

Special Random Variables

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Plan

- The binomial distribution
- The Poisson distribution
- Hypergeometric distribution

The Binomial Random Variable

The Bernoulli Random Variable

- Suppose that a trial has only two outcomes, which are classified as either a "success" or as a "failure", such trial is known as **Bernoulli trial**
 - Define a random variable X as

$$X = \begin{cases} 1 & \text{if the outcome is a success} \\ 0 & \text{if the outcome is a failure} \end{cases}$$

The Bernoulli Random Variable

- X follows a Bernoulli distribution with parameter p and the probability mass function

$$P(X = 1) = p$$

$$P(X = 0) = 1 - p$$

- The probability mass function

$$P(X = x) = p^x(1 - p)^{1-x}, \quad x = 0, 1; \quad 0 \leq p \leq 1$$

- $E(X) = p$ and $Var(X) = p(1 - p)$

The binomial random variable

- Let p be the probability of success of a coin
- Suppose a coin is rolled two times, what is the probability distribution of X , the number of successes in two trials
- Suppose a coin is rolled three times, what is the probability distribution of X , the number of successes in two trials

The binomial random variable

- Binomial random variable deals with the distribution of the number of successes in n independent Bernoulli trials
- The probability success p remain constant from trial to trial
- Let X_1, \dots, X_n are independent and each follow a Bernoulli distribution with parameter p , and define

$$X = X_1 + \dots + X_n$$

- X represents the number of success in n Bernoulli trials and X follows a binomial distribution with parameters n and p , i.e. $X \sim B(n, p)$

The binomial random variable

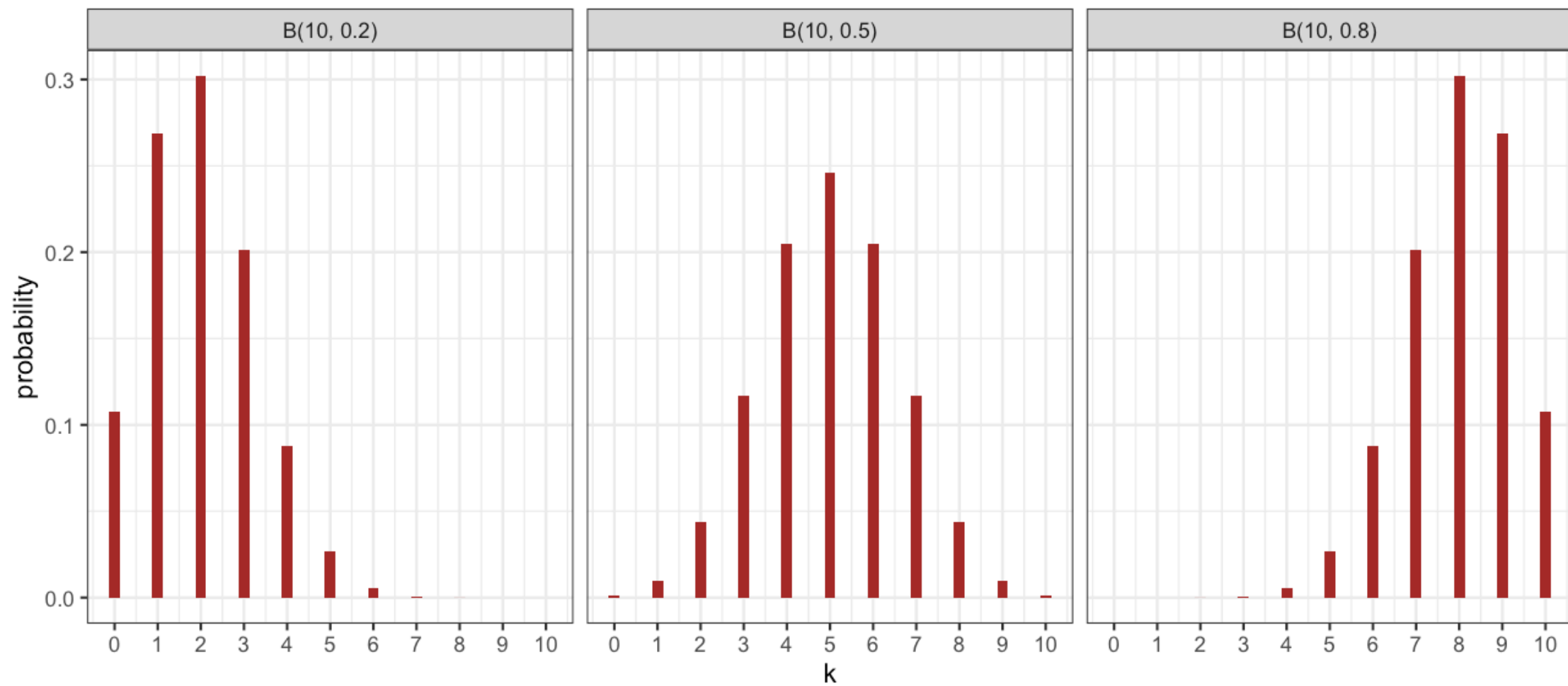
- The probability mass function of X

$$P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}, \quad x = 0, 1, \dots, n$$

- $\binom{n}{x}$ \rightarrow number of ways x success can be obtained in n Bernoulli trials

Example 5.1a

- It is known that disks produced by a certain company will be defective with probability $.01$ independently of each other.
- The company sells the disks in packages of 10 and offers a money-back guarantee that at most 1 of the 10 disks is defective.
- What proportion of packages is returned?
- If someone buys three packages, what is the probability that exactly one of them will be returned?



