-value is calculated as  $2 \times P(X > |t|)$ , where the random variable X has a t-distribution with n + m - 2 degrees of freedom, and one-sided p-values are P(X > t) and P(X < t). A size  $\alpha$  two-sided hypothesis test accepts the null hypothesis if

$$|t| \leq t_{\alpha/2,n+m-2}$$

and rejects the null hypothesis when

$$|t| > t_{\alpha/2, n+m-2}$$

and size  $\alpha$  one-sided hypothesis tests have rejection regions  $t > t_{\alpha,n+m-2}$  or  $t < -t_{\alpha,n+m-2}$ .

These procedures are known as two-sample *t*-tests *with* a pooled variance estimate.

## **Independent Samples**

- Pooled or un-pooled?
  - General procedure is without a pooled variance estimate (this is the safer choice)
  - However, the two population variances are known to be equal, then the pooled variance provides a powerful analysis (e.g., shorter confidence interval)
  - If the results from with and without pooled variance estimates are different, that is likely because the two sample standard deviations are different

#### Two Sample z-Procedures

- If we assume that the population variances  $\sigma_A^2$  and  $\sigma_B^2$  are known, then we can use the standard normal distribution instead of a *t*-distribution
- Similar to the one-sample case, two-sample *t*-tests with large sample sizes are essentially equivalent to the two-sample *z*-tests
  - Thus, two-sample *t*-tests can be considered as small-sample tests

#### Two Sample z-Procedures

A two-sided  $1 - \alpha$  level confidence interval for the difference in population means  $\mu_A - \mu_B$  is

$$\mu_A - \mu_B \in \left(\bar{x} - \bar{y} - z_{\alpha/2}\sqrt{\frac{\sigma_A^2}{n} + \frac{\sigma_B^2}{m}}, \bar{x} - \bar{y} + z_{\alpha/2}\sqrt{\frac{\sigma_A^2}{n} + \frac{\sigma_B^2}{m}}\right)$$

The appropriate z-statistic for the null hypothesis  $H_0: \mu_A - \mu_B = \delta$  is

$$z = \frac{\bar{x} - \bar{y} - \delta}{\sqrt{\frac{\sigma_A^2}{n} + \frac{\sigma_B^2}{m}}}$$

A two-sided *p*-value is calculated as  $2 \times \Phi(-|z|)$ , and one-sided *p*-values are  $1 - \Phi(z)$  and  $\Phi(z)$ . A size  $\alpha$  two-sided hypothesis test accepts the null hypothesis if

$$|z| \leq z_{\alpha/2}$$

and rejects the null hypothesis when

$$|z| > z_{\alpha/2}$$

and size  $\alpha$  one-sided hypothesis tests have rejection regions  $z > z_{\alpha}$  or  $z < -z_{\alpha}$ .

# Summary of Chapter 9

- Comparison of two population means
  - Paired and unpaired samples
  - Mainly use *t*-distributions
- Paired samples
  - Take the difference and create a one-sample dataset
  - Then, perform analyses using methods in Chapter 8
- Unpaired samples
  - Unpooled variance: assume that the two datasets have different variances
  - Pooled variance: assume that the two datasets have the same variance
- Use z-procedures if population variances are known

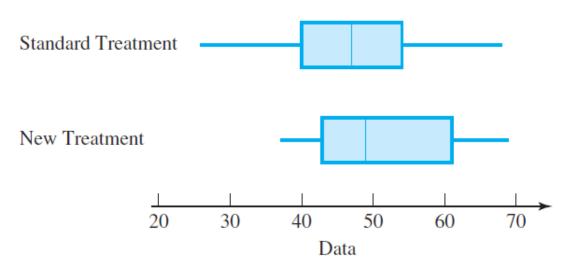
# **Independent Samples**

More examples ...

