#### A CASE STUDY ON CLASS JOINING PATTERN OF STUDENTS DURING COVID

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#### INTRODUCTION:

2020 started as was expected. We were back into the loop of alarm – running to college – classes – exhausted – dozing off – alarm. The *CORONA VIRUS* had started to spread around the world and soon was declared a pandemic by **WHO** in March 2020.

On March 14, our college declared a two week holiday owing to the ongoing *CORONA VIRUS* pandemic. We were quite excited to have a break to the loop of the monotonous day and also we would get some time for our upcoming end semester examination.

This two week holiday has now been extended to almost a year. Physical classes had to be replaced by online classes.

We have collected data on the number of students who have joined 5 minutes before the start of the first class of the day and the first class of the second half (i.e. post lunch break) for the three years of our department (STATISTICS) and the third years of two other departments, viz. MICROBIOLOGY and COMPUTER SCIENCE for the third week of February, 2021.

In our article, we have worked on the following:

- Comparison between daily average number of students joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the 1<sup>st</sup> class of second half (i.e. post lunch break)
  - O Students of 3<sup>rd</sup> year of the Statistics department
  - Students of the Statistics department (all the three years taken together)
  - Students of 3<sup>rd</sup> year of the three departments considered in our experiment (viz. Statistics, Microbiology and Computer Science)
- Test for homogeneity over the three years of Statistics department
  - o Joining time for the 1<sup>st</sup> class of the day
  - O Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)
- Test for homogeneity over three departments (viz. Statistics, Microbiology and Computer Science)
  - o Joining time for the 1st class of the day
  - o Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)
- Test for independence of year and joining time in online class
  - o Joining time for the 1st class of the day
  - o Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)
- Test for independence of department and joining time in online class
  - o Joining time for the 1st class of the day
  - o Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

#### **TESTING PROBLEMS:**

1.1 Comparison between daily average number of students of 3<sup>rd</sup> year of the Statistics department joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the 1<sup>st</sup> class of second half (i.e. post lunch break):

Let X be the random variable denoting the number of students who have joined 5 minutes before the 1<sup>st</sup> class of the day and let Y be that for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break).

Let us assume that  $X \sim \text{Poisson}(\lambda_1)$  independently of  $Y \sim \text{Poisson}(\lambda_2)$ .

To test,  $H_0: \lambda_1 = \lambda_2$  against  $H_1: \lambda_1 \neq \lambda_2$ .

Let random samples of sizes  $n_1$  and  $n_2$  be drawn respectively from the distributions of X and Y, independent of each other.

We have,

 $X_i \sim \text{Poisson}(\lambda_1)$ , i = 1 (1)  $n_1$ , independently of

 $Y_i \sim \text{Poisson}(\lambda_2), i = 1 (1) n_2.$ 

Define, 
$$\overline{X} = \frac{1}{n_1} \sum_{i=1}^{n_1} X_i$$
 and  $\overline{Y} = \frac{1}{n_2} \sum_{i=1}^{n_2} Y_i$ .

By Variance Stabilizing Transformation (Square Root Transformation of Poisson mean), we have,

$$\sqrt{n_1} \left\{ \sqrt{\overline{X}} - \sqrt{\lambda_1} \right\} \stackrel{d}{\to} N(0, \frac{1}{4})$$
, independently of

$$\sqrt{n_2} \left\{ \sqrt{\overline{Y}} - \sqrt{\lambda_2} \right\} \stackrel{d}{\to} N\left(0, \frac{1}{4}\right)$$
, for moderately large  $n_1, n_2$ .

<u>Test statistic</u>: Define  $T = \sqrt{\overline{X}} - \sqrt{\overline{Y}}$ 

$$E(T) = \sqrt{\lambda_1} - \sqrt{\lambda_2}$$

$$V(T) = \frac{1}{4n_1} + \frac{1}{4n_2}$$

$$\therefore \frac{(\sqrt{\bar{X}} - \sqrt{\bar{Y}}) - (\sqrt{\lambda_1} - \sqrt{\lambda_2})}{\sqrt{\frac{1}{4n_1} + \frac{1}{4n_2}}} \xrightarrow{d} N(0,1), \text{ for moderately large } n_1, n_2.$$

Under H<sub>0</sub>, 
$$Z = \frac{(\sqrt{\bar{X}} - \sqrt{\bar{Y}})}{\sqrt{\frac{1}{4n_1} + \frac{1}{4n_2}}} \sim N(0,1)$$
 asymptotically for moderately large  $n_1, n_2$ .