Expectation of binomial distribution

$$\begin{aligned} E(X) &= \sum_{x=0}^n x \binom{n}{x} p^x (1-p)^{n-x} \\ &= \sum_{x=1}^n x \frac{n!}{x!(n-x)!} p^x (1-p)^{n-x} \\ &= \sum_{x=1}^n \frac{n!}{(x-1)!(n-x)!} p^x (1-p)^{n-x} \\ &= np \sum_{x=1}^n \frac{(n-1)!}{(x-1)!(n-x)!} p^{x-1} (1-p)^{n-x} = np \end{aligned}$$

Variance of binomial distribution

$$\begin{aligned} E[X(X-1)] &= \sum_{x=0}^n x(x-1) \binom{n}{x} p^x (1-p)^{n-x} \\ &= \sum_{x=2}^n x(x-1) \binom{n}{x} p^x (1-p)^{n-x} = n(n-1)p^2 \end{aligned}$$

$$\begin{aligned} Var(X) &= E(X^2) - [E(X)]^2 \\ &= E[X(X-1)] + E(X) - [E(X)]^2 \\ &= n(n-1)p^2 + np - n^2p^2 \\ &= np(1-p) \end{aligned}$$

Binomial cumulative distribution function

- The cumulative distribution function of $X \sim B(n, p)$

$$P(X \leq a) = \sum_{x=a}^n \binom{n}{x} p^x (1-p)^{n-x}$$

- Following relationship is helpful to calculate cumulative distribution of binomial distribution

$$P(X = x + 1) = \frac{p}{1-p} \frac{n-x}{x+1} P(X = x)$$

Problems

- 1, 3, 5, 6, 7, 9

Geometric distribution

- Let p be the probability of success of a coin
- What is the probability distribution of X , the number of trials needed to get the success for the first time?

Geometric distribution

- Let X be the number of trials needed to get the success for the first time and p be the probability of success of the Bernoulli trial
- X follows a geometric distribution with parameter p , $X \sim G(p)$ and the corresponding probability mass function

$$p(x) = (1 - p)^{x-1}p, \quad x = 1, 2, \dots$$

- $E(X) = 1/p$ and $V(X) = (1 - p)/p^2$
- $P(X \leq a) = 1 - (1 - p)^a$

Negative binomial distribution

- Let p be the probability of success of a coin
- What is the probability distribution of X , the number of trials needed to get the success for the r th ($r \geq 1$) time?

Negative binomial distribution

- Let X be the number of trials needed to get the r th ($r \geq 1$) success for the first time and p be the probability of success of the Bernoulli trial
- X follows a negative binomial distribution with parameters r and p , $X \sim NB(r, p)$ and the corresponding probability mass function

$$p(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, \quad x = r, r+1, \dots$$

- $E(X) = r/p$ and $V(X) = r(1-p)/p^2$

Exercise

- An archer hits a bull's-eye with a probability of 0.09, and the results of different attempts are assumed to be independent
- If a archer shoot a series of arrows, what is the probability that the first bull's-eye is scored with the fourth arrow?
- What is the probability that the third bull's-eye is scored with the tenth arrow?
- What is the expected number of arrows shot before the first bull's-eye is scored?
- What is the expected number of arrows shot before the third bull's-eye is scored?

The Poisson Random Variable

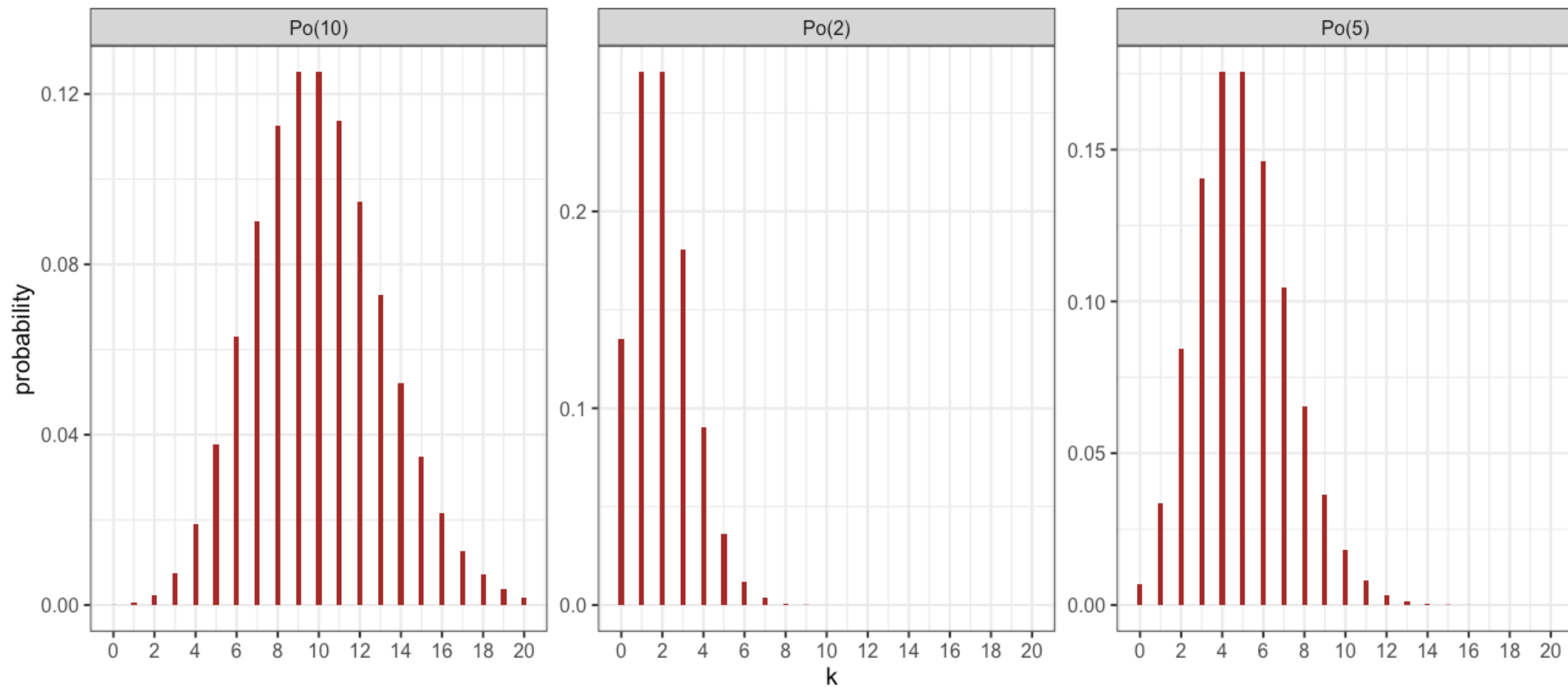
The Poisson Random Variable

- A random variable X is said to be a Poisson random variable with parameter $\lambda > 0$, if its probability mass function is given by

