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MASTER'S THESIS

# **Wasserstein gradient flows: modeling and applications**

Master's Educational Program: Data Science

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МАГИСТЕРСКАЯ ДИССЕРТАЦИЯ

**Градиентные потоки Вассерштайна: методы  
моделирования и применение в приложениях**

Магистерская образовательная программа: Науки о Данных

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# **Wasserstein gradient flows: modeling and applications**

Petr Mokrov

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## **Abstract**

Wasserstein gradient flows provide a powerful means of understanding and solving many diffusion equations. Specifically, Fokker-Planck equations, which model the diffusion of probability measures, can be understood as gradient descent over entropy functionals in Wasserstein space. This equivalence, introduced by Jordan, Kinderlehrer and Otto, inspired the so-called JKO scheme to approximate these diffusion processes via an implicit discretization of the gradient flow in Wasserstein space. Solving the optimization problem associated to each JKO step, however, presents serious computational challenges. We introduce a scalable method to approximate Wasserstein gradient flows, targeted to machine learning applications. Our approach relies on input-convex neural networks (ICNNs) to discretize the JKO steps, which can be optimized by stochastic gradient descent. Unlike previous work, our method does not require domain discretization or particle simulation. As a result, we can sample from the measure at each time step of the diffusion and compute its probability density. We demonstrate our algorithm's performance by computing diffusions following the Fokker-Planck equation and apply it to unnormalized density sampling as well as nonlinear filtering. The work is based on our article [47].

Keywords: optimal transport, Fokker-Planck equation, Wasserstein gradient flows.

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# **Градиентные потоки Вассерштайна: методы моделирования и применение в приложениях**

**Петр Мокров**

Представлено в Сколковский институт науки и технологий  
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## **Реферат**

Градиентные потоки Вассерштайна являются мощным средством, помогающим понять и решить некоторые диффузионные уравнения. В частности, уравнение Фоккера-Планка, которое описывает достаточно широкий класс диффузионных процессов, можно представить как градиентный спуск, минимизирующий энтропийный функционал в пространстве вероятностных мер с метрикой Вассерштайна. В пионерской работе Jordan, Kinderlehrer и Otto была предложена так называемая ЖКО схема для аппроксимации диффузионных процессов с помощью неявной временной дискретизации градиентного потока в пространстве Вассерштайна. Вместе с тем, задачи оптимизации, возникающие на каждом шаге ЖКО схемы, вычислительно трудозатратны. В настоящей работе предложен масштабируемый метод аппроксимации градиентных потоков Вассерштайна, который можно применять в приложениях машинного обучения. Для параметризации ЖКО шагов используются выпуклые нейронные сети (ICNNs), оптимизируемые с помощью стохастического градиентного спуска. В отличие от предшествующих работ, предложенный подход не требует дискретизации пространства и симуляции частиц. Описан алгоритм создания выборки из распределения градиентного потока в каждый момент времени и метод вычисления функции плотности этого распределения. В качестве вычислительного эксперимента было произведено моделирование некоторых диффузионных процессов, подчиняющихся уравнению Фоккера-Планка а также были решены практические задачи, связанные с семплированием из ненормируемой плотности и нелинейной фильтрацией. Работа основана на статье [47] автора настоящей диссертации с соавторами.

Ключевые слова: оптимальный транспорт, уравнение Фоккера-Планка, градиентные потоки Вассерштайна

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# 1 Introduction

*Gradient flows*, also known as *steepest descent curves* are very common field in evolution equations. Given a functional  $F : X \rightarrow \mathbb{R}$  defined on a vector space  $X$  and initial point  $x_0 \in X$  we seek for a curve  $x(t)$ , starting at  $x_0$  and minimizing  $F$  as fast as possible (Thus we are to solve the equation  $x'(t) = -\nabla F(x(t))$  for  $t \in \mathbb{R}_+$ ). Thorough gradient flows overview can be found in [60]. In the Euclidean case  $X = \mathbb{R}^n$  gradient flows are equivalent to ODE. Discretization of this flow leads to the gradient descent minimization algorithm. The steepest descent curves could be considered in more general metric spaces [4]. Of the particular interest of our work are gradient flows on the probability measures space equipped with 2-Wasserstein metric also known as Wasserstein gradient flows. We introduce basic notations and statements in the next subsections.

## 1.1 Basic Notations and Statements

$\mathcal{P}_2(\mathbb{R}^D)$  denotes the set of Borel probability measures on  $\mathbb{R}^D$  with finite second moment.  $\mathcal{P}_{2,ac}(\mathbb{R}^D)$  denotes its subset of probability measures absolutely continuous with respect to Lebesgue measure. For  $\rho \in \mathcal{P}_{2,ac}(\mathbb{R}^D)$ , we denote by  $\frac{d\rho}{dx}(x)$  its density with respect to the Lebesgue measure.  $\Pi(\mu, \nu)$  denotes the set of probability measures on  $\mathbb{R}^D \times \mathbb{R}^D$  with marginals  $\mu$  and  $\nu$ . For measurable  $T : \mathbb{R}^D \rightarrow \mathbb{R}^D$ , we denote by  $T\#$  the associated push-forward operator between measures. The (squared) Wasserstein-2 metric  $\mathcal{W}_2$  between  $\mu, \nu \in \mathcal{P}_2(\mathbb{R}^D)$  is [67]:

$$\mathcal{W}_2^2(\mu, \nu) \stackrel{\text{def}}{=} \min_{\pi \in \Pi(\mu, \nu)} \int_{\mathbb{R}^D \times \mathbb{R}^D} \|x - y\|_2^2 d\pi(x, y), \quad (1)$$

For  $\mu \in \mathcal{P}_{2,ac}(\mathbb{R}^D)$ , there exists a  $\mu$ -unique map  $\nabla\psi^* : \mathbb{R}^D \rightarrow \mathbb{R}^D$  that is the gradient of a convex function  $\psi^* : \mathbb{R}^D \rightarrow \mathbb{R} \sqcup \{\infty\}$  satisfying  $\nabla\psi^*\#\mu = \nu$  [45]. From Brenier's theorem [12], it follows that  $\pi^* = [\text{id}_{\mathbb{R}^D}, \nabla\psi^*]\#\mu$  is the unique minimizer of (1), i.e.,

$$\mathcal{W}_2^2(\mu, \nu) = \int_{\mathbb{R}^D} \|x - \nabla\psi^*(x)\|_2^2 d\mu(x). \quad (2)$$

The equation (2) and Brenier's theorem play crucial role in our research, as they allow to recast the search for  $\mathcal{W}_2$  distance as minimization with respect to the set of convex functions.

## 1.2 Wasserstein Gradient Flows

The idea of gradient flow in Wasserstein space  $(\mathcal{P}_2(\mathbb{R}^D), \mathcal{W}_2)$  is similar to Euclidean case: the flow follows the steepest descent direction, but this time the notion of gradient is more complex.

A curve of measures  $\{\rho_t\}_{t \in \mathbb{R}_+}$  following the Wasserstein gradient flow of a functional  $\mathcal{F}$

solves the continuity equation

$$\frac{\partial \rho_t}{\partial t} = \operatorname{div}(\rho_t \nabla_x \mathcal{F}'(\rho_t)), \quad \text{s.t. } \rho_0 = \rho^0, \quad (3)$$

where  $\mathcal{F}'(\cdot)$  is the first variation of  $\mathcal{F}$  [4, Theorem 8.3.1]. The term on the right can be understood as the gradient of  $\mathcal{F}$  in Wasserstein space, a vector field perturbatively rearranging the mass in  $\rho_t$  to yield the steepest possible local change of  $\mathcal{F}$ . We refer the reader to [59, Chapter 8] for an accessible introduction.

Wasserstein gradient flows are used in various applied tasks. For example, gradient flows are applied in training [8, 42, 24] or refinement [7] of implicit generative models. In reinforcement learning, gradient flows facilitate policy optimization [55, 71]. Other tasks include crowd motion modelling [44, 58, 50], dataset optimization [2], and in-between animation [25].

Many applications come from the connection between Wasserstein gradient flows and Stochastic differential equations (SDEs). Consider an  $\mathbb{R}^D$ -valued stochastic process  $\{X_t\}_{t \in \mathbb{R}_+}$  governed by the following Itô SDE:

$$dX_t = -\nabla \Phi(X_t)dt + \sqrt{2\beta^{-1}}dW_t, \quad \text{s.t. } X_0 \sim \rho^0 \quad (4)$$

where  $\Phi : \mathbb{R}^D \rightarrow \mathbb{R}$  is the potential function,  $W_t$  is the standard Wiener process, and  $\beta > 0$  is the magnitude. The solution of (4) is called an *advection-diffusion* process. It arises in various applications including physics [62], finance [20, 52], population dynamics [34, 13] and molecular discovery [3]. In machine learning the SDE in the form (4) appears in applications filtering [33, 19] and unnormalized posterior sampling via a discretization of the Langevin diffusion [69].