<u>Test rule</u>: Reject H<sub>0</sub> iff  $|Z_{obs}| > \tau_{\frac{\alpha}{2}}$ , where  $\alpha$  is the level of significance of the test.

#### Computation:

$$n_1=6, n_2=6, \bar{X}=25.67, \bar{Y}=20.67, \alpha=0.05, au_{rac{lpha}{2}}=1.95996$$

$$Z_{obs} = 1.8019$$

Clearly, 
$$|Z_{obs}| < \tau_{\frac{\alpha}{2}}$$

Therefore, we accept H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that the daily average number of students of 3<sup>rd</sup> year of the Statistics department joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the 1<sup>st</sup> class of second half (i.e. post lunch break) are equal.

# 1.2 Comparison between daily average number of students of the Statistics department (all the three years taken together) joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the first class of second half (i.e. post lunch break):

Let X be the random variable denoting the number of students who have joined 5 minutes before the 1<sup>st</sup> class of the day and let Y be that for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break).

Let us assume that  $X \sim \text{Poisson}(\lambda_1)$  independently of  $Y \sim \text{Poisson}(\lambda_2)$ .

To test, 
$$H_0: \lambda_1 = \lambda_2$$
 against  $H_1: \lambda_1 \neq \lambda_2$ .

Proceeding similarly as in 1.1, we arrive at the following conclusions.

#### **Computation**:

$$n_1=6, n_2=6, \bar{X}=79.50, \bar{Y}=67.67, \alpha=0.05, au_{\frac{\alpha}{2}}=1.95996$$

$$Z_{obs} = 2.3913$$

Clearly, 
$$|Z_{obs}| > \tau_{\frac{\alpha}{2}}$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the daily average number of students of the Statistics department (all the three years taken together) joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the first class of second half (i.e. post lunch break) are equal.

1.3 Comparison between daily average number of students of 3<sup>rd</sup> year of the three departments considered in our experiment (viz. Statistics, Microbiology and Computer Science) joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the 1<sup>st</sup> class of second half (i.e. post lunch break):

Let X be the random variable denoting the number of students who have joined 5 minutes before the 1<sup>st</sup> class of the day and let Y be that for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break).

Let us assume that  $X \sim \text{Poisson}(\lambda_1)$  independently of  $Y \sim \text{Poisson}(\lambda_2)$ .

To test,  $H_0: \lambda_1 = \lambda_2$  against  $H_1: \lambda_1 \neq \lambda_2$ .

Proceeding similarly as in 1.1, we arrive at the following conclusions.

#### **Computation**:

$$n_1 = 6, n_2 = 6, \bar{X} = 75.17, \bar{Y} = 64.67, \alpha = 0.05, \tau_{\frac{\alpha}{2}} = 1.95996$$

$$Z_{obs} = 2.1765$$

Clearly,  $|Z_{obs}| > \tau_{\frac{\alpha}{2}}$ 

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the daily average number of students of 3<sup>rd</sup> year of the three departments considered in our experiment (viz. Statistics, Microbiology and Computer Science) joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the 1<sup>st</sup> class of second half (i.e. post lunch break) are equal.

## 2.1 Test for homogeneity over the three years of Statistics department:

#### 2.1.1 Joining time for the 1<sup>st</sup> class of the day

Consider the three years of Statistics department as 3 independent populations, each classified into 2 classes based on joining time.

Let  $p_{ij}$  denote the population proportion of members in the  $i^{th}$  class of the  $j^{th}$  population, i = 1, 2, j = 1, 2, 3.

To test,  $H_0: p_{i1} = p_{i2} = p_{i3}$ , i = 1,2 against  $H_1:$  not  $H_0$ 

Let a random sample of size  $n_i$  be drawn from the  $j^{th}$  population, j = 1,2,3.

Let  $f_{ij}$  denote the number of members in the  $i^{th}$  class of the  $j^{th}$  population, i = 1, 2, j = 1, 2, 3.

$$\therefore \sum_{i=1}^2 f_{ij} = n_j.$$

Test statistic: Define, 
$$\chi^2 = \sum_{i=1}^2 \sum_{j=1}^3 \frac{(f_{ij} - n_j p_{ij})^2}{n_j p_{ij}}$$

Under H<sub>0</sub>,  $\chi^2 = \sum_{i=1}^2 \sum_{j=1}^3 \frac{(f_{ij} - n_j p_{io})^2}{n_j p_{io}}$ , where  $p_{io}$  is the common value of  $p_{i1}, p_{i2}, p_{i3}$  under H<sub>0</sub>, i = 1, 2.