$$P(X = x) = \frac{e^{-\lambda} \lambda^x}{x!}, \quad x = 0, 1, 2, \dots$$

- Using the relationship $e^\lambda = \sum_{x=0}^{\infty} \lambda^x / x!$, it can be shown that $\sum_{x=0}^{\infty} P(X = x) = 1$ for a Poisson distribution
- For $X \sim Po(\lambda)$

$$E(X) = V(X) = \lambda$$



Examples of Poisson random variables

- The number of misprints on a page of a book
- The number of people in a community living to 100 year of age
- The number of wrong telephone numbers that are dialed in a day
- The number of transistors that fail on their first day of use
- The number of customers entering a post office on a given day

Expectation and variance of a Poisson random variable

$$E(X) = \sum_{x=0}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x!} = \sum_{x=1}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x!} = \sum_{x=1}^{\infty} \frac{e^{-\lambda} \lambda^x}{(x-1)!} = \lambda$$

- Similarly

$$E[X(X-1)] = \sum_{x=0}^{\infty} x(x-1) \frac{e^{-\lambda} \lambda^x}{x!} = \lambda^2$$

$$Var(X) = E[X(X-1)] + E(X) - [E(X)]^2 = \lambda$$

The Poisson distribution

- The moment generating function of $X \sim Po(\lambda)$

$$\begin{aligned} M_X(t) &= E[e^{tX}] = \sum_{x=0}^{\infty} e^{tx} e^{-\lambda} \lambda^x / (x!) \\ &= e^{-\lambda} \sum_{x=0}^{\infty} (\lambda e^t)^x / (x!) = e^{-\lambda} e^{\lambda e^t} = e^{-\lambda(1-e^t)} \end{aligned}$$

- Obtain $E(X)$ and $V(X)$

Poisson approximation of binomial distribution

- Suppose $X \sim B(n, p)$, for a large n and small p

$$P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x} \simeq \frac{e^{-\lambda} \lambda^x}{x!}$$

- $\lambda = np$

- **Examples:** $X \sim B(70, .1)$

$$P(X = 5) = \begin{cases} 0.1284 & \text{for binomial} \\ 0.1277 & \text{for Poisson} \end{cases}$$

Example 5.2a

- Suppose that the average number of accidents occurring weekly on a particular stretch of a highway equals 3.
- Calculate the probability that there is at least one accident this week.

Example 5.2b

- Suppose the probability that an item produced by a certain machine will be defective is .1.
- Find the probability that a sample of 10 items will contain at most one defective item.
- Assume that the quality of successive items is independent.

Example 5.2d

- If the average number of claims handled daily by an insurance company is 5, what proportion of days have less than 3 claims?
- What is the probability that there will be 4 claims in exactly 3 of the next 5 days?
- Assume that the number of claims on different days is independent.

Distribution of sum of two independent Poisson random variables

- If $X_1 \sim Po(\lambda_1)$ and $X_2 \sim Po(\lambda_2)$, then

$$Y = X_1 + X_2 \sim Po(\lambda_1 + \lambda_2)$$

- Moment generating function of X_1

$$M_{X_1}(t) = E[e^{tX}] = \exp [\lambda_1(e^t - 1)]$$

- Moment generating function of $Y = X_1 + X_2$

$$M_Y(t) = M_{X_1}(t)M_{X_2}(t) = \exp [(\lambda_1 + \lambda_2)(e^t - 1)]$$

Computation of Poisson distribution function

- For $X \sim Po(\lambda)$

$$\frac{P(X = x + 1)}{P(X = x)} = \frac{\lambda}{x + 1}, x = 1, 2, \dots$$

Example 5.2f

- It has been established that the number of defective stereos produced daily at a certain plant is Poisson distributed with mean 4.
- Over a 2-day span, what is the probability that the number of defective stereos does not exceed 3?

Problems

- 13, 14, 18

Hypergeometric Distribution

Hypergeometric Distribution

- A bin contains $N + M$ batteries, of which N are of acceptable quality and M are defective
- A sample of size n is randomly chosen (without replacement) and all possible sampled subsets are equally likely
- The X denotes the number of acceptable batteries in the sample of size n and its probability mass function

$$P(X = x) = \frac{\binom{N}{x} \binom{M}{n-x}}{\binom{N+M}{n}}, \quad k = 0, 1, \dots, \min(n, N)$$

- X follows a hypergeometric distribution with parameters N , M , and n

Example 5.3a

- The components of a 6-component system are to be randomly chosen from a bin of 20 used components.
- The resulting system will be functional if at least 4 of its 6 components are in working condition.
- If 15 of the 20 components in the bin are in working condition, what is the probability that the resulting system will be functional?
 - X , the number of components with working condition out of six selected components, follows a hypergeometric distribution with parameters $N = 20$, $M = 15$, and $n = 6$

Expectation and variance of hypergeometric distribution

- Let $X = \sum_{i=1}^n X_i$ be the number of acceptable batteries in n selection, where

$$X_i = \begin{cases} 1 & \text{if the } i\text{th selection is acceptable} \\ 0 & \text{otherwise} \end{cases}$$

- $E(X) = E\left(\sum_i X_i\right) = \sum_i E(X_i) = \sum_i P(X_i = 1) = \frac{nN}{N+M}$
- $Var(X) = Var\left(\sum_i X_i\right) = \sum_i Var(X_i) + 2 \sum_{i>j} Cov(X_i, X_j)$