The marginal measure  $\rho_t$  of  $X_t$  which follows (4) at each time satisfies the *Fokker-Planck equation* with fixed diffusion coefficient:

$$\frac{\partial \rho_t}{\partial t} = \operatorname{div}(\nabla \Phi(x)\rho_t) + \beta^{-1}\Delta \rho_t, \quad \text{s.t. } \rho_0 = \rho^0. \quad (5)$$

Equation (5) turns out to be the Wasserstein gradient flow (3) for  $\mathcal{F}$  given by the Fokker-Planck free energy functional [31]

$$\mathcal{F}_{\text{FP}}(\rho) = \mathcal{U}(\rho) - \beta^{-1}\mathcal{E}(\rho), \quad (6)$$

where  $\mathcal{U}(\rho) = \int_{\mathbb{R}^D} \Phi(x)d\rho(x)$  is the *potential energy* and  $\mathcal{E}(\rho) = -\int_{\mathbb{R}^D} \log \frac{d\rho}{dx}(x)d\rho(x)$  is the *entropy*. As the result, to solve the SDE (4), one may compute the Wasserstein gradient flow of the Fokker-Planck equation with the free-energy functional  $\mathcal{F}_{\text{FP}}$  given by (6).

### 1.3 JKO Scheme

Modelling of the processes which satisfy (5) (and correspondingly (4)) is the primal goal of our work. However, the closed-form solutions are generally intractable, necessitating numerical approximation techniques. Jordan, Kinderlehrer, and Otto proposed a method — later abbreviated as *JKO integration* — to approximate the dynamics of  $\rho_t$  in (5) [31]. It consists of a time-discretization update of the continuous flow given by:

$$\rho^{(k)} \leftarrow \arg \min_{\rho \in \mathcal{P}_2(\mathbb{R}^n)} \left[ \mathcal{F}(\rho) + \frac{1}{2h} \mathcal{W}_2^2(\rho^{(k-1)}, \rho) \right] \quad (7)$$

where  $\rho^{(0)} = \rho^0$  is the initial condition and  $h > 0$  is the time-discretization step size. The discrete time gradient flow converges to the continuous one as  $h \rightarrow 0$ , i.e.,  $\rho^{(k)} \approx \rho_{kh}$ . The method was further developed in [4, 60], but performing JKO iterations in practice remains challenging thanks to the minimization with respect to  $\mathcal{W}_2$ . Our main contributions are summarized as follows.

PM: можно тут ещё добавить что-нибудь про icnn

**Contributions.** We propose a scalable parametric method to approximate Wasserstein gradient flows with Fokker-Planck functional  $\mathcal{F}_{FP}$  (6) via JKO stepping using input-convex neural networks (ICNNs) [6]. Specifically, we leverage Brenier’s theorem to bypass the costly computation of the Wasserstein distance, and parametrize the optimal transport map as the gradient of an ICNN. Given sample access to the initial measure  $\rho_0$ , we use stochastic gradient descent (SGD) to sequentially learn time-discretized JKO dynamics of  $\rho_t$ . The trained model can sample from a continuous approximation of  $\rho_t$  and compute its density  $\frac{d\rho_t}{dx}(x)$ . We demonstrate performance by computing diffusion following the Fokker-Planck equation and applying it to unnormalized density sampling as well as nonlinear filtering.

## 2 Literature Review

One way to compute the diffusion which satisfies (5) is to use a fixed discretization of the domain and apply standard numerical integration methods [17, 49, 14, 16, 39] to get  $\rho_t$ . For example, [50] proposes a method to approximate the diffusion based on JKO stepping and entropy-regularized optimal transport. However, these methods are limited to small dimensions since the discretization of space grows exponentially.

An alternative to domain discretization is stochastic particle simulation which exploits the SDE form (4). It involves drawing random samples (particles) from the initial distribution and simulating their evolution via standard methods such as Euler-Maruyama scheme [35, §9.2]. After convergence, the particles are approximately distributed according to the stationary distribution,



but no density estimate is readily available.

Another way to avoid discretization is to parameterize the density of  $\rho_t$ . Most methods approximate only the first and second moments  $\rho_t$ , e.g., via Gaussian approximation. Kalman filtering approaches can then compute the dynamics [33, 38, 32, 61]. More advanced Gaussian mixture approximations [64, 1] or more general parametric families have also been studied [63, 68]. In [48], variational methods are used to minimize the divergence between the predictive and the true density.

The closest to our approach is a line of works related to JKO scheme. A common method to perform JKO steps is to discretize the spatial domain. For support size  $\lesssim 10^6$ , (7) can be solved by standard optimal transport algorithms [51]. Another original idea combining spatial discretization with set of convex functions discretization proposed in [10]. However, in dimensions  $D \geq 3$ , discrete supports can hardly approximate continuous distributions and hence the dynamics of gradient flows. To tackle this issue, [23] propose a stochastic parametric method to approximate the density of  $\rho_t$ . The authors regularize the Wasserstein distance in the JKO step to ensure strict convexity and solve the unconstrained dual problem via stochastic program on a finite linear subset of basis functions. The method is biased and yields *unnormalized* probability density without direct sample access.

PM: Можно целую субсекцию сделать про применение ICNNs

### 3 Computational Methodology

We now describe our approach to compute Wasserstein gradient flows via JKO stepping with ICNNs.

#### 3.1 JKO Reformulation via Optimal Push-forwards Maps

Our key idea is to replace the optimization (7) over probability measures by an optimization over convex functions, an idea inspired by [10]. Thanks to Brenier's theorem, for any  $\rho \in \mathcal{P}_{2,ac}$  there exists a unique  $\rho^{(k-1)}$ -measurable gradient  $\nabla\psi : \mathbb{R}^D \rightarrow \mathbb{R}^D$  of a convex function  $\psi$  satisfying  $\rho = \nabla\psi\# \rho^{(k-1)}$ . We set  $\rho = \nabla\psi\# \rho^{(k-1)}$  and rewrite (7) as an optimization over convex  $\psi$ :

$$\psi^{(k)} \leftarrow \arg \min_{\text{Convex } \psi} \left[ \mathcal{F}(\nabla\psi\# \rho^{(k-1)}) + \frac{1}{2h} \mathcal{W}_2^2(\rho^{(k-1)}, \nabla\psi\# \rho^{(k-1)}) \right]. \quad (8)$$

To proceed to the next step of JKO scheme, we define  $\rho^{(k)} \stackrel{\text{def}}{=} \nabla\psi^{(k)}\# \rho^{(k-1)}$ .

Since  $\rho$  is the pushforward of  $\rho^{(k-1)}$  by the gradient of a convex function  $\nabla\psi$ , the  $\mathcal{W}_2^2$  term

in (8) can be evaluated explicitly, simplifying the Wasserstein-2 distance term in (8):

$$\psi^{(k)} \leftarrow \arg \min_{\text{Convex } \psi} \left[ \mathcal{F}(\nabla \psi \# \rho^{(k-1)}) + \frac{1}{2h} \int_{\mathbb{R}^D} \|x - \nabla \psi(x)\|_2^2 d\rho^{(k-1)}(x) \right]. \quad (9)$$

This formulation avoids the difficulty of computing Wasserstein-2 distances. An additional advantage is that we can *sample* from  $\rho^{(k)}$ . Since  $\rho^{(k)} = [\nabla \psi^{(k)} \circ \dots \circ \nabla \psi^{(1)}] \# \rho^0$ , one may sample  $x_0 \sim \rho^{(0)}$ , and then  $\nabla \psi^{(k)} \circ \dots \circ \nabla \psi^{(1)}(x_0)$  gives a sample from  $\rho^{(k)}$ . Moreover, if functions  $\psi^{(\cdot)}$  are strictly convex, then gradients  $\nabla \psi^{(\cdot)}$  are invertible. In this case, the *density*  $\frac{d\rho^{(k)}}{dx}$  of  $\rho^{(k)} = \nabla \psi^{(k)} \circ \dots \circ \nabla \psi^{(1)} \# \rho^0$  is computable by the change of variables formula (assuming  $\psi^{(\cdot)}$  are twice differentiable)

$$\frac{d\rho^{(k)}}{dx}(x_k) = [\det \nabla^2 \psi^{(k)}(x_{k-1})]^{-1} \dots [\det \nabla^2 \psi^{(1)}(x_0)]^{-1} \cdot \frac{d\rho^{(0)}}{dx}(x_0), \quad (10)$$

where  $x_i = \nabla \psi^{(i)}(x_{i-1})$  for  $i = 1, \dots, k$  and  $\frac{d\rho^{(0)}}{dx}$  is the density of  $\rho^{(0)}$ .

### 3.2 Stochastic Optimization for JKO via ICNNs

In general, the solution  $\psi^{(k)}$  of (9) is intractable since it requires optimization over all convex functions. To tackle this issue, [10] discretizes the space of convex function. The approach also requires discretization of measures  $\rho^{(k)}$  limiting this method to small dimensions.

We propose to parametrize the search space using input convex neural networks (ICNNs) [6] satisfying a universal approximation property among convex functions [18]. ICNNs are parametric models of the form  $\psi_\theta : \mathbb{R}^D \rightarrow \mathbb{R}$  with  $\psi_\theta$  convex w.r.t. the input. ICNNs are constructed from neural network layers, with restrictions on the weights and activation functions to preserve the input-convexity, see [6, §3.1] or [36, §B.2]. The parameters are optimized via deep learning optimization techniques such as SGD.