We estimate  $p_{i0}$  as,  $\widehat{p_{io}} = \frac{\sum_{j=1}^{3} f_{ij}}{\sum_{j=1}^{3} n_j}$ .

$$\therefore \chi^2 = \sum_{i=1}^2 \sum_{j=1}^3 \frac{(f_{ij} - n_j \widehat{p_{lo}})^2}{n_j \widehat{p_{lo}}} \sim \chi^2_{(3-1)(2-1)}.$$

<u>Test rule</u>: Reject H<sub>0</sub> iff  $\chi_{obs}^2 > \chi_{\alpha;(3-1)(2-1)}^2$ , where  $\alpha$  is the level of significance of the test.

#### **Computation**:

Joining time \ Year	1 <sup>st</sup>	2 <sup>nd</sup>	3 <sup>rd</sup>	Total
Joined 5 mins before	135	188	154	477
Not joined 5 mins before	195	124	206	525
Total	330	312	360	1002

$$\alpha = 0.05, \chi^2_{\alpha;(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 29.3166$$

Clearly, 
$$\chi_{obs}^2 > \chi_{\alpha;(3-1)(2-1)}^2$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the three years of Statistics department are homogeneous with respect to their joining time for the 1<sup>st</sup> class of the day.

# 2.1.2 Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Proceeding similarly as in **2.1.1**, we arrive at the following conclusions.

#### Computation:

Joining time \ Year	1 <sup>st</sup>	2 <sup>nd</sup>	3 <sup>rd</sup>	Total
Joined 5 mins before	98	184	124	406
Not joined 5 mins before	232	128	236	596

Total	330	312	360	1002
1 Otal	330	314	300	1002

$$\alpha = 0.05, \chi^2_{\alpha;(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 65.6395$$

Clearly, 
$$\chi_{obs}^2 > \chi_{\alpha:(3-1)(2-1)}^2$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the three years of Statistics department are homogeneous with respect to their joining time for the 1<sup>st</sup> class of the 2<sup>nd</sup> half (i.e. post lunch break).

Based on the data that we have collected it seems that there is not enough evidence to say that the three years of Statistics department are homogeneous with respect to their joining time in class.

# 2.2 Test for homogeneity over three departments (viz. Statistics, Microbiology and Computer Science) considered in our experiment:

#### 2.2.1 Joining time for the 1st class of the day

Consider the three departments as 3 independent populations, each classified into 2 classes based on joining time.

Let  $p_{ij}$  denote the population proportion of members in the  $i^{th}$  class of the  $j^{th}$  population, i = 1,2, j = 1,2,3.

To test,  $H_0: p_{i1} = p_{i2} = p_{i3}$ , i = 1,2 against  $H_1:$  not  $H_0$ 

Proceeding similarly as 2.1.1, we arrive at the following conclusions.

#### **Computation**:

Joining time \ Dept.	STSA	MCBA	CMSA	Total
Joined 5 mins before	154	97	200	451
Not joined 5 mins before	206	185	208	599
Total	360	282	408	1050

$$\alpha = 0.05, \chi^2_{\alpha;(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 14.5574$$

Clearly, 
$$\chi^2_{obs} > \chi^2_{\alpha;(3-1)(2-1)}$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the three departments considered in our experiment are homogeneous with respect to their joining time for the 1<sup>st</sup> class of the day.

# 2.2.2 Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Proceeding similarly as **2.2.1**, we arrive at the following conclusions.

#### **Computation**:

Joining time \ Dept.	STSA	MCBA	CMSA	Total
Joined 5 mins before	124	75	189	388
Not joined 5 mins before	236	207	219	662
Total	360	282	408	1050

$$\alpha = 0.05, \chi^2_{\alpha:(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 29.3349$$

Clearly, 
$$\chi_{obs}^2 > \chi_{\alpha;(3-1)(2-1)}^2$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the three departments considered in our experiment are homogeneous with respect to their joining time for the 1<sup>st</sup> class of the 2<sup>nd</sup> half (i.e. post lunch break).

Based on the data that we have collected it seems that there is not enough evidence to say that the three departments considered in our experiment are homogeneous with respect to their joining time in class.