Expectation and variance of Hypergeometric distribution

$$Var(X) = Var\left(\sum_i X_i\right) = \sum_i Var(X_i) + 2 \sum_{i>j} Cov(X_i, X_j)$$

$$Var(X_i) = P(X_i = 1)[1 - P(X_i = 1)] = \frac{N}{N+M} \frac{M}{N+M}$$

$$\begin{aligned} Cov(X_i, X_j) &= E(X_i X_j) - E(X_i)E(X_j) \\ &= P(X_i X_j = 1) - \left(\frac{N}{N+M}\right)^2 \\ &= P(X_i = 1, X_j = 1) - \left(\frac{N}{N+M}\right)^2 \end{aligned}$$

Expectation and variance of Hypergeometric distribution

$$P(X_i = 1, X_j = 1) = P(X_i = 1)P(X_j = 1 \mid X_i = 1) = \frac{N}{N+M} \frac{N-1}{N+M-1}$$

$$\begin{aligned} Cov(X_i, X_j) &= \frac{N(N-1)}{(N+M)(N+M-1)} - \left(\frac{N}{N+M} \right)^2 \\ &= \frac{-NM}{(N+M)^2(N+M-1)} \end{aligned}$$

$$\begin{aligned} Var(X) &= \frac{nNM}{(N+M)^2} - \frac{n(n-1)NM}{(N+M)^2(N+M-1)} \\ &= \frac{nNM}{(N+M)^2} \left(1 - \frac{n-1}{N+M-1} \right) \end{aligned}$$

Hypergeometric Distribution

- Let $X \sim B(n, p)$ and $Y \sim B(m, p)$ then

$$P(X = i \mid X + Y = k) = \frac{\binom{n}{i} \binom{m}{k-i}}{\binom{n+m}{k}}$$

Multinomial distribution

- Consider a sequence of n independent trials where each individual trial can have k outcomes that occur with probability p_1, \dots, p_k , respectively, where $\sum_{i=1}^k p_i = 1$
- Let X_i be the number of i th-type of occurrences, $X_i \sim B(n, p_i)$
- The counts (X_1, \dots, X_k) follows a multinomial distribution with parameters n and p_1, \dots, p_k and their joint probability mass function

$$p(X_1 = x_1, \dots, X_k = x_k) = \frac{n!}{x_1! \cdots x_k!} p_1^{x_1} \cdots p_k^{x_k}$$

- $\sum_{i=1}^k x_i = n$

Multinomial distribution

- Let X_i be the number of i th-type of occurrences, $X_i \sim B(n, p_i)$
- $E(X_i) = np_i$ and $V(X_i) = np_i(1 - p_i)$
- X_i 's are not independent, $Cov(X_i, X_j) \neq 0$

The Uniform Random Variable

The Uniform Random Variable

- A random variable X is said to be uniformly distributed over the interval $[\alpha, \beta]$ if its pdf is given by

$$f(x) = \begin{cases} \frac{1}{\beta - \alpha} & \text{if } \alpha \leq x \leq \beta \\ 0 & \text{otherwise} \end{cases}$$

- Show that

$$E(X) = \frac{\alpha + \beta}{2} \text{ and } Var(X) = \frac{(\beta - \alpha)^2}{12}$$

The Normal Random Variable

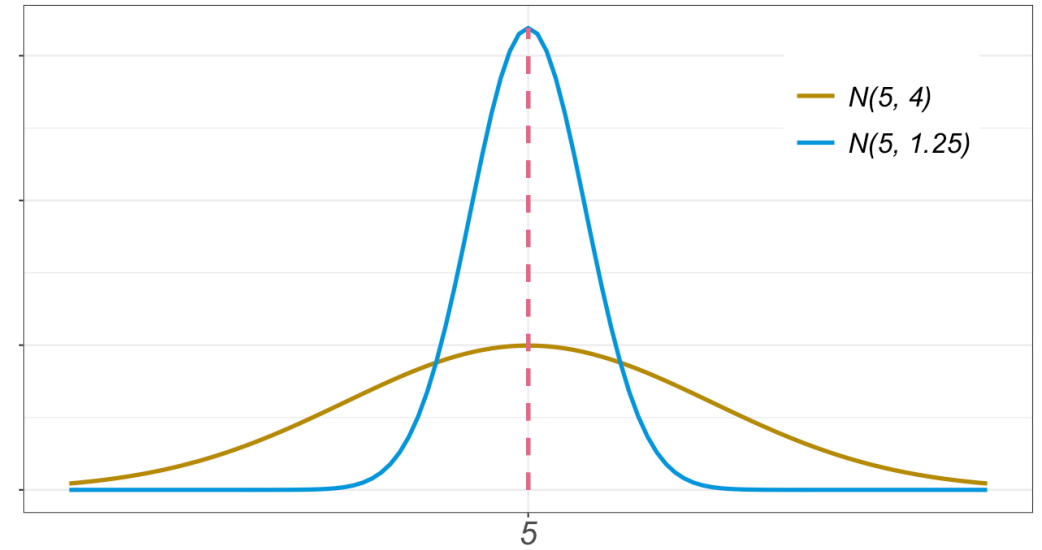
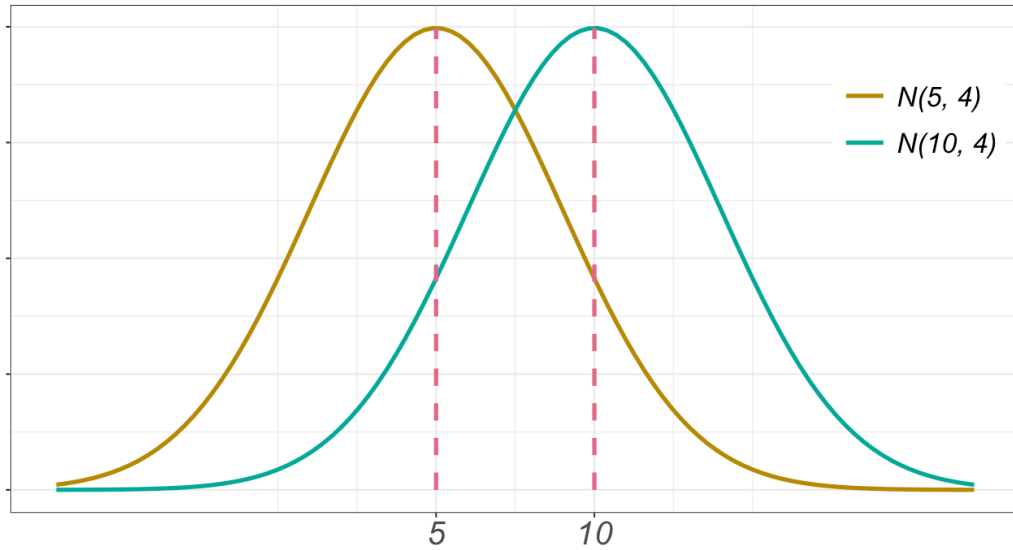
The Normal Random Variable