The JKO step then becomes finding the optimal parameters  $\theta^*$  for  $\psi_\theta$ :

$$\theta^* \leftarrow \arg \min_{\theta} \left[ \mathcal{F}(\nabla \psi_\theta \# \rho^{(k-1)}) + \frac{1}{2h} \int_{\mathbb{R}^D} \|x - \nabla \psi_\theta(x)\|_2^2 d\rho^{(k-1)}(x) \right]. \quad (11)$$

If the functional  $\mathcal{F}$  can be estimated stochastically using random batches from  $\rho^{(k-1)}$ , then SGD can be used to optimize  $\theta$ .  $\mathcal{F}_{\text{FP}}$  given by (6) is an example of such a functional:

**Theorem 1** (Estimator of  $\mathcal{F}_{\text{FP}}$ ). *Let  $\rho \in \mathcal{P}_{2,ac}(\mathbb{R}^D)$  and  $T : \mathbb{R}^D \rightarrow \mathbb{R}^D$  be a diffeomorphism. For*

a random batch  $x_1, \dots, x_N \sim \rho$ , the expression  $[\widehat{\mathcal{U}}_T(x_1, \dots, x_N) - \beta^{-1} \widehat{\Delta \mathcal{E}}_T(x_1, \dots, x_N)]$ , where

$$\begin{aligned}\widehat{\mathcal{U}}_T(x_1, \dots, x_N) &\stackrel{\text{def}}{=} \frac{1}{N} \sum_{n=1}^N \Phi(T(x_n)) \text{ and} \\ \widehat{\Delta \mathcal{E}}_T(x_1, \dots, x_N) &\stackrel{\text{def}}{=} \frac{1}{N} \sum_{n=1}^N \log |\det \nabla T(x_n)|,\end{aligned}$$

is an estimator of  $\mathcal{F}_{\text{FP}}(T\sharp\rho)$  up to constant (w.r.t.  $T$ ) shift given by  $\beta^{-1}\mathcal{E}(\rho)$ .

*Proof.*  $\widehat{\mathcal{U}}_T$  is a straightforward unbiased estimator for  $\mathcal{U}(T\sharp\rho)$ . Let  $p$  and  $p_T$  be the densities of  $\rho$  and  $T\sharp\rho$ . Since  $T$  is a diffeomorphism, we have  $p_T(y) = p(x) \cdot |\det \nabla T(x)|^{-1}$  where  $x = T^{-1}(y)$ . Using the change of variables formula, we write

$$\begin{aligned}\mathcal{E}(T\sharp\rho) &= - \int_{\mathbb{R}^D} p_T(y) \log p_T(y) dy \\ &= - \int_{\mathbb{R}^D} p(x) \cdot |\det \nabla T(x)|^{-1} \log \left[ p(x) \cdot |\det \nabla T(x)|^{-1} \right] \cdot |\det \nabla T(x)| dx \\ &= - \int_{\mathbb{R}^D} p(x) \log p(x) dx + \int_{\mathbb{R}^D} p(x) \log |\det \nabla T(x)| dx \\ &= \mathcal{E}(\rho) + \int_{\mathbb{R}^D} p(x) \log |\det \nabla T(x)| dx, \\ \implies \Delta \mathcal{E}_T(\rho) &\stackrel{\text{def}}{=} \mathcal{E}(T\sharp\rho) - \mathcal{E}(\rho) = \int_{\mathbb{R}^D} \log |\det \nabla T(x)| d\rho(x)\end{aligned}$$

which explains that  $\widehat{\Delta \mathcal{E}}_T$  is an unbiased estimator of  $\Delta \mathcal{E}_T(\rho)$ . As the result,  $\widehat{\mathcal{U}}_T - \beta^{-1} \widehat{\Delta \mathcal{E}}_T$  is an estimator for  $\mathcal{F}_{\text{FP}}(T\sharp\rho) = \mathcal{U}(T\sharp\rho) - \beta^{-1} \mathcal{E}(T\sharp\rho)$  up to a shift of  $\beta^{-1} \mathcal{E}(\rho)$ .  $\square$

To apply Theorem 1 to our case, we take  $T \leftarrow \nabla \psi_\theta$  and  $\rho \leftarrow \rho^{(k-1)}$  to obtain a stochastic estimator for  $\mathcal{F}_{\text{FP}}(\nabla \psi_\theta \sharp \rho^{(k-1)})$  in (11). Here,  $\beta^{-1} \mathcal{E}(\rho^{(k-1)})$  is  $\theta$ -independent and constant since  $\rho^{(k-1)}$  is fixed, so the offset of the estimator plays no role in the optimization w.r.t.  $\theta$ .

Algorithm 1 details our stochastic JKO method for  $\mathcal{F}_{\text{FP}}$ . The training is done solely based on random samples from the initial measure  $\rho^0$ : its density is not needed.

This algorithm assumes  $\mathcal{F}$  is the Fokker-Planck diffusion energy functional. However, our method admits straightforward generalization to any  $\mathcal{F}$  that can be stochastically estimated; studying such functionals is a promising avenue for future work.

### 3.3 Computing the Density of the Diffusion Process

Our algorithm provides a *computable density* for  $\rho^{(k)}$ . As discussed in §3.1, it is possible to sample from  $\rho^{(k)}$  while simultaneously computing the density of the samples. However, this approach does not provide a direct way to evaluate  $\frac{d\rho^{(k)}}{dx}(x_k)$  for arbitrary  $x_k \in \mathbb{R}^D$ . We resolve this issue below.

---

**Algorithm 1:** Fokker-Planck JKO via ICNNs

---

**Input** : Initial measure  $\rho^0$  accessible by samples;  
JKO discretization step  $h > 0$ , number of JKO steps  $K > 0$ ;  
target potential  $\Phi(x)$ , diffusion process temperature  $\beta^{-1}$ ;  
batch size  $N$ ;  
**Output** : trained ICNN models  $\{\psi^{(k)}\}_{k=1}^K$  representing JKO steps  
**for**  $k = 1, 2, \dots, K$  **do**  
     $\psi_\theta \leftarrow$  basic ICNN model;  
    **for**  $i = 1, 2, \dots$  **do**  
        Sample batch  $Z \sim \rho^0$  of size  $N$ ;  
         $X \leftarrow \nabla \psi^{(k-1)} \circ \dots \circ \nabla \psi^{(1)}(Z)$ ;  
         $\widehat{\mathcal{W}}_2^2 \leftarrow \frac{1}{N} \sum_{x \in X} \|\nabla \psi_\theta(x) - x\|_2^2$ ;  
         $\widehat{\mathcal{U}} \leftarrow \frac{1}{N} \sum_{x \in X} \Phi(\nabla \psi_\theta(x))$ ;  
         $\widehat{\Delta \mathcal{E}} \leftarrow \frac{1}{N} \sum_{x \in X} \log \det \nabla^2 \psi_\theta(x)$ ;  
         $\widehat{\mathcal{L}} \leftarrow \frac{1}{2h} \widehat{\mathcal{W}}_2^2 + \widehat{\mathcal{U}} - \beta^{-1} \widehat{\Delta \mathcal{E}}$ ;  
        Perform a gradient step over  $\theta$  by using  $\frac{\partial \widehat{\mathcal{L}}}{\partial \theta}$ ;  
     $\psi^{(k)} \leftarrow \psi_\theta$

---

If a convex function is strongly convex, then its gradient is bijective on  $\mathbb{R}^D$ . By the change of variables formula for  $x_k \in \mathbb{R}^D$ , it holds  $\frac{d\rho^{(k)}}{dx}(x_k) = \frac{d\rho^{(k-1)}}{dx}(x_{k-1}) \cdot [\det \nabla^2 \psi^{(k)}(x_{k-1})]^{-1}$  where  $x_k = \nabla \psi^{(k)}(x_{k-1})$ . To compute  $x_{k-1}$ , one needs to solve the convex optimization problem:

$$x_k = \nabla \psi^{(k)}(x_{k-1}) \quad \Longleftrightarrow \quad x_{k-1} = \arg \max_{x \in \mathbb{R}^D} [\langle x, x_k \rangle - \psi^{(k)}(x)]. \quad (12)$$

If we know the density of  $\rho^0$ , to compute the density of  $\rho^{(k)}$  at  $x_k$  we solve  $k$  convex problems

$$x_{k-1} = \arg \max_{x \in \mathbb{R}^D} [\langle x, x_k \rangle - \psi^{(k)}(x)] \quad \dots \quad x_0 = \arg \max_{x \in \mathbb{R}^D} [\langle x, x_1 \rangle - \psi^{(1)}(x)]$$

to obtain  $x_{k-1}, \dots, x_0$  and then evaluate the density as

$$\frac{d\rho_k}{dx}(x_k) = \frac{d\rho^0}{dx}(x_0) \cdot \left[ \prod_{i=1}^k \det \nabla^2 \psi^{(i)}(x_{i-1}) \right]^{-1}.$$

Note the steps above provide a general method for tracing back the position of a particle along the flow, and density computation is simply a byproduct.

### 3.4 Nonlinear Filtering Theory

In this subsection we give a theoretical introduction to the Nonlinear Filtering as well as describe our practical implementation details.

Consider a diffusion process  $X_t$  governed by the Fokker-Planck equation (5). At times  $t_1 < t_2 < \dots < t_K$  we obtain noisy observations of the process  $Y_k = X_{t_k} + v_k$  with  $v_k \sim \mathcal{N}(0, \sigma)$ . The goal is to compute the predictive distribution  $p_{t,X}(x|Y_{1:K})$  for  $t \geq t_K$  given observations  $Y_{1:K}$ .

