MATH 588 HW6

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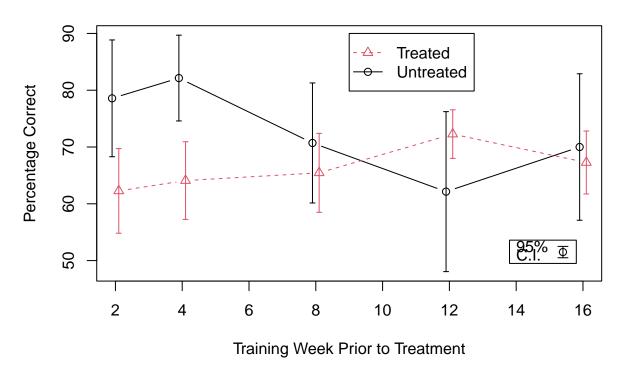
3/27/2022

Question 1

 \mathbf{a}

```
library(Sleuth3)
mem = case1601
names(mem) = casefold(names(mem))
# Group means and SEs as 2 by 5 matrices
means = as.matrix(aggregate(mem[,3:7], list(mem$treatment), mean)[,-1])
sds = as.matrix(aggregate(mem[,3:7], list(mem$treatment), sd)[,-1])
ns = as.matrix(aggregate(mem[,3:7], list(mem$treatment), length)[,-1])
SEs = sds/sqrt(ns-1)
multiplier = matrix(qt(0.975, ns-1), ncol=5, nrow=2)
upper = means + multiplier*SEs
lower = means - multiplier*SEs
# Start with an empty plot that will hold everything
lim = c(min(lower), max(upper))
{plot(c(2,16), lim, type="n", xlab="Training Week Prior to Treatment",
ylab="Percentage Correct",
main="Monkey Memory Study")
# Means and CIs by treatment
x = c(2,4,8,12,16)
delta = 0.1
epsilon = 0.075
for (i in 1:2) {
  xa = x+c(-1,1)[i]*delta
  lines(xa, means[i,], col=i, lty=i, pch=i, type="b")
  segments(xa, lower[i,], xa, upper[i,], col=i)
  segments(rep(xa-epsilon,2), c(lower[i,],upper[i,]),
  rep(xa+epsilon,2), c(lower[i,],upper[i,]), col=i)
}
# Legend
legend(9, 90, c("Treated", "Untreated"), lty=2:1, col=2:1, pch=2:1)
# Optional "CI" box:
text(c(14,14), c(52.5,51), c("95%", "C.I."), adj=0)
segments(c(15.25,15.4,15.25), c(50.5,50.5,52.5), c(15.55,15.4,15.55), c(50.5,52.5,52.5))
points(15.4,51.5)
rect(13.8, 49.5, 15.8, 53.6)}
```

Monkey Memory Study



b

```
CTRL=mem$treatment=="Control"
round(diag(cov(mem[CTRL,3:7])),3)
                    week8 week12 week16
            week4
## 105.952 57.143 111.905 198.810 166.667
round(diag(cov(mem[!CTRL,3:7])),3)
##
    week2
            week4
                    week8 week12 week16
## 111.818 94.091 97.273
                           36.818
                                  61.818
## OR just sds^2 from above
apply(sds^2, 2, function(x) max(x)/min(x))
      week2
              week4
                       week8
                               week12
## 1.055363 1.646591 1.150423 5.399765 2.696078
round(abs(cor(mem[!CTRL,3:7]) - cor(mem[CTRL,3:7])) ,3)
##
         week2 week4 week8 week12 week16
## week2 0.000 0.809 0.227 0.352
                                  0.203
## week4 0.809 0.000 0.260 0.352
                                   0.844
## week8 0.227 0.260 0.000 0.231
                                   0.525
## week12 0.352 0.352 0.231 0.000
## week16 0.203 0.844 0.525 0.325 0.000
```

Week 12 has a highest variance between controls and treated group which is 5.39 times higher for controls. The highest correlation values between pairs week 4 and 16 and it is 0.844. (between control and treated group)