- A random variable X is said to be normally distributed random variable with parameters μ and σ^2 ,
 - i.e. $X \sim N(\mu, \sigma^2)$, if its density is

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}, \quad -\infty < x < \infty$$

- The normal density function is bell-shaped and symmetric about its mean μ
- The maximum value of the density function is $(\sigma\sqrt{2\pi})^{-1} \simeq 0.399/\sigma$ attains at $x = \mu$

Parameters of normal distribution and its density function



The normal random variable

- The expected value

$$E(X - \mu) = \int_{-\infty}^{\infty} (x - \mu) \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2} dx$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} y e^{-y^2/2} dy = 0$$

$$E(X) = \mu$$

The normal random variable

- The variance

$$\begin{aligned} \text{Var}(X) &= E(X - \mu)^2 = \int_{-\infty}^{\infty} (x - \mu)^2 \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2} dx \\ &= \sigma^2 \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} y^2 e^{-y^2/2} dy = \sigma^2 \end{aligned}$$

- Integration by parts

$$\int uv \, dx = u \int v \, dx - \int \frac{du}{dx} \left(\int v \, dx \right) dx$$

Distribution of a linear combination of a normal distribution

- If $X \sim N(\mu, \sigma^2)$, what is the distribution of $Y = a + bX$, where a and b are constants?
- If $M(t)$ is the moment generating function (mgf) of X , the mgf of $Y = a + bX$

$$M_Y(t) = \int e^{ty} f(x) dx = \int e^{t(a+bx)} f(x) dx = e^{ta} M_X(t)$$

Distribution of a linear combination of a normal distribution

- The mgf of $X \sim N(\mu, \sigma^2)$

$$M_X(t) = e^{t\mu + (t^2\sigma^2/2)}$$

- The mgf of $Y = a + bX$

$$\begin{aligned} M_Y(t) &= e^{ta} M_X(tb) \\ &= e^{ta} e^{t\mu b + (t^2 b^2 \sigma^2 / 2)} \\ &= e^{t(b\mu + a) + (t^2 b^2 \sigma^2 / 2)} \end{aligned}$$

- So, $Y \sim N(a + b\mu, \sigma^2 b^2)$

Standard normal distribution

- A normal distribution with mean 0 and variance 1 is known as the standard normal distribution, which is denoted by Z

$$Z \sim N(0, 1)$$

- The mgf of standard normal distribution: $M_Z(t) = e^{t^2/2}$
- The probability density function of a standard normal distribution

$$\phi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad -\infty < x < \infty$$

Standard normal distribution

- If $X \sim N(\mu, \sigma^2)$ then

$$Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$$

- Proof this relationship!
- It can be shown that

$$E(Z) = E\left[\frac{X - \mu}{\sigma}\right] = 0 \text{ and } Var(Z) = Var\left[\frac{X - \mu}{\sigma}\right] = 1$$

Cumulative distribution of the standard normal distribution

- The cdf of the standard normal distribution

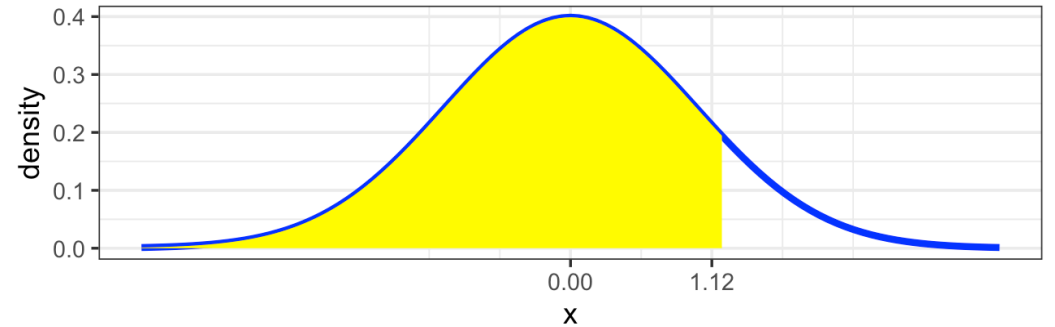
$$\Phi(x) = P(Z < x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy$$

- This integration cannot be evaluated algebraically
 - Most of the mathematical statistics book has Z table, which provides $\Phi(x)$ values for different x (see page 642 of the textbook for the Z table)
- An important relationship