For each  $k$  and  $t \geq t_k$  predictive distribution  $p_{t,X}(x|Y_{1:k})$  follows the diffusion process on time interval  $[t_k, t]$  with initial distribution  $p_{t_k,X}(x|Y_{1:k})$ . If  $t_k = t$  then

$$p_{t_k,X}(x|Y_{1:k}) \propto p(Y_k|X_{t_k} = x) \cdot p_{t_k,X}(x|Y_{1:k-1}). \quad (13)$$

For  $k = 1, \dots, K$ , we sequentially obtain the predictive distribution  $p_{t_k,X}(x|Y_{1:k})$  by using the previous predictive distribution  $p_{t_{k-1},X}(x|Y_{1:k-1})$ . First, given access to  $p_{t_{k-1},X}(x|Y_{1:k-1})$ , we approximate the diffusion on interval  $[t_{k-1}, t_k]$  with initial distribution  $p_{t_{k-1},X}(x|Y_{1:k-1})$  by our Algorithm 1 to get access to  $p_{t_k,X}(x|Y_{1:k-1})$ . In particular, we perform  $n_k$  JKO steps of size  $h_k = \frac{t_k - t_{k-1}}{n_k}$  and obtain ICNNs  $\psi_1^{(k)}, \dots, \psi_{n_k}^{(k)}$  (approximately) satisfying

$$\mu_{p_{t_k,X}(x|Y_{1:k-1})} = [\nabla \psi_{n_k}^{(k)} \circ \dots \circ \nabla \psi_1^{(k)}] \# \mu_{p_{t_{k-1},X}(x|Y_{1:k-1})} \quad (14)$$

Here  $\mu_{p(\cdot)}$  is the measure with density  $p(\cdot)$ . We define  $B_k \stackrel{\text{def}}{=} \nabla \psi_{n_k}^{(k)} \circ \dots \circ \nabla \psi_1^{(k)}$ .

Let  $x_k \in \mathbb{R}^D$  and sequentially define  $x_{i-1} = B_i^{-1}(x_i)$  for  $i = k, \dots, 1$ . We derive

$$\begin{aligned} p_{t_k,X}(x_k|Y_{1:k}) &\stackrel{(13)}{\propto} \\ p(Y_k|X_{t_k} = x_k) \cdot p_{t_k,X}(x_k|Y_{1:k-1}) &\stackrel{(14)}{=} \\ p(Y_k|X_{t_k} = x_k) \cdot [\det \nabla B_k(x_{k-1})]^{-1} \cdot p_{t_{k-1},X}(x_{k-1}|Y_{1:k-1}) &\stackrel{(13)}{\propto} \\ &\dots \\ \prod_{i=1}^k p(Y_i|X_{t_i} = x_i) \cdot [\prod_{i=1}^k \det \nabla B_i(x_{i-1})]^{-1} \cdot p_{t_0,X}(x_0) &\quad (15) \end{aligned}$$

where we substitute (13) sequentially for  $k, k-1, \dots, 1$ . As the result, from (15) we obtain the unnormalized density of predictive distribution  $p_{t_k,X}(x_k|Y_{1:k})$ . To sample from the predictive distribution (to train the next step  $k+1$ ) we use Metropolis-Hastings algorithm [57]. For completeness, we recall the Algorithm 2 below. The algorithm builds a chain  $x^{(1)}, x^{(2)}, \dots$  converging to the distribution given by unnormalized density  $\pi(\cdot)$ . As input, the algorithm also takes a family of proposal distributions  $q_x(\cdot)$  for  $x \in \mathbb{R}^D$ . The value  $\alpha(\cdot, \cdot)$  is called the acceptance probability.

---

**Algorithm 2:** Metropolis-Hastings algorithm

---

**Input** : Unnormalized density  $\pi(\cdot)$ ; family of proposal distributions  $q_x(\cdot)$  ( $x \in \mathbb{R}^D$ )

**Output** : Sequence  $x^{(1)}, x^{(2)}, x^{(3)}, \dots$  of samples from  $\pi$

Select  $x^{(0)} \in \mathbb{R}^D$

**for**  $j = 1, 2, \dots$  **do**

Sample  $y \sim q_{x^{(j-1)}}$ ;

Compute  $\alpha(x^{(j-1)}, y) = \min \left( 1, \frac{\pi(y)q_y(x^{(j-1)})}{\pi(x^{(j-1)})q_{x^{(j-1)}}(y)} \right)$

With probability  $\alpha(x^{(j-1)}, y)$  set  $x^{(j)} \leftarrow y$ ; otherwise set  $x^{(j)} \leftarrow x^{(j-1)}$

---

To sample from  $p_{t_k, X}(x_k | Y_{1:k})$  we use Algorithm 2 with  $\pi$  equal to unnormalized density (15). We note that computing  $\pi(x_k)$  for  $x_k \in \mathbb{R}^D$  is not easy since it requires computing pre-images  $x_{k-1}, \dots, x_0$  by inverting  $B_k, B_{k-1}, \dots, B_1$ . As the consequence, this makes computation of acceptance probability  $\alpha(\cdot, \cdot)$  hard. To resolve this issue, we choose special  $x$ -independent proposals

$$q = q_x \stackrel{\text{def}}{=} (B_k \circ B_{k-1} \circ \dots \circ B_1) \# \mu_{p_{0,X}}. \quad (16)$$

In this case, all det terms in  $\alpha(x, y)$  vanish simplifying the computation (we write  $x = x_k$ ,  $y = y_k$ ):

$$\begin{aligned} \frac{\pi(y)q_y(x)}{\pi(x)q_x(y)} &= \frac{\pi(y)q(x)}{\pi(x)q(y)} = \\ \frac{p_{0,X}(y_0) \prod_{i=1}^k p_{t_i, Y}(Y_i | X_{t_i} = y_i) \prod_{i=1}^k \det \nabla B_i(x_{i-1}) \cdot p_{0,X}(x_0) \prod_{i=1}^k \det \nabla B_i(y_{i-1})}{p_{0,X}(x_0) \prod_{i=1}^k p_{t_i, Y}(Y_i | X_{t_i} = x_i) \prod_{i=1}^k \det \nabla B_i(y_{i-1}) \cdot p_{0,X}(y_0) \prod_{i=1}^k \det \nabla B_i(x_{i-1})} &= \\ \frac{\prod_{i=1}^k p_{t_i, Y}(Y_i | X_{t_i} = y_i)}{\prod_{i=1}^k p_{t_i, Y}(Y_i | X_{t_i} = x_i)} & \quad (17) \end{aligned}$$

To compute (17) one needs to know preimages  $x_{k-1}, \dots, x_0$  and  $y_{k-1}, \dots, y_0$  of points  $y = y_k$  and  $x = x_k$  respectively. They can be straightforwardly computed when sampling from  $q$  happens (16). The resulting Algorithm could be found in Appendix I.

## 4 Experiments and Results

In this section, we evaluate our method on toy and real-world applications. Our code is written in PyTorch and is publicly available at

<https://github.com/PetrMokrov/Large-Scale-Wasserstein-Gradient-Flows>

The experiments are conducted on a GTX 1080Ti. In most cases, we performed several random restarts to obtain mean and variation of the considered metric. As the result, experiments require about 100-150 hours of computation.

## 4.1 General Experimental Details

**Neural network architectures.** In all experiments, we use the DenseICNN [36, Appendix B.2] architecture for  $\psi_\theta$  in Algorithm 1 with *SoftPlus* activations. The network  $\psi_\theta$  is twice differentiable w.r.t. the input  $x$  and has bijective gradient  $\nabla\psi_\theta : \mathbb{R}^D \rightarrow \mathbb{R}^D$  with positive semi-definite Hessian  $\nabla^2\psi_\theta(x) \succeq 0$  at each  $x$ . We use automatic differentiation to compute  $\nabla\psi_\theta$  and  $\nabla^2\psi_\theta$ . Throughout our experiments we set the number of hidden layers to be equal 2 and vary the width of the model depending on the task. We use Adam optimizer with learning rate decreasing with the number of JKO steps. We initialize the ICNN models either via pretraining to satisfy  $\nabla\psi_\theta(x) \approx x$  or by using parameters  $\theta$  obtained from the previous JKO step.

**Metric.** To qualitatively compare measures, we use the symmetric Kullback-Leibler divergence

$$\text{SymKL}(\rho_1, \rho_2) \stackrel{\text{def}}{=} \text{KL}(\rho_1 \parallel \rho_2) + \text{KL}(\rho_2 \parallel \rho_1), \quad (18)$$

where  $\text{KL}(\rho_1 \parallel \rho_2) \stackrel{\text{def}}{=} \int_{\mathbb{R}^D} \log \frac{d\rho_1}{d\rho_2}(x) d\rho_1(x)$  is the Kullback-Leibler divergence. For particle-based methods, we obtain an approximation of the distribution by kernel density estimation.

**Competitive methods details.** For [Dual JKO], we used the implementation provided by the authors with default hyper-parameters. For [EM PR] we implemented the Proximal Recursion operator following the pseudocode of [15] and used the default hyper-parameters but we increased the number of particles for fair comparison with the vanilla [EM] algorithm. Note we limited the number of particles to  $N = 10^4$  because of the high computational complexity of the method. For [SVGD], we used the official implementation available at

<https://github.com/dilinwang820/Stein-Variational-Gradient-Descent>

In particle-based simulations [EM], [BBF] and [EM PR] we used the particle propagation timestep  $dt = 10^{-3}$ .