 \mathbf{c}

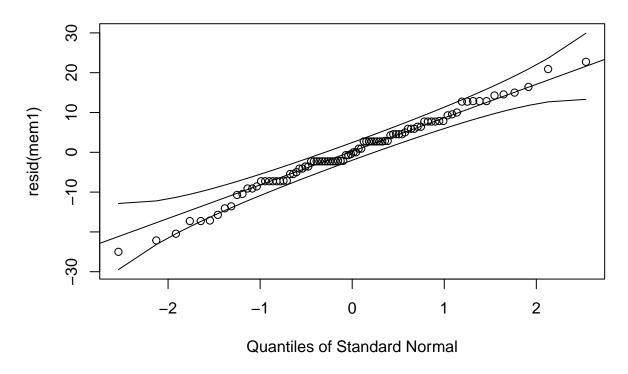
```
mem1 = aov(as.matrix(mem[,3:7])~treatment, data=mem)
anova(mem1, test="Hotelling")
## Analysis of Variance Table
##
              Df Hotelling-Lawley approx F num Df den Df
##
                                                            Pr(>F)
## (Intercept)
                          290.837
                                    698.01
                                                5
                                                      12 2.363e-14 ***
               1
                                      5.59
## treatment
               1
                            2.331
                                                5
                                                      12 0.006878 **
## Residuals
              16
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
anova(aov(cbind(week2,week4,week8,week12,week16)~treatment, data=mem))
## Analysis of Variance Table
##
##
              Df Pillai approx F num Df den Df
                                                   Pr(>F)
                                       5
## (Intercept)
               1 0.99657
                           698.01
                                             12 2.363e-14 ***
                                       5
## treatment
               1 0.69978
                             5.59
                                             12 0.006878 **
## Residuals
              16
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

The p-value is less than 0.05. So, the null hypothesis (means are equal) can be rejected at 5% level of significance.

\mathbf{d}

```
source("http://www.stat.cmu.edu/~hseltman/files/qqn.R")
qqn(resid(mem1), main="Monkey Memory Quantile Normal Plot")
```

Monkey Memory Quantile Normal Plot



The plot suggest, the normality assumption does not violated in this case.

Question 2

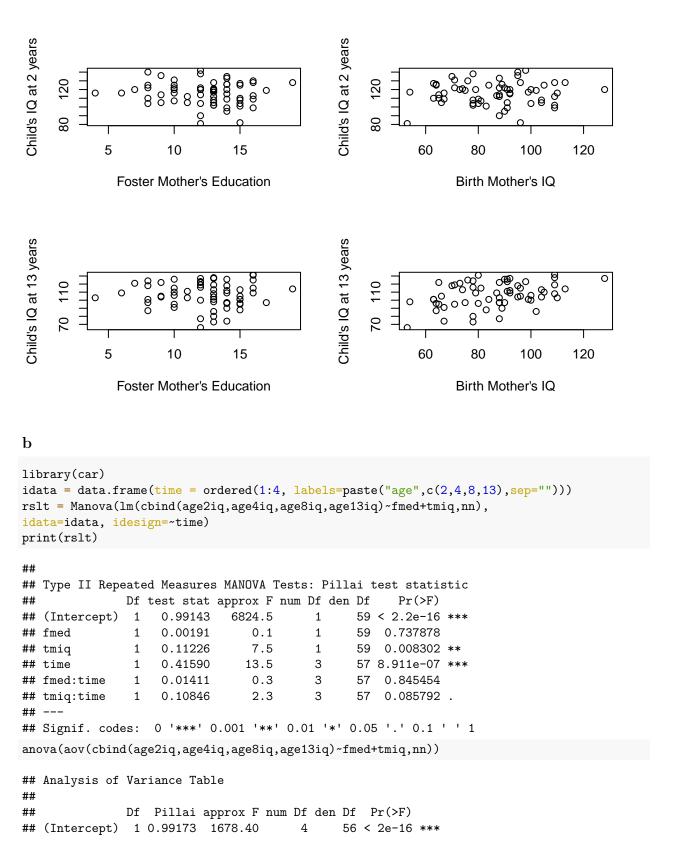
 \mathbf{a}

```
nn = ex1605
names(nn) = casefold(names(nn))
{par(mfrow=c(2,2), oma=c(0,0,1.5,0))}

with(nn, plot(fmed, age2iq, xlab="Foster Mother's Education",
ylab="Child's IQ at 2 years"))
with(nn, plot(tmiq, age2iq, xlab="Birth Mother's IQ",
ylab="Child's IQ at 2 years"))
with(nn, plot(fmed, age13iq, xlab="Foster Mother's Education",
ylab="Child's IQ at 13 years"))

with(nn, plot(tmiq, age13iq, xlab="Birth Mother's IQ",
ylab="Child's IQ at 13 years"))
mtext("Nature / Nurture EDA", outer=T, cex=1.3)}
```

Nature / Nurture EDA