#### 3.1 Test for independence of year and joining time in online class:

#### 3.1.1 Joining time for the 1<sup>st</sup> class of the day

Let the population of the students of Statistics department of the current year (2020 - 2021) be classified according to two attributes based on their year of study (1<sup>st</sup>, 2<sup>nd</sup> and 3<sup>rd</sup> years) (attribute A, say) and based on their joining time (attribute B, say).

Let A<sub>1</sub>, A<sub>2</sub> and A<sub>3</sub> denote classes corresponding to attribute A and let B<sub>1</sub> and B<sub>2</sub> be those corresponding to attribute B.

Let  $p_{ij}$  denote the population proportion of members who belong to the  $i^{th}$  class of A and the  $j^{th}$  class of B, i = 1,2,3, j = 1,2.

Define,  $p_{io} = \sum_{j=1}^{2} p_{ij}$ : proportion of members in the population who belong to the  $i^{th}$  class of A, i = 1,2,3, and,

 $p_{oj} = \sum_{i=1}^{3} p_{ij}$ : proportion of members in the population who belong to the  $j^{th}$  class of B, j = 1,2.

To test,  $H_0: p_{ij} = p_{io} * p_{oj}$ ,  $\forall (i,j)$  against  $H_1:$  not  $H_0$ 

Let a random sample of size n be drawn from the population. Let  $f_{ij}$  denote the number of members in the sample who belong to the  $i^{th}$  class of A and the  $j^{th}$  class of B, i = 1,2,3, j = 1,2.

Define, 
$$f_{io} = \sum_{j=1}^{2} f_{ij}$$
, and,  $f_{oj} = \sum_{i=1}^{3} f_{ij}$ .

Test statistic: Define, 
$$\chi^2 = \sum_{i=1}^3 \sum_{j=1}^2 \frac{(f_{ij} - n p_{ij})^2}{n p_{ij}}$$

Under H<sub>0</sub>, 
$$\chi^2 = \sum_{i=1}^3 \sum_{j=1}^2 \frac{(f_{ij} - n \, p_{io} p_{oj})^2}{n \, p_{io} p_{oj}}$$

We estimate  $p_{io}$  and  $p_{oj}$  as,

$$\widehat{p_{io}} = \frac{f_{io}}{n}$$
, i = 1,2,3, and,  $\widehat{p_{oJ}} = \frac{f_{oj}}{n}$ , j=1,2.

$$\therefore \chi^{2} = \sum_{i=1}^{3} \sum_{j=1}^{2} \frac{(f_{ij} - n \, \widehat{p_{io}} \widehat{p_{oj}})^{2}}{n \, \widehat{p_{io}} \widehat{p_{oj}}} \sim \chi^{2}_{(3-1)(2-1)}.$$

<u>Test rule</u>: Reject H<sub>0</sub> iff  $\chi^2_{obs} > \chi^2_{\alpha;(3-1)(2-1)}$ , where  $\alpha$  is the level of significance of the test.

#### Computation:

Year \ Joining time	Joined 5 mins before	Not joined 5 mins before	Total
1 <sup>st</sup>	135	195	330
2 <sup>nd</sup>	188	124	312
3 <sup>rd</sup>	154	206	360
Total	477	525	1002

$$\alpha = 0.05, \chi^2_{\alpha:(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 29.3163$$

Clearly, 
$$\chi_{obs}^2 > \chi_{\alpha;(3-1)(2-1)}^2$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the attributes 'year of study (Statistics department)' and 'joining time for the 1<sup>st</sup> class of the day' are independent.

# 3.1.2 Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Proceeding similarly as in 3.1.1, we arrive at the following conclusions.

#### **Computation**:

Dept. \ Joining time	Joined 5 mins before	Not joined 5 mins before	Total
1 <sup>st</sup>	98	232	330
2 <sup>nd</sup>	184	128	312
3 <sup>rd</sup>	124	236	360
Total	406	596	1002

$$\alpha = 0.05, \chi^2_{\alpha;(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 65.6401$$

Clearly, 
$$\chi^2_{obs} > \chi^2_{\alpha;(3-1)(2-1)}$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the attributes 'Year of Study (Statistics department)' and 'joining time for the 1<sup>st</sup> class of the 2<sup>nd</sup> half (i.e. post lunch break)' are independent.

Thus, based on the data that we have collected it seems that there is not enough evidence to say that the attributes 'year of study (Statistics department)' and 'joining time in class' are independent.