We estimate the SymKL (18) using Monte Carlo (MC) on  $10^4$  samples. In our method, MC estimate is straightforward since the method permits both sampling and computing the density. In particle-based methods, we use kernel density estimator to approximate the density utilizing `scipy` implementation of `gaussian_kde` with bandwidth chosen by Scott's rule. In [Dual JKO], we employ importance sampling procedure and normalization constant estimation as detailed in [23].

We set  $\beta$  to be equal to 1 throughout our experiments.

## 4.2 Convergence to Stationary Solution

Starting from an arbitrary initial measure  $\rho^0$ , an advection-diffusion process (5) converges to the unique stationary solution  $\rho^*$  [56] with density

$$\frac{d\rho^*}{dx}(x) = Z^{-1} \exp(-\beta\Phi(x)), \quad (19)$$

| $D$ | $M$ | $l$ | $w$  |
|-----|-----|-----|------|
| 2   | 5   | 10  | 256  |
| 4   | 6   | 10  | 384  |
| 6   | 7   | 10  | 512  |
| 8   | 8   | 10  | 512  |
| 10  | 9   | 10  | 512  |
| 12  | 10  | 10  | 1024 |
| 13  | 10  | 10  | 512  |
| 32  | 10  | 6   | 1024 |

Table 1: Hyper-parameters in the convergence exp.

where  $Z = \int_{\mathbb{R}^D} \exp(-\beta\Phi(x))dx$  is the normalization constant. This property makes it possible to compute the symmetric KL between the distribution to which our method converges and the ground truth, provided  $Z$  is known.

We use  $\mathcal{N}(0, 16I_D)$  as the initial measure  $\rho^0$  and a random Gaussian mixture  $\frac{1}{N_p} \sum_{m=1}^M \mathcal{N}(\mu_m, I_D)$ , where  $\mu_1, \dots, \mu_M \sim \text{Uniform}([- \frac{l}{2}, \frac{l}{2}]^D)$  as the stationary measure  $\rho^*$ . We set the width  $w$  of used ICNNs  $\psi_\theta$  depending on dimension  $D$ . The parameters are summarized in Table 1. In our method, we perform  $K = 40$  JKO steps with step size  $h = 0.1$ . We compare with a particle simulation method (with  $10^3, 10^4, 10^5$  particles) based on the Euler-Maruyama [EM] approximation [35, §9.2]. We repeat the experiment 5 times and report the averaged results in Figure 1.

Each JKO step uses 1000 gradient descent iterations of Algorithm 1. For dimensions  $D = 2, 4, \dots, 12$  the first 20 JKO transitions are optimized with  $lr = 5 \cdot 10^{-3}$  and the remaining steps use  $lr = 2 \cdot 10^{-3}$ . For qualitative experiments in  $D = 13, 32$  we perform 50 and 70 JKO steps with step size  $h = 0.1$ . The learning rate setup in these cases is similar to quantitative experiment setting but has additional stage with  $lr = 5 \cdot 10^{-4}$  on the final JKO steps. The batch size is  $N = 512$ .

In Figure 2, we present qualitative results of our method converging to the ground truth in  $D = 13, 32$ .

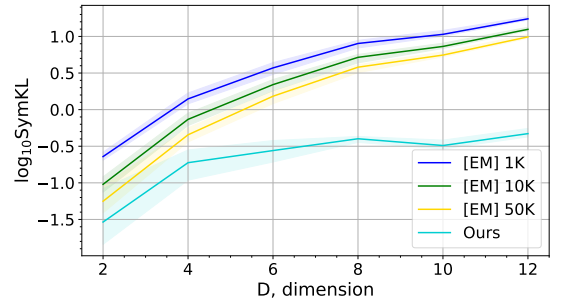


Figure 1: SymKL between the computed and the stationary measure in  $D = 2, 4, \dots, 12$

## 4.3 Modeling Ornstein-Uhlenbeck Processes

Ornstein-Uhlenbeck processes are advection-diffusion processes (5) with  $\Phi(x) = \frac{1}{2}(x-b)^T A(x-b)$  for symmetric positive definite  $A \in \mathbb{R}^{D \times D}$  and  $b \in \mathbb{R}^D$ . They are among the few examples where



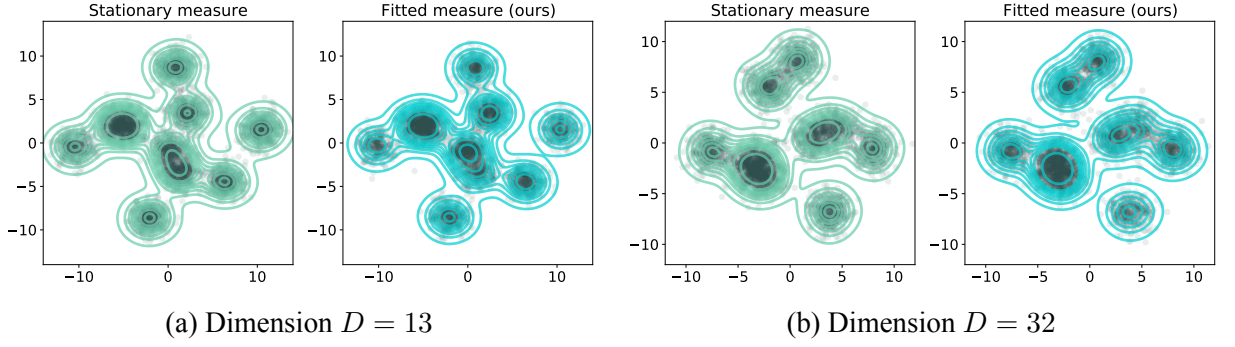


Figure 2: Projections to 2 first PCA components of the true stationary measure and the measure approximated by our method in dimensions  $D = 13$  (on the left) and  $D = 32$  (on the right).

we know  $\rho_t$  for any  $t \in \mathbb{R}^+$  in closed form, when the initial measure  $\rho^0$  is Gaussian [66]. This allows to quantitatively evaluate the computed dynamics of the process, not just the stationary measure.

We choose matrices  $A \in \mathbb{R}^{D \times D}$  to be randomly generated using `sklearn.datasets.make_spd_matrix`. Vectors  $b \in \mathbb{R}^D$  are sampled from standard Gaussian measure. All ICNNs  $\psi_\theta$  have  $w = 64$  and we train each of them for 500 iterations per JKO step with  $lr = 5 \cdot 10^{-3}$  and batch size  $N = 1024$ . For the entropy regularized JKO method we follow the recommendations by authors of [23]. The only difference is that for each dimension we select the support of the kernels which results in the best SymKL metric value. The details are given in our code.

We set  $\rho^0$  to be the standard Gaussian measure  $\mathcal{N}(0, I_D)$  and approximate the dynamics of the process by our method with JKO step  $h = 0.05$  and compute SymKL between the true  $\rho_t$  and the approximate one at time  $t = 0.5$  and  $t = 0.9$ . We repeat the experiment 15 times in dimensions  $D = 1, 2, \dots, 12$  and report the performance at in Figure 3. The baselines are [EM] with  $10^3, 10^4, 5 \times 10^4$  particles, EM particle simulation endowed with the Proximal Recursion operator [EM PR] with  $10^4$  particles [15], and the parametric dual inference method [23] for JKO steps [Dual JKO]. The detailed comparison for times  $t = 0.1, 0.2, \dots, 1$  is given in Appendix II.

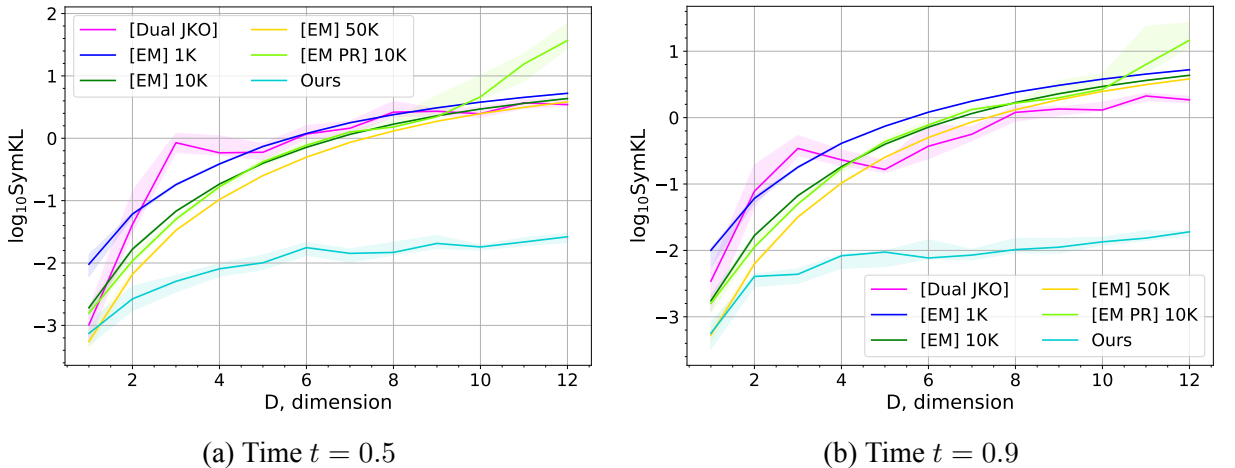


Figure 3: SymKL values between the computed measure and the true measure  $\rho_t$  at  $t = 0.5$  (on the left) and  $t = 0.9$  (on the right) in dimensions  $D = 1, 2, \dots, 12$ . Best viewed in color.