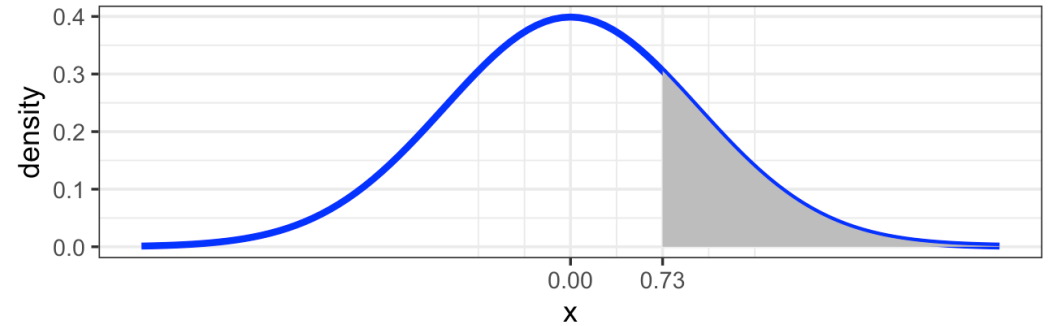
$$\Phi(x) + \Phi(-x) = 1$$

Probability calculation related to $Z \sim N(0, 1)$

$$\begin{aligned} P(Z < 1.12) &= \Phi(1.12) \\ &= 0.8686 \end{aligned}$$

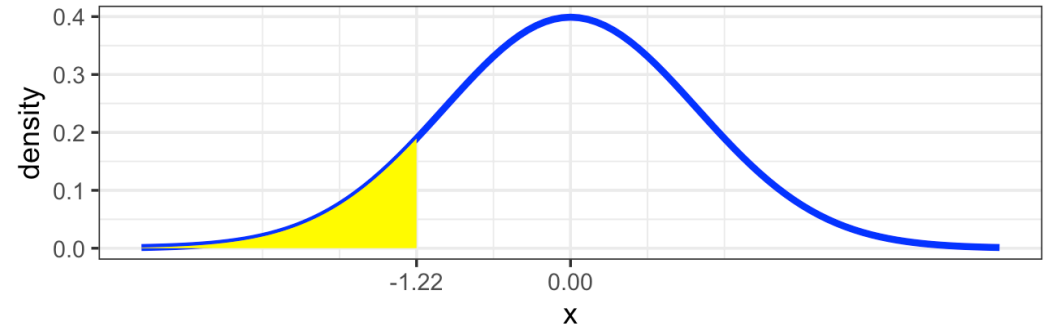


$$\begin{aligned} P(Z > .73) &= 1 - P(Z \leq .73) \\ &= 1 - \Phi(.73) \\ &= 1 - 0.7673 \\ &= 0.2327 \end{aligned}$$

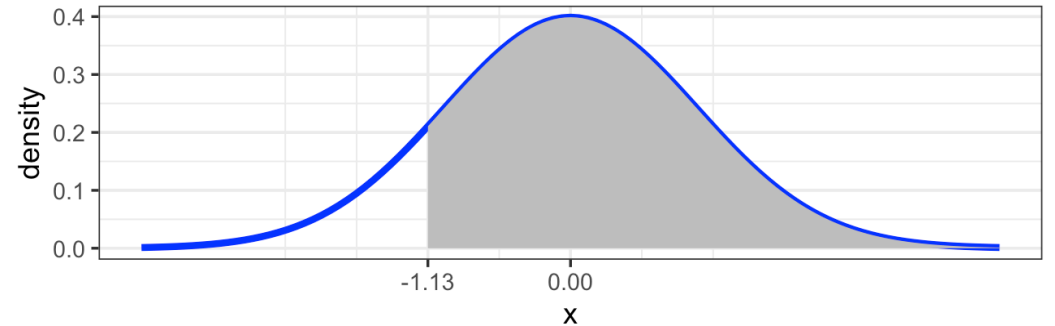


Probability calculation related to $Z \sim N(0, 1)$

$$\begin{aligned} P(Z < -1.22) &= \Phi(-1.22) \\ &= 1 - \Phi(1.22) \\ &= 1 - 0.8686 \\ &= 0.1112 \end{aligned}$$

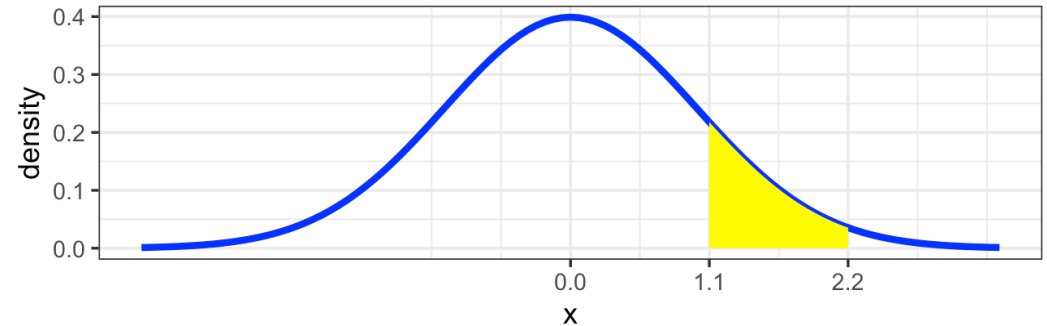


$$\begin{aligned} P(Z > -1.13) &= 1 - P(Z \leq -1.13) \\ &= 1 - \Phi(-1.13) \\ &= \Phi(1.13) \\ &= 0.8708 \end{aligned}$$

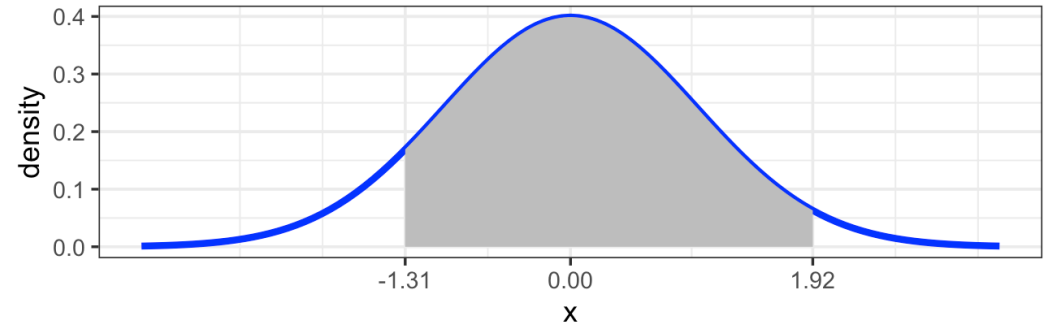


Probability calculation related to $Z \sim N(0, 1)$

$$\begin{aligned} P(1.1 < Z < 2.2) &= \Phi(2.2) - \Phi(1.1) \\ &= 0.9861 - 0.8643 \\ &= 0.1218 \end{aligned}$$



$$\begin{aligned} P(-1.31 < Z < 1.92) &= \Phi(1.92) - \Phi(-1.31) \\ &= \Phi(1.92) - 1 + \Phi(1.31) \\ &= 0.9726 - 1 + 0.9049 \\ &= 0.8775 \end{aligned}$$