#### 3.2 Test for independence of department and joining time in online class:

## 3.2.1 Joining time for the 1<sup>st</sup> class of the day

Let the population of the students of the three departments considered in our experiment, of the current year (2020 - 2021) be classified according to two attributes based on their department (viz. Statistics, Microbiology and Computer Science) (attribute A, say) and on their joining time (attribute B, say).

Let A<sub>1</sub>, A<sub>2</sub> and A<sub>3</sub> denote classes corresponding to attribute A and let B<sub>1</sub> and B<sub>2</sub> be those corresponding to attribute B.

Let  $p_{ij}$  denote the population proportion of members who belong to the  $i^{th}$  class of A and the  $j^{th}$  class of B, i = 1,2,3, j = 1,2.

Define,  $p_{io} = \sum_{j=1}^{2} p_{ij}$ : proportion of members in the population who belong to the  $i^{th}$  class of A, i = 1,2,3, and,

 $p_{oj} = \sum_{i=1}^{3} p_{ij}$ : proportion of members in the population who belong to the  $j^{th}$  class of B, j = 1,2.

To test,  $H_0: p_{ij} = p_{io} * p_{oj}$ ,  $\forall (i,j)$  against  $H_1:$  not  $H_0$ 

Proceeding similarly as in 3.1.1, we arrive at the following conclusions.

#### **Computation:**

Dept. \ Joining time	Joined 5 mins before	Not joined 5 mins before	Total
STSA	154	206	360
MCBA	97	185	282
CMSA	200	208	408
Total	451	599	1050

$$\alpha = 0.05, \chi^2_{\alpha:(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 14.5572$$

Clearly, 
$$\chi_{obs}^2 > \chi_{\alpha;(3-1)(2-1)}^2$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the attributes 'department' and 'joining time for the 1st class of the day' are independent.

# 3.2.2 Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Proceeding similarly as in **3.2.1**, we arrive at the following conclusions.

#### **Computation**:

Dept. \ Joining time	Joined 5 mins before	Not joined 5 mins before	Total
STSA	124	236	360
MCBA	75	207	282
CMSA	189	219	408
Total	388	662	1050

$$\alpha = 0.05, \chi^2_{\alpha;(3-1)(2-1)} = 5.99146$$

$$\chi_{obs}^2 = 29.3342$$

Clearly, 
$$\chi_{obs}^2 > \chi_{\alpha;(3-1)(2-1)}^2$$

Therefore, we reject H<sub>0</sub> at 5% level of significance.

Based on the data that we have collected it seems that there is not enough evidence to say that the attributes 'department' and 'joining time for the 1<sup>st</sup> class of the 2<sup>nd</sup> half (i.e. post lunch)' are independent.

Thus, based on the data that we have collected it seems that there is not enough evidence to say that the attributes 'department' and 'joining time in class' are independent.

#### **CONCLUSION:**

From the above tests based on our data we arrive at the following conclusions:

- The daily average number of students joining 5 minutes before the 1<sup>st</sup> class of the day and 5 minutes before the 1<sup>st</sup> class of second half (i.e. post lunch break)
  - Seems to be equal for the students of 3<sup>rd</sup> year of the Statistics department
  - Seems to be unequal for the students of the Statistics department (all the three years taken together)
  - Seems to be unequal for the students of 3<sup>rd</sup> year of the three departments considered in our experiment (viz. Statistics, Microbiology and Computer Science)
- The three years of Statistics department

- Does not seem to be homogeneous with respect to the joining time for the 1<sup>st</sup> class of the day
- Does not seem to be homogeneous with respect to the joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Thus, the three years of Statistics department does not seem to be homogeneous with respect to the joining time in class.

- The three departments (viz. Statistics, Microbiology and Computer Science)
  - Does not seem to be homogeneous with respect to the joining time for the 1<sup>st</sup> class of the day
  - Does not seem to be homogeneous with respect to the joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Thus, the three departments considered here does not seem to be homogeneous with respect to the joining time in class.

- The year of study (Statistics department) seems to be dependent on the
  - o Joining time for the 1st class of the day
  - o Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Thus, the year of study (Statistics department) seems to be dependent on the joining time in class.

- The attribute department seems to be dependent on the
  - o Joining time for the 1st class of the day
  - O Joining time for the 1<sup>st</sup> class of 2<sup>nd</sup> half (i.e. post lunch break)

Thus, the attribute department seems to be dependent on the joining time in class.