| Dataset  | Accuracy |        | Log-Likelihood |        |
|----------|----------|--------|----------------|--------|
|          | Ours     | [SVGD] | Ours           | [SVGD] |
| covtype  | 0.75     | 0.75   | -0.515         | -0.515 |
| german   | 0.67     | 0.65   | -0.6           | -0.6   |
| diabetis | 0.775    | 0.78   | -0.45          | -0.46  |
| twonorm  | 0.98     | 0.98   | -0.059         | -0.062 |
| ringnorm | 0.74     | 0.74   | -0.5           | -0.5   |
| banana   | 0.55     | 0.54   | -0.69          | -0.69  |
| splice   | 0.845    | 0.85   | -0.36          | -0.355 |
| waveform | 0.78     | 0.765  | -0.485         | -0.465 |
| image    | 0.82     | 0.815  | -0.43          | -0.44  |

Table 2: Comparison of our method with [SVGD] [41] for Bayesian log. regression.

| Dataset  | $w$ | $lr$              | $iter$ | batch | $K$ |
|----------|-----|-------------------|--------|-------|-----|
| covtype  | 512 | $2 \cdot 10^{-5}$ | $10^4$ | 1024  | 6   |
| german   | 512 | $2 \cdot 10^{-4}$ | 5000   | 512   | 5   |
| diabetis | 128 | $5 \cdot 10^{-5}$ | 6000   | 1024  | 16  |
| twonorm  | 512 | $5 \cdot 10^{-5}$ | 5000   | 1024  | 7   |
| ringnorm | 512 | $5 \cdot 10^{-5}$ | 5000   | 1024  | 2   |
| banana   | 128 | $2 \cdot 10^{-4}$ | 5000   | 1024  | 5   |
| splice   | 512 | $2 \cdot 10^{-3}$ | 2000   | 512   | 5   |
| waveform | 512 | $5 \cdot 10^{-5}$ | 5000   | 512   | 2   |
| image    | 512 | $5 \cdot 10^{-5}$ | 5000   | 512   | 5   |

Table 3: Hyper-parameters we use in Bayesian log. regression experiment.

#### 4.4 Unnormalized Posterior Sampling in Bayesian Logistic Regression

An important task in Bayesian machine learning to which our algorithm can be applied is sampling from an unnormalized posterior distribution. Given the model parameters  $x \in \mathbb{R}^D$  with the prior distribution  $p_0(x)$  as well as the conditional density  $p(\mathcal{S}|x) = \prod_{m=1}^M p(s_m|x)$  of the data  $\mathcal{S} = \{s_1, \dots, s_M\}$ , the posterior distribution is given by

$$p(x|\mathcal{S}) = \frac{p(\mathcal{S}|x)p_0(x)}{p(\mathcal{S})} \propto p(\mathcal{S}|x)p_0(x) = p_0(x) \cdot \prod_{m=1}^M p(s_m|x).$$

Computing the normalization constant  $p(\mathcal{S})$  is in general intractable, underscoring the need for estimation methods that sample from  $p(\mathcal{S}|x)$  given the density only up to a normalizing constant.

In our context, sampling from  $p(x|\mathcal{S})$  can be solved similarly to the task in §4.2. From (19), it follows that the advection-diffusion process with temperature  $\beta > 0$  and  $\Phi(x) = -\frac{1}{\beta} \log [p_0(x) \cdot p(\mathcal{S}|x)]$  has  $\frac{d\rho^*}{dx}(x) = p(x|\mathcal{S})$  as the stationary distribution. Thus, we can use our method to approximate the diffusion process and obtain a sampler for  $p(x|\mathcal{S})$  as a result.

The potential energy  $\mathcal{U}(\rho) = \int_{\mathbb{R}^D} \Phi(x) d\rho(x)$  can be estimated efficiently by using a trick similar to the ones in stochastic gradient Langevin dynamics [69], which consists in resampling samples in  $\mathcal{S}$  uniformly. For evaluation, we consider the Bayesian linear regression setup of [41]. We use the 8 datasets from [46]. The number of features ranges from 2 to 60 and the dataset size from 700 to 7400 data points. We also use the Covtype dataset<sup>1</sup> with 500K data points and 54 features. The prior on regression weights  $w$  is given by  $p_0(w|\alpha) = \mathcal{N}(w|0, \alpha^{-1})$  with  $p_0(\alpha) = \text{Gamma}(\alpha|1, 0.01)$ , so the prior on parameters  $x = [w, \alpha]$  of the model is given by  $p_0(x) = p_0(w, \alpha) = p_0(w|\alpha) \cdot p_0(\alpha)$ . To remove positiveness constraint on  $\alpha$  we consider  $[w, \log(\alpha)]$  as the regression model parameters instead of  $[w, \alpha]$ . To learn the posterior distribution  $p(x|S_{\text{train}})$  we

<sup>1</sup><https://www.csie.ntu.edu.tw/~cjlin/libsvmtools/datasets/binary.html>

use JKO step size  $h = 0.1$ . Let  $iter$  denote the number of gradient steps over  $\theta$  per each JKO step and  $K$  denote the overall number of JKO steps. The used hyper-parameters for each dataset are summarized in Table 3.

To estimate the log-likelihood and accuracy of the predictive distribution on  $\mathcal{S}_{test}$  based on  $p(x|\mathcal{S}_{train})$ , we use straightforward MC estimate on  $2^{12}$  random parameter samples. We randomly split each dataset into train  $\mathcal{S}_{train}$  and test  $\mathcal{S}_{test}$  ones with ratio 4:1 and apply the inference on the posterior  $p(x|\mathcal{S}_{train})$ . In Table 2, we report accuracy and log-likelihood of the predictive distribution on  $\mathcal{S}_{test}$ . As the baseline, we use particle-based Stein Variational Gradient Descent [41]. We use the author’s implementation with the default hyper-parameters.

## 4.5 Nonlinear Filtering

We demonstrate the application of our method to filtering a nonlinear diffusion. See §3.4 for the theoretical introduction and algorithmic details.

For evaluation, we consider the experimental setup of [23, §6.3]. We assume that the 1-dimensional diffusion process  $X_t$  has potential function  $\Phi(x) = \frac{1}{\pi} \sin(2\pi x) + \frac{1}{4}x^2$  which makes the process highly nonlinear. We simulate nonlinear filtering on the time interval  $t_{start} = 0$  sec.,  $t_{fin} = 5$  sec. and take the noise observations each 0.5 sec. The noise variance is  $\sigma^2 = 1$  and  $p(X_0) = \mathcal{N}(X_0|0, 1)$ . Specifically, to obtain the noise observations  $Y_k = X_{t_k} + v_k$  from the process, we simulate a particle  $X_0$  randomly sampled from the initial measure  $\mathcal{N}(0, 1)$  by using Euler-Maruyama method to obtain the trajectory  $X_t$ . At observation times  $t_1 = 0.5, \dots, t_9 = 4.5$  we add random noise  $v_k \sim \mathcal{N}(0, 1)$  to obtain observations  $Y_1, \dots, Y_9$ .

We predict the conditional density  $p_{t_{fin}, X}(x|Y_{1:9})$  and compare the prediction with ground truth obtained with numerical integration method by Chang and Cooper [17], who use a fine discrete grid. When implementing their method, we construct regular fine grid on the segment  $[-5, 5]$  with 2000 points and numerically solve the SDE with timestep  $dt = 10^{-3}$ . At observation times  $t_k$ ,  $k \in 1, \dots, 9$  we multiply the obtained probability density function  $p_{t_k, X}(x|Y_{1:k-1})$  by the density of the normal distribution  $p(Y_k|X_{t_k} = x)$  estimated at the grid which results in unnormalized  $p_{t_k, X}(x|Y_{1:k})$ . After normalization on the grid,  $p_{t_k, X}(x|Y_{1:k})$  can be used in the new diffusion round on time interval  $[t_k, t_{k+1}]$ . At final time  $t_{fin}$  we estimate SymKL between the true distribution and ones obtained via other competitive methods by numerically integrating (18) on the grid.

As the baselines, we use [Dual JKO] [23] as well as the Bayesian Bootstrap filter [BBF] [26], which combines particle simulation with bootstrap resampling at observation times. We implement [BBF] following the original article [26]. Particle propagation performed via Euler-Maruyama method with timestep  $dt = 10^{-3}$ . The final distribution  $p(X_{t_{fin}}|Y_{1:9})$  is estimated using kernel density estimator. For [Dual JKO] we use the code provided by the authors with the default

hyper-parameters.

In our method, we use JKO step size  $h = 0.1$  and model it by ICNN with width  $w = 256$ . Each JKO step takes 700 optimization iterations with  $lr = 5 \cdot 10^{-3}$  and batch size  $N = 1024$ . At observation times  $t_k$ ,  $k \in 1, 2, \dots, 9$  we use the Metropolis-Hastings algorithm 2 with acceptance probability  $\alpha$  calculated by (17). Starting from the randomly sampled  $x^{(1)}$  we skip the first 1000 values of the Markov Chain generated by the algorithm which allows the series to converge to the distribution of interest  $p_{t_k, X}(x|Y_{1:k})$ . We take each second element from the chain in order to decorrelate the samples. To simultaneously sample the batch of size  $N$ , we run  $N$  chains in parallel. To compute SymKL, we normalize the resulting distribution  $p(X_{t_{\text{fin}}}|Y_{1:9})$  on the Chang-Cooper support grid.