Probability calculation related to $X \sim N(\mu, \sigma^2)$

- For any $a < b$

$$\begin{aligned} P(a < X < b) &= P\left(\frac{a - \mu}{\sigma} < \frac{X - \mu}{\sigma} < \frac{b - \mu}{\sigma}\right) \\ &= P\left(\frac{a - \mu}{\sigma} < Z < \frac{b - \mu}{\sigma}\right) \\ &= \Phi\left(\frac{b - \mu}{\sigma}\right) - \Phi\left(\frac{a - \mu}{\sigma}\right) \end{aligned}$$

Probability calculation related to $X \sim N(\mu, \sigma^2)$

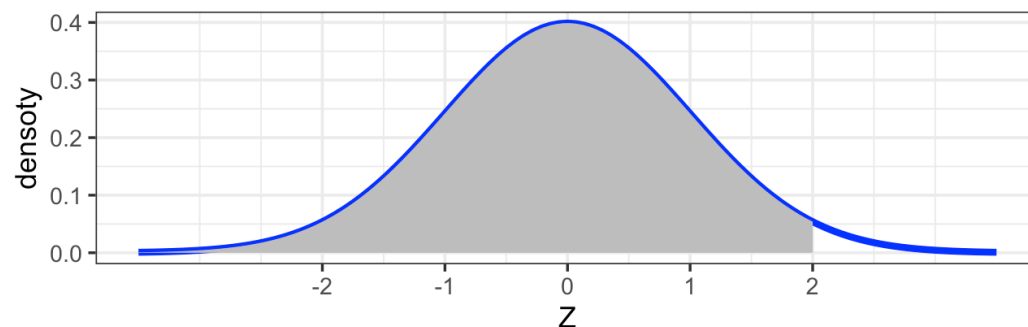
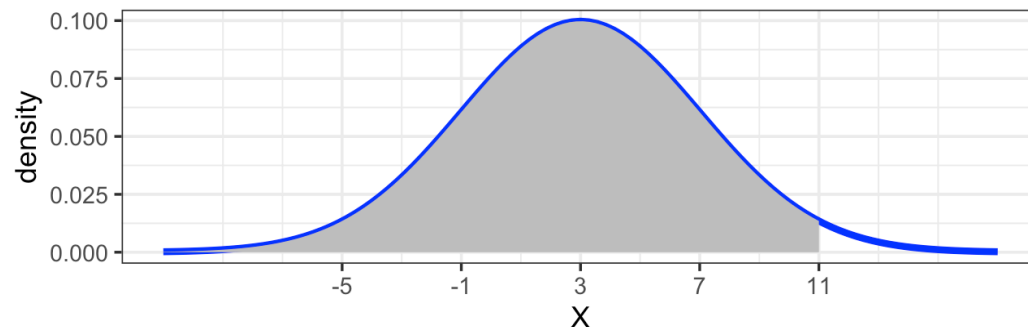
- For $X \sim N(3, 16)$

$$\begin{aligned} P(X < 11) &= \Phi\left(\frac{11 - 3}{4}\right) \\ &= \Phi(2) \\ &= 0.9772 \end{aligned}$$

- Calculate

$$P(X > -1)$$

$$P(2 < X < 7)$$



Sum of several independent normal variables

- Let X_1, \dots, X_n be n independent normal random variables, where

$$X_i \sim N(\mu_i, \sigma_i^2)$$

- The distribution of $Y = \sum_{i=1}^n X_i$ follows a normal distribution with

$$\text{mean} = \sum_{i=1}^n \mu_i \quad \text{and} \quad \text{variance} = \sum_{i=1}^n \sigma_i^2$$

Example 5.5d

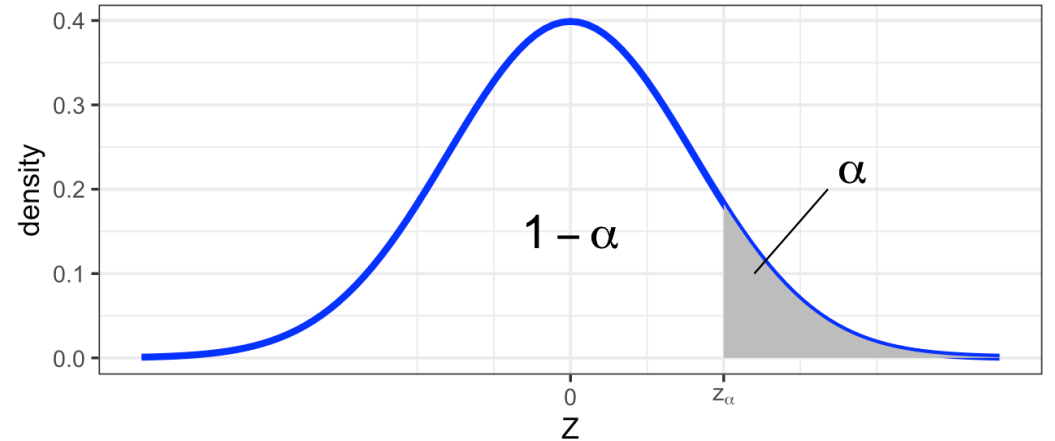
- Data from National Oceanic and Atmospheric Administration indicate that the yearly precipitation in Los Angeles is a normal random variable with a mean of 12.08 inches and standard deviation of 3.1 inches.
- Find the probability that total precipitation during the next two years will exceed 25 inches.