```
## fmed 1 0.01991 0.28 4 56 0.88693
## tmiq 1 0.17759 3.02 4 56 0.02504 *
## Residuals 59
## ---
## Signif. codes: 0 '*** 0.001 '** 0.05 '.' 0.1 ' ' 1
```

In the model we observed that the p-value for variable time "time" is less than 0.05. As a result we can interpret that at 5% level of significance we can reject the null hypothesis (average IQ similar over the time) considering all other factors are constant.

 \mathbf{c}

The null hypothesis that the "IQs do not change with the birth mothers IQ" can be rejected at 5% level of significance because the p-value is less than 0.05. We observed a positive correlation between mother's age and age 13 group, but other plots are not very clear to make a statement.

Question 3

```
library(readr)
fakebeet <- read_csv("E:/NMT MS/Spring 22/MATH 588/Home_Work/Spring-2022---MATH-588-01-Advanced-Data-An
with(fakebeet, t.test(P[species=="one"], P[species=="two"]))
##
##
   Welch Two Sample t-test
##
## data: P[species == "one"] and P[species == "two"]
## t = -0.80636, df = 31.684, p-value = 0.426
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -12.540680
                 5.429569
## sample estimates:
## mean of x mean of y
  181.5000 185.0556
with(fakebeet, t.test(Q[species=="one"], Q[species=="two"]))
##
##
   Welch Two Sample t-test
##
## data: Q[species == "one"] and Q[species == "two"]
## t = 1.1734, df = 33.969, p-value = 0.2488
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -1.992645 7.437089
## sample estimates:
## mean of x mean of y
   129.1667 126.4444
anova(aov(cbind(P,Q)~species, data=fakebeet), test="Hotelling")
## Analysis of Variance Table
##
##
               Df Hotelling-Lawley approx F num Df den Df Pr(>F)
## (Intercept)
                            357.11
                                     5892.3
                                                 2
                                                        33 < 2e-16 ***
                1
                                        3.8
                                                 2
## species
                              0.23
                                                        33 0.03386 *
```

```
## Residuals 34
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

The individual t-test are presenting a p-value of greater than 0.05, which interpreting that the null hypothesis can not be rejected. But the multivariate T square p-value is less than 0.05. So, it's suggest that we can reject the null hypothesis that individual species are equal. That's why multivariate t-test is important to consider.

Question 4

```
library(readr)
swiss1 <- read_table2("E:/NMT MS/Spring 22/MATH 588/Home_Work/Spring-2022---MATH-588-01-Advanced-Data-A
genuine <- swiss1[1:100,2:7]
genuine$Type <- rep("Genuine",100)</pre>
counterfeit <- swiss1[101:200,2:7]</pre>
counterfeit$Type <- rep("Counterfeit",100)</pre>
dat <- rbind(genuine,counterfeit)</pre>
mod <- manova(cbind(Length,Left,Right,Bottom,Top,Diagonal) ~ Type, data=dat)</pre>
summary(mod)
                                       Df Pillai approx F num Df den Df
                                                                                                                                                 Pr(>F)
                                          1 0.92415
                                                                            391.92
                                                                                                              6
                                                                                                                             193 < 2.2e-16 ***
## Type
## Residuals 198
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
We can reject the null hypothesis at 5% level of significance because the p-value is less than 0.05.
ci_ellip_plt <- function(var){</pre>
     library(car)
     #var = "Length"
     x = dat[dat$Type == "Genuine",var]
     y = dat[dat$Type == "Counterfeit",var]
     cid = data.frame(x,y)
     colnames(cid) <- c(paste(var, "Genuine"), paste(var, "Counterfeit"))</pre>
     n=100
     p=2
     Sd=cov(cid)
     dbar=c(colMeans(cid))
     # Plot the 95% confidence ellipse for the vector of mean differences
     library(ellipse)
     plot(ellipse(Sd, centre=dbar, t=sqrt(((n-1)*p/(n*(n-p)))*qf(0.95, p, n-p))), type="l", main = paste("95% CI plot(n-p)) | type="l", main = paste("95% CI 
par(mfrow=c(3,2))
```

```
ci_ellip_plt(var = "Length")
ci_ellip_plt(var = "Left")
ci_ellip_plt(var = "Right")
ci_ellip_plt(var = "Bottom")
ci_ellip_plt(var = "Top")
ci_ellip_plt(var = "Diagonal")
```