We repeat the experiment 15 times. In Figure 4a, we report the SymKL between predicted density and true  $p(X_{t_{\text{fin}}}|Y_{1:9})$ . We visually compare the fitted and true conditional distributions in Figure 4b.

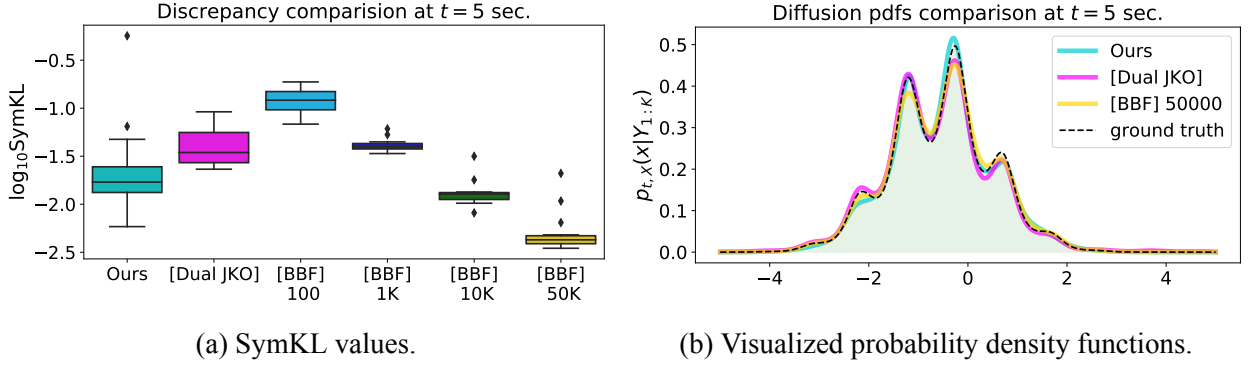


Figure 4: Comparison of the predicted conditional density and true  $p(X_{t_{\text{fin}}}|Y_{1:9})$ .

## 5 Discussion

### 5.1 Further applications

We apply our method to common Bayesian tasks such as unnormalized posterior sampling (§4.4) and nonlinear filtering (§4.5). Below we mention several other potential applications:

- **Population dynamics.** In this task, one needs to recover the potential energy  $\Phi(x)$  included in the Fokker-Planck free energy functional  $\mathcal{F}_{\text{FP}}$  based on samples from the diffusion obtained at timesteps  $t_1, \dots, t_n$ , see [28]. This setting can be found in computational biology, see §6.3 of [28]. A recent paper [13] utilizes ICNN-powered JKO to model population dynamics and successfully models single-cell RNA sequencing data.
- **Reinforcement learning.** Wasserstein gradient flows provide a theoretically-grounded way to optimize an agent policy in reinforcement learning, see [55, 71]. The idea of the method

is to maximize the expected total reward (see (10) in [71]) using the gradient flow associated with the Fokker-Planck functional (see (12) in [71]). The authors of the original paper proposed discrete particle approximation method to solve the underlying JKO scheme. Substituting their approach with our ICNN-based JKO can potentially improve the results.

- **Refining Generative Adversarial Networks.** In the GAN setting, given trained generator  $G$  and discriminator  $D$ , one can improve the samples from  $G$  by  $D$  via considering a gradient flow w.r.t. entropy-regularized  $f$ -divergence between real and generated data distribution (see [7], in particular, formula (4) for reference). Using KL-divergence makes the gradient flow consistent with our method: the functional  $\mathcal{F}$  defining the flow has only entropic and potential energy terms. The usage of our method instead of particle simulation may improve the generator model.
- **Molecular Discovery.** In [3], in parallel to our work the JKO-ICNN scheme is proposed. The authors consider the molecular discovery as an application. The task is to increase the *drug-likeness* of a given distribution  $\rho$  of molecules while staying close to the original distribution  $\rho_0$ . The task reduces to optimizing the functional  $\mathcal{F}(\rho) = \mathbb{E}_{x \sim \rho} \Phi(x) + \mathcal{D}(\rho, \rho_0)$  for a certain potential  $\Phi$  ( $V$  - in the notation of [3]) and a discrepancy  $\mathcal{D}$ . The authors applied the JKO-ICNN method to minimize  $\mathcal{F}$  on MOSES [53] molecular dataset and obtained promising results.

## 5.2 Complexity of training and sampling.

Let  $T$  be the number of operations required to evaluate ICNN  $\psi_\theta(x)$ , and assume that the evaluation of  $\Phi(x)$  in the potential energy  $\mathcal{U}$  takes  $O(1)$  time.

Recall that computing the gradient is a small constant factor harder than computing the function itself [40]. Thus, evaluation of  $\nabla \psi_\theta(x) : \mathbb{R}^D \rightarrow \mathbb{R}^D$  requires  $O(T)$  operations and evaluating the Hessian  $\nabla^2 \psi_\theta(x) : \mathbb{R}^D \rightarrow \mathbb{R}^{D \times D}$  takes  $O(DT)$  time. To compute  $\log \det \nabla^2 \psi_\theta(x)$ , we need  $O(D^3)$  extra operations. Sampling from  $\rho^{(k-1)} = \nabla \psi^{(k-1)} \circ \dots \circ \nabla \psi^{(1)} \# \rho_0$  involves pushing  $x_0 \sim \rho^0$  forward by a sequence of ICNNs  $\psi^{(\cdot)}$  of length  $k - 1$ , requiring  $O((k - 1)T)$  operations.

| Operation   | Time Complexity      |
|---|----------------------|
| Eval. $\psi_\theta, \nabla \psi_\theta, \nabla^2 \psi_\theta$                     | $T, O(T), O(DT)$     |
| Eval. $\log \det \nabla^2 \psi_\theta$  | $O(DT + D^3)$        |
| Sample $x \sim \rho^{(k)}$  | $O((k-1)T)$          |
| Eval. $\hat{\mathcal{L}}$ on $x \sim \rho^{(k)}$                                  | $O(DT + D^3)$        |
| Eval. $\frac{\partial \hat{\mathcal{L}}}{\partial \theta}$ on $x \sim \rho^{(k)}$ | $O(DT + D^3)$        |
| Sample $x \sim \rho^{(k)}$ and<br>Eval. $\frac{d\rho^{(k)}}{dx}(x)$               | $O((k-1)(TD + D^3))$ |

Table 4: Complexity of operations in our method for computing JKO steps via ICNNs.

The forward pass to evaluate the JKO step objective  $\hat{\mathcal{L}}$  in Algorithm 1 requires  $O(DT + D^3)$  operations, as does the backward pass to compute the gradient  $\frac{\partial \hat{\mathcal{L}}}{\partial \theta}$  w.r.t.  $\theta$ .

The *memory complexity* is more difficult to characterize, since it depends on the autodiff implementation. It does not exceed the time complexity and is linear in the number of JKO steps  $k$ .

**Wall-clock times.** All particle-based methods considered in §4 and [Dual JKO] require from several seconds to several minutes CPU computation time. Our method requires from several minutes to few hours on GPU, the time is explained by the necessity to train a new network at each step.

**Advantages.** Due to using continuous approximation, our method scales well to high dimensions, as we show in §4.2 and §4.3. After training, we can produce infinitely many samples  $x_k \sim \rho^{(k)}$ , together with their trajectories  $x_{k-1}, x_{k-2}, \dots, x_0$  along the gradient flow. Moreover, the densities of samples in the flow  $\frac{d\rho^{(k)}}{dx}(x_k), \frac{d\rho^{(k-1)}}{dx}(x_{k-1}), \dots, \frac{d\rho^{(0)}}{dx}(x_0)$  can be evaluated immediately.

In contrast, particle-based and domain discretization methods do not scale well with the dimension (Figure 3) and provide no density. Interestingly, despite its parametric approximation, [Dual JKO] performs comparably to particle simulation and worse than ours (see additionally [23, Figure 3]).

**Limitations.** To train  $k$  JKO steps, our method requires time proportional to  $k^2$  due to the increased complexity of sampling  $x \sim \rho^{(k)}$ . This may be disadvantageous for training long diffusions. In addition, for very high dimensions  $D$ , exact evaluation of  $\log \det \nabla^2 \psi_\theta(x)$  is time-consuming.

**Future work.** To reduce the computational complexity of sampling from  $\rho^{(k)}$ , at step  $k$  one may regress an invertible network  $H : \mathbb{R}^D \rightarrow \mathbb{R}^D$  [9, 30] to satisfy  $H(x_0) \approx \nabla \psi^{(k)} \circ \dots \circ \nabla \psi^{(1)}(x_0)$  and use  $H \# \rho_0 \rightarrow \rho^{(k)}$  to simplify sampling. An alternative is to use variational inference [11, 54, 70] to approximate  $\rho^{(k)}$ . These ideas have already been partially exploited by [22]. To mitigate the computational complexity of computing  $\log \det \nabla \psi_\theta(x)$ , fast approximation can be used [65, 27, 3]. More broadly, developing ICNNs with easily-computable exact Hessians is a critical avenue for further research as ICNNs continue to gain attention in machine learning [43, 36, 37, 29, 21, 5].