Quantiles/percentiles of standard normal distribution

- $z_\alpha \rightarrow$ the $100(1 - \alpha)$ percentile of standard normal distribution, i.e.,

$$P(Z < z_\alpha) = 1 - \alpha$$

- From the table, we can find
 - $z_{.05} = 1.645$
 - $z_{.025} = 1.96$
 - $z_{.01} = 2.33$



Exponential random variables

Exponential random variables

- An exponential random variable with parameter $\lambda > 0$ has the following probability density function

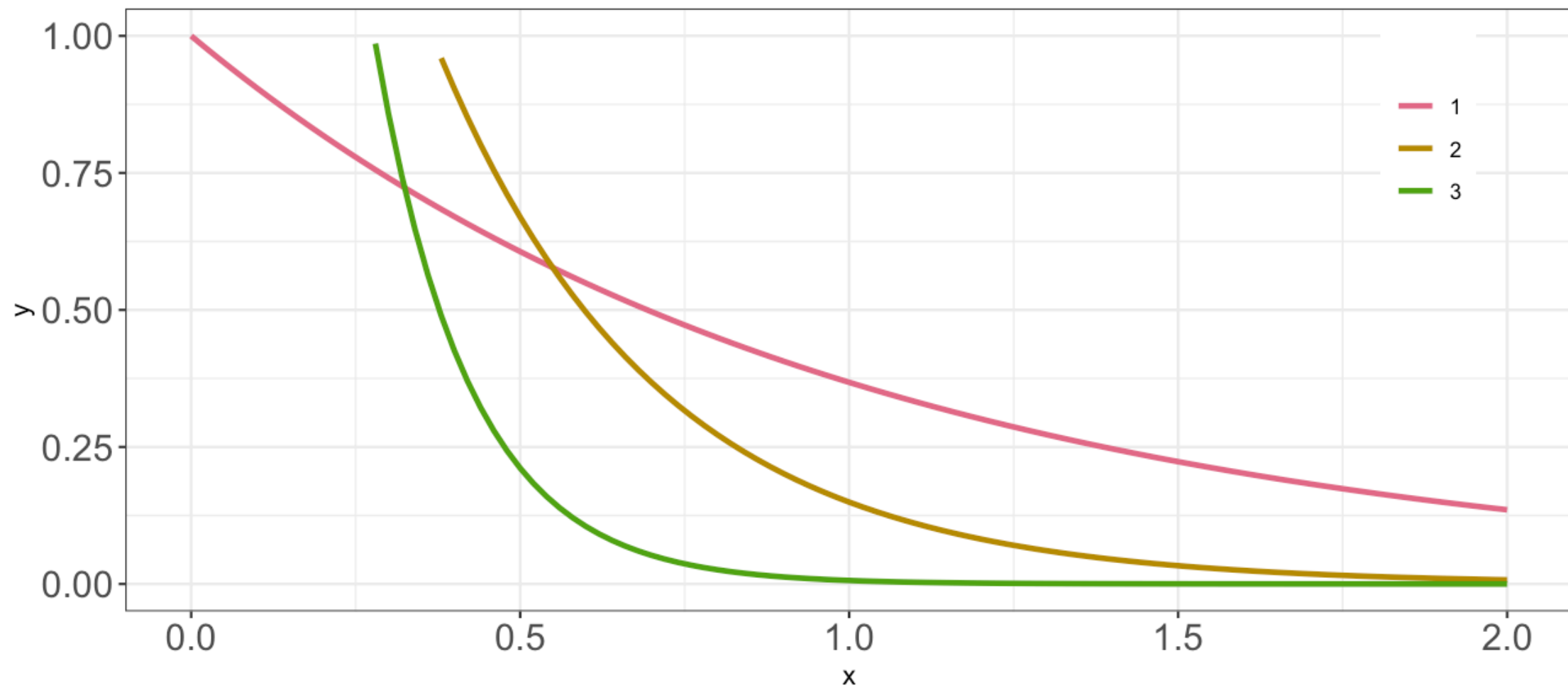
$$f(x) = \lambda e^{-\lambda x}, \quad x \geq 0$$

- The moment generating function of a exponential random variable

$$M_X(t) = \lambda/(\lambda - t)$$

- Show that

$$E(X) = 1/\lambda \text{ and } Var(X) = 1/\lambda^2$$



The cumulative distribution function

$$F(x) = \int_0^x \lambda e^{-\lambda y} dy = 1 - e^{-\lambda x}$$

- $P(a < X < b) = F(b) - F(a) = e^{-\lambda a} - e^{-\lambda b}$

Memoryless property of exponential random variable

- If X follows an exponential distribution with parameter λ

$$P(X > s + t \mid X > t) = P(X > s)$$

Exponential distribution

- If X_1, \dots, X_n are independent exponential random variables with parameters $\lambda_1, \dots, \lambda_n$, respectively
- Then $Y = \min\{X_1, \dots, X_n\}$ follows an exponential distribution with parameter $\sum_i \lambda_i$

The Gamma distribution

- A random variable is said to have a gamma distribution with parameters $\alpha > 0$ and $\lambda > 0$ if its density function is given by

$$f(x) = \frac{1}{\Gamma(\alpha)} \lambda e^{-\lambda x} (\lambda x)^{\alpha-1}, \quad x > 0$$

- $\Gamma(\alpha) = (\alpha - 1)\Gamma(\alpha - 1) = \int_0^\infty e^{-x} x^{\alpha-1} dx$

The Gamma distribution

- The moment generating function of $X \sim G(\alpha, \lambda)$

$$M_X(t) = \left(\frac{\lambda}{\lambda - t} \right)^\alpha$$

- $E(X) = \alpha/\lambda$ and $V(X) = \alpha/\lambda^2$
- Let X_1, \dots, X_n are independent exponentially distributed random variables with a common parameter λ , then

$$\sum_{i=1}^n X_i \sim G(n, \lambda)$$

Beta distribution

- A random variable X follows a beta distribution with parameters a and b if the probability density function of X is

$$f(x) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} x^{a-1} (1-x)^{b-1}, \quad 0 \leq x \leq 1$$

- $\int_0^1 x^{a-1} (1-x)^{b-1} dx = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)} = \text{Beta}(a, b)$

- Expectation and variance

- $E(X) = \frac{a}{(a+b)}$ and $V(X) = \frac{ab}{(a+b)^2(a+b+1)}$

Problems

- 1, 3, 5, 6, 7, 8, 10, 11, 13, 14, 23, 24, 25, 26, 33, 34, 35, 26, 37, 38