- Data from standard and new treatments are shown below
- Same sample size but unpaired data

| Standard<br>treatment<br>x <sub>i</sub> | New<br>treatment<br>y <sub>i</sub> |
|---|------------------------------------|
| 33                                      | 65                                 |
| 54                                      | 61                                 |
| 62                                      | 37                                 |
| 46                                      | 47                                 |
| 52                                      | 45                                 |
| 42                                      | 53                                 |
| 34                                      | 53                                 |
| 51                                      | 69                                 |
| 26                                      | 49                                 |
| 68                                      | 42                                 |
| 47                                      | 40                                 |
| 40                                      | 67                                 |
| 46                                      | 46                                 |
| 51                                      | 43                                 |
| 60                                      | 51                                 |

• If we simply analyze each data set independently, it is difficult to know whether the two are different with statistical significance



#### Standard Treatment

Sample size = 15 Sample mean = 47.47 Sample standard deviation = 11.40

#### New Treatment

Sample size = 15 Sample mean = 51.20 Sample standard deviation = 10.09

- First, let's perform the general (unpooled) procedure
- We want evidence that the new treatment is better than the standard treatment, thus we set the hypothesis test as:

$$H_0: \mu_A \geq \mu_B$$
 versus  $H_A: \mu_A < \mu_B$ 

- First, let's perform the general (unpooled) procedure
- We want evidence that the new treatment is better than the standard treatment, thus we set the hypothesis test as:

$$H_0: \mu_A \ge \mu_B$$
 versus  $H_A: \mu_A < \mu_B$ 

• Degrees of freedom:  $\left(\frac{11}{2}\right)$ 

$$v = \frac{\left(\frac{11.40^2}{15} + \frac{10.09^2}{15}\right)^2}{\frac{11.40^4}{15^2 \times 14} + \frac{10.09^4}{15^2 \times 14}} = 27.59$$

which can be rounded down to v = 27

- First, let's perform the general (unpooled) procedure
- We want evidence that the new treatment is better than the standard treatment, thus we set the hypothesis test as:

$$H_0: \mu_A \geq \mu_B$$
 versus  $H_A: \mu_A < \mu_B$ 

Degrees of freedom:  $v = \frac{\left(\frac{11.40^2}{15} + \frac{10.09^2}{15}\right)^2}{\frac{11.40^4}{15^2 \times 14} + \frac{10.09^4}{15^2 \times 14}} = 27.59$ 

which can be rounded down to v = 27

• t-statistic:  $t = \frac{47.47 - 51.20}{\sqrt{\frac{11.40^2}{15} + \frac{10.09^2}{15}}} = -0.949$ 

• Finally, the *p*-value is:

$$p$$
-value =  $P(X < -0.949) = 0.175$ 

where the random variable X has a t-distribution with  $\nu = 27$  degrees of freedom

→ Cannot reject the null hypothesis

• For a 99% confidence interval, we first find the critical point

$$t_{0.01,27} = 2.473$$

Then the interval becomes

$$\mu_A - \mu_B \in \left(-\infty, 47.47 - 51.20 + 2.473\sqrt{\frac{11.40^2}{15} + \frac{10.09^2}{15}}\right)$$

$$= (-\infty, 5.99)$$

→ The standard treatment may be 6 points better than the new treatment on average

• Suppose we decide to perform a pooled variance analysis

(assume 
$$s_x = 11.40$$
 and  $s_v = 10.09$  are similar)

- Suppose we decide to perform a pooled variance analysis
- For the hypothesis test:

$$H_0: \mu_A \ge \mu_B$$
 versus  $H_A: \mu_A < \mu_B$ 

Pooled variance is:

$$s_p^2 = \frac{(14 \times 11.40^2) + (14 \times 10.09^2)}{28} = 115.88$$

so that the pooled standard deviation is  $s_p = \sqrt{115.88} = 10.76$ 

- Suppose we decide to perform a pooled variance analysis
- For the hypothesis test:

$$H_0: \mu_A \geq \mu_B$$
 versus  $H_A: \mu_A < \mu_B$ 

Pooled variance is:

$$s_p^2 = \frac{(14 \times 11.40^2) + (14 \times 10.09^2)}{28} = 115.88$$

so that the pooled standard deviation is  $s_p = \sqrt{115.88} = 10.76$ 

• *t*-statistic becomes:

$$t = \frac{47.47 - 51.20}{10.76\sqrt{\frac{1}{15} + \frac{1}{15}}} = -0.946$$

• Then the *p*-value for the pooled variance procedure is:

$$p$$
-value =  $P(X < -0.946) = 0.175$ 

where the random variable X has a t-distribution with v = 28 degrees of freedom

- Both the unpooled and pooled analyses results in similar *p*-values
- One-sided 99% confidence interval also turns out to be very similar

## **Excel Examples**

• Next, let's show how to solve similar questions in Excel