## 6 Conclusions

The novel approach for modelling Fokker-Planck equation is proposed. It is based on JKO scheme combined with recently developed Input Convex Neural Networks which provide a rich set of convex functions amenable to optimization procedures. The experiments shows that the methods works competitively in several model and real-world tasks including Bayesian logistic regression

and nonlinear filtering. The advantages and limitations of the approach are analyzed and possible future directions are proposed.

## **7 Author Contribution**

The author implemented the code covering the research (the implementation of ICNN was taken from [36]) as well as contributed to the development of theoretical and practical methods underlining research methodology and applications. In particular, it was the author who proposed Nonlinear Filtering problem reformulation via ICNN powered JKO (see §3.4). Additionally, the author helped a lot with preparing the text of [47] which forms the basis of the work under consideration.

## **8 Publications**

The work is primarily based on the paper [47], written by the author of the very work with co-authors.

## **9 Acknowledgements**

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# I Nonlinear Filtering Algorithm

The Algorithm 3 represents a modified version of 1 with embedded MCMC sampling technique based on Metropolis-Hastings algorithm. Note, that in order to decorrelate the MCMC chain we introduce decorrelation parameter  $K_d$  and take each  $K_d$ -th sample. Also we warm-up the chain omitting the first  $N_w$  samples. See §3.4 for reference.

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## Algorithm 3: Nonlinear filtering with ICNN JKO

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**Input** : Initial diffusion distribution  $p_{0,X}$ , observations  $Y_{1:K}$ , observation times  $t_{1:K}$ , diffusion final time  $t_f$ , JKO discretization step  $h$ , target potential  $\Phi(x)$ , diffusion process temperature  $\beta^{-1}$ , ICNN model  $\psi_\theta$ , batch size  $N$ , MCMC decorrelation parameter  $K_d$ , MCMC warm-up parameter  $N_w$

**Output** :  $[B_1, B_2, \dots, B_{K+1}]$  : sequence of ICNN blocks representing gradient flow between observation time moments  $[0, t_1), [t_1, t_2), \dots [t_{K-1}, t_K)$  and  $[t_K, t_f)$

$t_c \leftarrow 0$  ;  $t_{K+1} \leftarrow t_f$ ;

Sample initial batch  $X \sim p_{0,X}$ ;

$i_{\text{model}} \leftarrow 1$  ;  $k \leftarrow 1$  ;  $n_{\text{omit}} \leftarrow N_w$ ;

**while**  $t_c < t_f$  **do**

**for**  $i = 1, 2, \dots$  **do**

**if**  $t_c < t_1$  **then**

            Sample batch  $X \sim p_{0,X}$

**else**

**for**  $i_{\text{batch}} = 1, 2, \dots, n_{\text{omit}}$  **do**

**for**  $i_{\text{samp}} = 1, 2, \dots, N$  **do**

$x_{k-1} \leftarrow X[i_{\text{samp}}]$ ;

                    Sample  $y_0 \sim \rho_{\mathcal{N}(0,1)}$ ;

$y_{k-1} \leftarrow B_{k-1} \circ \dots \circ B_1(y_0)$ ;      /\* store  $y_1, \dots, y_{k-2}$  with  $y_{k-1}$  \*/

$\alpha \leftarrow \min \left( 1, \frac{\rho_{\mathcal{N}(0,1)}(x_0)p_{0,X}(y_0)}{\rho_{\mathcal{N}(0,1)}(y_0)p_{0,X}(x_0)} \prod_{\tau=1}^{k-1} \frac{p_{\mathcal{N}(Y_\tau,1)}(y_\tau)}{p_{\mathcal{N}(Y_\tau,1)}(x_\tau)} \right)$ ;

                    With probability  $\alpha$  set  $X[i_{\text{samp}}] \leftarrow y_{k-1}$ ;

$n_{\text{omit}} \leftarrow K_d$

$\widehat{\mathcal{W}}_2^2 \leftarrow \frac{1}{N} \sum_{x \in X} \|\nabla \psi_\theta(x) - x\|_2^2$ ;

$\widehat{\mathcal{U}} \leftarrow \frac{1}{N} \sum_{x \in X} \Phi(\nabla \psi_\theta(x))$ ;

$\widehat{\Delta \mathcal{E}} \leftarrow -\frac{1}{N} \sum_{x \in X} \log \det \nabla^2 \psi_\theta(x)$ ;

$\mathcal{L} \leftarrow \frac{1}{2h} \widehat{\mathcal{W}}_2^2 + \widehat{\mathcal{U}} + \beta^{-1} \widehat{\Delta \mathcal{E}}$  ; Perform a gradient step over  $\theta$  by using  $\frac{\partial \mathcal{L}}{\partial \theta}$ ;

$\psi_{i_{\text{model}}}^{(k)} \leftarrow \psi_\phi$ ;

$i_{\text{model}} \leftarrow i_{\text{model}} + 1$ ;

$t_c \leftarrow t_c + h$ ;

**if**  $t_c = t_k$  **then**

$B_k \leftarrow \nabla \psi_{i_{\text{model}}}^{(k)} \circ \dots \circ \nabla \psi_1^{(k)}$ ;

$i_{\text{model}} \leftarrow 1$  ;  $n_{\text{omit}} \leftarrow N_w$  ;  $k \leftarrow k + 1$ ;

---

## II Additional Experiments

In Figure 5, we compare the true distribution  $\rho_t$  with the predicted distribution via the competitive methods when modelling Ornstein-Uhlenbeck processes (§4.3). The comparison is given for time  $t = 0.1, 0.2, \dots, 1.0$ .

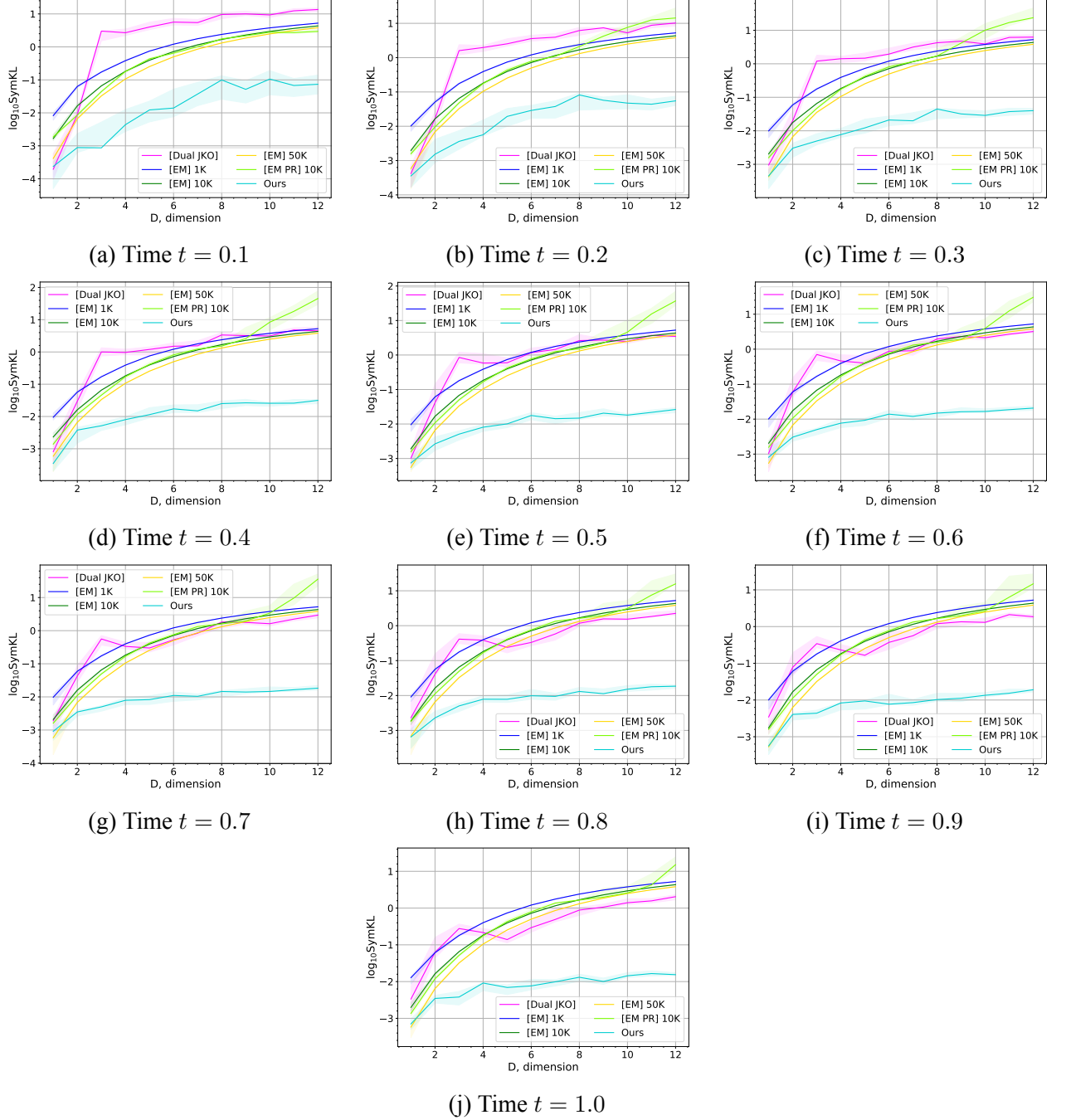


Figure 5: SymKL values between the computed measures and the true measure at  $t = 0.1, 0.2, \dots, 1$  in dimensions  $D = 1, 2, \dots, 12$ . Best viewed in color.